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A Bayesian Network Methodology for Infrastructure Seismic Risk Assessment and Decision Support

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A Bayesian Network Methodology for Infrastructure Seismic Risk Assessment and Decision Support

By

Michelle Terese Bensi

A dissertation submitted in partial satisfaction of the

requirements for the degree of

Doctor of Philosophy

in

Engineering – Civil and Environmental Engineering

and the Designated Emphasis in

Computational Science and Engineering

in the

Graduate Division

of the

University of California, Berkeley

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Professor Stephen A. Mahin

Professor Peter Bartlett

Fall 2010

#### A Bayesian Network Methodology

for Infrastructure Seismic Risk Assessment

and Decision Support

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by

Michelle Terese Bensi

#### Abstract

#### A Bayesian Network Methodology for Infrastructure Seismic Risk Assessment and Decision Support

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#### Michelle Terese Bensi

#### Doctor of Philosophy in Engineering - Civil and Environmental Engineering

and the Designated Emphasis in Computational Science and Engineering

University of California, Berkeley

Professor Armen Der Kiureghian, Chair

A Bayesian network methodology is developed for performing infrastructure seismic risk assessment and providing decision support with an emphasis on immediate postearthquake applications. The methodology consists of four major components: (1) a seismic demand model of ground motion intensity as a spatially distributed Gaussian random field accounting for multiple seismic sources with uncertain characteristics and including finite fault rupture and directivity effects; (2) a model of the performance of point-site and distributed components under seismic loading; (3) models of system performance as a function of component states; and (4) the extension of the Bayesian network to include decision and utility nodes to aid post-earthquake decision-making.

A Bayesian network is a probabilistic graphical model that represents a set of random variables and their probabilistic dependencies. The variables may represent demand or capacity values, or the states of components and systems. Bayesian networks are graphical and intuitive, facilitate information updating, can be used for identification of critical components within a system, and can be extended by decision and utility nodes to solve decision problems. The facility for information updating renders the Bayesian network an ideal tool for infrastructure seismic risk assessment and decision support, particularly in near-real time applications immediately following a destructive earthquake. Evidence on one or more variables (e.g. observed component capacities, demands, or damage states) can be entered into the Bayesian network and this information propagates throughout the network to provide an up-to-date probabilistic characterization of the performance of the infrastructure system under the uncertain and evolving state of information that is characteristic of the post-event period. Like most computational methods, Bavesian networks have limitations. In particular, calculations in Bayesian networks can be highly demanding of computer memory. The present study develops methodologies to minimize computational demands by optimizing network topology and, when necessary, making trade-offs between accuracy and computational efficiency.

The study begins with a brief introduction to Bayesian networks. Next, each of the aforementioned components of the methodology is described. The seismic demand model provides distributions of ground motion intensity at discrete points in the geographic domain of a spatially distributed infrastructure system. This model can be used to perform and go beyond conventional probabilistic seismic hazard assessment. In particular, the model provides a full random field characterization of the ground motion intensity, thus allowing assessment of seismic risk for spatially distributed systems. Equally important, the model enables updating of the distribution of intensity at any selected site upon observation of the intensity at other sites. The modeling of random fields via Bayesian network results in a densely connected topology that renders probabilistic inference computationally demanding and possibly intractable. To address this problem, several approaches for approximating the correlation structure of variables drawn from a random field are developed, which amount to selectively removing links and nodes in the Bayesian network. It is found that a method based on numerical optimization achieves the best trade-off of accuracy versus efficiency.

Bayesian network formulations are presented for modeling component performance as a function of seismic demand using fragility functions. The framework accounts for potential sources of correlation in component response. Models for point-site and distributed components are presented. The latter is based on an assumption that damages along a component occur according to a non-homogenous Poisson process. Five Bayesian network formulations for modeling system performance as a function of component states are developed. One approach uses a naïve topology, two formulations are based on an intuitive interpretation of system performance, and two approaches utilize minimal link and cut sets. The last two formulations are then adapted and refined with the goal of minimizing computational demands by arranging nodes in chain-like structures that reduce the size of conditional probability tables and, consequently, required computation time and memory.

The Bayesian network is extended by decision and utility nodes to create a new graphical construct known as an influence diagram. This diagram aids decision-making by specifying decision alternatives that maximize expected utility given all available evidence. The extension of the framework to include decision and utility nodes is demonstrated by application to a post-earthquake decision scenario involving inspection and shutdown decisions. A limited memory influence diagram is constructed to model this decision problem. A heuristic based on a value of information criterion is described for prioritizing component inspections following an earthquake.

Two example applications demonstrate the Bayesian network methodology for infrastructure seismic risk assessment and decision support. The second example employs a preliminary and hypothetical model of the proposed California high speed rail system.

To Dan. Thank you... for everything.

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### **Chapter 1: Introduction**

#### 1.1 Motivation

Civil infrastructure systems are the backbones of modern societies, yet they remain vulnerable to a wide-variety of natural and man-made hazards. Ensuring the reliability of infrastructure systems under normal operating conditions and resilience in the face of hazards is critical to the health, safety, and security of communities. In many countries around the world, civil infrastructure systems face increasing service demands while they deteriorate as a result of inadequate maintenance, material degradation, and increasing exposure to overloads. In addition, these systems are subject to a variety of natural and man-made hazards. If lifelines are not operational following an extreme event, such as an earthquake or hurricane, entire communities may have to be displaced. On the other hand, limited resources are available to address these deficiencies and there is need for infrastructure-specific decision support tools that will assist facility owners to optimally allocate scarce funds. Unfortunately, for many years, research has focused primarily on improving the performance and reliability of individual structures or components and relatively little effort has been dedicated to examining infrastructure systems as a whole.

Risk assessment and management of infrastructure systems differs substantially from the risk analysis traditionally performed for individual infrastructure components (Der Kiureghian 2009). These differences include:

*Hazard characterization*: Infrastructure systems are spatially distributed and thus their rate of hazard exposure is greater than that of single-site facilities. Furthermore, they are exposed to a wider-variety of failure modes. This is particularly true of infrastructure systems exposed to geographically distributed loads, such as earthquakes, which require the modeling of hazard as a spatially

distributed random field. Neglecting the correlation in demands arising from the random field nature may result in incomplete hazard characterization and an overestimation of the reliability of infrastructure systems that are generally redundant.

*System-level effects*: System-level management activities, such as the prioritization of component inspections and repairs after a hazard event, require that lifelines be analyzed using systems analysis techniques. Neglecting to properly account for the systems effect may result in inefficient allocation of resources to components that have little or no influence on overall system reliability and potential disregard of components that are crucial for the performance of the system.

*Infrastructure systems are complex and interacting*: Lifelines do not operate independently and the performance of one infrastructure system is often dependent on the operation of other systems. Failing to account for such dependency may result in overestimation of system reliability. For example, while the retrofit of a water distribution system may render it able to withstand loads imposed by a large hazard event, this investment is of little value if the electrical grid required to power the pumps is out-of-service following the event.

**Evolving nature of information**: Infrastructure systems are characterized by an environment in which the information available to managers and decision-makers is generally uncertain and continuously evolving in time. This is true under normal operating conditions when, for example, components go in and out of service due to degradation and scheduled repairs and as rehabilitation and retrofit actions change the system reliability. However, the evolving nature of information is particularly important in the aftermath of a hazard event. Immediately following an extreme event, information incrementally becomes available to decision-makers from a variety of sources such as structural health monitoring sensors, hazard reports, and human observation. A decision support system for managing an infrastructure system should be able to incorporate the flow of uncertain and evolving information in assessment of risk and analysis of optimal choices.

In many regions of the world earthquakes pose the dominant risk to infrastructure systems. However, through sound engineering and decision-making, it is possible to mitigate these risks. There are three distinct stages during which a decision-maker must act to assess and manage risks posed to an infrastructure system: normal operating conditions, immediate post-event, and longer-term post-event. Each stage is associated with acquisition of information that is both uncertain and evolving in time, and that must be incorporated into hazard and component/system performance models. The aforementioned three phases are further described below.

**Decision-making under normal operating conditions**: In the decision phase prior to the occurrence of an earthquake, it is necessary to assess earthquake risks by identifying critical components and bottlenecks in the system, and then take actions to minimize the risk exposure. Examples of activities to minimize risk exposure

include retrofit or replacement of components and modification of the system configuration. Additionally, during regular operations, actions must be taken to maximize system reliability under normal conditions. System reliability is maximized through optimization of inspection, maintenance and rehabilitation activities, e.g. by selecting types of inspection, maintenance and rehabilitation actions and their timing. Decisions that improve system reliability under normal operating conditions and resilience to extreme events should ideally be made in tandem because many failures occur when damaged/deteriorated structures are subjected to abnormal loads.

*Immediate post-event response*: In the phase immediately following an earthquake, decisions must be made to reduce the severity of post-event consequences. Often, at this stage, emphasis is placed on life safety and restoration of critical services. Post-event decisions include the dispatching of rescue, inspection, and repair crews as well as selecting to open or close facilities. These decisions are made under the competing demands to maintain operability (to prevent revenue loss) while not sacrificing safety (to avoid deaths and injury, as well as liability). This decision-phase is characterized by an environment in which information is both uncertain and quickly evolving in time. Sources of information at this phase may include ground motion intensity measurements, data received from structural health monitoring sensors, inspection results, and even observations reported by citizens and infrastructure facility employees. To aid decision-making, such information must be incorporated into hazard and component/system response models in near-real time to provide an up-to-date probabilistic characterization of the system performance.

**Longer-term post-event recovery and recalibration**: In the longer-term following an extreme event, it is necessary to take actions to minimize long-term economic consequences. Such actions may include optimal prioritization of replacement and retrofitting of damaged facilities and reconfiguration of the system. Additionally, following an event, information is gained about component/system performance as well as the hazard environment. This information should be used to update and refine the hazard model as well as the component fragilities and system characterizations that are used in decision support tools to minimize risk from future earthquakes.

#### 1.2 Goals, methodology, and scope

With the above motivation in mind, the goal of this study is to develop the framework for a probabilistic decision-support system for seismic infrastructure risk assessment and management. The study focuses primarily on developing a framework that is able to incorporate uncertain and time-evolving information from a variety of sources and to provide support to decision-makers in immediate post-event applications. However, the framework developed here can be applied to aid decision-making during each of the aforementioned phases.

The actions taken immediately following an earthquake can have a significant impact on the severity and extent of consequences as well as the efficiency and speed of recovery efforts. Despite the criticality of immediate post-event actions, decisions are often made in an ad hoc manner, under great uncertainty, and in the face of information that rapidly evolves in time. Decision makers, acting in a stressful and chaotic post-disaster environment, often only have limited and uncertain information as well as their personal experience to draw upon as resources. For near-real time applications, decision-makers require estimates of reliabilities of the system and its components conditioned on the available information, e.g. partially or precisely known magnitude and location of the earthquake, measurements of shaking intensity at recording stations, and observed performance of components and subsystems. Additionally, the output of any management tool must provide rankings of decision alternatives amidst competing objectives. Thus, post-event decision-makers are in need of a tool that (1) can be employed in near-real time following a major disaster; (2) is capable of synthesizing incoming information with existing probabilistic models; and (3) can properly account for uncertainties both in the incoming information as well as in the analytical and empirical models that are used to assess the demands on and the damage states of various components of an infrastructure system. Recent advances in technology, computer science, learning techniques, risk assessment, and hazard modeling have yielded knowledge and methodologies that can be integrated to create a unique tool to aid the post-event decision-making.

In view of the aforementioned requirements of a decision-support system, it has been determined that Bayesian networks provide the ideal tool to meet the described challenges. Characteristics that make Bayesian networks well-suited for the proposed application include:

- (1) A Bayesian network is an efficient and intuitive graphical tool for the representation and assessment of systems under uncertainty.
- (2) Bayesian networks can be used to model multiple hazards and their interdependencies resulting from common causes and characteristics.
- (3) Bayesian networks provide an efficient framework for probabilistic updating and the assessment of component/system performance in light of uncertain and evolving information, particularly for near-real time, post-event applications.
- (4) Through a max-propagation algorithm, a Bayesian network can identify critical components in a system and most probable hazard scenarios.
- (5) The graphical interface of the Bayesian network makes it an excellent tool for interaction with and use by practitioners and end-users, who may not be experts in probabilistic system analysis.
- (6) The ability to update information makes the Bayesian network an excellent tool for scenario evaluation for design purposes as well as for training/planning exercises for emergency response.
- (7) A Bayesian network can be extended to include utility and decision nodes, thus providing a decision tool for ranking alternatives based on expected utility.
- (8) Bayesian networks provide an ideal platform for linking interdisciplinary modules to provide a comprehensive decision-support framework.

Assessing the seismic risk to civil infrastructure systems subject to earthquake hazard requires: (1) characterization of earthquake sources and associated properties, e.g. fault geometries, earthquake location, earthquake magnitude; (2) an accurate and adaptable seismic demand model that provides estimates of ground motion intensities at distributed locations throughout the geographic domain of the infrastructure system as well as estimates of other attributes of the hazard, such as fault dislocation and liquefaction potential, while properly accounting for the relevant spatial correlation structures; (3) models of component performance under seismic loadings; and (4) models of the performance of the infrastructure system in terms of the component performances. Furthermore, given fully updateable models of seismic hazard and system response, it is necessary to have an explicit framework for providing decision support.

In the study, we develop a methodology for using Bayesian networks to perform seismic infrastructure risk assessment and provide decision support. We aim to lay the groundwork necessary to build a comprehensive framework by tackling issues such as the integration of existing knowledge related to seismic hazard modeling and assessment, component performance modeling, and system analysis into a Bayesian network. Furthermore, we devote considerable effort to addressing issues related to computational efficiency to improve the viability of the methodology for application to realistic infrastructure systems, which are large and complex. While development of a deliverable final product for end users is outside the scope of this study, the work presented herein will provide the foundation necessary to eventually develop applied tools.

#### **1.3 Existing Resources**

Before detailing the approach proposed in this study, we briefly describe the resources currently available to decision-makers. Available resources facilitate performance of tasks ranging from analysis of earthquake risk at the component level to portfolio risk management. The resources described below are primarily useful for risk assessment. As such, they are predominately used in pre-event applications for understanding seismic hazard and identifying potential weak points. Often, they provide guidance for considering representative earthquake scenarios to accomplish these tasks. No method described below is able to facilitate near-real time Bayesian updating, or provide decision-support while fully accounting for the time-evolving and uncertain nature of the available information and models.

#### 1.3.1 HAZUS<sup>MH</sup>

Hazards U.S. Multi-Hazard (HAZUS<sup>MH</sup>) (FEMA 2008) is a widely-used methodology and software program developed by the Federal Emergency Management Agency (FEMA) for estimating losses from multiple hazards, such as earthquakes, hurricane winds, and floods. Using Geographic Information Systems (GIS), the software provides hazard map data and allows the user to estimate the impact of hazards on buildings and individual components in an infrastructure system by mapping and displaying damage data and economic loss estimates. HAZUS<sup>MH</sup> is designed to aid decision-makers by providing loss estimates that

inform mitigation plans, emergency preparedness, and response/recovery planning. HAZUS<sup>MH</sup> involves a five step process: (1) identify hazards; (2) profile hazards; (3) inventory assets; (4) estimate losses; and (5) evaluate mitigation options. HAZUS<sup>MH</sup> is geared toward use *before* a disaster strikes so that mitigation actions can be taken prior to events occurring (FEMA 2008). In this sense, it is a "static" system, as it does not have a means of updating the model and estimates as information becomes available after an event. However, given its comprehensive nature, HAZUS<sup>MH</sup> is a valuable asset for infrastructure owners.

In spite of its comprehensive nature, HAZUS<sup>MH</sup> has several shortcomings in areas that have been a source of considerable focus in the work presented herein. HAZUS<sup>MH</sup> makes simplifying assumptions with regard to modeling seismic hazard. The effects of these simplifying assumptions are not detrimental when assessing single-site facilities, but become important when assessing the hazard on a spatially distributed system. Characterization of lifeline systems in HAZUS<sup>MH</sup> is "necessarily incomplete and oversimplified" (DHS 2003). As mentioned earlier, HAZUS<sup>MH</sup> also lacks an ability to update risk estimates based on evolving and uncertain information. We dedicate considerable effort to addressing these shortcomings in the framework described in this study. In some areas, HAZUS<sup>MH</sup> goes beyond the scope of the present study by addressing a wider range of hazards as well as providing specialized guidance for specific classes of infrastructure systems. To that end, HAZUS<sup>MH</sup> is a valuable risk assessment tool that can be viewed as a complementary system rather than a competing system to the one proposed in this report.

