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# THE GREENNESS OF CITIES: CARBON DIOXIDE EMISSIONS AND URBAN DEVELOPMENT

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#### **ABSTRACT**

Carbon dioxide emissions may create significant social harm because of global warming, yet American urban development tends to be in low density areas with very hot summers. In this paper, we attempt to quantify the carbon dioxide emissions associated with new construction in different locations across the country. We look at emissions from driving, public transit, home heating, and household electricity usage. We find that the lowest emissions areas are generally in California and that the highest emissions areas are in Texas and Oklahoma. There is a strong negative association between emissions and land use regulations. By restricting new development, the cleanest areas of the country would seem to be pushing new development towards places with higher emissions. Cities generally have significantly lower emissions than suburban areas, and the city-suburb gap is particularly large in older areas, like New York.

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#### I. Introduction

While there remains considerable debate about the expected costs of global warming, a growing scientific consensus believes that greenhouse gas emissions create significant risks of climate change. A wide range of experts have advocated reducing individual carbon footprints and investing billions to reduce the risks of a major change in the earth's environment (Stern, 2008). Almost 40 percent of total U.S. carbon dioxide emissions are associated with residences and cars, so changing patterns of urban development and transportation can significantly impact emissions. How do major cities differ with respect to their per-household emissions levels?

In Section II of this paper, we review the basic theory of spatial environmental externalities. If emissions are taxed appropriately, then private individuals will make appropriate decisions about location choices without any additional location-specific policies. When emissions are not taxed, then location decisions will be inefficient. The optimal location-specific tax on building in one place versus another equals the difference in emissions times the gap between the social cost of emissions and the current tax on these emissions. Even if there was an appropriate carbon tax, location decisions might still be sub-optimal if governments subsidize development in high emissions areas or artificially restrict development in low emissions areas.

In Section III of this paper, we measure household carbon dioxide emissions production in 66 major metropolitan areas within the United States.<sup>3</sup> For a standardized household, we predict this household's residential emissions and emissions from transportation use. We look at emissions associated with gasoline consumption, public transportation, home heating (fuel oil and natural gas) and electricity usage. We use data from the 2001 National Household Travel Survey to measure gasoline consumption. We use year 2000 household level data from the Census of Population and Housing to measure household electricity, natural gas and fuel oil

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<sup>&</sup>lt;sup>1</sup> See also the critical reviews in Weitzman (2007) and Nordhaus (2007).

<sup>&</sup>lt;sup>2</sup> See http://www.eia.doe.gov/oiaf/1605/ggrpt/carbon.html for sources of carbon dioxide emissions.

<sup>&</sup>lt;sup>3</sup> Our work parallels the findings of the Vulcan Project at Purdue University (<a href="http://www.purdue.edu/eas/carbon/vulcan/index.php">http://www.purdue.edu/eas/carbon/vulcan/index.php</a>) and the recent Brookings Institution study by Brown and Logan (2008) fall into this category. Our exercise is slightly different since we look at the impact of a standardized household and we focus on marginal, rather than average, homes. For an example of international analysis that disaggregates greenhouse gas emissions variation within a nation, see Auffhammer and Carson's (2008) study of China.

consumption. To aggregate gasoline, fuel oil and natural gas into a single carbon dioxide emissions index, we use conversion factors. To determine the carbon dioxide impact of electricity consumption in different major cities, we use regional average power plant emissions factors, which reflect the fact that some regions' power is generated by dirtier fuels such as coal while other regions rely more on renewable energy sources. We distinguish between the emissions of an area's average house and the emissions of a marginal house by looking particularly at homes built in the last twenty years.

Our estimates suggest a range of carbon dioxide emissions from about 19 tons per household per year in San Diego and Los Angeles to about 32 tons in Oklahoma City and Memphis. The older cities of the Northeast tend to lie within those extremes. While people in these older cities drive less, they need large amounts of heating and produce more emissions as a result. For illustrative purposes, we use a social cost figure of 43 dollars per ton of carbon dioxide, which implies that the social cost of a new home in Houston is \$550 dollars more per year than the social cost of a new home in San Francisco.

We also use our methodology to compare the emissions in central cities and suburbs for 48 major metropolitan areas. In general, central city residence is associated with lower levels of emissions, although there are a few places where that fact is reversed. Carbon dioxide emissions differences within metropolitan areas are smaller than the differences across metropolitan areas. The place with the most extreme emissions difference between central cities and suburbs is New York, where we estimate that suburban development causes more than 300 dollars more damage in carbon dioxide emissions than central city development.

Across metropolitan areas, we find a weak positive connection between the level of emissions and recent growth when we weight by initial population size. We find a strong negative correlation between emissions and the level of land use controls. Overall, the metro areas with the lowest per-household carbon dioxide emissions levels are also the most restrictive towards new development. This fact suggests that current land use restrictions may be doing exactly the opposite of what a climate change activist may have hoped. Those restrictions, often implemented for local environmental reasons (such as to preserve open space or reduce neighborhood traffic), seem to push new development towards the least environmentally friendly

urban areas (Fischel 1999, Glaeser and Tobio, 2007). We now turn to the basic economics of environmental externalities and urban development.

#### II. Urban Development and Environmental Externalities

This theory section makes three simple points. First, if emissions are actually taxed at the appropriate rate then there is no need for further spatial policy to improve private decisions about location. Second, if emissions are taxed below the optimal level, then it is appropriate to subsidize the areas that have less energy usage and tax the areas with more energy usage. Third, even with an optimal emissions tax, suboptimal public policies, such as zoning or transport subsidies, may still lead to suboptimal locations.

We outline a simple model where location choice interacts with environmental externalities. We assume that there is a fixed population of size N identical individuals that must choose between I communities. The population of community i is denoted  $N_i$ . Individuals choose their communities and their level of energy consumption. We let "E" denote the amount of energy selected by each individual. This energy choice is meant to include household and transportation-related energy use.

Individuals maximize  $Y_i - P_i^H - (P_i^E + t)E + t\hat{E} + V(E; Z_i) - C(N\hat{E})$ , a quasi-linear utility function where Y refers to income,  $P_i^H$  refers to housing costs that are specific to region i,  $P_i^E$  refers to energy costs which are specific to region i, t refers to an energy tax which is initially independent of region,  $\hat{E}$  refers to the average energy consumption in the world as a whole,  $Z_i$  refers to attributes of the area which individuals treat as exogenous, V(.;.) reflects the region-specific benefits from energy use and  $C(N\hat{E})$  represents the costs of global energy consumption that may be associated with climate change. Each individual pays an energy tax of tE, but then receives a tax refund of  $t\hat{E}$  so that the tax is revenue neutral. The function V(.;.) allows different area characteristics to influence the benefits from energy use, and is meant to capture the

possibility that air conditioning might be more valuable in hot, humid places. We assume that V(.;.) is increasing and concave in its first argument.

Income and housing costs are derived from labor markets and housing markets. Specifically, each region has  $Q_i^F$  identical employers who earn revenues, denoted f(.), that are increasing and concave in the number of people hired. Each region also has  $Q_i^B$  of builders whose costs, denoted k(.), are increasing and convex in the number of buildings produced. The employers and builders are owned equally by all of the people in the country. These assumptions enable us to write that wage income equals  $f'\left(\frac{N_i}{Q_i^F}\right)$ , the marginal product of labor, and the cost of housing equals  $k'\left(\frac{N_i}{Q_i^B}\right)$ , the marginal cost of supplying housing. Each person receives an equal share of all business profits throughout the world.

Equilibrium is then determined by two optimality conditions. First, individuals must be choosing their private energy consumption to maximize their utility levels which implies that  $P_i^E + t = V_1(E_i^*; Z_i)$ , where  $E_i^*$  denotes privately optimal energy consumption conditional on prices and taxes in area i, and  $V_1(E_i^*; Z_i)$  is the derivative of V(.;.) with respect to its first argument. Individuals must also be indifferent between the different locations, which means that  $f'\left(\frac{N_i}{Q_i^F}\right) - k'\left(\frac{N_i}{Q_i^B}\right) - (t + P_i^E)E_i^* + V(E_i^*; Z_i) \text{ must be constant across space}.$ 

Since everyone is essentially identical, we focus on an additive social welfare function:

$$(1) \sum_{i} Q_{i}^{F} f\left(\frac{N_{i}}{Q_{i}^{F}}\right) - Q_{i}^{B} k\left(\frac{N_{i}}{Q_{i}^{B}}\right) + N_{i}\left(V(E_{i}; Z_{i}) - P_{i}^{E} E_{i} - C(N\hat{E})\right).$$

This yields first order condition for energy consumption:  $P_i^E + NC'(N\hat{E}) = V_1(E_i; Z_i)$ , which gives the standard result that the private optimality condition will be equivalent to the social optimality condition if  $t = NC'(N\hat{E})$ . The first order condition for social optimality locations is

that  $f'\left(\frac{N_i}{Q_i^F}\right) - k'\left(\frac{N_i}{Q_i^B}\right) + V(E_i; Z_i) - E_i\left(P_i^E + NC'(N\hat{E})\right)$  is constant across space. This condition is satisfied if  $t = NC'(N\hat{E})$ . There is no need for any added spatial policies if energy is properly taxed.

If  $t \neq NC'(N\hat{E})$ , then the spatial equilibrium is not Pareto optimal because people don't fully internalize the externalities associated with their energy use when they change locations. If energy use in an area is independent of the number of people in that area, then a location specific tax of  $E_i^*(NC'(N\hat{E})-t)$  transforms the private location decision into a second best social optimum, where people make the socially optimal location decisions conditional upon their socially suboptimal energy decisions. In comparing any two areas, the difference in tax payment for area i versus area j should equal  $(E_i^* - E_2^*)(NC'(N\hat{E}) - t)$ , the difference in energy usage times the difference between the optimal tax and the current tax. Our primary empirical exercise will be to calculate these quantities for different areas.

We can use the same model to ask when local environmentalism is good environmentalism. We model local environmentalism by assuming that a location imposes a location specific tax,  $\tau_i$  on energy usage in that state, and that revenues from this tax are rebated to the residents of the state. The first order condition for individual energy consumption is now  $P_i^E + t + \tau_i = V_1(E_i^{**}(\tau_i); Z_i), \text{ where } E_i^{**}(\tau_i) \text{ is a function mapping local energy taxes into local energy use.}$  The concavity of V(.;.) implies that  $E_i^{**}(\tau_i) < 0$ . Higher taxes will lead to local energy decisions that are better from a global perspective as long as  $t + \tau_i < NC'(N\hat{E})$ , but they will not necessarily increase welfare because these taxes also impact migration decisions.