#### **1.3.2 SHAKEMAP/SHAKECAST**

Immediately following a significant earthquake, the United States Geological Survey (USGS) Earthquake Hazards Program, in conjunction with regional seismic networks, produce a ShakeMap. A ShakeMap is a map providing estimates of ground motion and shaking intensity in a region due to an earthquake (USGS 2008). ShakeMaps are developed for use in response and recovery efforts and information dissemination to scientists, the public, and news agencies. Scenario earthquakes can be also considered for disaster preparedness exercises and planning. ShakeMaps are determined by combining measurements of ground motion intensity collected during an earthquake with knowledge about earthquake source characteristics and local site conditions (USGS 2008). However, with regard to the current application to seismic infrastructure risk assessment, ShakeMap has several shortcomings: (1) Interpolations between points for which measurements are available require use of empirical relations; however ground motions/intensities can vary significantly over short distances. Thus, at small scale and away from observation points, results obtained from ShakeMap can be unreliable. (2) A ShakeMap is not a complete or certain descriptor of ground motion at a site and thus may not be sufficient for predicting the severity of damage at the site (USGS 2008). (3) ShakeMap only minimally addresses issues of uncertainty and neglects random field effects. (4) ShakeMap provides only maps of seismic intensities and does not provide information about component or system performance.

ShakeCast (short for ShakeMap Broadcast) is a "fully automated system for delivering specific ShakeMap products to critical users and for triggering established post-earthquake response protocols" (USGS 2007). Used primarily for emergency response, loss estimation, and public information, ShakeCast provides automatic notifications of earthquake shaking information (e.g. by cell phone, pager, email) at specific facilities based on preset thresholds (USGS 2007). Also, using pre-established damage thresholds, ShakeCast provides estimates of the likelihood of damage at locations affected by an earthquake; it also converts HAZUS structure type information to structural damage levels. (USGS 2008, 2007). ShakeCast goes beyond what ShakeMap and HAZUS<sup>MH</sup> are capable of doing with regard to near-real time applications; however, it falls short of the objectives outlined in the research discussed in this study. For example, it does not permit Bayesian updating nor does it model system performance. It is believed that integration of the methodology described in this study within an existing platform such as ShakeMap/Cast could prove extremely useful in near-real time post-earthquake applications.

#### 1.3.3 REDARS

REDARS (Risks from Earthquake DAmage to Roadway Systems) is a seismic risk analysis tool designed to assess risks posed specifically to highway systems (Werner et al. 2006). REDARS is a scenario-based analysis tool. For a given earthquake scenario (deterministic assessment) or for a simulated suite of earthquake realizations in which uncertainty is included (probabilistic assessment), REDARS estimates: (1) seismic hazards due to ground shaking, liquefaction, and fault rupture; (2) predicted damage to components in the system; (3) damage consequences including repair costs, down time, and ability of the system to carry traffic as repairs progress post-event; and (4) estimates of economic losses and travel time increases. Network analysis procedures are applied to each damage scenario to determine how closures affect system wide travel times and traffic flows. REDARS is designed for use in pre- and post-earthquake applications and to provide design guidance. Prior to the occurrence of an earthquake, REDARS estimates the effect of seismic improvement options based on the ability of the retrofit to reduce expected economic losses. It compares expected costs and benefits of available improvement options and provides these results to the decision-maker. In post-earthquake applications, REDARS can incorporate data corresponding to realized damages (e.g. field observations) and uses the information to assess the relative abilities of repair strategies/priorities as well as a options for traffic management. It however does not facilitate Bayesian updating. As with the earlier methods, the treatment of uncertainties in REDARS is less comprehensive and rigorous than in the approach developed in this study. However, REDARS is more sophisticated than the present study in its approach to modeling traffic and economic consequences.

#### 1.3.4 Additional resources

Internationally, methodologies similar to, or based on, the HAZUS<sup>MH</sup> framework are available. These include:

- RADIUS (Risk Assessment Tools for Diagnosis of Urban Areas Against Seismic Disasters): The goal of the RADIUS project is to help decision-makers and other leaders in developing countries understand earthquake risk and to use public awareness as a means for reducing it. As part of the project, nine city case studies were selected to develop damage scenarios and action plans as well as to perform a comparative study on understanding urban seismic risk around the world, but particularly in developing countries. Furthermore, the project developed tools for risk management and facilitating inter-city information exchange. Additional information is available in (GHI 2009).
- RISK-UE (An advanced approach to earthquake risk scenarios with applications to different European towns): RISK-UE is a methodology similar to HAZUS<sup>MH</sup> and RADIUS. It provides similar guidance but for specific application to European cities, which have characteristics that differentiate them from other locations. These include: the presence of historical and monumental structures, uniqueness of city layouts (e.g. centered around old historical centers), cultural impacts, and functional/social organization. RISK-UE provides a "modular methodology" that includes assessment of seismic hazard, system exposure, as well as the vulnerabilities of lifelines and modern and historical structures. Handbooks are available corresponding to each module (Mouroux and Brun 2006).

#### 1.4 Study Outline

This study develops a comprehensive Bayesian network methodology for performing seismic infrastructure risk assessment and providing decision support. Components of the Bayesian network framework include: (1) a seismic demand model of ground motion intensity as a spatially distributed Gaussian random field accounting for multiple seismic sources with unique and potentially uncertain properties as well as finite fault rupture and directivity effects; (2) models of performance of point-site and distributed components; (3) models of system performance in terms of connectivity and flow; and (4) the extension of the Bayesian network to include decision and utility nodes to aid post-earthquake decision-making. In support of the aforementioned objectives, and to fill needs not met by existing tools, the Bayesian network methodology developed for seismic infrastructure risk assessment and decision support is described in the following chapters:

**Chapter 2** provides an introduction to Bayesian networks. It begins with a brief review of Bayesian statistics and emphasizes the value of the Bayesian approach for facilitating information updating. In particular, Bayesian networks facilitate information updating in large and complex problems. The chapter gives motivation for using Bayesian networks for the proposed application. It also provides a brief introduction to constructing and performing inference in Bayesian networks. The chapter is not intended to be a replacement for more comprehensive texts. Instead, it strives to provide the background necessary to understand the Bayesian network models outlined in the remainder of the report. As a result, wherever possible, complex theoretical descriptions are replaced by intuitive and example-based explanations.

**Chapter 3** describes a Bayesian network-based model that provides distributions of ground motion intensity at discrete points in the domain of a geographically distributed infrastructure system. The focus of the chapter is primarily on hazard due to ground-shaking, but the framework is extended to other hazards with less detail and rigor. Both point-source and finite rupture model formulations are considered. The formulation presented in Chapter 3 accounts of random field and directivity effects.

**Chapter 4** provides Bayesian network models for efficiently modeling correlated random variables drawn from a Gaussian random field. Modeling random variables drawn from a random field results in a Bayesian network topology that is densely connected. For such a topology, the computational and memory demands rapidly grow with the number of points drawn from the random field. In Chapter 4, we develop approximate methods to overcome this difficulty by reducing the density of the Bayesian network model. The aim is to minimize the number of links in the Bayesian network while limiting the error in the representation of the correlation structure of the random field.

**Chapter 5** presents Bayesian network formulations for modeling the performance of components in an infrastructure system as a function of the seismic demands placed upon them. First, a generic model is presented for describing component performance while accounting for potential sources of correlation. In many instances, seismic fragility research is not sufficiently mature, nor in possession of adequate data, to facilitate construction of the generic formulation. As a result, simplified Bayesian network-based models for point-site and distributed components are developed that reflect the current state of practice. The chapter ends with a brief description of fragility functions available for modeling the performance of components in a variety of infrastructure systems.

**Chapter 6** presents several Bayesian network formulations for modeling the performance of infrastructure systems as a function of the states of the system components. The chapter begins with a brief overview of conventional methods of modeling system performance and compares them with the Bayesian network approach. Five formulations for modeling system performance by Bayesian network are described, including a naïve approach, two approaches based on an intuitive interpretation of system performance, and two approaches that utilize minimal link and cut sets. The last two formulations are then adapted with the goal of minimizing computational demands using a heuristically augmented optimization-based procedure. The procedure is developed to automate the construction of efficient Bayesian network formulations for modeling system performance. Chapter 6 ends with two example applications.

**Chapter 7** extends the Bayesian network by decision and utility nodes to create a new graphical structure known as an influence diagram. This allows the Bayesian network framework described in this study to be used to solve decision problems based on the principles of expected utility maximization. The extension of the Bayesian network to solve decision problems is illustrated by an example application involving post-earthquake component inspection and shutdown decisions.

**Chapter 8** provides two example applications. The first example is based on a simple hypothetical transportation system. The example is used to show how Bayesian networks can be used to perform and go beyond conventional probabilistic seismic hazard assessment. The second application is based on the proposed California high speed rail system. Because the system is still under design, the example is based on assumptions and idealizations. The application is used to demonstrate the value of using Bayesian networks for near-real time, post-earthquake hazard assessment and decision support.

*Chapter 9* wraps up the report providing a summary of the study and an outlook for continued research and development in this area.

### Chapter 2: Introduction to Bayesian Networks

#### 2.1 Background

Statistical inference is the process of drawing a conclusion, e.g. estimating the value of an uncertain quantity, using data or observations. More formally, it is the process of distinguishing systematic patterns in a set of observations from the random noise intrinsic to almost all data sets (Gill 2002). There are two primary schools of thought relating to statistical inference: the frequentist approach and the Bayesian approach. The frequentist approach is based on the frequency notion of probability. This view specifies that the probability of an event is the limit of its relative frequency of occurrence as the number of event trials becomes large, theoretically infinite. It is thus an empirical approach to computing probabilities. Unfortunately, in many research disciplines (including earthquake engineering), a given data-generating process will rarely produce a sufficiently large sample. For this reason, Bayesian methods, which are generally more robust in situations of data scarcity (Koop 2003), have grown in popularity since the mid-20<sup>th</sup> century. Under the frequentist philosophy, unknown parameters are treated as fixed deterministic values to be estimated rather than as random variables. The Bayesian approach instead treats unknown parameters as random variables and argues that uncertainty can be expressed using the rules of probability, i.e., as a probability distribution that reflects the relative likelihoods of outcomes. Specifically, Bayesian statistics is concerned with the determination of the conditional probability of an unknown quantity given observations. It uses observations to contradict or validate assumptions made *prior* to statistical analysis, e.g. based on subjective expert opinion, engineering judgment, or physical models. Bayesian statistical models can be updated continuously as new observations become available and thus represent the probabilistic modeling of an accumulation of knowledge

(Congdon 2003). As the amount of available data increases, conclusions based on a Bayesian analysis will approach the results obtained using a frequentist approach because the influence of the prior assumptions diminishes.

The use of the term *Bayesian* to describe this approach comes from the well known Bayes' rule (attributed to the 18<sup>th</sup> century mathematician and philosopher Thomas Bayes):

$$\Pr(A|B) = \frac{\Pr(AB)}{\Pr(B)} = \frac{\Pr(B|A)}{\Pr(B)} \Pr(A)$$
(2.1)

where Pr(AB) is the joint probability of events *A* and *B*, Pr(A) is the marginal probability of event *A*, Pr(A|B) is the conditional probability of event *A* given that event *B* has occurred, and Pr(B) is the marginal probability of event *B*. The quantity Pr(B|A) is known as the *likelihood* of the observed event *B*. Note that the probability of event *A* appears on both sides of (2.1). Bayes' rule describes how the probability of event *A* changes given information gained about the occurrence of event *B*. Typically, Pr(A) is known as a *prior probability* and represents a belief about the probability of *A prior* to knowing anything about *B*. Analogously, Pr(A|B) is called the *posterior probability* of *A* because it represents a belief about the probability of *B*. The dominator Pr(B) is a normalizing factor obtained using the Theorem of Total Probability (Box and Tiao 1992)

$$Pr(B) = Pr(BA) + Pr(B\overline{A}) = Pr(B|A)Pr(A) + Pr(B|\overline{A})Pr(\overline{A})$$
(2.2)

where  $\overline{A}$  is the complement *A*. It is seen that Bayes' rule facilitates the *updating* of beliefs about event *A* given evidence about event *B*.

More generally, let *X* be a random variable and  $\mathbf{E}_{\mathbf{x}} = \{E_{x,1}, \dots, E_{x,n}\}$  be a set of observations that provides information about *X*, e.g.  $X_i = x_i, X_i \le x_i$ , or  $X_i > x_i$ . Let  $\Theta$  be an unknown parameter or set of parameters describing the probability distribution of *X*. Equation (2.1) is adapted to yield an expression for the updated conditional probability distribution of the unknown parameter given the vector of observations:

$$f\left(\theta \middle| \bigcap E_{x,i}\right) = \frac{\Pr(\bigcap E_{x,i} \middle| \theta)}{\Pr(\bigcap E_{x,i})} f(\theta)$$
(2.3)

In (2.3),  $f(\theta)$  is the prior probability density function (PDF) representing the analyst's belief about the relative likelihoods of the possible outcomes  $\theta$  of the unknown parameter  $\Theta$  in the absence of observations  $\mathbf{E}_{\mathbf{x}}$ . The prior distribution does not depend on the observations.  $f(\theta | \bigcap E_{x,i})$  is referred to as the *posterior distribution* of  $\theta$  and  $\Pr(\bigcap E_{x,i} | \theta)$  is once again referred to as the likelihood function. In this case, the likelihood gives the probability of observing the data given that  $\theta$  is the true value of the unknown parameter  $\Theta$ . Assuming the observed events  $E_{x,i}$  are statistically independent, the likelihood can be expressed as

$$\Pr\left(\bigcap E_{x,i} \left| \theta\right.\right) = \prod_{i} \left[\Pr(E_{x,i} \left| \theta\right.)\right]$$
(2.4)

Thus, the likelihood is based on observations while the prior comes from other, typically subjective, sources. The denominator in (2.3) is obtained via normalization using the Theorem of Total Probability:

$$\Pr\left(\bigcap E_{x,i}\right) = \int_{\theta} \Pr\left(\bigcap E_{x,i} \left|\theta\right.\right) f(\theta) d\theta$$
(2.5)

(Box and Tiao 1971). The normalizing factor in (2.5) is often difficult to compute. However, the availability of conjugate priors (Fink 1997) facilitates convenient solutions for certain known distributions. Markov Chain Monte Carlo methods (Gamerman and Lopes 2006) have also improved the tractability of the Bayesian updating problem.

Bayesian methods have become increasingly popular in a wide variety of applications requiring efficient and practical solutions. They have also given rise to a class of graphical models known as Bayesian networks, which facilitate efficient Bayesian updating in large and complex problems. This chapter provides motivation for using Bayesian networks to model and manage seismic hazard for infrastructure systems and gives a brief introduction to constructing and performing inference in Bayesian networks.

#### 2.1.1 Motivation for using Bayesian networks

A Bayesian network (BN) is a probabilistic graphical model that represents a set of random variables and their probabilistic dependencies. Within the context of this study, the variables may represent demand or capacity values, other variables that are used to compute the demands and capacities, or the states of components and systems. As will be demonstrated throughout this study, BNs are graphical, facilitate information updating, can be used for identification of critical components within a system, and can be extended by decision and utility nodes to solve decision problems. In particular, the facility for information updating renders the BN an ideal tool for infrastructure risk assessment and decision support. Evidence on one or more variables, e.g. observed component capacities, demands, or component/system states, is entered into the BN and this information propagates throughout the network to provide an up-to-date probabilistic characterization of the performance of the infrastructure system in light of the new observations, following the principles of Bayes' rule. This dynamic system characterization is particularly critical for post-event, near-real time applications when the available information is uncertain and rapidly evolving in time.

The use of graphical models for encoding dependence relations has an intuitive appeal, as they provide a transparent language for the communication of relations among variables (Kjaerulff and Madsen 2008) and the flow of information between them. Thus, they facilitate communication of model structure among interested parties, e.g. modelers and decision-makers. Because the BN is a graphical representation of random variables and dependencies, it permits probability and graph theories to be linked and is thus an intuitive means for dealing with uncertainties in complex systems such as civil infrastructures (Friis-Hansen 2004).

#### 2.2 Brief introduction to Bayesian networks

This section provides a brief introduction to BNs with the goal of providing knowledge sufficient to understand models and methods presented in the subsequent chapters of this study. Comprehensive coverage of BNs is available in textbooks (Jensen and Nielson 2007; Kjaerulff and Madsen 2008).