To make this point, we reduce the world to only two regions and assume that there is no energy tax in region 2. We further assume that  $t + \tau_i \leq NC'(N\hat{E})$ . Differentiating the spatial equilibrium yields:

$$(2) \frac{\partial N_{1}}{\partial \tau_{1}} = \frac{E_{1}^{**'}(\tau_{1}) \left(V_{1}(E_{1}^{**}(\tau_{1}); Z_{1}) - (t + P_{1}^{E})E_{1}^{**}(\tau_{1})\right)}{\frac{1}{Q_{1}^{B}} k'' \left(\frac{N_{1}}{Q_{1}^{B}}\right) + \frac{1}{Q_{2}^{B}} k'' \left(\frac{N_{2}}{Q_{2}^{B}}\right) - \frac{1}{Q_{1}^{F}} f'' \left(\frac{N_{1}}{Q_{1}^{F}}\right) - \frac{1}{Q_{2}^{F}} f'' \left(\frac{N_{2!}}{Q_{2}^{F}}\right)} < 0,$$

so a tax on energy use in region one reduces the population of region one. This effect might be quite small, especially if the tax is modest, because the tax impacts migration behavior only by inducing people in area one to consume too little energy relative to the privately optimal level of energy consumption in the absence of this tax.

The tax in region one improves overall social welfare if and only if:

(3) 
$$\frac{NC'(N\hat{E}) - t - \tau_1}{NC'(N\hat{E}) - t} \left( -N_1 E_1^{**'}(\tau_1) \right) > \left( E_1^{**}(\tau_1) - E_2 \right) \frac{\partial N_1}{\partial \tau_1}$$

The left hand side of the equation is positive; the right hand side is negative if  $E_1^{**}(\tau_1) > E_2$ . If energy usage in region one is greater than energy usage in region two, then the impact of added energy taxes in that region must have a positive effect on welfare. In that case, the tax reduces both energy consumption, and the number of people in region one, which is desirable since it is the high energy use region.

If region one is using less energy than region two, then the situation is more ambiguous. If the migration margin is very large then it is at least conceivable that this tax will make the energy problem more problematic. A local tax that sets  $t + \tau_1 = NC'(N\hat{E})$  is certainly sub-optimal, since in that case the gains from reducing the tax on the migration margin will exceed the costs of reducing the tax in terms of increased energy usage in region one.

In many cases, this result may be more of an economic curiosity than a real concern. Many energy taxes seem too small to really impact migration behavior, at least if the taxes are rebated to residents in some way. However, environmentally inspired land use restrictions seem more likely to have counterproductive results. To model these interventions, we assume that location one has imposed a tax on new construction equal to  $z_1$  which is meant to refer to a "zoning tax." With this tax, the equilibrium first order condition for builders in location one

satisfies  $P_1^H = z_1 + k' \left( \frac{N_1}{Q_1^B} \right)$ . We assume that the tax either goes to infra-marginal residents of the community or that it is shared across both communities.<sup>4</sup>

Unlike the place-specific energy tax, the zoning tax does not impact energy use directly, but it does reduce the number of people in location one. Specifically:

$$(4) \frac{\partial N_1}{\partial z_1} = \frac{-1}{\frac{1}{Q_1^B} k'' \left(\frac{N_1}{Q_1^B}\right) + \frac{1}{Q_2^B} k'' \left(\frac{N_2}{Q_2^B}\right) - \frac{1}{Q_1^F} f'' \left(\frac{N_1}{Q_1^F}\right) - \frac{1}{Q_2^F} f'' \left(\frac{N_2}{Q_2^F}\right)} < 0.$$

The overall impact of zoning on social welfare is  $(E_2 - E_1)(NC'(N\hat{E}) - t) + z_1)\frac{\partial N_1}{\partial z_1}$ , which is positive as long as  $(E_1 - E_2)(NC'(N\hat{E}) - t) > z_1$ . If the area with the high zoning tax is also the high energy user, then the zoning tax will improve welfare, at least until the point where the tax equals the difference in energy usage times the difference between the social cost of energy use and the current tax. If the zoning tax is imposed in areas that have particularly low energy use, then it is counterproductive. This motivates our empirical exercise examining whether areas with extensive land use restrictions are also areas that have high levels of energy use.

#### III. Greenhouse Gas Emissions Across Metropolitan Areas

We now turn to estimating the quantity of carbon dioxide emissions that households produce in 66 major metropolitan areas.<sup>5</sup> Our goal is to calculate the marginal impact of an extra household in location j on the total carbon dioxide emissions of that location. The marginal household and the average household need not be the same, and we will try to create marginal estimates by comparing the emissions of an average household and the emissions associated with

<sup>&</sup>lt;sup>4</sup> If the tax is rebated only to new homeowners then the tax will be completely irrelevant.

<sup>&</sup>lt;sup>5</sup> Our sample includes 66 metropolitan areas with at least 250,000 households based on year 2000 Census IPUMS. In the year 2000, 72% of all metropolitan area residents live in one of these 66 metropolitan areas. We use the IPUMS definitions of metropolitan areas to assign households to metropolitan areas (see http://usa.ipums.org/usa-action/variableDescription.do?mnemonic=METAREA). Table Two lists the set of metropolitan areas that we study.

more recent development. Ideally, we would also be able to address the possibility that marginal emissions associated with more electricity generation are different from the average emissions, but we have no way of doing this well. In principle, the marginal resident could foster the development of a new lower polluting electric power plant, or the marginal megawatt of electricity could involve more harmful energy uses.<sup>6</sup>

We consider four main sources of carbon dioxide emissions: private within-city transport, public transportation, residential heating (natural gas and fuel oil) and residential electricity consumption. Car usage and home heating involves a relatively simple translation from energy use to carbon dioxide emissions. Household electricity use and public rail transit requires us to convert megawatt hours of usage into carbon dioxide emissions by using information about the carbon dioxide emissions associated with electricity production in different regions of the country. We are not considering the impact of shifting people on the energy emissions associated with moving goods and we are not considering the impact of shifting people on industrial output. The problem of figuring out how industrial location and the transport network changes with different urban development patterns is beyond the scope of this paper.<sup>7</sup>

One natural concern with our approach is that households in areas that spend more on energy have less income to spend on other things that also involve greenhouse gas emissions. If people in Texas are spending a lot on air conditioning and gas at the pump, then perhaps they are spending less on other things that are equally environmentally harmful. We cannot fully address this concern, since it would require a complete energy accounting for every form of consumption, but we do not believe our omissions fatally compromise our empirical exercise. After all, few forms of consumption involve nearly as much energy use as the direct purchase and use of energy. Moreover, areas that tend to have high levels of energy use are generally low cost areas like far flung suburbs or the Sunbelt, where people have more, not less, money available for other things. One can argue that the high land costs in expensive cities represent a transfer to earlier property owners who use their property-related revenues to buy more energy, but tracing through this chain of money and emissions is far too complicated a task for us.

<sup>&</sup>lt;sup>6</sup> To the extent that all regions have a similar relationship between marginal and average usage, then the implications of this work for inter city comparisons, may not be terribly effected by our inability to measure true marginal impacts.

<sup>&</sup>lt;sup>7</sup> Since much of modern industry is capital intensive and has low transport costs, we suspect it might not move that much in response to a population shift.

#### Car Usage and Emissions

We begin with estimating gasoline usage across metropolitan areas. Our primary data source is the 2001 National Household Transportation Survey (NHTS). This data source contains information on household characteristics and reported annual miles driven. The NHTS uses information on the types of vehicles the household owns to estimate annual gasoline consumption. The survey also reports the population density of the household's census tract, and zip code identifiers that enable us to use zip code characteristics to predict gasoline usage. We use these zip code identifiers to calculate each household's distance to the metropolitan area's Central Business District.

Our primary approach is to use the NHTS to predict gasoline usage based on individual and zip code level characteristics. We regress:

(5) 
$$Gasoline = \sum_{j} \beta_{j} Z_{k}^{j} + \sum_{q} \gamma_{q} X_{i}^{q} + \mu_{k} + \varepsilon_{i}$$

where  $Z_k^j$  refers to the value of zip code characteristic j in zip code k,  $\beta_j$  reflects the impact of those variables,  $X_i^q$  refers to the value of individual level q for person i,  $\gamma_q$  is the coefficient on that characteristic and the other two terms are individual level and zip code level error terms. Since there are a significant number of truly extraordinary outliers, and since we are running this regression in levels rather than logs, we top code the top one percent of the sample. The results of this equation are shown in Table 1.

The overall r-squared of the equation is 30 percent. Family size and income strongly increase gas consumption, so it is important to control for these characteristics. The area-level characteristics have the predicted signs. Population density, whether at the tract, zip code or metropolitan area level, reduces gasoline usage (see Golob and Brownstone 2008). Distance to the metropolitan's central business district also increases average gasoline consumption. We also

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<sup>&</sup>lt;sup>8</sup> For an analysis of how urban form affects vehicle miles traveled based on the 1990 version of this micro data set Bento et. al. (2005).

interact census tract density with region dummies and find that the density-gas consumption relationship is weaker in the West.

We then take these coefficients and predict gasoline usage for a family with an income of 62,500 dollars and 2.62 members for each census tract located within 66 major metropolitan areas. Specifically, our predicted value for a census tract with characteristics  $Z_k^j$  is  $\sum_j \beta_j Z_k^j + \sum_q \gamma_q X_{Ave}^q$ , where  $X_{Ave}^q$  denoted the individual characteristics of a standardized individual. We then form metropolitan area averages by aggregating up from the tract level using the tract's household count as the weight. 10

These estimates control for household level income and size, but they are, of course, imprecise. We are only using two primary characteristics for each tract, its proximity to downtown and its population density. As such, there will be an almost automatic relationship between urban sprawl and gasoline usage since gasoline usage decreases with density and increases with distance from downtown. There is a less automatic connection between gasoline consumption and metropolitan area population size, which is shown in Figure 1. On average, a .1 log point increase in MSA population size is associated with a 7.3 gallon reduction in the consumption of gallons of gas.

An alternative approach is to run regression (5) using metropolitan area fixed effects instead of region fixed effects, and then use those metropolitan area fixed effects as our measure of gasoline usage. In that case, we would have had to restrict our work to the small number of metropolitan areas with reasonably large data samples. We have estimated metropolitan area gasoline usage in this alternative manner, and the correlation between our measure and the measure estimated using metropolitan area fixed effects is high.

To estimate the gasoline related emissions of a marginal household, we again start with the gasoline consumption predicted at the tract level using our coefficients shown in Table 1. We then aggregate census tract gasoline usage up to the metropolitan area, by averaging across census tracts, weighting not by current population levels, but instead by the amount of housing

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<sup>&</sup>lt;sup>9</sup> These demographic statistics are based on the sample means for the 66 metropolitan areas from the year 2000 Census IPUMS.

<sup>&</sup>lt;sup>10</sup> We include all census tracts within thirty miles of the metropolitan area's CBD.

built between 1980 and 2000. If the location of housing in the near future looks like the location of housing in the near past, then the location of recent construction gives us some idea about where new homes will go.

On average, homes built in the last 20 years are associated with 47 more gallons of gasoline per household per year than average homes, which reflects the tendency to build on the urban edge. While we believe that focusing on recent housing patterns adjustment makes sense, it makes little difference to the cross-metropolitan area rankings. The correlation between estimated metropolitan area gasoline consumption using the total population of each census tract and the estimate based on the number of houses built since 1980 is .96.

To convert gallons of gasoline into carbon dioxide emissions, we multiply first by 19.564, which is a standard factor used by the Department of Energy.<sup>11</sup> This conversion factor includes only the direct emissions from a gallon of gasoline, not the indirect emissions associated with refining and delivering gas to the pump, which typically increase the energy use associated with a gallon of gas by 20 percent.<sup>12</sup> To reflect this, we assume that each gallon of gas is associated with 23.46 pounds of carbon dioxide emissions.