#### 2.2.1 Bayesian network terminology

A BN is a directed acyclic graphical model consisting of a set of *nodes* (circles) representing random variables and a set of *directed links* (arrows) representing probabilistic dependencies. Links often represent causal relationships . Consider the simple BN shown in Figure 2.1 which models five random variables  $\mathbf{X} = \{X_1, \dots, X_5\}$  and their probabilistic relationships. For example, the random variable  $X_3$  is probabilistically dependent on variables  $X_1$  and  $X_2$ , as represented by arrows going from nodes  $X_1$  and  $X_2$  to node  $X_3$ . In the BN terminology,  $X_3$  is a *child* of  $X_1$  and  $X_2$ , while  $X_1$  and  $X_2$  are the *parents* of  $X_3$ . Likewise,  $X_4$  is defined conditionally on its parent node  $X_1$  and  $X_5$  is defined conditionally on  $X_4$ . For discrete  $X_i$ , each node is associated with a set of mutually exclusive and collectively exhaustive states. To be able to utilize exact inference algorithms, it is generally necessary to discretize all continuous random variables in the BN (with the exception of linear functions of continuous Gaussian nodes without discrete children). BNs with continuous nodes are briefly addressed in this section; additional details can be found in Lauritzen (1992), Lauritzen & Jensen (2001), Shenov (2006), Madsen (2008) and Langseth et al. (2009). For discrete nodes, a *conditional probability table* (CPT) is attached to each node that provides the conditional probability mass function (PMF) of the random variable represented by the node, given each of the mutually exclusive combinations of the states of its parents. For nodes without parents (e.g.  $X_1$  and  $X_2$  in Figure 2.1), known as root nodes, a marginal probability table is assigned.

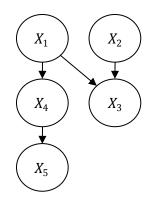


Figure 2.1: A simple BN

#### 2.2.2 Constructing Bayesian networks

The construction of a BN generally consists of two steps: (1) definition of the graphical model representing the probabilistic dependence structure of the problem, i.e. its *d*-separation properties (see below); and (2) the construction of the CPTs which together define the joint distribution over all random variables in the BN (Friis-Hansen 2004).

When constructing a BN, care must be taken to ensure that models are not misleading, unverifiable, unnecessarily complex, or computationally intractable. Accurate modeling via BN requires thorough understanding of the problem and the ability to identify the primary elements that influence it. Modeling complex problems and systems via BN may require trade-offs between transparency (verifiability), computational complexity, and detail of modeling (Friis-Hansen 2004). In many applications, statistical data is available to develop robust models to relate random variables in a BN. In civil engineering, including earthquake engineering and infrastructure system analysis, data is often scarce or unreliable. Thus, dependence relations between parents and children and the probabilities of root nodes may be based instead on theoretical models, empirical relations, expert opinion, engineering judgment, or any mixture of these.

When constructing BNs, it is not required that links represent causal relations, though typically such a modeling approach is preferred. Rather, the conditional relations must be specified such that the *d-separation properties* of the model correspond to those of the "real world" problem being modeled. Conditional (in)dependencies should not be included in the model if they do not hold in reality (Jensen and Nielson 2007). Because, the concept of d-separation is important when working with BNs, a brief qualitative description is provided below.

Consider two distinct sets of variables *X* and *Y* in a BN. These two variable sets are considered *d-separated* by a third set of variables *Z* if *X* and *Y* are independent given the variables in *Z*, i.e. if p(x, y|z) = p(x|z)p(y|z). D-separation corresponds to a blockage of the flow of information between *X* and *Y* (Bayes Nets 2007). Figure 2.2 shows three types

of connections that occur in BNs: (1) *serial* (head-to-tail); (2) *diverging* (tail-to-tail); and (3) *converging* (head-to-head). For each type of connection, the path from *X* to *Y* is or is not blocked (d-separated) based on whether or not *Z* is *instantiated*, i.e. its value has been observed. Instantiated nodes are shaded in Figure 2.2. The path between *X* and *Y* is blocked for serial and diverging connections when *Z* is instantiated. That is, information about *X* only updates a belief about *Y* when *Z* is unobserved, and vice versa. A converging connection is blocked when *Z* is not instantiated, i.e. information about *X* provides information about *Y* only when *Z* is observed. A more extensive discussion of d-separation can be found in Pearl (2000) and Jensen & Nielson (2007).

d-connected (not blocked)

<u>d-separated (blocked)</u>

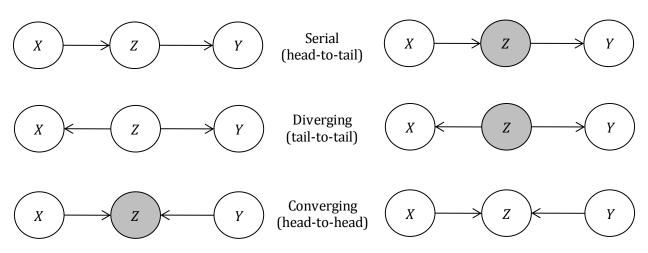


Figure 2.2: Graphical representation of d-separation properties

The state space of the random variables in a BN (i.e. the set of all possible combinations of variable outcomes) grows exponentially as the number of variables in the BN increases. However, through the use of conditional relations, the joint distribution is factored into the product of local conditional PMFs (CPTs), which are simpler to specify and facilitate more efficient calculations. The joint PMF of all random variables **X**, in the BN is constructed as the product of the conditional PMFs

$$p(\mathbf{x}) = \prod_{i=1}^{n} p(x_i | \operatorname{Pa}(x_i))$$
(2.6)

where  $Pa(x_i)$  is the set of parents of node  $X_i$ ,  $p(x_i|Pa(x_i))$  is the CPT of  $X_i$ , and n is the number of random variables (nodes) in the BN. Thus, for the BN in Figure 2.1, the joint PMF is

$$p(x_1, x_2, x_3, x_4, x_5) = p(x_5 | x_4) p(x_4 | x_1) p(x_3 | x_1, x_2) p(x_1) p(x_2)$$
(2.7)

->

Having the ability to model a problem using conditional distributions is convenient in civil engineering applications, where often only conditional relationships are available. For example, physical models or empirical relations establish the probability of a load exceeding a specific value *given* an event of a certain size. Similarly, fragility curves specify the probability of damage of a structure *given* a particular demand value.

#### 2.2.2.1 Generation of CPTs required by the BN using Monte Carlo simulation

In this study, CPTs of nodes are typically generated using Monte Carlo simulation. To demonstrate the generation of CPTs, consider a node *Y* with three parent nodes,  $X_1$ ,  $X_2$  and  $X_3$ , as shown in Figure 2.3. Let node  $X_1$  be a discrete node with  $m_1$  categorical states and an arbitrary number of parents. Node  $X_2$  is an interval node without parents. The term *interval node* is used here to describe a node that results from discretization of a continuous random variable. Node  $X_2$  has  $m_2$  states, which correspond to mutually exclusive collectively exhaustive intervals over the entire range on which the continuous random variable is defined.  $X_3$  is also an interval node but with  $m_3$  states and an arbitrary number of parents.

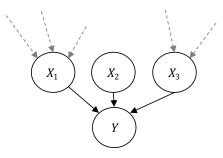


Figure 2.3: Example used to demonstrate the construction of CPTs by Monte Carlo simulation

If *Y* is defined as a real-valued function of  $X_1$ ,  $X_2$  and  $X_3$ , then the discrete interval states of *Y* must be defined over a range which contains all admissible values of *Y* for all combinations of the states of its parent nodes. Once the admissible range of *Y* has been determined, it must then be discretized into states representing mutually exclusive, collectively exhaustive intervals within this range. In many instances, it will be found that certain interval states within the range of admissible values of *Y* will be associated with a large amount of the posterior probability mass of *Y*. In such instances, it is preferable to use a more refined discretization over the portions of the range associated with high probability mass. Note that the posterior PMF of *Y* will differ based on the evidence case considered. Because the posterior distribution of *Y* is not known a priori, it may be necessary to modify the discretization of *Y* after constructing the initial BN, by considering representative evidence cases and studying the resulting posterior distributions of *Y*. Additional information relating to dynamic discretization is contained in other references, e.g. (Neil et al. 2007).

If *Y* is defined using a classification function on  $X_1, X_2$ , and  $X_3$ , then the states of *Y* are defined according to the classes associated with the classification function. An example of a classification function is a component damage function (e.g. a fragility model), which specifies a damage state given values of  $X_1, X_2$ , and  $X_3$  that represent, for example, the demands placed on the component. Note that, it may be necessary to consolidate multiple classes, if computational demands are of concern.

To demonstrate the generation of the CPT of *Y*, let *Y* be a deterministic function of random variables  $X_2$  and  $X_3$  with the form of the function defined by the categorical states of  $X_1$ , i.e.  $Y = f(X_1, X_2, X_3) = g_i(X_2, X_3)$  when  $X_1 = x_1^i$  (where  $x_1^i$  indicates the  $i^{th}$  categorical state of  $X_1$ ). Note that the range of Y must include all admissible values of Y when considering all possible functional forms associated with the categories of  $X_1$ . The conditional distribution of *Y* must be defined for each combination of the  $m_1 * m_2 * m_3$  states of its parents. We use the convention that a combination of parents states corresponds to a column of the CPT associated with Y. Let  $x_i^j$  indicate the  $j^{\text{th}}$  state of  $X_i$ , i = 1,2,3, and for i = 2,3, let that state define the interval  $[x_{i,L}^j, x_{i,U}^j]$ . Consider the construction of the CPT of Y for a single combination of the states  $X_1 = x_1^i$ ,  $X_2 = x_2^j$  (which corresponds to an outcome of  $X_2$  in the interval  $[x_{2,L}^j, x_{2,U}^j]$  of the domain of the continuous random variable), and  $X_3 = x_3^k$  (which corresponds to an outcome of  $X_3$  in the interval  $[x_{3,L}^k, x_{3,U}^k]$ ). To obtain the distribution of Y for this combination,  $N_{sim}$  draws are taken from distributions with PDFs defined  $f_{X_2}^j(x_2)$ and  $f_{X_3}^k(x_3)$  within each of the respective intervals  $[x_{2,L}^j, x_{2,U}^j]$  and  $[x_{3,L}^k, x_{3,U}^k]$ . For each draw, Y is computed using the functional form associated with  $x_1^i$ . Then, a normalized histogram is computed to obtain the required column of the CPT, with bins defined corresponding to the discrete interval states of Y specified based on its admissible range. Because  $X_2$  is a root node, the form of its distribution is known. Therefore, the  $N_{sim}$  draws from within a particular interval for this variable can be taken from a normalized distribution with the correct shape. For example, suppose  $X_2$  is distributed according to the PDF  $f_{X_2}(x_2)$ . The  $N_{sim}$  draws for state  $x_2^j$  are taken from the renormalized distribution

$$f_{X_{2}}^{j}(x_{2}) = \frac{f_{X_{2}}(x_{2})}{\int_{x_{2,L}^{j}}^{x_{2,U}^{j}} f_{X_{2}}(z)dz} \qquad x_{2,L}^{j} < x_{2} \le x_{2,U}^{j}$$
(2.8)

On the other hand, because  $X_3$  has arbitrary parents, the functional form of its distribution is not known. As a result, it is suggested to sample uniformly from the interval  $[x_{3,L}^k, x_{3,U}^k]$ . An exception to this rule is when an interval is located in the tail of the distribution of  $X_3$ and extends to  $-\infty$  or  $+\infty$ . In that case, an exponential distribution over the infinite interval may be used to better capture the tail behavior. The reader is referred to Straub (2009) for additional details.

#### 2.2.3 Inference in Bayesian Networks

There are three principle tasks that can be performed using BNs: (1) probabilistic inference of unobserved values given evidence; (2) parameter learning; and (3) structure learning. Probabilistic inference in BNs begins with a complete model of random variables and their probabilistic dependencies, typically defined based on expert knowledge of causal relations, engineering judgment, physical models, and/or predefined empirical relationships. Probabilistic inference may be viewed conceptually as efficient application of Bayes' rule on a large scale when a problem contains complex dependencies. That is, the distributions of a subset of variables in the BN are updated given observations on another subset of variables in the model. The use of the word *inference* comes from the notion that "the probability of a cause can be inferred when its effect has been observed" (Kjaerulff and Madsen 2008).

Parameter and structure learning are data-driven processes. In parameter learning, conditional distributions in the BNs have unknown parameters that are estimated from data so as to maximize their likelihood, e.g. using the expectation-maximization algorithm (Moon 1996). In structure learning, the topological form of the BN is not known a priori, as is the case when BNs are used for probabilistic inference or parameter learning. In structure learning, algorithms are used to determine the topology of the BN, i.e. where and in which direction arrows exist between nodes (Heckerman 2008).

In the applications considered in this report, BNs are used for probabilistic inference with probabilistic relationships between variables determined using physical and empirical models as well as engineering judgment. Observations are then used to update these prior assumptions.

Probabilistic inference in BNs takes two forms: forward (predictive) analysis and backward (diagnostic) analysis. Forward analysis calculates the probability distribution of any node in the BN based on the assumed prior marginal PMFs of the root nodes and the conditional PMFs of all other nodes. Backward analysis involves the computation of the posterior probability distribution of any node given observations on one or more nodes in the BN (i.e. instantiation of any subset of the variables to one or more of their admissible values) (Bobbio et al. 2001). While many analysis techniques are able to perform forward analysis, the true power in using BNs comes from the ease with which they facilitate information updating, i.e. backward analysis.

As described qualitatively above, BNs efficiently compute the conditional distribution of any subset  $\mathbf{X}'$  of the variables in the BN given evidence about the states of any other subset  $\mathbf{X}^e$  of the variables. The ease with which BNs facilitate the calculation of the conditional distribution  $p(\mathbf{X}'|\mathbf{X}^e = \mathbf{x}^e)$  is the main feature that makes the BN framework ideally suited for the proposed application to near-real time seismic risk assessment and decision support. For example, suppose observations have been made on nodes  $X_3$  and  $X_5$  in Figure 2.1 and that the conditional distribution  $p(x_2|X_3 = x_3, X_5 = x_5)$  is of interest. This posterior

distribution can be computed by first marginalizing the joint distribution in (2.7) to obtain the joint distributions over subsets of the variables:

$$p(x_2, x_3, x_5) = \sum_{x_1, x_4} p(x_1, \dots, x_5)$$

$$p(x_3, x_5) = \sum_{x_1, x_2, x_4} p(x_1, \dots, x_5)$$
(2.9)

The desired conditional distribution is then obtained as

$$p(x_2|X_3 = x_3, X_5 = x_5) = \frac{p(x_2, x_3, x_5)}{p(x_3, x_5)}$$
(2.10)

While it is possible to obtain updated distributions as shown above, this is not a computationally efficient approach to performing probabilistic inference in non-trivial BNs. Instead, *inference engines/algorithms* are available that efficiently perform calculations in BNs. A variety of algorithms is available, including exact and approximate methods. In general, exact inference in BNs is an NP-hard (nondeterministic polynomial-time hard) task (Cooper 1990). Provision of a rigorous explanation of NP-hard is outside the scope of this study. Instead, we limit ourselves to emphasizing that classifying a problem as NP-hard implies it does not have a structure which will, in general, lead to a solution in a reasonable (i.e. non-exponential) amount of time. Heuristics often make NP-hard problems tractable by looking for patterns or structure. When such patterns are found, computational complexity can be reduced to something below exponential growth. Tools for solving NPhard problems often include multiple heuristics because it is uncommon that a single heuristic will work for all classes of problems. Furthermore, heuristics will often slow down the search for a solution (Wilson 2010). Many BN algorithms (e.g. the default Hugin (Andersen et al. 1989) algorithm) employ heuristics to find a structure in the BN and aid the discovery of a (near) optimal elimination ordering. The reader is referred to Wilson (2010) for an intuitive explanation of the concept of NP-hard.

Exact inference in a complex and densely-connected BN with many nodes may require an intractably large amount of computer memory and/or computation time. In response to the computational demands of exact inference, approximate algorithms have been developed. These methods include: approximations to exact algorithms, e.g. bounded conditioning (Horvitz et al. 1989) and mini-bucket elimination (Dechter and Rish 2003); variational approaches (Jordan et al. 1999); and sampling-based procedures such as likelihood weighting (Fung and Chang 1990; Shachter and Peot 1990), Markov-Chain Monte Carlo methods (MacKay 1999) and importance-sampling (Yuan and Druzdzel 2006). Variational approaches provide bounds on probabilities of interest using approximation procedures that take advantage of known phenomena to simplify graphical models. While guidance is available for applying these approaches to some generic classes of problems, for general BNs with unique and complex dependencies, derivations of variational approximations may be difficult (Jordan et al. 1999). More common approximate algorithms are based on

stochastic sampling. In theory, sampling-based procedures will converge to the true solution given a sufficiently large sample. In general, sampling methods offer good performance for predictive/forward analysis. However, using many sampling-based algorithms, diagnostic/backward analysis is often associated with poor convergence, particularly when evidence is entered into the BN which is a priori unlikely. Recent advances in the use of importance sampling for inference in BNs has increased the accuracy of sampling algorithms for both forward and backward analysis without significant degradation in convergence times. Their improved performance has been demonstrated through examples in Cheng & Druzdzel (2000) and Yuan & Druzdzel (2006).