#### Public Transportation

We now turn to the emissions associated with public transportation. There are no adequate individual surveys that can inform us about energy usage by bus and train commuters. Instead, we turn to aggregate data for each of the nation's public transit systems from the National Transit Database<sup>13</sup>. For all of the nation's public transit systems, this data source provides us with information about energy used, which takes the form of gasoline in the case of buses and electricity in the case of rail. The data does not tell us about private forms of public transit, such as private bus lines or taxis or the Las Vegas monorail.

For each bus or rail system, the data set provides us with the zip code of their headquarters. We then assign each zip code to the relevant metropolitan area and sum up all of the gasoline and

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<sup>&</sup>lt;sup>11</sup> See http://www.eia.doe.gov/oiaf/1605/factors.html.

<sup>&</sup>lt;sup>12</sup>A typical energy efficiency figure for gasoline is 83 percent: <a href="http://frwebgate.access.gpo.gov/cgi-bin/getdoc.cgi?dbname=2000">http://frwebgate.access.gpo.gov/cgi-bin/getdoc.cgi?dbname=2000</a> register&docid=00-14446-filed.pdf

<sup>13</sup> http://www.ntdprogram.gov/ntdprogram/

electricity used by public transit systems within each metropolitan area. This provides us with total energy usage by public transit for each metropolitan area.

To convert energy use into carbon dioxide emissions, we continue to use a factor of 19.546 for gasoline. We again increase that factor by 20 percent to reflect the energy used in refining and distribution. The conversion for electricity is somewhat more difficult, since electricity is associated with different levels of emissions in different regions of the country. We will therefore be using different conversion factors for electricity in different places, and we will discuss those at length when we get to home electricity usage. By combing emissions from gas and emissions from electricity, we estimate a total emissions figure within the metropolitan area. To convert this to a household-level figure, we divide by the number of households in the metropolitan area.

There are two reasons why the marginal emissions from a new household might not be the same as the average emissions for an existing household. First, the marginal household might be more or less inclined to use public transportation. Second, even if the marginal household uses public transport, we do not know how much extra energy this will entail. Typically, we think of some public transit technologies as having large fixed costs, which could mean that the marginal costs are quite low. However, in some cases, new development may mean that a new bus line is extended to a newer, lower density area, and in this case, the marginal costs might be quite high. Since we lack the data to make an effective estimate of the marginal effect, we will use the average emissions from public transit throughout this paper. Since the emissions from private automobiles are on average fifty times higher than the emissions from driving, the benefits to our overall estimates of improving the accuracy of our public transit emissions measures are likely to be small.

#### Household Heating

We now turn to the emissions from the two primary household heating sources: fuel oil and natural gas. Fuel oil use is rare in the United States outside of the Northeast, and is an important

source of home heating in only a few metropolitan areas. Natural gas is the more common source of home heat. In some areas, electricity also provides heat, but we will deal with electricity separately in the next section.

For our purposes we need a large representative sample that provides information by metropolitan area on household heating. The Department of Energy's Residential Energy Consumption Survey<sup>14</sup> is too small of a data set to address our needs. This data set also does not provide each survey respondent's metropolitan area. Instead, we use data from the 2000 Census five percent sample (IPUMS). This data set provides information for each household on its expenditure on electricity, natural gas and fuel oil.

The key problem with the IPUMS data is that we are interested in household energy use, not energy spending. Conveniently, the Department of Energy provides data on prices for natural gas<sup>15</sup> and fuel oil<sup>16</sup> for the year 2000. These prices are at the state level, so we miss variation in prices within the state. We use these prices to convert household energy expenditure to household energy consumption.

One particular problem with the expenditure data is that some renters do not pay for energy directly, but are charged implicitly through their rents (Levinson and Niemann 2004). These renters will report zero energy expenditures, when they are indeed using electricity and some home heating fuel. Indeed, when we look at the frequency of reported zero expenditure in different metropolitan areas, we find that these tend to be disproportionate among renters and other residents of multi-family houses. In these cases, it is impossible to know whether a zero value for expenditure truly indicates that the household does not consume this particular fuel or whether the household just doesn't pay directly for that energy. As such, we have the most confidence in the IPUMS data for measuring actual household energy consumption for owners of single family homes.

We use the IPUMS 2000 data to estimate a separate regression for each of the 66 metropolitan areas using the subsample of owners of single family homes:

16 http://tonto.eia.doe.gov/dnav/pet/pet sum mkt a EPD2 PRT cpgal a.htm

<sup>14</sup> http://www.eia.doe.gov/emeu/recs/recs2001/publicuse2001.html

<sup>15</sup> http://www.eia.doe.gov/emeu/states/ seds.html

ittp://www.cia.doc.gov/cincu/states/\_scus.ittili

(6) Energy Use=a\*Log(Income)+b\*Household Size +c\*Age of Head+ MSA Effects.

In the case of natural gas in the New York City area, for example, we estimate:

(6') Natural Gas = 
$$-138 + 13 \bullet Log(Income) + 9.8 \bullet Size + .81 \bullet Age$$
.

Standard errors are in parentheses. In this regression, there are 28,757 single owner occupied housing unit observations and the r-squared is .02. For each metropolitan area, we estimate similar regressions for fuel oil and electricity consumption. We then use metropolitan area specific regression coefficients to predict the natural gas and fuel oil consumption for a household with an income of 62,500 dollars and 2.62 members.

We try to correct for individual characteristics, but we do not correct for housing characteristics. After all, we are not attempting to estimate emissions assuming that people in Houston live in New York City apartment buildings. The building sizes in an area are a key component in emissions and we want to include that. Our approach allows for the fact that a household with a fixed set of demographics is likely to live in a larger, newer home if it lived in Houston than it would have chosen if it lived in Boston or New York City, since land prices are higher in the latter cities. Our approach captures the fact that a standardized household will live its life differently depending on the relative prices that it faces in different cities.

To estimate energy consumption for renters and owners in multifamily units for each of the 66 metropolitan areas, we adjust our metropolitan area specific predictions that were based on estimates of equation (6). For example, we will estimate equation (6) using Census IPUMS data for Los Angeles owners of single family homes. This yields a prediction of average electricity consumption for Los Angeles home owners of single family homes for a household with standardized demographics. We still need to impute what this household's electricity consumption would have been if it had lived in Los Angeles as a renter of a single family home, an owner of a unit in a multi-family unit, or as a renter in a multi-family unit. To impute these last three categories, we use a second micro data set called the 2001 Residential Energy Consumption Survey (RECS).<sup>17</sup> This data set is a national sample with 4,392 households that

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<sup>17</sup> http://www.eia.doe.gov/emeu/recs/recs2001/publicuse2001.html

includes actual household energy consumption data. We use this energy consumption data to estimate national level OLS regressions; of the form:

(6") 
$$Log(Energy Consumption) = Controls + b_1*Owner + b_2*Multi-Family + U$$

Using the OLS estimates of b<sub>1</sub> and b<sub>2</sub> from these regressions, we adjust our metropolitan area specific predictions of energy consumption. For example, suppose that we estimate using the national data that b<sub>2</sub> equals zero and that b<sub>1</sub> equals .1. If based on our Los Angeles regression for owners of single family homes, we predict that the average home owner (with standardized demographics) consumes 9 megawatt/hours of electricity per year, then we would impute that the average renter consumes 8.18 megawatt/hours of electricity per year.<sup>18</sup>

This procedure allows us to predict a standardized household's consumption of energy for each metropolitan area, if it lived in four different housing categories. We then calculate a weighted average across these four categories by metropolitan area. The weights, which vary by metropolitan area, are based on the IPUMS data's frequency count of each of these four housing types. This multi-step method allows us to impute the energy consumption for renters and all residents in multi-family buildings, where we are concerned that reported energy expenditure does not accurately measure household consumption. Our correction procedure is especially important in a metropolitan area such as New York City, and is much less important in places like Houston where most of the households are single family owners.

Natural gas consumption is driven primarily by climate. Figure 2 shows the correlation between our estimated natural gas consumption and January temperature. We do not find the correlation coefficient of -.81surprising, but it does suggest that our results are reasonable.

For fuel oil and natural gas, there are again conversion factors that enable us to move from energy use to carbon dioxide emissions. In the case of fuel oil the factor is 22.38 pounds of carbon dioxide per gallon of fuel oil<sup>19</sup>. We again increase this number by 20 percent to reflect the energy used in refining and distributing. According to the same source, there are 120.59 pounds of carbon dioxide emissions per 1,000 cubic feet of natural gas. In this case, there is much less energy involved in distribution so we use this conversion factor without any

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<sup>&</sup>lt;sup>18</sup> The REX regressions are available on request.

<sup>19</sup> http://www.eia.doe.gov/oiaf/1605/factors.html

adjustment. We combine the emissions from natural gas and fuel oil to form an estimate of total home heating emissions.

To examine the impact of a marginal home, we repeat this procedure using only homes built between 1980 and 2000. Since older homes are less fuel efficient, the average home will overstate true energy use, especially in older areas of the country. We use only homes built within the last 20 years to minimize this effect. In principle, we could have used only homes built in the last five or ten years, but our sample sizes become too small if we limit our samples in this way. We will refer to these estimates as our estimates of marginal heating emissions.

#### Household Electricity

In the case of electricity consumption, we begin with the same IPUMS-based procedure used for fuel oil and natural gas. We use state-wide price data to convert electricity expenditure into consumption in megawatt hours<sup>20</sup>. We then regress estimated electricity consumption on household characteristics by metropolitan area, just as we did for home heating. We also follow the same imputation procedure for owners of multi-family units and all renters. Following this strategy, we predict household annual electricity consumption for each metropolitan area for a standardized household with 2.62 people earning an annual income of \$62,500.

In the case of electricity, consumption rises most sharply with July temperatures, as shown in Figure 3. The correlation is relatively strong (.61) but there are some significant outliers in the Pacific Northwest, namely Tacoma and Seattle. These places have particularly high electricity usage, relative to July temperatures, which reflects, in part, the low costs of electricity in that region.

The conversion between electricity usage and carbon dioxide emissions is considerably more complicated than the conversion between natural gas or petroleum usage and emissions. If we had a national market for electricity, then it would be appropriate to use a uniform conversion factor, but since electricity markets are regional, we must allow for different conversion factors in different areas of the country. There is considerable heterogeneity in the emissions for

<sup>&</sup>lt;sup>20</sup> http://www.eia.doe.gov/cneaf/electricity/epm/table5 6 b.html

megawatt hour of electricity between areas that rely on coal, like the Northeast, and areas that use more hydroelectric energy, like the West.

What geographic area should we use to calculate the emissions related to electricity usage? In principle, one could calculate anything from a national average of emissions per megawatt hour to a block specific figure. Using smaller levels of geography certainly increases the accuracy with which emissions are allocated to electricity usage. However, if electricity is perfectly substitutable between two places, then this precision is somewhat misleading, and irrelevant for estimating the marginal emissions associated with new construction. The relevant consideration is not the actual greenness of the particular area's supplier, but rather the average emissions of the entire area.