Due to the shortcomings of many approximate methods, particularly in lieu of unlikely evidence scenarios, we focus herein on exact inference algorithms. Exact inference algorithms are a predominant means for performing inference in multi-purpose BN software (Hugin Expert A/S 2008), though sampling algorithms have also been implemented in some packages (DSL 2007). References are available that describe and compare BN software packages (Murphy 2007, 2010). Though we focus on exact algorithms here, it is anticipated that sampling algorithms will become more robust and more commonly used as an inference tool in the future. The work contained in this study, though focused primarily on exact inference, is also useful when dealing with approximate inference procedures, as some of the same problems are encountered. For example, nodes with many parents result in large CPTs that must be stored in the computer memory. This poses a problems to exact or many approximate inference algorithms. This problem, which becomes especially acute when working with large infrastructure systems, is one of the issues addressed in this study.

#### 2.2.3.1 Overview of the elimination algorithm

One of the most basic exact inference algorithms is the *elimination algorithm* (also known as the *variable* or *bucket elimination algorithm*). While not used heavily in practice, it forms a basis for other, more efficient, algorithms and is thus described here. The elimination algorithm is used to determine the distribution of a subset of random variables in the BN by incrementally eliminating nodes from the BN that do not belong to the subset of interest. The elimination of a node from the BN has both a mathematical and a graphical interpretation.

Mathematically, elimination of a node corresponds to summing the joint distribution of the random variables over all states of the node to be eliminated. Consider the BN in Figure 2.1 with the joint distribution in (2.7) and suppose that the joint distribution of  $X_3$  and  $X_4$  is of interest. The nodes  $X_1, X_2$ , and  $X_5$  must be eliminated from the BN to obtain this quantity. Thus, the joint distribution must be summed over all states of  $X_1, X_2$  and  $X_5$ . The elimination of  $X_1$  results in the joint distribution of the remaining variables  $X_2, ..., X_5$ :

$$p(x_2, x_3, x_4, x_5) = \sum_{x_1} p(x_5 | x_4) p(x_4 | x_1) p(x_3 | x_1, x_2) p(x_1) p(x_2)$$
(2.11)

$$= p(x_5|x_4)p(x_2)\sum_{x_1} p(x_4|x_1)p(x_3|x_1,x_2)p(x_1)$$
$$= p(x_5|x_4)p(x_2)\phi(x_4,x_3,x_2)$$

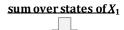
Note that the summation operator in the second line of (2.11) has been moved as far to the right as possible. This means that the sum need only be performed over the product of CPTs which include the variable  $X_1$ . The result of the sum is a table or *potential* over the remaining variables:  $\phi(x_4, x_3, x_2) = p(x_3, x_4 | x_2)$ . Figure 2.4 shows an example of the mechanics of computing the product of CPTs and performing the summation required to determine the potential  $\phi(x_4, x_3, x_2)$  in (2.11). For the sake of simplicity, each variable is assigned only two states.

conditional probability tables

$p(x_4 x_1)$				p(x	( <sub>1</sub> )	$p(x_3 x_1,x_2)$				
<i>x</i> <sub>1,1</sub> <i>x</i> <sub>1,2</sub>			<i>x</i> <sub>1,1</sub>	0.35	x <sub>2,1</sub>		X	x <sub>2,2</sub>		
$x_{4,1}$	0.2	0.3		x <sub>1,2</sub>	0.65		<i>x</i> <sub>1,1</sub>	x <sub>1,2</sub>	<i>x</i> <sub>1,1</sub>	x <sub>1,2</sub>
<i>x</i> <sub>4,2</sub>	0.8	0.7				x <sub>3,1</sub>	0.4	0.1	0.25	0.45
			-			X <sub>3.2</sub>	0.6	0.9	0.75	0.55

#### compute product of CPTs containing X<sub>1</sub>

	$\checkmark$										
	$p(x_4 x_1)^*p(x_1)^*p(x_3 x_1,x_2)$										
I			x	2,1		X 2,2					
		x <sub>3,1</sub>		x	x <sub>3,2</sub>		x <sub>3,1</sub>		3,2		
		<i>x</i> <sub>4,1</sub>	<i>x</i> <sub>4,2</sub>	<i>x</i> <sub>4,1</sub>	<i>X</i> <sub>4,2</sub>	<i>X</i> <sub>4,1</sub>	X 4,2	<i>X</i> <sub>4,1</sub>	<i>X</i> 4,2		
	<i>x</i> <sub>1,1</sub>	0.028	0.112	0.042	0.168	0.018	0.070	0.053	0.210		
	<i>x</i> <sub>1,2</sub>	0.020	0.046	0.176	0.410	0.088	0.205	0.107	0.250		



Ŷ									
$\phi(x_2, x_3, x_4)$									
x <sub>2,1</sub> x <sub>2,2</sub>									
		2,1		<u> </u>					
X	3,1	X <sub>3,2</sub>		X <sub>3,1</sub>		X 3,2			
<i>x</i> <sub>4,1</sub>	<i>x</i> <sub>4,2</sub>	<i>x</i> <sub>4,1</sub>	<i>x</i> <sub>4,2</sub>	<i>x</i> <sub>4,1</sub>	X 4,2	<i>X</i> <sub>4,1</sub>	<i>X</i> <sub>4,2</sub>		
0.048	0.158	0.218	0.578	0.105	0.275	0.160	0.460		

<b>Figure 2.4: Mechanics</b>	of computing $\phi(x_2, x_3, x_4)$
------------------------------	------------------------------------

Next, consider the graphical interpretation of the elimination algorithm. The first step in graphical elimination is to create a *moral graph* through a process known as *moralization*. In the moral graph, nodes that share a common child (e.g. nodes  $X_1$  and  $X_2$  in Figure 2.1) are connected by adding an undirected link, in a process referred to as *marrying parents*.

Then, all directed links in the graph are turned into undirected links. The moral graph corresponding to the BN in Figure 2.1 is shown in Figure 2.5a. The link introduced due to moralization is shown by a dotted line.

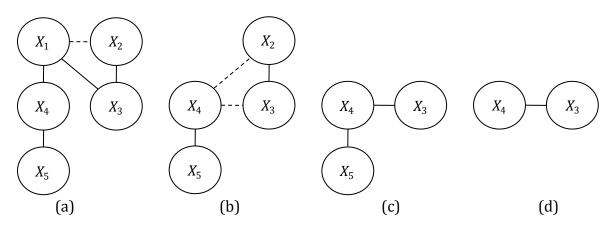


Figure 2.5: Graphs obtained during elimination algorithm: (a) moral graph corresponding to BN in Figure 2.1; graphs after elimination of (b) node X<sub>1</sub>; (c) node X<sub>2</sub>; and (d) node X<sub>5</sub>

Next, node  $X_1$  is eliminated from the BN as shown in Figure 2.5b. The graphical elimination of a node requires that all neighbors of the eliminated node be connected by undirected links (if they are not already connected). In Figure 2.5a, node  $X_1$  has nodes  $X_2$ ,  $X_3$ , and  $X_4$  as neighbors. When  $X_1$  is eliminated, undirected links are introduced between nodes  $X_2$  and  $X_4$  as well as  $X_3$  and  $X_4$  as shown by dotted lines in Figure 2.5b. This connectivity is reflected mathematically in the potential in (2.11) which contains variables  $X_2$ ,  $X_3$ , and  $X_4$ , though they did not exist together in a single CPT in the original joint distribution in (2.7).

The above process is repeated for the elimination of  $X_2$  and  $X_5$  resulting in the graphs in Figure 2.5c and Figure 2.5d, respectively. The corresponding mathematical operations are given below.

$$p(x_{2}, x_{3}, x_{4}, x_{5}) = p(x_{5}|x_{4})p(x_{2})\phi(x_{4}, x_{3}, x_{2})$$

$$p(x_{3}, x_{4}, x_{5}) = p(x_{5}|x_{4})\sum_{x_{2}} p(x_{2})\phi(x_{4}, x_{3}, x_{2})$$

$$= p(x_{5}|x_{4})\phi(x_{3}, x_{4})$$

$$p(x_{3}, x_{4}) = \phi(x_{3}, x_{4})\sum_{x_{5}} p(x_{5}|x_{4})$$

$$= \phi(x_{3}, x_{4})(\phi(x_{4}) = 1)$$
(2.12)

Note that in performing the elimination algorithm above, the product over all CPTs defining the joint distribution (as in (2.7)) was not required. By requiring local products only over CPTs containing the variable to be eliminated, memory demands are reduced. The order of computational complexity of the elimination algorithm is a function of the dimension of the largest potential that must be computed at any step and as reflected, in part, by the number of links added at any step. The larger the dimensions of the CPTs/potentials that must be multiplied, the larger the memory demand. Thus, to reduce computational demand it is necessary to keep the dimension of CPTs as small as possible. Small CPTs are achieved by ensuring that nodes have as few parents as possible and/or by reducing the number of states associated with each node. More generally, the graphical structure of the BN can also result in large potentials being created during the elimination process.

The order in which the variables are eliminated has a significant impact on the size of the CPTs/potentials that must be multiplied at each stage of the elimination algorithm. In the above example, the largest memory demand comes from the product of three CPTs (involving four nodes) when eliminating  $X_1: p(x_4|x_1)p(x_3|x_1,x_2)p(x_1)$ .

If, instead,  $X_5$  is eliminated first, followed by  $X_2$  and then  $X_1$ , the number of variables in the products involved is no more than three at any step, as shown below:

$$p(x_{1}, x_{2}, x_{3}, x_{4}) = \sum_{x_{5}} p(x_{5}|x_{4})p(x_{4}|x_{1})p(x_{3}|x_{1}, x_{2})p(x_{1})p(x_{2})$$

$$= p(x_{4}|x_{1})p(x_{3}|x_{1}, x_{2})p(x_{2})p(x_{1})\sum_{x_{5}} p(x_{5}|x_{4})$$

$$= p(x_{4}|x_{1})p(x_{3}|x_{1}, x_{2})p(x_{2})p(x_{1})(1)$$

$$p(x_{1}, x_{3}, x_{4}) = p(x_{4}|x_{1})p(x_{1})\sum_{x_{2}} p(x_{3}|x_{1}, x_{2})p(x_{2})$$

$$= p(x_{4}|x_{1})p(x_{1})\phi(x_{3}, x_{1})$$

$$p(x_{3}, x_{4}) = \sum_{x_{1}} p(x_{4}|x_{1})p(x_{1})\phi(x_{3}, x_{1})$$

$$= \phi(x_{3}, x_{4})$$
(2.13)

The graphical interpretation of the calculations in (2.13) is shown in Figure 2.6. Only two additional links are introduced, one of which is introduced in the moralization step. Note that links introduced during the moralization step are governed by the structure of the BN and not by the elimination order. This demonstrates that the order in which variables are eliminated has a significant impact on the computational demand. However, for complex BNs, determination of the optimal elimination order is NP-hard. For this reason, in practice,

heuristics and sampling based procedures are used to determine the best elimination order. The reader is referred to Dechter (1996) for additional details on the elimination algorithm.

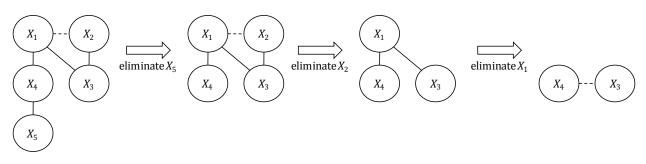


Figure 2.6: Illustration of the elimination algorithm for the BN in Figure 2.1 using the elimination order  $X_5, X_2, X_1, X_3, X_4$ 

The elimination algorithm is "query sensitive," i.e. the entire algorithm must be re-run for each quantity of interest. This is because the node(s) corresponding to this quantity must be the only nodes not eliminated and therefore must be the last nodes in the elimination order. In the above example, if one is next interested in the joint distribution of  $X_1$  and  $X_5$  given  $X_4 = x_4$ , the algorithm would need to be performed again. The calculations used to compute the joint distribution of  $X_3$  and  $X_4$  are of no use for determining the joint distribution of  $X_1$  and  $X_5$ . Thus, the elimination algorithm is efficient from the viewpoint of computer memory, but it is inefficient from a computational time perspective because it does not reuse computations when considering different combinations of evidence and/or desired posterior distributions. An alternative option for performing exact inference that facilitates such reuse is the junction tree algorithm, which is described next.

#### 2.2.3.2 Overview of the junction tree algorithm

The junction tree algorithm (JTA) can be thought of as a generalization of the elimination algorithm. The algorithm creates a data structure known as a junction tree, which contains subsets of random variables in the BN known as *cliques*. These cliques are re-used to facilitate efficient inference for many quantities of interest without re-running the algorithm. The JTA is illustrated here by way of an example.

Like the elimination algorithm, the JTA begins with a moral graph and specification of the elimination order. In the JTA, nodes are eliminated *schematically* and links are introduced to record induced relations, but summations are *not* performed. Instead, cliques and clique potentials are recorded as each node is eliminated (as illustrated by the example below). The recorded cliques and associated potentials are used to create the junction tree. When a node is eliminated, the clique that is formed consists of the eliminated node and its neighbors. A single junction tree is constructed to facilitate inference for all quantities of interest. However, the structure of this junction tree will differ depending on the elimination order that is used to construct it.

Once again, consider the simple BN in Figure 2.1 and specify the elimination order:  $X_4, X_1, X_5, X_3, X_2$ . First consider the elimination of  $X_4$ . In eliminating  $X_4$ , a clique is formed containing  $X_4$  and its neighbors  $X_1$  and  $X_5$ , as shown by a box with dashed lines in the second diagram from the left in Figure 2.7. Assuming all nodes have ten states, the size of the potential associated with this clique is  $10^3$ . As in the elimination algorithm, the neighbors of  $X_4$  are connected after it is eliminated from the BN. Next,  $X_1$  is eliminated, creating a clique consisting of itself and its neighboring nodes. The clique comprised of nodes  $X_1, X_2, X_3$ , and  $X_5$  is shown in the third diagram in Figure 2.7. The size of the associated clique potential is  $10^4$  (once again assuming ten states per node). In eliminating the remaining nodes, no additional cliques must be recorded because they contain nodes which all occur together as a subset of a previously created clique.

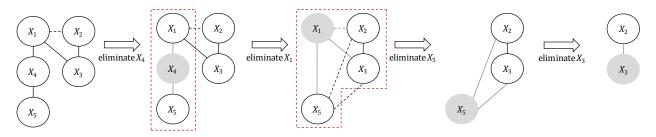


Figure 2.7: Illustration of clique creation using the BN in Figure 2.1 with elimination order  $X_{4,}X_{1,}X_{5,}X_{3,}$  $X_{2}$ 

The junction tree resulting from the above elimination order consists of two cliques and is shown in Figure 2.8. Each clique is shown in an oval with a *separator set* shown as a rectangle. The separator set assigned to each branch of a junction tree includes nodes that are in common to the cliques at the ends of the branch. This junction tree can now be used to answer any probabilistic inquiry on the variables within the BN.

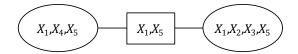


Figure 2.8: Junction tree for BN in Figure 2.1 with elimination order X<sub>4</sub>, X<sub>1</sub>, X<sub>5</sub>, X<sub>3</sub>, X<sub>2</sub>

If, instead, the elimination order  $X_5$ ,  $X_4$ ,  $X_2$ ,  $X_3$ ,  $X_1$  is used to create an alternate junction tree, the size of the largest clique created is smaller. This elimination order corresponds to the optimal order, which can be found because the dimension of the problem is small. Figure 2.9 illustrates the cliques that are formed as nodes are eliminated and Figure 2.10 shows the corresponding junction tree.