For example, consider a setting where there is a clean and a dirty electricity producer in a region, with identical costs of production and plenty of consumers who don't care about the source of their electricity. In equilibrium, both producers will generate the same amount of electricity. A new consumer who buys only from the clean producer will still be associated with the average level of emissions. Since these two providers are perfect substitutes, if a new resident buys only from the clean provider, then someone else will be buying from the dirty provider. For this reason, it makes sense to consider the average emissions within the market not the individual emissions of one particular place.

The North American Electric Reliability Corporation (NERC) has divided the U.S. into eight electricity markets. While electricity within these regions is not perfectly fungible and there is some leakage across NERC regions, there is much more substitutability of electricity within NERC regions than across regions. The difficulties involved in transmitting electricity over long distances mean that electricity in one region cannot readily substitute for electricity in another region. We therefore feel comfortable treating these markets as more or less closed systems (Holland and Mansur 2008).

We calculate NERC region average emissions data using power plant level data from the Environmental Protection Agency's eGRID, or Emissions & Generation Resource Integrated

Database data base<sup>21</sup>. The eGRID data base contains the emissions characteristics of virtually all electric power in the United States and includes emissions and resource mix data for virtually every electricity-generating power plant in the U.S. eGRID uses data from 24 different federal data sources from three different federal agencies: EPA, the Energy Information Administration (EIA), and the Federal Energy Regulatory Commission (FERC). Emissions data from EPA are integrated with generation data from EIA to create the key conversion factor of pounds of carbon dioxide emitted per megawatt hour of electricity produced (lbs/MWh).

Using eGRID, we calculate the emissions for megawatt hour for each of the NERC regions. There is remarkable heterogeneity across these regions (Holland and Mansur, 2008). For example, San Francisco is located in a NERC region that generates 1000 pounds of carbon dioxide for each megawatt hour of electricity. In contrast, Philadelphia is located in a NERC region where the average power plant in the region generates 1600 pounds of carbon dioxide for each megawatt hour.

We then use these conversion factors to turn household electricity usage into carbon dioxide emissions for each metropolitan area. We use the same conversion factor to handle the electricity consumption of commuter rails. To consider the impact of the marginal home, as above, we restrict our IPUMS estimates to homes built only between 1980 and 2000.

#### Overall Household Rankings

We finally turn to an overall ranking of metropolitan areas based on carbon dioxide emissions. Table 2 lists the 66 largest metropolitan areas for which we have data. The first column shows carbon dioxide emissions from predicted gasoline consumption within each metropolitan area.<sup>22</sup> There is considerable range in the consumption of gasoline at the metropolitan area level. The New York metropolitan area is estimated to use the least gasoline,

<sup>&</sup>lt;sup>21</sup> see http://www.epa.gov/cleanenergy/egrid/index.htm

<sup>&</sup>lt;sup>22</sup> These predictions are based on predicting gasoline consumption in each census tract for a standardized household. Within a metropolitan area, census tracts differ with respect to their population density and their distance to the City Center. Across metropolitan areas, census tracts differ with respect to their MSA's region and overall density. We exploit this variation as well to predict each tract's annual gasoline consumption per household. We then use census data on household counts to weight this tract level data into a metropolitan area level average prediction.

which reflects its high degree of employment and population concentration and its relatively heavy use of public transportation. Greenville, South Carolina, is estimated to have the most gasoline consumption. The gasoline-related emissions in Greenville are almost twice as high as the gasoline-related emissions in the New York area.

The second column reports our results on per household energy emissions due to public transportation. This column adds together rail and bus emissions and converts both by appropriate factors to arrive at carbon dioxide emissions. There is, of course, considerable heterogeneity. Emissions from public transportation in New York City are more than three tons of carbon dioxide from public transit per capita.<sup>23</sup> However, even in New York, these emissions are relatively modest relative to the contributions of cars, since public transportation shares infrastructure, like buses, and uses electricity.

The third column gives our results on fuel oil and natural gas. Again, the results show a fair amount of regional disparity. Detroit leads the country in home heating emissions and Boston is a close second. Much of the West has almost no emissions from home heating. In general, places that use fuel oil have much higher emissions than places that use only natural gas, which explains why emissions from this source are much lower in Chicago than in Detroit.

The fourth column shows electricity consumption and the fifth column shows the NERC-based conversion factor for converting electricity into emissions. To calculate electricity related emissions in each area, the fourth and fifth columns need to be multiplied together.<sup>24</sup> We show these columns separately to illustrate the role of electricity usage versus the role of clean electricity production. New Orleans is the leader in electricity usage, while residents of Buffalo consume the least electricity. San Francisco has the second lowest electricity usage in our data.

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<sup>&</sup>lt;sup>23</sup> We do not have data on energy consumption from public transit in Las Vegas.

<sup>&</sup>lt;sup>24</sup> Households use electricity not only at home but also where they shop and work. In results that are available on request, we have used the 2003 Commercial Building Energy Survey. This building level data set collects information on roughly 5000 buildings across the United States. While this data set does not have metropolitan area identifiers, it does provide information on the heating degree days and cooling degree days at the location of each of the buildings. We regress building energy consumption per worker on building type dummies and these climate measures. Using city level data from Burchfield et. al. (2005), we predict commercial building energy consumption per worker for each metropolitan area. The cross-metropolitan area correlation between commercial energy consumption prediction and our residential energy consumption measure is .65. On average across the metropolitan areas, commercial energy consumption per worker is 30% higher than residential consumption per household.

The sixth column sums together all of the different sources of carbon dioxide emissions. The table is ordered by the amount of these emissions. California's cities are blessed with a temperate climate and they use particularly efficient appliances and produce electricity in particularly clean ways. Four of the five cities with the lowest emissions levels are all in California. Providence, Rhode Island ranks in the top five due to its low electricity use.

The high emissions cities are almost all in the South. These places have large amounts of driving and very high electricity usage. Their electricity usage is also not particularly clean. Texas is particularly well represented among the places with the highest levels of emissions. Memphis has the absolute highest level among our 66 metropolitan areas. Indianapolis and Minneapolis are the northernmost places among our ten highest emission metropolitan areas.

New construction in the Northeast is generally between those extremes. These places use moderate amounts of electricity. They drive less than Californians, but use large amounts of fuel oil. The Midwest looks generally similar to the Northeast, but larger amounts of driving push gasoline emissions up.

In column seven, we multiply total emissions by 43 dollars per ton to find the total emissions-related externality associated with an average home in each location. The 43 dollar number is somewhat arbitrary, and we are using it purely for illustrative purposes. It is conservative relative to the Stern report (2008), which suggests a cost of carbon dioxide that is twice this amount, but it is considerably more aggressive than the numbers used by Nordhaus (2007). Tol (2005) is one meta-study that also suggests that this number may be somewhat too high while our number is in the middle of the range in Metcalf (2007). Using this figure, the range of costs associated with each home goes from \$1,148 dollars in San Diego to more than \$2,015 dollars in Memphis. This \$867 dollar gap is an annual flow, and at a discount rate of 5 percent, this would suggest a tax of 17,340 dollars on every new home in Memphis relative to San Diego. The last column gives standard errors for these cost estimates. The procedure for

<sup>&</sup>lt;sup>26</sup> It is relevant to note that carbon tax policy proposals have suggested taxes per ton of carbon dioxide roughly in this range. Metcalf (2007) proposes a bundled carbon tax and a labor tax decrease. As shown in his Figure Six, he proposes that the carbon tax start at \$15 per ton (in year 2005 dollars) now and rise by 4% a year. Under this proposal, the carbon tax per ton of carbon dioxide would equal \$60 per ton (in year 2005 dollars) by 2050.

estimating these standard errors is detailed in the statistical appendix.<sup>27</sup> The standard errors of the carbon emissions (measured in tons) equal the standard errors of the emissions costs divided by 43.

Table 3 shows the 66 Metropolitan Area ranking based on the subset of households who live in homes built between 1980 and 2000. Table 3's structure is identical to Table 2 but Table 3 provides an estimate of how average emissions vary across metropolitan areas for a standardized household who lives in housing built between 1980 and 2000. This is useful information for determining whether within MSA growth patterns are shrinking the city's average footprint.

The differences between the two tables tend to offset each other. People who live in new homes consume more gasoline, which reflects the tendency of new growth to be in the suburbs. However, new homes are more energy efficient and therefore have lower emissions from home heating. In general, we find that this ranking based on recent growth is highly positively correlated with the average rankings reported in Table 2.

The energy use differences between metropolitan areas are quite large. Our estimate is that a new house in coastal California is associated with two-thirds or less of the emissions associated with a new house in Houston or Oklahoma City. These differences suggest that changing urban development patterns can have potentially large impacts on total carbon emissions. Since residential and personal transportation are associated with about 40 percent of total emissions, a 33 percent reduction in these sources would reduce total U.S. emissions by 13 percent. Of course, any policy interventions would impact the flow of new housing, rather than the stock, so changes in urban development patterns would only reduce emissions gradually.

Our cost estimates suggest optimal location-specific taxes on development, in the absence of other carbon emission taxes. The six hundred dollar difference in emissions costs between the coastal California areas and Memphis suggests a flow tax of six hundred dollars per year for each household in Memphis. This is not a small number. If the tax were paid in a single lump sum payment, of perhaps \$12,000, then this would represent a sizable increase in the cost of living in

<sup>&</sup>lt;sup>27</sup> The standard errors for the predictions are based on the sampling variation in the 2001 NHTS data set as reported in Table 1. We are assuming that the large sample sizes in the IPUMS data set minimize the sampling error in our predictions of the other entries in Table 2.

Memphis. The U.S. Census tells us that the median value of a home in Memphis in 2006 was 91,000 dollars. Of course, the model suggests that a direct carbon tax would improve social welfare more than any location tax, so we believe that the main value of our results is only to suggest the external costs associated with moving to places like Memphis.

To study the cross-MSA correlates of greenhouse gas production, we present five separate OLS regressions in Table 4. In each of these regressions, the explanatory variables include the logarithm of average city income, the logarithm of city population, average January temperature and average July temperature. We also include a measure of the share of city centralization: the share of the population within five miles of the city center. The first column shows the correlates of private transportation related emissions. Income is uncorrelated with gasoline usage at the metropolitan level. At the individual level, there is a strong connection between gasoline consumption and income, but these estimates are supposed to correct for that relationship and they seem to do that. Larger metropolitan areas have somewhat less driving, which reflects the fact that these cities are somewhat denser. As the share of population within five miles of the city center increases by 10 percent, carbon dioxide emissions from driving decreases by 1300 pounds. Finally, places with warm Januarys have less driving, but places with hot Julys have more driving. These correlations are presumably spurious, and reflect other variables, like the degree of sprawl, associated with these weather variables.

The next regression shows the correlates of public transit emissions. In this case, city population is the only variable that is strongly correlated with emissions. Bigger cities are more likely to have extensive public transit systems. There is also a weak correlation between this outcome and the concentration of population within five miles of the city center.