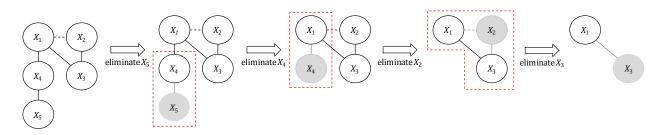


Figure 2.9: Illustration of cliques created using the BN in Figure 2.1 with elimination order  $X_{5,}X_{4,}X_{2,}X_{3,}$  $X_{1}$ 

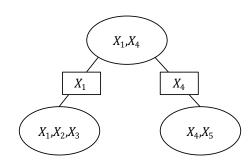


Figure 2.10: Junction tree for BN in Figure 2.1 with elimination order X<sub>5</sub>, X<sub>4</sub>, X<sub>2</sub>, X<sub>3</sub>, X<sub>1</sub>

A junction tree is constructed such that three properties are maintained: (1) the graph is singly connected, i.e. there is only one path between each pair of cliques; (2) each clique created when performing elimination must exist in the tree or be a subset of another clique; and (3) if a pair of cliques  $C_i$  and  $C_j$  both contain node  $X_k$ , then each clique on the path between  $C_i$  and  $C_j$  must also contain  $X_k$  (Paskin 2003).

Propagation of probabilistic information in a junction tree to obtain posterior distributions on all variables utilizes a message passing algorithm. This algorithm involves the bidirectional exchange of information between cliques by performing operations over the variables/nodes in the separator sets. Specifically, the joint distribution of the variables that exist in two adjacent cliques is equal to the product of the clique potentials divided by the separator potentials (Friis-Hansen 2004). For the purposes of this study, it is not necessary to fully understand the mechanics of this propagation algorithm. A detailed explanation can be found in textbooks (e.g. Jensen and Nielson 2007). However, the reader should understand how the growth in computational complexity is related to the sizes of cliques. In particular, two measures of computational complexity are used in later chapters of this study: (1) the largest clique table size, which is the size of the potential associated with the largest clique in the junction tree; and (2) the total clique table size, which is the sum of the sizes of all potentials associated with cliques in a junction tree. For the junction tree in Figure 2.10, the largest clique contains  $X_1, X_2, X_3$ . Assuming ten-state nodes, its potential has size  $10^3$ . The total clique table size for the junction tree is  $10^3 + 10^2 + 10^2$ , where  $10^2$  is the size of the potentials of the two-node cliques. Note that the size of largest

clique table has a domineering contribution to the total clique table size. The memory demand associated with the junction tree based on the optimal ordering (Figure 2.10) is an order of magnitude smaller than the demand associated with the suboptimal ordering in Figure 2.7, which has a total clique table size of  $10^4 + 10^3$ . The order of computational complexity of the JTA is exponential in the size of the largest clique table for BNs with discrete nodes; for BNs with continuous Gaussian nodes, the order of computational complexity of the JTA is cubic in the size of the largest clique table (Paskin 2003).

# 2.3 Use of BNs in civil engineering applications

The use of BNs for engineering risk and reliability analysis has been limited, but growing in recent years. The range of topics to which BNs have been applied is varied and the frequency with which they are used differs depending on the field of application. For example, in a study on the growth of literature relating to risk analysis, dependability, and maintenance modeling via BN, Medina Oliva et al. (2009) found that the use of BNs for dependability analysis comprised the largest portion of the literature surveyed (64%), while references focusing on risk analysis and maintenance applications were fewer (23% and 13%, respectively). Furthermore, they found that the use of BNs for dependability analysis increased by 800% between 2001 and 2007, while the use of BNs for risk analysis increased by a factor of four and their application to maintenance increased by a factor of three over the same period.

The use of BNs specifically for natural hazard assessment has likewise gained momentum in recent years. For example, Straub (2005) presents a generic framework for assessing risks associated with natural hazards using BN and applies the framework to an example involving rock-fall hazard. BNs have also been applied to the modeling of risks due to typhoon (Nishijima and Faber 2007) and avalanches (Grêt-Regamey and Straub 2006). The use of BNs for management of water resources and ecological risk has also grown significantly and was the subject of a special issue of the journal *Environmental Modeling and Software* (Castelletti and Soncini-Sessa 2007). Topics in the special issue include sustainable withdrawal of water from aquifers, use of financial incentives to induce water-saving, and effects of environmental changes on fish populations. Smith (2006) used BNs to analyze geotechnical and hydrological risks posed to a single embankment dam. BNs have also been used for assessing infrastructure risks posed by man-made hazards (Jha 2006).

The use of BNs in seismic applications is relatively limited. A general framework for earthquake risk management using BNs is outlined in Bayraktarkli et al. (2005). The framework is comprised of three components: (1) an exposure model that is an indicator of hazard potential (e.g. ground motion intensity and associated return periods); (2) a vulnerability model which gives an indicator of direct/immediate consequences of a particular hazard (e.g. distribution of component damage given an earthquake of a particular intensity); and (3) a robustness model to quantify indirect consequences. Because the paper focuses on outlining a generic framework, it does not address many of the details which complicate the application of BNs to seismic hazard assessment and which are the subject of a considerable portion of this study, e.g. efficient modeling of

random fields, as well as directivity and finite rupture effects. A continuation of this work is presented in Bayraktarli et al. (2006), which integrates the BN framework with a GIS-based system for considering large inventories of buildings. Because Bayraktarli et al. (2006) focuses on structural modeling, it also does not consider random field effects or issues associated with modeling system performance. Thus, the approach is not directly applicable to seismic risk assessment of infrastructure systems. However, the integration of the BN with GIS systems represents a novel and valuable technological integration that is particularly useful in near-real time applications. Kuehn et al. (2009) use a learning algorithm to determine the "best" BN to model the joint probability distribution of ground motion, site, and earthquake parameters, directly using data from the Next Generation Attenuation (NGA) database (PEER 2005). The work results in a BN that models groundmotion intensity at a single site as a function of source and site characteristics. It represents an alternative to regression-based methods such as those developed under the NGA project (Abrahamson et al. 2008). Additional applications of BNs to seismic hazard include: liquefaction modeling (Bayraktarli 2006); tsunami early warning (Blaser et al. 2009); modeling stochastic dependence between earthquake occurrences (Agostinelli and Rotondi 2003); and post-earthquake consequence assessment (Faizian et al. 2005). Details relating to the above applications of BNs for modeling, assessing, and/or managing seismic hazards are described, as pertinent, in later chapters of this report.

## 2.4 Summary

This chapter presents a brief introduction to BNs. The goal is to provide knowledge about BNs that is necessary for understanding the models and methods described in the remainder of this study. As such, detailed theoretical descriptions are avoided in favor of intuitive and example-based explanations, wherever possible. This chapter introduces the terminology of BNs, presents the basics of their construction, and explains how network topology (e.g. nodes with many parents) and elimination orders affect computational demands via large clique tables that must be stored in computer memory. A brief review of BN applications relevant to the current study is presented.

# Chapter 3: BN-Based Seismic Demand Models

#### 3.1 Introduction

Infrastructure systems, such as utility networks and transportation systems, are typically distributed over large geographic regions. As a result, they have a higher rate of exposure to seismic hazard than single-site facilities and are subject to a wider range of hazards, e.g. ground shaking as well as ground deformation due to liquefaction, landslide, or fault rupture. Some of these effects have spatially distributed natures and must be considered as random fields. For example, the ground motion intensity at distributed points throughout a geographic region constitutes a random field. Random field effects must be included to account for the statistical dependence between the seismic demands on components of the infrastructure system, particularly when they are in close proximity to one another.

This chapter describes BN-based models for selected seismic hazards at discrete points in the domain of a geographically distributed infrastructure system. The focus of the chapter is primarily on the hazard due to ground-shaking. Intensity of ground shaking can be measured using a variety of metrics including: spectral acceleration (SA), peak ground acceleration (PGA), peak ground velocity (PGV), and peak ground displacement (PGD). Typically the performance of components in an infrastructure system will be expressed as a function of one of these quantities. For example, the seismic performance of structural systems (e.g. buildings, bridges) will often be expressed as a function of SA. Liquefaction hazard is typically expressed as a function of PGA. The performance of distributed components in an infrastructure system (e.g. pipelines, highway embankments) is typically defined based on PGV. While other metrics are occasionally used (e.g. Arias intensity), such measures are not considered in this study. The focus of this chapter is on hazard due to

ground shaking. However, preliminary BN models for liquefaction-induced ground deformation as well as displacement due to fault rupture are also developed, but with less detail and rigor.

This chapter begins with a description of BN models for hazard due to ground shaking. A brief summary of existing works is provided. A more general model of ground motion intensity at different points in an infrastructure system is next developed. Proposed models for point-source and finite rupture idealizations are presented, including formulations for including directivity effects. Finally, preliminary BN-based models for liquefaction and hazard due to fault rupture are developed.

## 3.2 Hazard due to ground shaking

This section describes the construction of a BN model of ground motion intensity using physical relations and empirically derived ground motion prediction equations, also known as attenuation laws. These prediction equations are based on regressions of observed data that relate ground motion intensity at a site to earthquake source and site characteristics, e.g. see Abrahamson et al. (2008). Ground motion prediction equations typically have the form

$$\ln S_i = \ln \overline{S}_i + \varepsilon_m + \varepsilon_{r,i} \tag{3.1}$$

where  $\ln \overline{S_i} = f(M, R_i, \Theta_{st,i}, \Theta_{sc})$  is the natural logarithm of the median ground motion intensity (e.g. SA, PGA, PGV, PGD) at site *i* expressed as a function of the earthquake magnitude, *M*; the distance between the site and earthquake source,  $R_i$ ; a vector of site properties  $\Theta_{st,i}$  (e.g. site shear wave velocity, depth to significant impedance contrast); and a vector of source characteristics,  $\Theta_{sc}$  (e.g. faulting mechanism). In (3.1),  $\varepsilon_m$  is an interevent error term and  $\varepsilon_{r,i}$  is an intra-event error term. Both errors describe variability in the logarithmic intensity value  $\ln S_i$  relative to the median  $\ln \overline{S_i}$ . The inter-event error term captures the variability from earthquake-to-earthquake that is common to all sites, while the intra-event error term captures the variability from site-to-site for a single earthquake. Both error terms are zero-mean and normally distributed. Furthermore, the  $\varepsilon_{r,i}$  for different sites are correlated due to the random-field nature of the ground motion.

In developing the proposed BN, several challenges are encountered, including the need to account for the correlation structure of a random field and the effects of finite rupture and directivity. These issues are not addressed in the limited literature that exists on the application of BNs to seismic hazard assessment. In the following sections, existing works related to modeling hazard due to ground shaking are reviewed, followed by the development of new and more general models applicable to modeling ground motion intensity at sites distributed over a geographic region.

#### 3.2.1 Existing models

Bayraktarli et al. (2005) present a general BN framework for management of risks due to earthquakes. The framework consists of three-components: (1) an exposure model representing the hazard potential (e.g. ground motion intensity associated with a particular return period); (2) a vulnerability model representing immediate consequences (e.g. structural damage) conditioned on an exposure event; and (3) a robustness model representing indirect consequences (e.g. loss of life) dependent on the vulnerability. The associated BN model is shown in Figure 3.1. Pertinent to this chapter is the exposure model. The seismic demand measure considered in the study is spectral displacement. It is computed as a function of earthquake magnitude, source-to-site distance (modeled by the node labeled Earthquake Distance), soil type, and the fundamental period of the structure. using a software application that provides time histories consistent with a response spectrum obtained using the Boore et al. (1997) attenuation relationship. The authors make several simplifying assumptions in developing the exposure model, e.g. they do not address finite rupture and directivity effects nor do they explicitly account for the uncertainties in the attenuation models. Their exposure model can thus be considered as a special case of the more comprehensive seismic demand model developed in this chapter.

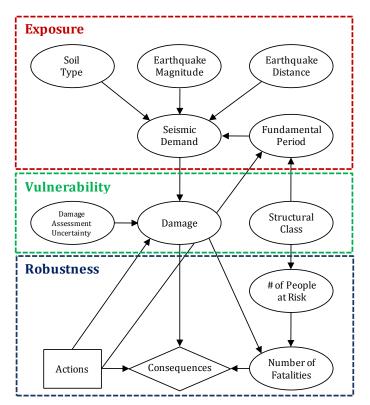
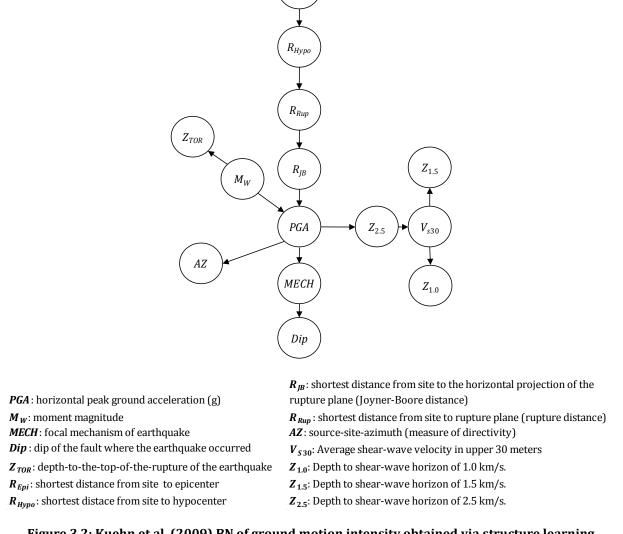


Figure 3.1: Bayraktarli et al. (2005) BN framework for seismic risk management

Bayraktarli et al. (2006) expand upon Bayraktarli et al. (2005). The exposure model considered in the 2006 study quantifies seismic demands using two intensity measures:

spectral displacement (for estimating vulnerability of buildings) and peak ground acceleration (for estimating vulnerability of soils). The spectral displacement at a site is specified as in the 2005 study as a function of earthquake magnitude, source-to-site distance, fundamental period, and site class. The node *Earthquake Magnitude* is discretized into four states, each corresponding to one discrete magnitude value, with the probability of each state specified using the Gutenberg-Richter magnitude-recurrence relationship (Gutenberg and Richter 1944). The node *Earthquake Distance* has four states, each corresponding to a single distance value. Distances are computed by assuming the earthquake source is a point. Peak ground acceleration is obtained from simulated time histories that are generated for combinations of magnitudes and distances.

Kuehn et al. (2009) use structure learning to determine the "best BN topology" for modeling dependencies between ground motion intensity and source/site characteristics. The BN is learned directly from data available in the PEER NGA database (PEER 2005). Their approach offers an alternative to regression-based methods, such as those developed under the NGA project (Abrahamson et al. 2008). The resulting BN is shown in Figure 3.2 along with the definition of variables used in the model. Consistent with the NGA models, PGA is causally dependent on earthquake magnitude and source-to-site distance (lovner-Boore distance). Furthermore, there is an induced relationship between all source-to-site distance measures that are available in the PEER NGA database and which appear in ground motion prediction equations. While important intuitive similarities exist between the empirically and theoretically derived ground motion prediction equations and the learned BN, the directions of links in Figure 3.2 are not always consistent with the directions suggested by the mathematical form of the NGA models. For example, the BN in Figure 3.2 has a link going from PGA to  $Z_{2.5}$ , indicating that the depth to shear-wave velocity horizon of 2.5 km/s is probabilistically dependent on PGA. This is not consistent with a causal interpretation. However, it is important to note that, while BNs developed using a priori knowledge (e.g. physical models, empirical relations, and expert opinion) are typically constructed using a causal interpretation, BNs developed using algorithms designed to find the "best" topology need not reflect such an intuitive approach. The structural learning algorithms find dependencies, which are likely to be consistent with the d-separation properties of the problem, as supported by data, rather than reflecting direct causality. Thus, the directed links between  $Z_{2.5}$ ,  $Z_{1.5}$ ,  $Z_{1.0}$  and  $V_{s30}$  indicate probabilistic dependencies rather than causal relations.



R<sub>Epi</sub>

Figure 3.2: Kuehn et al. (2009) BN of ground motion intensity obtained via structure learning

Several important observations follow from the form of the BN in Figure 3.2, which provide insight into variables affecting ground motion intensity (PGA) at a site. Notice that  $Z_{2.5}$ mediates between PGA and  $V_{s30}$ . The d-separation properties of the learned BN indicate that, when the value of  $Z_{2.5}$  is known, knowledge of  $V_{s30}$  supplies no additional information about the PGA. Kuehn et al. (2009) suggest that this observation supports claims that  $V_{S30}$ is not a good proxy for site effects. Similarly, the authors note that variables DIP and  $Z_{TOR}$ , which have been incorporated in some of the most recent NGA models, are only indirectly related to PGA. As seen in Figure 3.2, the mediating variables  $M_w$  and MECH, when observed, d-separate  $Z_{TOR}$  and DIP from PGA. Last, notice that AZ, which is a measure of directivity, is a child of node *PGA*. As indicated by the authors, this suggests that directivity effects are present in the records of the NGA database. Kuehn et al. (2009) also show how the results obtained from the BN can be used to generate hazard curves. However, as with Bayraktarli et al. (2005/2006), the BN in Kuehn et al. (2009) is not directly applicable to describing correlated seismic demands on the components of a spatially distributed infrastructure system because it is formulated for performing hazard assessment for a single site. To facilitate information updating in near-real time such that observations made at one location in the system update the distributions of the ground motion intensity at other sites, it is necessary to include a model of the correlation of seismic demands. Furthermore, for highly redundant systems, neglecting correlation in demands results in an overestimation of the system reliability. These issues are addressed in the formulations developed in this chapter.