The third regression looks at the relationship between home heating related emissions and the area-level variables. There is an extraordinarily strong negative correlation between this variable and January temperature, which was discussed above (also see Ewing and Rong 2008). Lower July temperatures also weakly increase home heating emissions. None of the other variables are strongly correlated with this outcome variable. The power of temperature to predict home heating emissions explains why the r-squared for this regression is higher than for any of the other regressions in this table.

The fourth regression correlates electricity related emissions with our independent variables. Areas that are more geographically concentrated have lower levels of electricity usage and lower emissions. The strongest determinant of home electricity usage in this regression, unsurprisingly, is July temperature. Still, the ability of the weather to explain electricity is weaker than the ability of the weather to explain home heating emissions.

Finally, the fifth regressions look at the correlates of total emissions. In this case, all of the variables except for city income are statistically significant. More populous cities have lower emissions, and this is being driven both by less electricity usage and by less driving. More decentralized cities have higher emissions, and this reflects less electricity and less driving. Places with milder Januarys have lower emissions, which is the result of less use of artificial heat. Places with hotter Julys have higher emissions, reflecting the electricity needed to run air conditioners.

As such, these regressions suggest that there are several different variables associated with lower levels of emissions at the city level. Older dense cities have lower emissions, but not if they are particularly cold. The temperate Sunbelt uses little electricity, but not the places with particularly hot summers.

What is the connection between low greenhouse gas emissions and city growth? Figure 4 shows the correlation between these marginal cost estimates and development in the area since 2000. Our dependent variable is the ratio between average annual housing permits in the area since the year 2000 and the total stock of housing in these places in the year 2000. This measure captures the extent to which the area is building new homes.

The overall relationship is basically flat, which suggests that current development patterns are neutral towards emissions. Unfortunately, that conclusion may be a bit optimistic because the correlation becomes significantly positive if we weight by the initial population of the area. The flat relationship that we see is driven primarily by Las Vegas and Phoenix, two areas that have high levels of growth and low levels of emissions. Without those areas, the relationship between growth and emissions becomes more strongly positive.

If moderate temperatures lower emissions and expenditure, then why aren't people moving to places with more temperate climates? One possible reason for the weak relationship

between new construction and per household emissions is land use regulations. As Figure 5 shows, there is a strong negative association between the Wharton Land Use Regulation Index and carbon dioxide emissions. This Regulatory Index is discussed in detail in Gyourko, Saiz, and Summers (2008).<sup>28</sup> Places with the least emissions tend also to regulate most heavily. This relationship is strongly statistically significant.

The negative connection between land use regulation and emissions is ironic but unsurprising. Environmentalists have fought both to reduce emissions and to restrict new development. In California, they have been successful in both fights. The result of this combination of activities is that the places with the lowest emissions in the country are also the places that have made it most difficult to build. We do not believe that California's small perhousehold footprint is caused by land use regulation. Californians' heavy reliance on driving and not using public transit is well documented (Kahn 2006). Instead, as documented in Table 2, California's relative greenness reflects a temperate climate and relatively clean electric utilities. California's regulatory authorities have been the nation's leader in enacting anti-pollution regulation. The state enacted more stringent vehicle emissions and earlier than the rest of the nation and now is pursuing the ambitious AB32 legislation signed into law by Governor Schwarzenegger in 2006. California's current low per-capita electricity consumption levels are a relatively new trend. In 1968, per-capita electricity consumption in California roughly equaled the nation's per-capita electricity consumption. Today, California's per-capita electricity consumption.

#### IV. Greenhouse Gas Emissions within Metropolitan Areas

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<sup>&</sup>lt;sup>28</sup> Gyourko, Saiz and Summers (2008) describe their index; "This aggregate measure is comprised of eleven subindexes that summarize information on the different aspects of the regulatory environment. Nine pertain to local characteristics, while two reflect state court and state legislative/executive branch behavior. Each index is designed so that a low value indicates a less restrictive or more *laissez faire* approach to regulating the local housing market. Factor analysis is used to create the aggregate index, which then is standardized so that the sample mean is zero and the standard deviation equals one."

In the previous section, we focused on cross-metropolitan area implications of greenhouse gas emissions. We now look within metropolitan areas, and focus on energy use differences between central cities and suburbs. After all, locating in central cities generally involves far less driving and living in smaller apartments. Since these choices are associated with fewer greenhouse gas emissions, they should also be seen as having fewer negative externalities.

Our approach is again to estimate the average energy consumption associated with locating in different areas, holding an individual's income and size constant, but not controlling for other choices like housing characteristics. Living in a larger house is a major part of moving to the suburbs for many people, and that should be captured in the environmental impact of suburbanization. We will use the same data sources and the same methodology as above, but we now focus on the differences between central city and suburban locations.

To keep definitions constant across data sources, we use the Census definition of Central City status, which we have for both census tracts and in the IPUMS. We exclude those data points that do not provide us with a central city identifier. This reduces our set of metropolitan areas down to 48. Sample sizes are unfortunately too small for us to provide robust estimates of emissions for the marginal home within metropolitan areas. As a result, we look only at the emissions associated with an average home.

To provide estimates of gasoline consumption in central cities and suburbs, we continue to use the regression results reported in Table 1 based on the 2001 National Household Travel Survey. This regression enables us to estimate the level of gasoline usage that a standardized household would purchase in each census tract. We then average all of the predicted gas usage numbers in census tracts that are in Central City PUMAs to form our estimate of Central City gasoline consumption. We do the same thing for suburban census tracts to form our estimate of suburban gasoline consumption. We continue to multiply gasoline usage by 23.47 to get total emissions.

As before, we compute gasoline usage for both marginal and average houses. We calculate average household gas consumption by averaging across census tracts using the total number of households in each census tract. We calculate marginal household gas consumption by averaging across census tracts, weighting them by the number of households built in the last ten years.

In the case of public transportation, we again calculate the total amount of emissions in the metropolitan area. We then allocate those emissions on the basis of public transportation usage. We calculate the total number of households in the central city and suburban commuters who use public transportation. We divide the total public transit emissions by this quantity to find the average public emissions per household that commutes using public transportation. We then multiply this number by the share of households in the suburbs and central city respectively that commute using public transportation to estimate the amount of public transit emissions associated with central city and suburban households.

For fuel oil and natural gas, we continue to use our IPUMS methodology of converting spending into energy use. In this case, the methodology is very dependent on central city and suburban residents facing the same fuel prices. We estimate our regressions separately for each metropolitan area, and in this case we also estimate an indicator variable that takes on a value of one if the household is in the suburbs. This indicator variable provides us an estimate of how much extra fuel is being consumed in suburban areas. We continue to multiply fuel and gas usage by the standard conversion measures to turn them into emissions.

We use the same procedure for electricity. We regress estimated electricity consumption on personal characteristics and a dummy variable that indicates a suburban location. We use the coefficient on that dummy variable as our estimate of the extra electricity associated with suburban living. We multiply this dummy variable by the NERC electric utility emissions factor to calculate the total emissions difference associated with electricity in the central cities and the suburbs. As discussed in the heating section above, we perform a correction using the 2001 RECS data to address the problem that renters and owners in multi-family units may not pay for their own electricity or home heating.

The suburban versus center city differentials are reported in Table 5. This table reports estimates for major metropolitan areas for which the IPUMS reports within metro area geography such that both center city residents and suburban residents can be identified. This yields a sample of 48 metropolitan areas. The first column shows the results for gasoline consumption. The city-suburb gap, in Table 5, ranges from 691 pounds of carbon dioxide (about 30 gallons of gas) in Los Angeles to ten times that amount in Philadelphia. Interestingly, there are large gaps in gas emissions both in older cities, where people in the central city take public

transportation, and in newer cities, where everyone drives but people in the suburbs drive much more.

In the second columns of Table 5, we turn to public transportation related emissions. Hartford has the largest central city-suburb gap in these emissions (2900 pounds of carbon dioxide), followed by Chicago, Seattle, then New York. Riverside has almost no gap. While public transportation made little difference to the metropolitan area figures, it does matter here. Since the central city populations tend to be the big users of public transportation and those populations are sometimes much smaller than the overall populations, the emissions that we credit to those people can be reasonably high. For example, in the case of New York City, more than one-third of the gains in reducing car-related emissions that are associated with central city residents are offset by higher emissions from public transit.

In the third column of Table 5, we turn to heating-related emissions. In this case, there is considerable heterogeneity across metropolitan areas. In New York, central city residents emit more than 6000 pounds of carbon dioxide less than suburbanites. In Detroit, central city residents emit more than 6000 pounds of carbon dioxide more than suburbanites.

The fourth column in Table 5 shows our result for electricity emissions. This column multiplies the NERC factor with the electricity usage gap. Almost everywhere, smaller urban homes mean lower electricity usage. Suburban electricity usage is lower in five cases when we consider average homes and in eight cases when we look at newer homes. Central city electricity usage does not always decline when we focus on newer homes, because while those homes may be more efficient, they are also more likely to have air conditioning.

The fifth column of Table 5 combines the results to show the total emissions gap between central cities and suburbs by metropolitan area. The sixth columns multiply this quantity by 43 dollars to find the total emissions cost, which ranges from -77 dollars, in Detroit, to 289 dollars, in New York. New York has the biggest gap between central city and suburbs. There are only two areas where suburbs have lower emissions than central cities. The seventh column shows the standard errors of the difference in costs which are again fairly small.

Table 6 regresses these differences on the same urban characteristics that we used in Table 4 to explain cross area differences in total carbon dioxide emissions. The dependent variable is the difference in emissions between the suburbs and the central city. The first regression shows that in bigger cities, suburbanites are more likely to drive longer distances

relative to central city residents. The suburb-central city driving gap also gets larger in places with warm Julys and shorter in places with warm Januarys.

The second regression shows that the impact of population on emissions is reversed when we look at public transit. In this case, big city residence is particularly likely to be associated with high levels of public transit emissions, which is, after all, what we saw in New York City in Table 5. In richer cities, the gap also increases.

In the third regression, we see that the heating gap between central cities and suburbs is larger for bigger, richer and more centralized cities. Interestingly, there is no connection between temperature and the city-suburb heating gap. The fourth regression shows that temperature and income, but not city population or centralization, predict the difference in electricity emissions.

The fifth regression looks at the correlates of the total suburb-city emissions gap. The gap is larger in cities with more income and more people. It is also larger when January temperatures are high and when July temperatures are high.

#### V. Conclusion

Past research has investigated how greenhouse gas emissions vary as a function of the scale of population and income. This paper has documented that holding population and income constant, that the spatial distribution of the population is also an important determinant of greenhouse gas production. If the urban population lived at higher population density levels closer to city centers in regions of the country with warmer winters and cooler summers in areas whose electric utilities used less coal for producing power, then greenhouse gas production would be lower.

If carbon dioxide emissions are taxed appropriately, then individuals will make appropriate decisions about their locations without any further government interventions. However, if we believe that current carbon taxes, which are essentially zero, do not charge people for the full use of their energy consumption, then location decisions will fail to internalize environmental costs. In this paper, we have quantified the greenhouse gas externality that a standardized household

would create if it were located in one of 66 major metropolitan areas and if it were located in the center city or suburb of 48 major metropolitan areas.

We estimate that costs per household range from 830 dollars per year in San Diego to almost 1410 dollars per year in Memphis. Across areas, emissions are positively associated with July temperature, negatively associated with January temperature, and negatively associated with both city population and centralization. New York has the biggest suburban minus central city gap of 289 dollars while Detroit has a central city minus suburban gap of 77 dollars.

Our work has many limitations. To translate our quantity estimates into dollar cost estimates, we have relied on the index weight of \$43 of damage per ton of carbon dioxide, and this number lies within a large confidence interval. Our estimates are based on regressions that can provide only a very imperfect estimate of gasoline usage or electricity consumption in particular areas. We restricted ourselves to household energy use.

That being said, this paper does provide what we consider to be reasonable estimates of the emissions-related externalities associated with homes built in different areas. However, we would be skeptical about actually using these numbers as the basis for a tax on development in Oklahoma or a subsidy for development in San Diego. There are surely much better ways, like a direct carbon tax, to get people to internalize the social costs of their actions. Perhaps, the clearest public policy-related conclusion that comes out of this analysis is that current land use controls seem to operate in a way that increases, rather than decreases, carbon dioxide emissions.

#### **Statistical Appendix for Calculating Standard Errors**

If we estimate a regression of the form  $Y = X'\beta + \varepsilon$ , then by Slutsky's Theorem and the Law of Large Numbers (LLN) the resulting coefficients will converge to a normally distributed random variable as follows:

$$\hat{\beta} \rightarrow^d N(\beta, \sigma^2_{\beta})$$

Where  $\sigma^2_{\beta}$  is an appropriately defined standard error.

Note that in general the resulting predicted values may be written as

$$\hat{Y} = \overline{X}' \hat{\beta}$$

Where  $\overline{X}$  is any  $x \in Support(X)$ . Therefore by a separate application of Slutsky's theorem  $\hat{Y}$  has the following asymptotic distribution:

$$\overline{X}'\hat{\beta} \to^d N(\overline{X}\beta, \overline{X}'\sigma^2{}_{\beta}\overline{X})$$

Here we define

$$\overline{X}_i = \begin{pmatrix} \overline{x}_1 \\ \overline{x}_2 \\ x_{i3} \\ \vdots \\ x_{ik} \end{pmatrix}$$

Further we collapse each estimated  $\hat{Y}$  by MSA. This is equivalent to multiplying each  $\hat{Y}$  by a vector  $\gamma'_{MSAj}$  where  $\gamma_{MSAj}$  contains  $\frac{1}{n_j}$  for every observation in MSA "j", a 0 otherwise, and  $n_{j_j}$  is the number of observations in MSA "j".

$$\gamma_{MSAj} := \begin{pmatrix} \frac{1}{n_j} \\ \vdots \\ \frac{1}{n_j} \\ 0 \\ \vdots \\ 0 \end{pmatrix}$$

Define  $\overline{Y}_{MSAj} = \gamma'_{MSAj} \hat{Y}$ . By a third application of Slutsky's theorem the appropriate asymptotic distribution of  $\overline{Y}_{MSAj}$  can be written as

$$\overline{Y}_{MSAj} \rightarrow^d N(\gamma'_{MSAj}\overline{X}\beta, \gamma'_{MSAj}\overline{X}'\sigma^2_{\beta}\overline{X}\gamma_{MSAj})$$

Now define

$$\Omega := \gamma'_{MSAj} \overline{X}' \sigma^2_{\beta} \overline{X} \gamma_{MSAj}$$

and further define  $C_{\overline{Y}_{MSAj}}$  to be an lpha confidence interval for  $\overline{Y}_{MSAj}$  .

$$C_{\overline{Y}_{MSAj}} = \left[ \overline{Y}_{MSAj} - |\Omega^{-1/2} Z_{\alpha}|, \overline{Y}_{MSAj} + |\Omega^{-1/2} Z_{\alpha}| \right]$$

$$Z_{\alpha} := \Phi^{-1}(0.95)$$

Thus if we estimate  $\Omega$  as

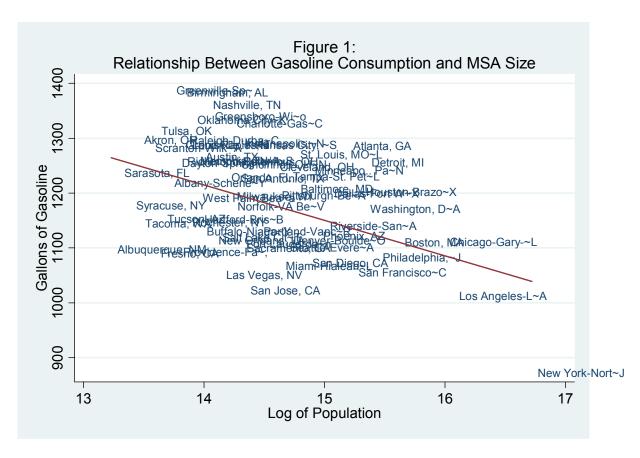
$$\hat{\Omega} = \gamma'_{MSAi} X' \hat{\sigma}^{2}_{\beta} (\gamma'_{MSAi} X')$$

then we must only replace  $Z_{\alpha}$  with its appropriate counterpart from the t-distribution.

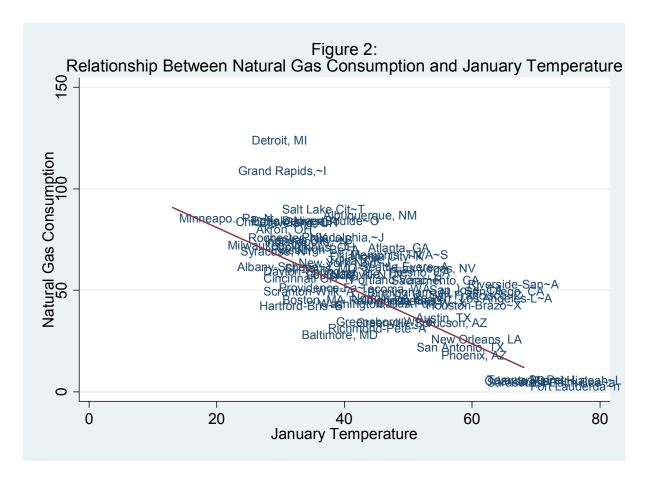
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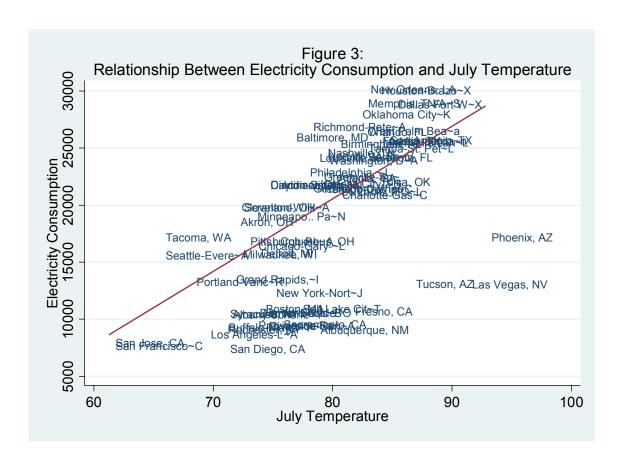
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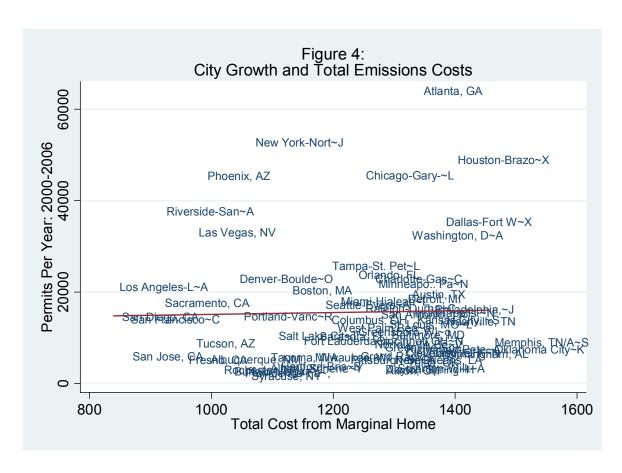
Notes: Gasoline consumption was estimated using the 2001 National Household Transportation Survey. Population is from the U.S. Census



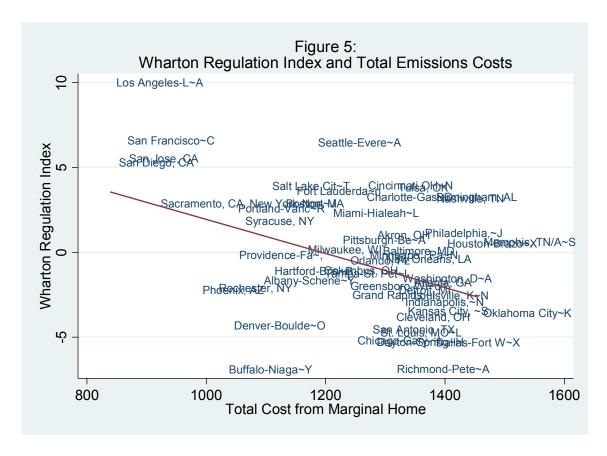
Notes: Natural Gas Consumption was estimated using the Integrated Public Use Microdata Series from the 2000 Census, the Department of Energy prices for natural gas, and the Department of Energy's Residential Energy Consumption Survey (RECS) for 2001. January Temperature is from the National Oceanic and Atmospheric Administration.



Notes: Electricity Consumption was estimated using the Integrated Public Use Microdata Series from the 2000 Census, the Department of Energy prices for electricity, and the Department of Energy's Residential Energy Consumption Survey (RECS) for 2001. Electricity Consumption was estimated using July Temperature is from the National Oceanic and Atmospheric Administration.



Notes: Housing permit data is from the U.S. Census. Total Cost from Marginal Home was estimated using data from the Integrated Public Use Microdata Series from the 2000 Census, the 2001 National Household Transportation Survey (NHTS), the Department of Energy, the National Transit Database, the 2001 Residential Energy Consumption Survey (RECS), and the Environmental Protection Agency's Emissions & Generation Resource Integrated Database (eGRID).



Notes: The Wharton Regulation Index is discussed in detail in Gyourko, Saiz, and Summers (2008). Total Cost from Marginal Home was estimated using data from the Integrated Public Use Microdata Series from the 2000 Census, the 2001 National Household Transportation Survey (NHTS), the Department of Energy, the National Transit Database, the 2001 Residential Energy Consumption Survey (RECS), and the Environmental Protection Agency's Emissions & Generation Resource Integrated Database (eGRID).