# 3.2.2 Proposed model

Next, we describe a BN-based formulation for modeling seismic demands on an infrastructure system. This formulation provides the distribution of ground motion intensity at discrete points in the geographic domain of a spatially distributed infrastructure system for a future earthquake of unknown characteristics on one of multiple potential seismic sources. Models are developed for characterizing the source as well as accounting for correlation in levels of ground motion intensity at spatially distributed locations. We begin with geometric derivations necessary to define the location of an earthquake on a fault and the consequent source-to-site distance for each site in the infrastructure system. The ground motion intensity at each location is then defined as a function of the source-to-site distance, earthquake magnitude, and other factors. Finally, we extend the seismic demand model to account for directivity effects.

#### 3.2.2.1 Source characterization and source-to-site distance

As described in 3.2.1, existing works do not explicitly address issues related to source characterization and geometry. This section develops physically derived representations for modeling earthquake source characteristics and idealized geometry, leading to conditional definitions of the source-to-site distance. For all formulations, a fault is idealized as a straight line or as a collection of contiguous straight line segments. A point-source model is presented first, followed by a finite-rupture source formulation.

# 3.2.2.1.1 Point-source model and formulation

The use of a point-source assumption results in a topologically simple BN. Furthermore, the mathematical relations necessary to conditionally define the distance between the point source and each site are relatively simple for this formulation, as shown in the following.

Consider the BN in Figure 3.3, which models source characteristics and source-to-site distances assuming an earthquake occurs as a point in space. This BN model is valid conditional on the occurrence of an earthquake. The highest level root node in the BN,  $S_C$ , represents the source on which the earthquake occurs. This node is an ancestor to all other nodes in the BN. The marginal probability table assigned to node  $S_C$  defines the probability

that an earthquake in the region occurs on each of the sources included in the model. This probability is obtained as a function of the mean rate of occurrence of earthquakes of engineering significance on each source. Let  $v_i$  define this mean rate for source *j*. Then

$$P(\text{earthquake on source } j | \text{earthquake has occurred}) = \frac{\nu_j}{\sum_{k=1}^{N_{sc}} \nu_k}$$
(3.2)

where  $N_{sc}$  denotes the number of seismic sources included in the model.

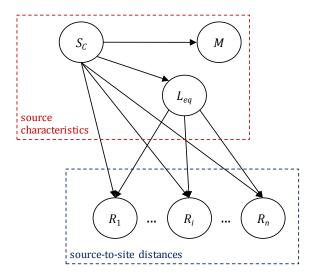


Figure 3.3: BN model of point-source characteristics and source-to-site distances

Node *M*, representing earthquake magnitude, is specified as a child of the source node  $S_{c_i}$ indicating that the probability of experiencing an earthquake of a certain magnitude differs depending on the source. Source dependence of earthquake magnitude is consistent with most earthquake magnitude recurrence relationships. Furthermore, the largest earthquake magnitude generated by a source depends on its size. Common magnitude-recurrence relations include the truncated GR law and characteristic earthquake model (Gutenberg and Richter 1944; Youngs and Coppersmith 1985). Throughout this study, all references to the magnitude of an earthquake should be taken to mean the moment magnitude. While node *M* has no children in Figure 3.3, it is included in the BN because it will be used in later descriptions. Care should be taken when discretizing *M* because the energy released by an earthquake exponentially increases with the magnitude. Specifically, an increase in moment magnitude by a value of one is associated with approximately 30 times more energy; an increase by a value of two is associated with a three-order of magnitude increase in energy. Thus, the difference in energy released by earthquakes with magnitudes M = 8.5 and M = 9.0 is much larger than the difference in energy released by earthquakes with magnitudes M = 4.5 and M = 5.0. For this reason, a constant discretization interval

for the magnitude is not ideal. Instead it is recommended to use smaller intervals at larger magnitude values.

In Figure 3.3, node  $L_{eq}$  represents the point-source location of the earthquake.  $L_{eq}$  is a child of node  $S_c$ , indicating the dependence of the earthquake location on the source on which it occurs. In lieu of contrary information, it is common to assume that the uncertain location of the earthquake is uniformly distributed within each source. For the analysis of earthquakes on each source, we find it convenient to use a source-specific coordinate system. The location of seismic sources is often defined using latitude, longitude, and altitude. This information is easily converted to the Cartesian Earth-center Earth-fixed (ECEF) coordinate system (Clynch 2006). Then, a transformation into the local coordinate system for each source is necessary. This transformation is described in Appendix 3.1 at the end of this chapter. Consider a line-fault idealization of the *j*<sup>th</sup> earthquake source. Using a source-specific coordinate system with the fault lying on the *x*-axis and the origin centered at one end of the fault (see Figure 3.4), the conditional distribution of the location of the earthquake is defined uniformly on the interval  $[0, L_F]$ , as described below.

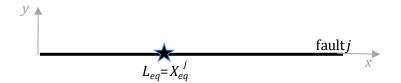


Figure 3.4: Fault-specific local coordinate system for line-fault idealization and point-source model

When considering multiple sources, the states of  $L_{eq}$  must be defined to span the range  $[0, maxL_F]$ , where  $maxL_F = \max_{i=1,...,N_{sc}} L_F^i$  and  $L_F^i$  is the length of fault *i*. The conditional PMF of  $L_{eq}$  for all sources with  $L_F^i < maxL_F$  will have states, corresponding to  $L_{eq} > L_F^i$ , with zero probabilistic weight. For illustration, consider two faults with lengths of 5 and 8 (in an arbitrary unit system). Crudely discretize  $L_{eq}$  into four states: [0,2], [2,4], [4,6], and [6,8]. Assuming that  $L_{eq}$  is equally likely to occur anywhere within each fault, the source dependent distributions of  $L_{eq}$  are shown in Figure 3.5. Note that, for source 1, the last state, [6,8], is associated with zero probability mass. Furthermore, for the same source, the probability mass associated with state [4,6] is half of the mass associated with lower states. This is because the length of source 1 (5 units) falls halfway within the third interval state.

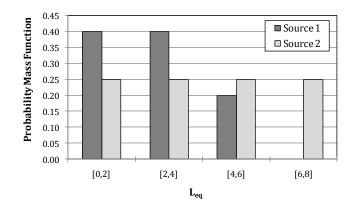


Figure 3.5: Source-dependent distribution of *L<sub>eq</sub>* for example

To account for differences in fault length, the source dependent PMF for  $L_{eq}$  is defined using the expression

$$P_{L_{eq}}^{(j,k)} = \left\{ \mathbb{I} \Big[ 0 < L_{eq,U}^{(k)} \cap L_{eq,U}^{(k)} < L_F^j \Big] \Big( \frac{L_{eq,U}^{(k)}}{L_F^j} \Big) + \mathbb{I} \Big[ L_F^j < L_{eq,U}^{(k)} \Big] \right\}$$

$$- \left\{ \mathbb{I} \Big[ 0 < L_{eq,L}^{(k)} \cap L_{eq,L}^{(k)} < L_F^j \Big] \Big( \frac{L_{eq,L}^{(k)}}{L_F^j} \Big) + \mathbb{I} \Big[ L_F^j < L_{eq,L}^{(k)} \Big] \right\}$$
(3.3)

where  $P_{L_{eq}}^{(j,k)}$  is the probability mass associated with the  $k^{\text{th}}$  state of node  $L_{eq}$ , given an earthquake on source j. The  $k^{\text{th}}$  state of  $L_{eq}$  is associated with the interval  $\left[L_{eq,L}^{(k)}, L_{eq,U}^{(k)}\right]$  consistent with the notation convention used in section 2.2.2.1.  $\mathbb{I}[a]$  is an indicator variable defined such that  $\mathbb{I}[a] = 1$  if a is true and  $\mathbb{I}[a] = 0$  if a is false. (3.3) is based on the assumption that the earthquake is equally likely to occur anywhere along a fault.

The distance between site *i* and the location of the point-source earthquake on fault *j*,  $R_i^j$ , is easily computed in the source-specific coordinate system as

$$R_{i}^{j} = \sqrt{\left(x_{s,i}^{j} - X_{eq}^{j}\right)^{2} + \left(y_{s,i}^{j}\right)^{2} + \left(z_{s,i}^{j}\right)^{2}}$$
(3.4)

where  $x_{s,i}^{j}$ ,  $y_{s,i}^{j}$  and  $z_{s,i}^{j}$  are the *x*-, *y*-, and *z*-coordinates of site *i* in the coordinate system defined by source *j*.  $X_{eq}^{j}$  is the coordinate of the earthquake source ( $L_{eq}$ ) along the *x*-axis in the *j*<sup>th</sup> source-specific coordinate system (see Figure 3.4). Nodes representing the distance between site *i* and an earthquake on source *j* are shown in Figure 3.3. Nodes corresponding

to the site coordinates are not shown because, for a given source, they are considered deterministic. The CPTs of nodes  $R_i$  are defined based on the above relation using Monte Carlo simulation.

For infrastructure systems distributed over large areas, sometimes it is necessary to model bends in faults. For this purpose we extend the above formulation and idealize the fault as a continuous set of multiple line segments. The BN in Figure 3.3 does not change for such a fault. Only the relationships required to construct the CPTs must be revised.

Figure 3.6a shows a three-segment fault defined by the four points  $\mathbf{P}_{f,i} = (x_{f,i}, y_{f,i}, z_{f,i}), i = 0, ..., 3$ , in three-dimensional space. It is numerically convenient to, once again, convert latitude/longitude/altitude data to ECEF coordinates, then shift the origin of the coordinate system to one end of the source, and finally orient the coordinate system such that the *x*-axis lies along the first fault segment. For convenience, the superscript *j*, which previously denoted the source index, has been dropped in the following description. It is assumed that all points are defined within the coordinate system associated with the fault on which the earthquake occurs.

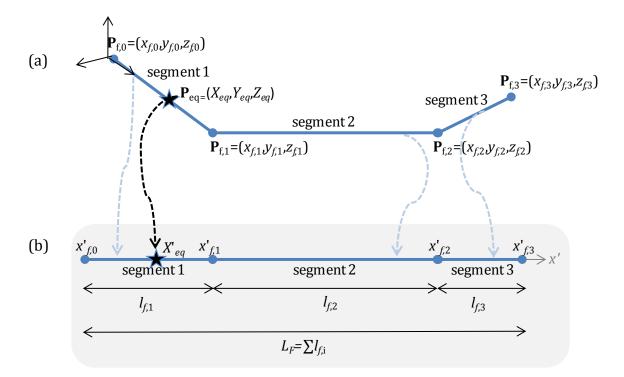


Figure 3.6: Illustration of (a) a multi-segment fault in three-dimensional space; and (b) the alignment of all fault segments consecutively along the *x*'-axis

Define  $N_{seg}$  as the number of discrete segments comprising the fault and  $l_{f,i}$ ,  $i = 1, ..., N_{seg}$ , as the length of the  $i^{\text{th}}$  segment computed using the coordinates of the segment ends:

$$l_{f,i} = \sqrt{\left(x_{f,i} - x_{f,i-1}\right)^2 + \left(y_{f,i} - y_{f,i-1}\right)^2 + \left(z_{f,i} - z_{f,i-1}\right)^2}$$
(3.5)

It follows that the total length of the fault is  $L_F = \sum_{i=1}^{N_{seg}} l_{f,i}$ . Consider that the segments of the fault are aligned to form a single straight line of length  $L_F$ , as illustrated in Figure 3.6b. A one-to-one mapping exists between the location of the point-source in space,  $\mathbf{P}_{eq} = (X_{eq}, Y_{eq}, Z_{eq})$ , and the coordinate  $X'_{eq}$  on the x'-axis. This coordinate is represented by node  $L_{eq}$  in the BN in Figure 3.3.  $X'_{eq}$  is assumed distributed uniformly on the interval  $[0, L_F]$  unless information is available to support a different assumption. For each state of  $L_{eq}$  (i.e. each value of  $X'_{eq}$ ), the source-to-site distance,  $R_i$ , is computed as follows. First, compute the coordinates of the earthquake point-source  $\mathbf{P}_{eq}$  as a function of  $X'_{eq}$ :

$$\mathbf{P}_{eq} = \sum_{i=1}^{N_{seg}} \mathbb{I}[x'_{f,i-1} \le X'_{eq} < x'_{f,i}] * \left(\mathbf{P}_{f,i-1} + (X'_{eq} - x'_{f,i-1}) * \frac{\mathcal{D}(\mathbf{P}_{f,i-1}, \mathbf{P}_{f,i})}{|\mathcal{D}(\mathbf{P}_{f,i-1}, \mathbf{P}_{f,i})|}\right)$$
(3.6)

where  $\mathbb{I}[a]$  is once again an indicator variable defined such that  $\mathbb{I}[a] = 1$  if a is true and  $\mathbb{I}[a] = 0$  otherwise,  $x'_{f,i} = \sum_{j=1}^{i} l_{f,j}$  and  $x'_{f,0} = 0$ .  $\mathcal{D}(\mathbf{P}_{f,i-1}, \mathbf{P}_{f,i})$  is a direction vector defined as

$$\mathcal{D}(\mathbf{P}_{f,i-1},\mathbf{P}_{f,i}) = \begin{bmatrix} x_{f,i} - x_{f,i-1} \\ y_{f,i} - y_{f,i-1} \\ z_{f,i} - z_{f,i-1} \end{bmatrix}$$
(3.7)

Given the location of the earthquake point-source in three-dimensional space, the sourceto-site distance is computed as

$$R_{i} = \sqrt{\left(x_{s,i} - X_{eq}\right)^{2} + \left(y_{s,i} - Y_{eq}\right)^{2} + \left(z_{s,i} - Z_{eq}\right)^{2}}$$
(3.8)

The CPTs of the nodes in Figure 3.3 are defined based on the above relationships using Monte Carlo simulation as described in section 2.2.2.1 (or, if preferred, using other methods for computing probabilities from known relations).

The above relations permit the use of multi-segment fault idealizations without complicating the topology of the BN. Because the topology of the BN need not be changed when including bends in the fault, there is no computational disadvantage, from the perspective of the BN, associated with accurately modeling the fault trace using a large number of discrete but connected line segments. An exception to the above statement may be encountered if the number of states associated with any node needs to be increased as a result of including bends in faults.

#### 3.2.2.1.2 Finite rupture model and formulation

Consider the more realistic idealization of an earthquake as a finite rupture along the fault rather than as a single point. We begin again with the special case in which a fault is idealized as a straight line. The source-specific local coordinate system with the fault oriented along the x -axis and its left end located at the origin is used, see Figure 3.7. In this coordinate system, the epicenter is defined as a point along the fault. The rupture is assumed to occur anywhere along the fault with uniform likelihood, while containing the epicenter and not extending beyond the known ends of the fault. In the formulations defined below it is assumed that all coordinates and source parameters are defined within the j<sup>th</sup> fault specific coordinate system. The superscript j has again been omitted for convenience and clarity.

The following notation is used for the finite rupture formulation and is illustrated in Figure 3.7:

 $X_e$  = location of the epicenter along the *x*-axis in the fault specific coordinate system

 $x_{f,1}$ ;  $x_{f,2} = x$ -direction coordinates of the ends of the fault in the source-specific coordinate system. It follows from Figure 3.7 that  $x_{f,1} = 0$  and  $x_{f,2} = L_F$ .

 $A_L$ ;  $A_R$  = the amount of the rupture that propagates to the left (towards the origin) and right (away from the origin) of the epicenter, respectively.

 $X_r = x$ -direction coordinate of the end of the rupture closest to the origin, hereafter referred to as the *rupture reference coordinate*.

 $R_L$  = length of the rupture =  $A_L + A_R$ . Note that, in general,  $R_L < L_F$ .

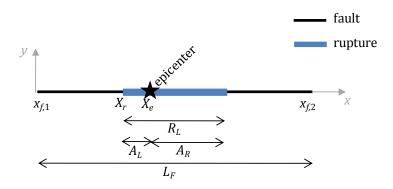


Figure 3.7: Fault and rupture in the local coordinate system

The uncertain location of the epicenter is often assumed to be distributed uniformly along the fault. While this assumption is used in this study, the derivations that follow are applicable for any probability distribution used to define the uncertain location of the epicenter. The distribution of  $X_e$  is defined by adapting (3.3).

For a given epicenter location, the rupture propagates to the left and right along the fault in uncertain proportions but constrained by the finite fault geometry. Assuming that the amount of rupture to the left is distributed uniformly within the physical boundaries (any other distribution can be used if available information supports it, but the following expressions would require adaptation), the distribution of  $A_L$  is

$$A_L \sim unif(\max(0, R_L - L_F + X_e), \min(R_L, X_e))$$
(3.9)

The above expression bounds the amount of rupture propagating to the left of the epicenter at a maximum of the length that can "fit" to the left and at a minimum of the length that cannot fit to the right. The amount of rupture that propagates to the right is simply the "leftover" portion of the rupture length, i.e.,  $A_R = R_L - A_L$ .

The rupture reference coordinate is then equal to  $X_e - A_L$  and it follows that the conditional distribution of  $X_r$  is

$$X_{r} \sim unif \left[ X_{e} - \min(R_{L}, X_{e}), X_{e} - \max(0, R_{L} - (L_{F} - X_{e})) \right]$$
(3.10)

Figure 3.8 shows a BN model of the above relationships. For each fault, there is a sourcedependent distribution of the earthquake magnitude and epicenter location. The rupture length on the fault is a function of the earthquake magnitude and source characteristics (Wells and Coppersmith 1994). The distribution of  $R_L$  is typically truncated at a maximum considered value, often one-half the fault length. In the BN, nodes M and  $X_e$  are specified as children of node  $S_C$ . Node  $R_L$  is a child of nodes M and  $S_C$  modeling the source-dependent magnitude-rupture length relationship. Consistent with (3.10),  $X_r$  is a child of nodes  $X_e$  and  $R_L$  as well as node  $S_C$ , which captures the dependence of  $X_r$  on fault length and the maximum rupture length considered on the source. The construction of the CPT of  $X_r$  is described next.

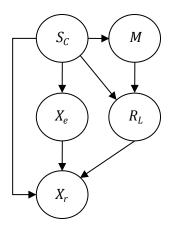


Figure 3.8: BN modeling rupture length and location as a function of earthquake source, epicenter location, and magnitude

To compute the CPT of  $X_r$  for each combination of the states of  $X_e$  and  $R_L$ ,  $N_{sim}$  draws are taken from the associated intervals. For each simulation, a value of  $X_r$  is drawn from the distribution in (3.10) based on the drawn values of  $X_e$  and  $R_L$ . Then, a normalized histogram is computed to obtain the required column of the CPT for  $X_r$ , with bins defined corresponding to the discrete interval states specified based on the admissible range of  $X_r$ . The range associated with  $X_r$  will typically be the same as the admissible range of  $X_e$ . The sampling intervals for  $X_e$  and  $R_L$  must account for source-specific differences in the fault length as well as the assumed maximum considered rupture length on each fault. Define  $maxR_L = max_{i=1,\dots,N_{S_c}} maxR_{L,i}$ , where  $maxR_{L,i}$  is the maximum considered rupture length on source *i*. Define  $R_L^{(k)}$  as the  $k^{\text{th}}$  state associated with node  $R_L$ , which corresponds to the interval  $[R_{L,L}^{(k)}, R_{L,U}^{(k)}]$ . Similarly, define  $X_e^{(k)}$  as the  $k^{\text{th}}$  state representing the interval  $[X_{e,L}^{(k)}, X_{e,U}^{(k)}]$ . To account for fault-specific geometric constraints given an earthquake on  $R_L^{(k)}$ should sampled source *i*. be uniformly from the interval  $\left[\min\left(\max_{L,i}, R_{L,i}^{(k)}\right), \min\left(\max_{L,i}, R_{L,i}^{(k)}\right)\right].$  Similarly  $X_e^{(k)}$  should be sampled uniformly from the interval  $\left[\min\left(L_F^i, X_{e,L}^{(k)}\right), \min\left(L_F^i, X_{e,U}^{(k)}\right)\right].$  For example, consider two faults of lengths 5 and 8 (in arbitrary units). Assume that the maximum considered rupture length for each fault is equal to one-half of the fault length:  $maxR_{L,1} = 2.5$  and  $maxR_{L,2} = 4$ . Crudely discretize  $R_L$  to have states [0,2] and [2,4] and  $X_e$  to have states [0,2], [2,4], [4,6], and [6,8]. Given an earthquake on source 1, the sampling interval for the first state of  $R_L$  is  $[\min(2.5,0), \min(2.5,2)] = [0,2]$ ; the sampling interval for the second state is  $[\min(2.5,2), \min(2.5,4)] = [2,2.5]$ . Note that, for the second state, values above the maximum considered rupture length (2.5 units) are not included. For an earthquake on the same source, the sampling intervals for  $X_e$  are  $[\min(0,5), \min(2,5)] = [0,2]$ ;  $[\min(2,5), \min(4,5)] = [2,4]$ ;  $[\min(4,5), \min(6,5)] = [4,5]$ ; and  $[\min(6,5), \min(8,5)] =$ [5,5]. Note that, for the third state of  $X_e$ , epicenter locations beyond the length of the fault, i.e. greater than 5 units, are not considered. The chosen sampling interval for the last state

of  $X_e$  (i.e. [6,8]) is irrelevant given an earthquake on source 1 because it is associated with zero probability.

For a given rupture length, the source-to-site distance (the shortest distance between the site and the rupture),  $R_i$ , is defined for each site within the source-specific coordinate system. See Figure 3.9 for examples of this distance for various site-rupture configurations. The distance between the  $i^{\text{th}}$  site and a rupture on the fault, is computed as

$$R_{i} = \sqrt{\left(X_{d,i}\right)^{2} + \left(y_{s,i}\right)^{2} + \left(z_{s,i}\right)^{2}}$$
(3.11)

where  $y_{s,i}$  and  $z_{s,i}$  are the coordinates of site *i* in the fault-specific coordinate system;  $X_{d,i}$  is the *x*-direction distance between site *i* and the nearest point on the rupture on the source.

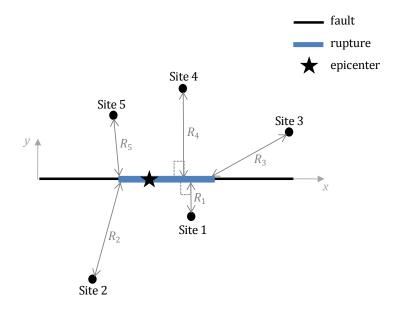


Figure 3.9: Geographically distributed sites in vicinity of a fault

For the straight-line idealization of a fault,  $X_{d,i}$  is expressed using simple geometry as

$$X_{d,i} = \max[X_r - x_{s,i}, x_{s,i} - \min(X_r + R_L, L_F), 0]$$
(3.12)

where  $x_{s,i}$  is the *x*-direction coordinate of site *i*. The term  $\min(X_r + R_L, L_F)$  specifies the coordinate of the right end of the rupture. From a geometric perspective, including the expression  $\min(X_r + R_L, L_F)$  instead of  $(X_R + R_L)$  is redundant because (3.10) has been derived such that the rupture reference coordinate can only assume values that ensure the rupture is contained within the fault. However, in the BN, a conditional distribution of  $X_{d,i}$  must be defined for all combinations of its parent nodes, even for physically impossible

combinations of  $X_r + R_L$ . To avoid introducing additional states, the value  $L_F$  is assigned to  $X_{d,i}$  for combinations in which  $X_r + R_L > L_F$ . Note that physically impossible combinations of  $X_r + R_L$  have probability zero, so that this assignment has no effect on the results. It should be clear that node  $R_i$  must be a child of nodes  $S_C$ ,  $R_L$ , and  $X_r$  and a function of the site coordinates as well.

In computing the CPT of  $R_i$  using the above expressions on conjunction with Monte Carlo simulation, errors may arise due to inconsistencies resulting from discretization, as illustrated below by example. Consider a fault of length 30 (in an arbitrary unit system) and a site located at coordinate (8, -5) as shown in Figure 3.10. Let the location of the epicenter be a deterministic quantity (i.e. node  $X_e$  has only one state) with value equal to 15. Additionally, let the rupture length be a deterministic quantity with a value of 10 units, but uncertain location. Consider that we are interested in the *x*-direction distance between the site and the nearest point on the rupture,  $X_{d,i}$ . The BN computing this quantity using the formulation in (3.12) is shown in Figure 3.11a.

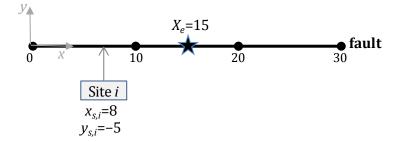


Figure 3.10: Example used to illustrate geometric inconsistencies due to discretization

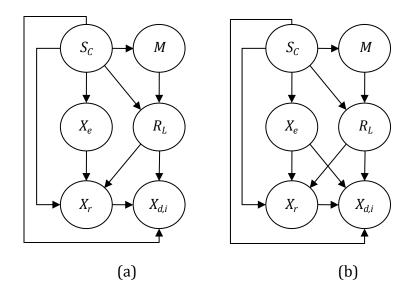


Figure 3.11: (a) BN modeling relationship in (3.12); and (b) BN modeling relationship in after link added between *X<sub>e</sub>* and *X<sub>d</sub>* to account for geometric inconsistencies due to discretization

For this example, assume  $X_r$  is crudely discretized into 3 states: [0,10]; [10,20]; [20,30], and  $X_{d,i}$  is arbitrarily discretized into 3 states: [0,5]; [5,10]; [10,15]. (In actual applications, the states of  $X_{d,i}$  should be carefully defined based on the geometry of the problem.) Looking at the geometry shown in Figure 3.10, it is obvious that the maximum value  $X_{d,i}$  can assume is  $X_e - x_{s,i} = 7$ . Thus, the probabilistic weight of the last state assigned to  $X_{d,i}$ , [10,15], should always be zero.

For the known epicenter and rupture length, and using the expression in (3.10),  $X_r$  will fall into the range [5,15] according to the uniform distribution. Half of this admissible range falls into state 1 (i.e. [0,10]), and the other half falls into state 2 (i.e. [10,20]). Therefore, there is a 0.5 probability that  $X_r$  will fall into each of its first two states and a 0.0 probability of falling into the last state.

Generating the CPT of  $X_{d,i}$  using Monte Carlo simulation based on the expression in (3.12), results in the CPT shown in Table 3.1. The first column of this table corresponds to a true result:  $X_{d,i}$  is certain to fall in the range [0,5] because  $X_r$  is located in the range [0,10]. The last column of this table is inconsequential because there is zero probability that  $X_r$  will fall into the state [20,30]. However, the distribution contained in the second column is erroneous due to the non-zero probability that  $X_{d,i}$  is in the state [10,15], which conflicts with our earlier observation that  $X_{d,i}$  can be at most equal to  $X_e - x_{s,i} = 7$ . This inconsistency arises due to discretization of the random variables describing the geometry of the problem.

Recall from section 2.2.2.1, that when generating the CPT for a node that has interval nodes as parents, values are drawn *uniformly* from within the intervals associated with each

combination of its parent nodes states. Thus, when generating the second column of the CPT associated with node  $X_{d,i}$ , values of  $X_r$  are drawn uniformly from within the range [10,20] despite the fact that values of  $X_r > 15$  are not geometrically admissible given the known epicenter location. Because a coarse refinement has been used to discretize  $X_r$ , the inconsistencies due to discretization result in large errors as seen in Table 3.1. When a more refined discretization is used, e.g. [0,5]; [5,10]; [10,15]; [15,20]; [20,25]; [25,30], the effect of geometric discretization inconsistency is negligible and no remedial action need be taken. However, if sufficiently large errors are encountered, and additional refinement of discretization is not an option, the inconsistency problem can be addressed by adding a link from node  $X_e$  to  $X_{d,i}$  and adapting the Monte Carlo simulation scheme to compute the CPT of  $X_{d,i}$  and  $R_i$  as described in Figure 3.12. The resulting BN is shown in Figure 3.11b. The procedure in Figure 3.12 is a conceptual and generic scheme, which has not been optimized for computational efficiency. Also note that, in this algorithm,  $X_{i,U}^{m_j}$  and  $X_{i,U}^{m_j}$  denote the lower and upper bounds of the  $j^{\text{th}}$  state of interval node  $X_i$ .  $m_{X_r}, m_{X_r}$ , and  $m_{R_r}$  are defined as the number of states associated with nodes  $X_e, X_R$ , and  $R_L$  respectively.

R <sub>L</sub>	10						
Xr	0-10	10-20	20-30				
0-5	1	0.28	0				
5-10	0	0.52	0				
10-15	0	0.2	0.28				
15-20	0	0	0.52				
20-25	0	0	0.2				

for  $i = 1, ..., m_{X_e}, j = 1, ..., m_{X_r}, k = 1, ..., m_{R_L}$ for  $n = 1, ..., N_{sim}$  (number of Monte Carlo simulations used to generate the CPT) 1. Generate  $X_e \sim unif[X_{e,L}^i, X_{e,U}^i]$ 2. Generate  $X_r \sim unif[j, \min(X_{r,U}^j, X_e)]$ 3. Generate  $X_L \sim unif[R_{L,L}^k, R_{L,U}^k]$ 4. Calculate  $X_{d,i}$  using (2.3) (3.12) 5. Calculate  $R_i$  using (3.11) end Determine normalized histogram of  $X_{d,i}$ end

# Figure 3.12: Algorithm for generating CPTs of $X_{d,i}$ and $R_i$ to address geometric inconsistencies due to discretization

The BN computing source-to-site distance must likewise be adapted to account for the dependence of  $X_{d,i}$  on  $X_e$  (which implies  $R_i$  is dependent on  $X_e$ ). The resulting BN is shown in Figure 3.13. The algorithm in Figure 3.12 describes the construction of the CPTs of  $R_i$  via

Monte Carlo simulation. Note that, in Figure 3.11b and Figure 3.13, nodes  $X_{d,i}$  and  $R_i$  have four parents. As a general rule, the number of parents to a node should be at most three, so potential computational inefficiencies may arise from this formulation.

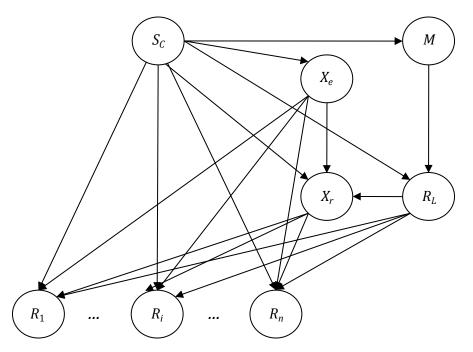


Figure 3.13: BN modeling of source-to-site distance as a function of earthquake source and location

As with the point-source case, we next extend the above formulation to the case in which a fault is idealized as a set of connected line segments. The multi-segment fault formulation only affects the computation of CPTs, not the topology of the BN in Figure 3.13. To facilitate the required derivations, we define the following additional variables:

 $\mathbf{P}_{s,i} = [x_{s,i}, y_{s,i}, z_{s,i}] =$  vector containing the *x*-, *y*-, and *z*-coordinates of site *i* in threedimensional space

 $r_{i,k}$  = distance between site *i* and the rupture on segment *k* of the source, if the rupture crosses the segment

 $\mathbf{P}_{f,k} = [x_{f,k}, y_{f,k}, z_{f,k}] =$  vector containing the *x*-, *y*-, and *z*-coordinates of a point defining the location of the end of a segment of the fault

 $\mathbf{P}_r = [X_r, Y_r, Z_r]$  = vector defining the location of the rupture reference coordinate in threedimensional space

For the finite rupture case, we use the same method as used in for the point-source formulation in which the  $N_{seg}$  fault segments are aligned to form a single line of length  $L_F$  lying on the x'-axis, as illustrated in Figure 3.6. The earthquake epicenter and rupture

reference coordinate are both mapped from the three-dimensional space onto a coordinate along the x'-axis. Define  $X'_e$  and  $X'_r$  to be the coordinates of the epicenter and rupture reference coordinate on this axis.