Table 1:
Gallons of Gasoline Consumed Per Year

	<b>Household's Annual</b>
	<b>Total Gasoline</b>
	Consumption
	(Gallons)
Midwest	78.998
	(0.41)
South Dummy	31.799
	(0.2)
West Dummy	-417.605
	(4.32)
Log(Zip Code Distance to CBD)	64.118
	(5.28)
Log(Census Tract Density)	-116.851
	(17.25)
Log(Metropolitan Area Density)	-38.4
	(2.12)
Log(Census Tract Density)*Midwest	3.67
	(0.17)
Log(Census Tract Density)*South	10.215
	(0.53)
Log(Census Tract Density)*West	59.899
	(5.37)
Log(household income)	299.264
	(16.4)
household size	163.704
	(28.8)
household head age	3.127
	(6.44)
Constant	-1731.832
	(6.92)
Observations	11728
R2	0.30

## Notes.

- (1) Data is from the 2001 National Highway Travel Survey (NTHS).
- (2) The unit of analysis is a household.
- (3) A dummy variable indicating that the head of household's age is missing is included.
- (4) Top 1% set as topcode
- (5) Standard errors are clustered by metropolitan area and are reported below the regression coefficient estimate.
- (6) The omitted category is a household in the Northeast region.

Table 2: Annual Standardized Household CO 2 Emissions

MSA	Emissions from Driving (Lbs of CO2)	Emissions from Public Transportation (Lbs of CO2)	Emissions from Home Heating (Lbs of CO2)	Electricity (Megawatt Hrs)	NERC Power Plant Emissions Factor (Lbs of CO2 per Megawatt Hrs)	Carbon Dioxide Emissions Cost	Standard Error (\$ per Year)
	<del></del>					(\$ per Year)	
San Diego, CA	24,774	689	5,994	7.18	1,007	1,148	14.87
San Francisco, CA	23,970	1,675	6,784	7.03	1,007	1,152	17.17
San Jose, CA	23,649	2,058	7,030	7.75	1,007	1,175	16.03
Providence, RI	22,562	1,273	12,965	7.35	1,185	1,177	13.24
Los Angeles, CA	23,553	1,062	6,439	8.43	1,007	1,188	17.91
Sacramento, CA Hartford, CT	25,534 23,092	458 1,539	6,875 13,752	9.07 8.09	1,007 1,185	1,237 1,239	13.96 17.71
Riverside, CA	26,380	42	6,461	9.27	1,007	1,239	13.64
Boston, MA	22,870	2,276	14,019	7.92	1,185	1,240	11.91
Tucson, AZ	26,363	616	4,535	12.25	1,007	1,270	14.65
Buffalo, NY	24,400	1,124	11,481	6.97	1,185	1,277	14.04
Las Vegas, NV	24,257	0	6,714	13.25	1,007	1,280	15.05
Albuquerque, NM	25,229	648	10,741	8.92	1,007	1,296	14.83
Fresno, CA	25,662	951	7,634	10.48	1,007	1,304	14.62
Rochester, NY	25,732	902	11,377	7.27	1,185	1,306	16.64
Phoenix, AZ	25,543	75	2,627	16.39	1,007	1,307	13.71
Denver, CO	25,159	1,374	10,494	10.10	1,007	1,336	13.76
Portland, OR	25,915	2,098	5,854	13.58	1,007	1,347	13.90
Syracuse, NY	26,744	574	11,588	8.07	1,185	1,347	17.78
Albany, NY	26,277	1,054	11,653	8.05	1,185	1,352	17.21
New York, NY	18,081	6,386	12,503	7.83	1,400	1,379	6.33
Salt Lake City, UT	25,491	3,104	11,146	10.15	1,007	1,406	13.90
Tacoma, WA	26,169	430	5,942	18.36	1,007	1,422	14.48
Seattle, WA	25,234	5,948	6,762	15.50	1,007	1,477	13.79
Pittsburgh, PA	25,591	2,093	12,313	9.67	1,614	1,600	15.88
Cleveland, OH	26,784	1,733	10,980	10.90	1,614	1,633	13.00
Akron, OH	28,604	768	10,652	10.91	1,614	1,644	16.94
Scranton, PA	27,611	282	13,173	10.19	1,614	1,651	18.64
Fort Lauderdale, FL	25,392 22,784	1,124	539	17.47	1,427	1,695	18.26
Philadelphia, PA Sarasota, FL		3,993 510	13,688 532	12.30 16.29	1,614 1,427	1,698 1,701	9.59
Milwaukee, WI	28,155 26,315	1,291	10,117	9.35	1,427	1,701	14.55 12.73
Columbus, OH	27,997	278	9,291	10.14	1,614	1,720	14.94
St. Louis, MO	28,105	1,267	8,749	13.54	1,472	1,737	13.98
West Palm Beach, FL	27,233	616	677	17.82	1,427	1,738	15.40
Tampa, FL	28,034	742	673	17.36	1,427	1,743	14.88
Cincinnati, OH	27,537	770	8,784	12.83	1,543	1,764	12.09
Miami, FL	24,187	4,689	896	17.92	1,427	1,768	21.69
Chicago, IL	24,278	5,221	10,374	9.83	1,614	1,781	15.21
Orlando, FL	28,174	1,361	734	18.48	1,427	1,789	14.44
Norfolk, VA	27,091	1,078	5,561	16.01	1,472	1,792	15.43
New Orleans, LA	24,899	663	4,964	19.05	1,472	1,795	19.18
Raleigh-Durham, NC	29,922	495	5,797	14.55	1,472	1,798	13.96
Greensboro, NC	31,300	216	4,747	14.53	1,472	1,799	15.09
Grand Rapids, MI	29,248	572	14,362	8.23	1,614	1,811	18.57
Charlotte, NC	30,820	1,084	5,963	14.25	1,472	1,825	14.58
Kansas City, MO	28,763	644	10,319	13.50	1,561	1,830	16.51
San Antonio, TX	27,694	1,929	4,110	15.74	1,555	1,832	15.24
Washington, DC	25,918	4,729	5,674	13.72	1,543	1,832	18.46
Baltimore, MD	26,540	2,135	5,405	13.78	1,614	1,835	17.57
Richmond, VA	29,459	771	4,101	16.87	1,472	1,835	15.23
Louisville, KY	27,880	884	8,538	14.92	1,543	1,837	13.19
Greenville, SC	32,169	130	4,964	15.25	1,472	1,841	16.90
Dayton, OH	28,888	986	9,027	12.69	1,614	1,847	17.89
Tulsa, OK	29,091	353	8,729	13.69	1,561	1,855	14.80
Detroit, MI	27,403	889	16,511	9.23	1,614	1,862	13.22
Atlanta, GA Minneapolis-St. Paul, MN	29,425	1,121	8,851	14.63	1,472	1,866	13.96
1	27,427	143	10,990	10.12	1,819	1,866	13.51
Indianapolis, IN	29,222	534	10,665	12.80	1,614	1,888	16.62
Austin, TX	29,134	1,595	4,613	16.58	1,555	1,892	15.02
Dallas, TX Houston, TX	27,323	1,723	6,100 5,344	17.81	1,555	1,926	16.30
Birmingham, AL	27,333 30,041	1,447 227	5,344 7,759	18.74 16.64	1,555 1,472	1,932 1,937	15.98 14.82
Nashville, TN	30,495	473	6,699	17.21	1,472	1,954	14.82
Oklahoma City, OK	28,953	332	8,710	16.41	1,649	2,005	14.71
Memphis, TN	28,440	1,073	8,438	18.70	1,472	2,005	14.73
ршо, 111	20,770	1,075	0,750	10.70	1,7/4	2,013	14.43

<sup>(1)</sup> Data is from the Integrated Public Use Microdata Series from the 2000 Census, the 2001 National Household Transportation Survey (NHTS), the Department of Energy, the National Transit Database, the 2001 Residential Energy Consumption Survey (RECS), and the Environmental Protection Agency's Emissions & Generation Resource Integrated Database (eGRID).

(2) See text for detailed descriptions of the data calculations

Table 3: Annual Standardized Household CO2 Emissions for Households Living in Homes Less Than Twenty Years Old

MSA	Emissions from Driving	Emissions from Public Transportation	Emissions from Home Heating	Electricity	NERC Power Plant Emissions Factor (Lbs of CO2 per	Carbon Dioxide Emissions Cost	Standard Error
	(Lbs of CO2)	(Lbs of CO2)	(Lbs of CO2)	(Megawatt Hrs)	Megawatt Hrs)	(\$ per Year)	(\$ per Year)
Los Angeles, CA	23,766	1,062	5,558	8.60	1,007	840	17.88
San Diego, CA	25,183	689	5,975	7.34	1,007	844	16.20
San Franciso, CA	24,777	1,675	5,765	7.62	1,007	858	17.13
San Jose, CA Sacramento, CA	24,004 25,827	2,058 458	6,055 6,636	7.85 9.50	1,007 1,007	860 913	19.96 13.93
Riverside, CA	26,761	42	6,413	9.34	1,007	916	13.65
Fresno, CA	25,587	951	7,126	10.60	1,007	953	14.46
Tucson, AZ	27,062	616	4,106	13.02	1,007	965	14.65
Las Vegas, NV	24,667	0	7,347	12.97	1,007	969	15.08
Phoenix, AZ	26,339	75	2,168	17.04	1,007	983	14.68
Albuquerque, NM	25,764	648	10,500	9.03	1,007	989	34.98
Rochester, NY	26,920	902	10,084	7.67	1,185	1,011	16.73
Buffalo, NY	26,539	1,124	10,866	7.84	1,185	1,028	14.05
Denver, CO Portland-Vancouver, OR	26,147 26,520	1,374 2,098	10,152 6,665	10.47 13.16	1,007 1,007	1,037 1,044	13.82 14.11
Providence, RI	25,648	1,273	12,213	8.00	1,185	1,044	14.11
Syracuse, NY	27,637	574	10,441	8.82	1,185	1,045	17.80
New York, NY	20,480	6,386	10,258	8.78	1,400	1,062	6.34
Albany, NY	28,618	1,054	10,462	8.78	1,185	1,087	17.24
Tacoma, WA	26,877	430	6,120	17.05	1,007	1,088	14.48
Salt Lake City, UT	26,282	3,104	10,870	10.84	1,007	1,100	13.46
Hartford, CT	27,047	1,539	12,245	8.89	1,185	1,105	17.71
Boston, MA	26,062	2,276	13,023	9.18	1,185	1,123	11.75
Fort Lauderdale, FL	25,992	1,124	354	17.97	1,427	1,142	18.14
Milwaukee, WI	28,020	1,291	8,957	9.71	1,614	1,160	12.93
Sarasota, FL	29,037	510 5,948	584	16.94	1,427	1,168	14.92 13.88
Seattle, WA Columbus, OH	25,838 29,515	3,948 278	7,425 8,602	15.45 10.43	1,007 1,614	1,177 1,187	14.15
Tampa, FL	28,885	742	707	17.48	1,427	1,189	14.79
West Palm Beach, FL	27,963	616	583	18.58	1,427	1,197	16.08
Miami, FL	25,056	4,689	697	17.86	1,427	1,203	21.58
Pittsburgh, PA	28,075	2,093	9,713	10.40	1,614	1,219	15.89
Orlando, FL	28,838	1,361	665	18.50	1,427	1,231	14.36
Greensboro, NC	31,442	216	4,169	14.58	1,472	1,232	14.74
Grand Rapids, MI	30,236	572	13,141	8.36	1,614	1,235	17.96
Raleigh, NC	30,433	495	5,446	14.56	1,472	1,243	14.54
Chicago, IL Norfolk, VA	26,088 27,606	5,221 1,078	10,113 5,587	10.17 16.47	1,614 1,472	1,243 1,258	15.53 15.40
Charlotte, SC	31,159	1,078	5,439	14.19	1,472	1,259	14.40
Cincinnati, OH	29,407	770	6,742	14.09	1,543	1,261	11.50
Minneapolis-St. Paul, MN	29,151	143	10,513	10.45	1,819	1,264	13.63
San Antonio, TX	28,810	1,929	2,638	16.49	1,555	1,269	15.26
Greenville, SC	32,568	130	4,104	15.28	1,472	1,275	19.71
Dayton-Springfield, OH	29,463	986	7,155	13.47	1,614	1,276	17.09
Akron, OH	30,454	768	9,662	11.47	1,614	1,277	15.22
St. Louis, MO-IL	29,811	1,267	7,366	14.39	1,472	1,282	12.68
Baltimore, MD	28,347	2,135	3,391	16.07	1,614	1,286	17.64
New Orleans, LA Scranton, PA	26,164 30,117	663 282	3,129 10,117	20.45 12.24	1,472 1,614	1,291 1,296	19.25 18.71
Cleveland, OH	29,298	1,733	10,117	12.27	1,614	1,309	13.28
Richmond, VA	29,521	771	3,790	18.23	1,472	1,310	14.69
Tulsa, OK	30,841	353	7,805	14.12	1,561	1,312	14.85
Detroit, MI	29,473	889	14,957	9.75	1,614	1,313	13.05
Washington, DC	27,511	4,729	5,228	15.48	1,543	1,319	18.65
Austin, TX	29,717	1,595	4,419	16.49	1,555	1,320	15.07
Indianapolis, IN	30,299	534	8,949	13.48	1,614	1,323	26.66
Kansas City, MO	30,235	644	9,042	14.01	1,561	1,328	17.23
Louisville, KY	30,231	884	6,965 8 555	15.63	1,543	1,337	18.52
Atlanta, GA Philadelphia, PA	30,192 25,426	1,121 3,993	8,555 10,831	15.20 14.16	1,472 1,614	1,338 1,357	14.01 9.66
Dallas, TX	28,155	1,723	5,253	18.53	1,555	1,357	16.31
Birmingham, AL	32,491	227	5,920	17.21	1,472	1,376	15.96
Nashville, TN	31,959	473	7,006	16.69	1,472	1,376	15.65
Houston, TX	28,216	1,447	5,148	19.30	1,555	1,394	16.54
Oklahoma City, OK	31,312	332	8,058	16.94	1,649	1,454	15.12
Memphis, TN	29,547	1,073	8,166	19.63	1,472	1,455	14.41

Notes:

<sup>(1)</sup> Data is from the Integrated Public Use Microdata Series from the 2000 Census, the 2001 National Household Transportation Survey (NHTS), the Department of Energy, the National Transit Database, the 2001 Residential Energy Consumption Survey (RECS), and the Environmental Protection Agency's Emissions & Generation Resource Integrated Database (eGRID).

(2) See text for detailed descriptions of the data calculations

Table 4: Regression Table

	(1)	(2)	(3)	(4)	(5)
	Emissions from Driving	Emissions from Public Transportation	from Home	Emissions from Electricity	Total Emissions
Log(Income)	1373	1123	-1901	3745	4340
	(1945)	(1351)	(2215)	(6452)	(7051)
Log(Population)	-2514	1193	1084	-2337	-2573
	(411)	(285)	(468)	(1362)	(1489)
Share of MSA Employment within 5 Miles of the City Center	-13079	2587	4215	-21618	-27896
	(2432)	(1690)	(2770)	(8068)	(8817)
January Mean Temperature	-71	-4	-191	-15	-280
	(17)	(12)	(19)	(56)	(62)
July Mean Temperature	107	-11	-92	612	615
	(39)	(27)	(44)	(128)	(140)
Constant	46439	-27651	25966	-30688	14066
	(20068)	(13940)	(22857)	(66569)	(72749)
Observations	66	66	66	66	66
R-squared	0.56	0.38	0.73	0.41	0.41

## Notes.

<sup>(1)</sup> Dependent variables are the total pounds of  $\mathrm{CO}_2$  emissions from the listed source.

<sup>(2)</sup> The dependent variables are the marginal emissions, that is, emissions calculated for housing built between 1980 and 2000. See Table 3.

<sup>(3)</sup> The unit of analysis is a metropolitan area.

<sup>(4)</sup> Standard errors are reported in parentheses.

<sup>(5)</sup> Data is from the Integrated Public Use Microdata Series from the 2000 Census, the 2001 National Household Transportation Survey (NHTS), the Department of Energy, the National Transit Database, the 2001 Residential Energy Consumption Survey (RECS), the Environmental Protection Agency's Emissions & Generation Resource Integrated Database (eGRID), and the National Oceanic and Atmospheric Administration.

Table 5: Suburb-City Differences in CO2 Output Emissions

MSA	Suburb-City Difference in Emissions from Driving	Suburb-City Difference in Emissions from Public Transportation	Suburb-City Difference in Emissions from Home Heating	Suburb-City Difference in Electricity	Suburb- City Difference in Carbon Dioxide Emissions	Suburb- City Standard Error (\$ per
	(Lbs of CO2)	(Lbs of CO2)	(Lbs of CO2)	(Lbs of CO2)	(\$ per Year)	Year)
New York, NY	6,150	-2,367	5,650	4,015	289.16	9.68
Nashville, TN	7,880	-649	986	3,911	260.74	25.15
Atlanta, GA	6,593	-1,242	958	5,676	257.69	23.37
Boston, MA	6,691	-1,091	4,460	1,837	255.82	17.73
Philadelphia, PA	6,884	-2,286	838	4,926	222.78	13.96
Washington, DC	5,436	-2,280	140	5,757	194.64	29.18
Hartford, CT	5,392	-2,905	3,926	1,689	174.21	25.76
San Francisco, CA	4,246	-939	2,678	2,078	173.35	25.66
Minneapolis-St. Paul, MN	5,314	-105	-225	2,960	170.77	21.11
Houston, TX	2,794	-561	676	4,726	164.13	23.33
Raleigh-Durham, NC	3,011	-182	-1,839	5,996	150.19	21.48
Memphis, TN	3,559	-423	252	3,529	148.72	21.91
Tulsa, OK	4,959	-161	-771	2,755	145.80	23.13
Milwaukee, WI	4,624	-860	140	2,466	136.96	20.18
Baltimore, MD	6,248	-1,647	-3,674	5,417	136.40	28.36
Dallas, TX	4,040	-986	-884	4,009	132.86	23.96
Providence, RI	4,427	-982	1,615	1,067	131.74	20.89
Portland, OR	2,965	-553	169	3,362	127.76	20.46
Richmond, VA	4,475	-995	-3,478	5,873	126.29	22.04
Cincinnati, OH	2,848	-383	-2,281	5,424	120.58	17.40
Syracuse, NY	2,043	-204	1,335	2,091	113.18	25.27
Cleveland, OH	4,396	-1,002	-2,113	3,864	110.60	19.62
Buffalo, NY	4,245	-813	124	1,558	109.95	20.10
Seattle, WA	2,894	-2,608	1,282	3,309	104.87	20.58
Norfolk, VA	2,997	-295	-83	2,226	104.17	22.48
Charlotte, SC	2,937	-604	-248	2,671	102.26	21.27
San Antonio, TX	3,589	-388	-911	2,331	99.34	23.54
Austin, TX	4,106	-784	-293	1,415	95.53	23.92
St. Louis, MO	4,296	-1,378	-1,377	2,742	92.06	19.39
Akron, OH	3,661	-369	-1,022	1,707	85.51	23.20
Sacramento, CA	2,185	-101	201	1,681	85.27	20.56
Phoenix, AZ	3,675	-101 -94	-1,497	1,835	84.25	21.34
Chicago, IL	5,577	-2,624	-1,497 -219	1,833	82.48	24.03
Greensboro, NC	2,199	-60	-3,340	4,220	64.91	21.06
Denver, CO		-641	150	934	63.34	20.65
	2,503		-192			
Oklahoma City, OK	1,086	-115		1,726	53.86	21.02
Fresno, CA	1,438 2,705	-92 -542	267	785	51.55 49.03	20.56
Kansas City, MO			-1,625	1,743		25.54
Rochester, NY	2,662	-554 182	-1,001	1,162	48.80	23.85
Grand Rapids, MI	1,528	-183	-1,172	1,870	43.94	25.65
New Orleans, LA	3,391	-474	-1,507	407	39.06	27.77
Riverside, CA	1,176	-8 527	685	-695 1.524	24.88	19.49
Dayton, OH	2,918	-527	-2,893	1,534	22.20	24.05
Pittsburgh, PA	5,824	-1,819	-3,744	318	12.43	23.14
Tampa, FL	2,931	-560	-873	-1,239	5.57	22.52
Tacoma, WA	3,043	-134	-365	-2,428	2.49	21.51
Los Angeles, CA	691	-229	-119	-2,455	-45.42	25.36
Detroit, MI	4,475	-1,214	-6,800	-48	-77.12	19.62

<sup>(1)</sup> Data is from the Integrated Public Use Microdata Series from the 2000 Census, the 2001 National Household Transportation Survey (NHTS), the Department of Energy, the National Transit Database, the 2001 Residential Energy Consumption Survey (RECS), and the Environmental Protection Agency's Emissions & Generation Resource Integrated Database (eGRID).

<sup>(2)</sup> See text for detailed descriptions of the data calculations

Table 6: Regression Table

	(1) City-Suburb Difference in Emissions from Driving	(2) City-Suburb Difference in Emissions from Public Transportation	(3) City-Suburb Difference in Emissions from Home Heating	(4) City-Suburb Difference in Electricity	(5) City-Suburb Difference in Carbon Dioxide Total Emissions
Log(Income)	4190	-1894	4641	9532	16468
	(2098)	(924)	(2732)	(2547)	(4557)
Log(Population)	665	-478	1070	-876	381
	(394)	(174)	(514)	(479)	(857)
Share of MSA Employment within 5 Miles of the City Center	3068	-1911	12772	-2792	11137
	(2623)	(1155)	(3415)	(3184)	(5697)
January Mean Temperature	-59	17	53	-61	-50
	(20)	(9)	(26)	(24)	(43)
July Mean Temperature	93	-7	-22	158	222
*	(38)	(17)	(50)	(47)	(83)
Constant	-57430	27215	-69873	-98891	-198979
	(21624)	(9526)	(28159)	(26251)	(46972)
Number of Observations	48	48	48	48	48
$R^2$	0.37	0.44	0.36	0.39	0.37

<sup>(1)</sup> Dependent variables are the suburb-city difference of total pounds of CO<sub>2</sub> emissions from the listed source. See Table 5.

<sup>(2)</sup> The unit of analysis is a metropolitan area.(3) Standard errors are reported in parentheses.

<sup>(4)</sup> Data is from the Integrated Public Use Microdata Series from the 2000 Census, the 2001 National Household Transportation Survey (NHTS), the Department of Energy, the National Transit Database, the 2001 Residential Energy Consumption Survey (RECS), the Environmental Protection Agency's Emissions & Generation Resource Integrated Database (eGRID), and the National Oceanic and Atmospheric Administration.