Define  $K_{rup} \subset \{1, ..., N_{seg}\}$  as the set of segment indices corresponding to segments crossed by the rupture. The rupture will cross fault segment k if  $X'_r < x'_{f,k}$  and  $(X'_r + R_L) > x'_{f,k-1}$ , where  $x'_{f,k}$  is the x'-coordinate of the fault corresponding to  $\mathbf{P}_{f,k}$  and is computed as

$$x'_{f,k} = \sum_{i=1}^{k} l_{f,i} \tag{3.13}$$

It follows that the rupture reference coordinate in the three-dimensional source-specific coordinate system is

$$\mathbf{P}_{r} = \sum_{k=1}^{N_{seg}} \mathbb{I}[x_{f,k-1}' \le X_{r}' \le x_{f,k}'] \left(\mathbf{P}_{f,k-1} + (X_{r}' - x_{f,k-1}') * \frac{\mathcal{D}(\mathbf{P}_{f,k-1}, \mathbf{P}_{f,k})}{|\mathcal{D}(\mathbf{P}_{f,k-1}, \mathbf{P}_{f,k})|}\right)$$
(3.14)

The coordinate of the other end of the rupture, i.e. the point in the three-dimensional source-specific coordinate system that corresponds to the point  $(X'_r + R_L)$  on the *x*'-axis, is

$$\mathbf{P}_{r+R_{L}} = \sum_{k=1}^{N_{seg}} \mathbb{I} \Big[ x_{f,k-1}' \leq (X_{r}'+R_{L}) \leq x_{f,k}' \Big] \Big( \mathbf{P}_{f,k-1} + (X_{r}'+R_{L}-x_{f,k-1}') \\ * \frac{\mathcal{D} \big( \mathbf{P}_{f,k-1}, \mathbf{P}_{f,k} \big)}{|\mathcal{D} \big( \mathbf{P}_{f,k-1}, \mathbf{P}_{f,k} \big)|} \Big)$$
(3.15)

For each fault segment crossed by the rupture define  $\tilde{\mathbf{P}}_{k,1}$  and  $\tilde{\mathbf{P}}_{k,2}$  as the ends of the rupture within each segment. These quantities are illustrated in Figure 3.14. Notice, in this figure, that the rupture crosses segments 2, 3, and 4. Points  $\tilde{\mathbf{P}}_{k,1}$  and  $\tilde{\mathbf{P}}_{k,2}$  are associated with the ends of the portion of the rupture that lies on each segment  $k \in K_{rup}$ . For example, the left end of the rupture lies on segment 2 and propagates past the "right end" of the segment. Therefore,  $\tilde{\mathbf{P}}_{2,1} = \mathbf{P}_r$  and  $\tilde{\mathbf{P}}_{2,2} = \mathbf{P}_{f,2}$  as shown in Figure 3.14. In general,  $\tilde{\mathbf{P}}_{k,1}$  and  $\tilde{\mathbf{P}}_{k,2}$  are defined mathematically as a function of the rupture and fault coordinates on the x'-axis:

$$\widetilde{\mathbf{P}}_{k,1} = \mathbf{P}_{r} \text{ if } \left( x'_{f,k-1} \le X'_{r} \le x'_{f,k} \right) \cap \left( (X'_{r} + R_{L}) > x'_{f,k} \right) = \mathbf{P}_{f,k-1} \text{ if } \left( X'_{r} \le x'_{f,k-1} \right) \cap \left( x'_{f,k-1} \le (X'_{r} + R_{L}) \le x'_{f,k} \right)$$
(3.16)

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$$= \mathbf{P}_{f,k-1} \text{ if } \left( X'_r < x'_{f,k-1} \right) \cap \left( (X'_r + R_L) > x'_{f,k} \right)$$
$$= \mathbf{P}_r \text{ if } X'_r \ge x'_{f,k-1} \cap \left( (X'_r + R_L) \le x'_{f,k} \right)$$
$$= \emptyset \text{ if } \left( X'_r > x'_{f,k} \right) \cup \left( (X'_r + R_L) < x'_{f,k-1} \right)$$

$$\begin{aligned} \widetilde{\mathbf{P}}_{k,2} &= \mathbf{P}_{f,k} \text{ if } \left( x'_{f,k-1} \le X'_{r} \le x'_{f,k} \right) \cap \left( (X'_{r} + R_{L}) > x'_{f,k} \right) \\ &= \mathbf{P}_{r+R_{L}} \text{ if } \left( X'_{r} \le x'_{f,k-1} \right) \cap \left( x'_{f,k-1} \le (X'_{r} + R_{L}) \le x'_{f,k} \right) \\ &= \mathbf{P}_{f,k} \text{ if } \left( X'_{r} < x'_{f,k-1} \right) \cap \left( (X'_{r} + R_{L}) > x'_{f,k} \right) \\ &= \mathbf{P}_{r+R_{L}} + \text{ if } X'_{r} \ge x'_{f,k-1} \cap \left( (X'_{r} + R_{L}) \le x'_{f,k} \right) \\ &= \emptyset \text{ if } \left( X'_{r} > x'_{f,k} \right) \cup \left( (X'_{r} + R_{L}) < x'_{f,k-1} \right) \end{aligned}$$

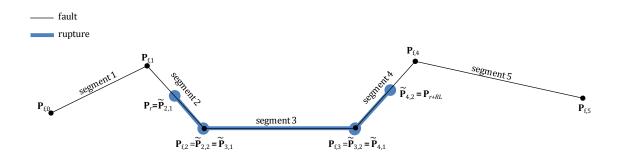


Figure 3.14: Illustration of parameters required for multi-segment finite rupture length formulation

The source-to-site distance,  $R_i$ , is determined by computing the shortest distance between site *i* and each fault segment crossed by the rupture and then selecting the minimum length amongst these values, i.e.

$$R_i = \min_{k \in K_{rup}} r_{i,k} \tag{3.17}$$

where  $r_{i,k}$  is the minimum distance between site *i* and the rupture on each segment  $k \in K_{rup}$ . This quantity is computed using different formulae depending on the location of the site relative to the ends of the rupture on the fault segment. The three possible configurations are shown in Figure 3.15. For configuration (a),  $r_{i,k}$  is equal to the distance between  $\mathbf{P}_{s,i}$  and  $\tilde{\mathbf{P}}_{k,1}$ ; for configuration (b),  $r_{i,k}$  is equal to the distance between  $\mathbf{P}_{s,i}$  and the point corresponding to the projection of  $\mathbf{P}_{s,i}$  onto the rupture segment; and for configuration (c),  $r_{i,k}$  is equal to the distance between  $\mathbf{P}_{s,i}$  and  $\tilde{\mathbf{P}}_{k,2}$ .

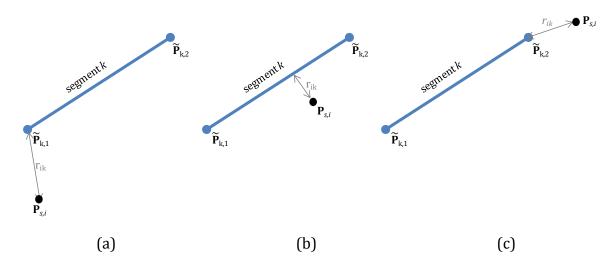


Figure 3.15: Possible configurations of a site location relative to a fault segment crossed by the rupture

The above description is written mathematically as

$$\begin{aligned} r_{ik} &= \left| \mathbf{P}_{s,i} - \widetilde{\mathbf{P}}_{k,1} \right| \text{ if } \vartheta_{i,k}^{j} \leq 0 \\ &= \left| \mathbf{P}_{s,i} - \left( \widetilde{\mathbf{P}}_{k,1} + \vartheta_{i,k} * \mathcal{D} \big( \widetilde{\mathbf{P}}_{k,1} , \widetilde{\mathbf{P}}_{k,2} \big) \big) \right| \text{ if } 0 < \vartheta_{i,k}^{j} < 1 \\ &= \left| \mathbf{P}_{s,i} - \widetilde{\mathbf{P}}_{k,2} \right| \text{ if } 1 \leq \vartheta_{i,k}^{j} \end{aligned}$$
(3.18)

where  $\vartheta_{i,k}$  is the ratio of the length of the projection of  $(\mathbf{P}_{s,i} - \widetilde{\mathbf{P}}_{k,1})$  on the vector  $\mathcal{D}(\widetilde{\mathbf{P}}_{k,1}, \widetilde{\mathbf{P}}_{k,2})$  and the length of  $\mathcal{D}(\widetilde{\mathbf{P}}_{k,1}, \widetilde{\mathbf{P}}_{k,2})$ :

$$\vartheta_{i,k}^{j} = \frac{\mathcal{D}(\widetilde{\mathbf{P}}_{k,1}, \widetilde{\mathbf{P}}_{k,2}) \cdot (\mathbf{P}_{s,i} - \widetilde{\mathbf{P}}_{k,1})}{\mathcal{D}(\widetilde{\mathbf{P}}_{k,1}, \widetilde{\mathbf{P}}_{k,2}) \cdot \mathcal{D}(\widetilde{\mathbf{P}}_{k,1}, \widetilde{\mathbf{P}}_{k,2})}$$
(3.19)

 $\vartheta_{i,k}$  will take on negative values for case (a) in Figure 3.15; values between 0 and 1 for case (b); and values greater than or equal to 1 for case (c). The above relations are used to populate the CPTs of nodes  $R_i$  in Figure 3.13.

#### 3.2.2.2 Ground motion intensity

The ground motion intensity at a site can be estimated using ground motion prediction equations. These equations express ground motion intensity at a site as a function of the median ground motion intensity (a function of magnitude, source-to-site distance, and site/source characteristics), as well as inter- and intra-event error terms. Nodes are added to the BNs in Figure 3.3 (point-source model) and Figure 3.13 (finite rupture model) that

represent the ground motion intensity at each site as well as the site-specific median ground motion and the inter- and intra-event deviations. The resulting BNs are shown in Figure 3.16 and Figure 3.17 for point-source and finite rupture formulations, respectively. These formulations are described below.

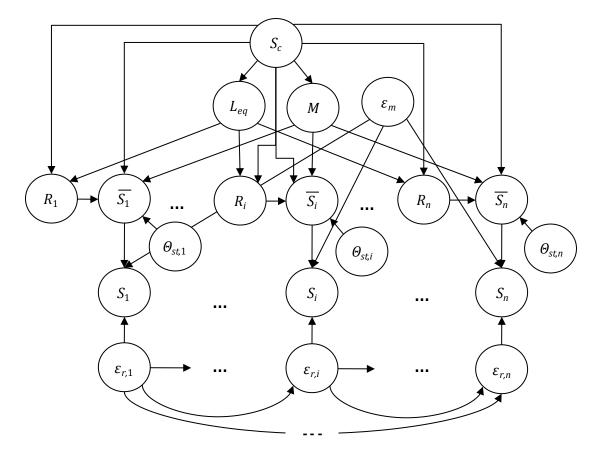


Figure 3.16: BN model of ground motion intensity at each site using a point-source assumption

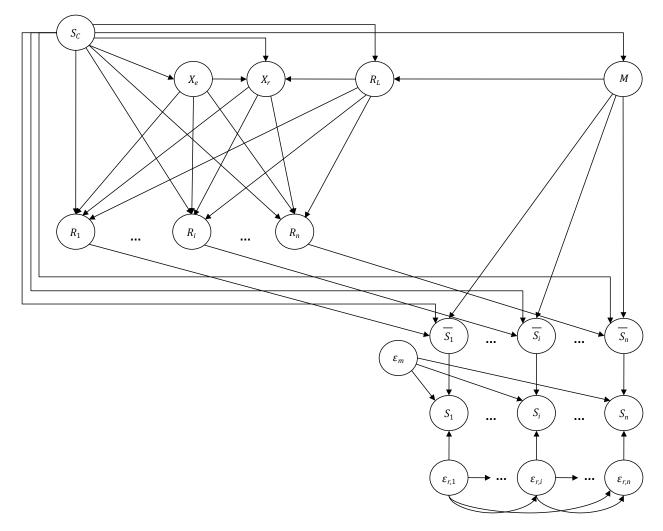


Figure 3.17: BN model of ground motion intensity at each site using a finite rupture model

In the BN in Figure 3.16, nodes labeled  $S_i$ , i = 1, ..., n, represent the ground motion intensity at *n* discrete sites in the region, assuming the earthquake is idealized as a point on the fault. For reference, the sites for which ground motion intensity nodes are included in the BN are referred to as *ground motion prediction points* (GMPPs). Nodes representing magnitude (*M*) and the location of the earthquake ( $L_{eq}$ ) are modeled as children of the source node ( $S_c$ ) in the manner described previously. The distance between the source and site *i* ( $R_i$ ) is expressed as function of the site coordinates (not shown in the BN because they are assumed deterministic), the source on which the earthquake occurs, and the location of the earthquake. The median ground motion intensity at site *i* is represented by node  $\bar{S}_i$ , which is a child of nodes M,  $R_i$ , and source and site characteristics. Nodes corresponding to site characteristics need not be explicitly shown if they are treated as deterministic or stochastic but unobservable; they are shown in Figure 3.16 for illustrative purposes. The actual ground motion intensity at a site,  $S_i$ , is modeled as a function of the median ground motion intensity as well as the inter- and intra-event error terms, consistent with (3.1). This dependence is shown by links from nodes  $\bar{S}_i$ ,  $\varepsilon_m$ , and  $\varepsilon_{r,i}$  to  $S_i$ .

The inter-event deviations ( $\varepsilon_m$ ) arise from inaccuracy or idealization in the characterization of the source and represent the earthquake to earthquake variability in actual ground motion intensity from the predicted median. For a given earthquake,  $\epsilon_m$  is the same for all sites. As a result, a single node  $\varepsilon_m$  is introduced that is a parent of all sitespecific ground motion intensity nodes. The intra-event deviations ( $\varepsilon_{r,i}$ ) arise from uncertain wave propagation and site effects. For a particular earthquake, intra-event deviations represent the variability in the ground motion intensity from site to site. The intra-event errors are site-specific and represent random variables drawn at discrete points in the domain of a spatially-correlated Gaussian random field. The correlation arises because, for a given earthquake, the magnitude of the difference between the actual and predicted median ground motion intensity is similar for sites in close proximity to one another but can be less similar as the distance between the sites increases. A node representing  $\varepsilon_{r,i}$  is added to the BN for each site. The correlation among  $\varepsilon_{r,i}$  is modeled with links between all pairs of the corresponding nodes, as shown in Figure 3.16. BNs with densely connected nodes, as arises when modeling points drawn from a random field, are highly demanding of computer memory. Methods for efficiently modeling random variables drawn from a Gaussian random field via BN are the subject of Chapter 4.

The seismic demand BN corresponding to a finite rupture model is shown in Figure 3.17. The conditional relations used to construct this BN are the same as in the point-source case, except that nodes are added to account for the effect of the finite rupture model on source-to-site distance consistent with the development in the previous section.

# 3.2.2.3 Rupture directivity effects

The intensity and characteristics of ground motions at sites located in close proximity to a fault may differ substantially from those of sites located at farther distances due to rupture directivity effects. Rupture directivity effects cause variation in ground motion amplitude and duration at a site based on whether the direction of rupture propagation is toward (forward directivity) or away (backward directivity) from the site.

Forward directivity effects are generated when the velocity of the rupture front is only slightly less than the shear wave velocity. As the rupture propagates from the epi/hypocenter toward the site, a shear wave front is generated as waves accumulate in front of the rupture. The waves generated between the epi/hypocenter and the site will arrive at the site at approximately the same time and, at long periods, the waves will constructively interfere. The result is a high amplitude, long period and short duration (i.e. pulse-like) ground motion. Evidence of forward directivity effects are most easily observed in the velocity time history.

Conversely, a site may experience backward directivity when located in the direction opposite to the direction of propagation of the rupture. Such a situation arises when the

rupture propagates in one direction from the epi/hypocenter. Ground motions associated with backward directivity effects are longer duration and lower amplitude. Figure 3.18 provides a graphical representation of directivity effects (adapted from Somerville et al. (1997)) and illustrates how the characteristics of the motion can differ in the forward and backward directivity regions.

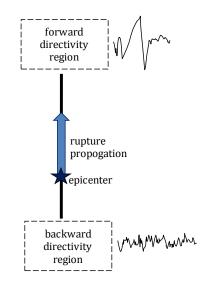


Figure 3.18: Illustration of directivity effects

The geometry of a fault relative to the location of a site is the best predictor of the potential to experience a pulse-like ground motion as a result of directivity. The angle between the direction of rupture and the ray path from the epi/hypocenter to the site and the amount of the rupture that lies between the epi/hypocenter and the site are good predictors of directivity potential. However, directivity is not a deterministic phenomenon. Directivity effects may not be felt at a site for which the geometry suggests it is likely. Conversely, pulse-like motions have been measured at sites when the geometric configuration is not favorable to directivity (Shahi and Baker 2010).

Large pulse-like ground motions can result from ruptures on both strike-slip and dip-slip faults. For strike-slip faults, directivity effects are normally experienced by sites located near to the fault in the direction of propagation of the rupture. Such sites experience largest effects on the strike-normal component of ground motion. For dip slip faults, sites located on the projection of the rupture plane in the up-dip direction will experience pulse-like motions (Bray and Rodriguez-Marek 2004).

The pulse-like motions that result from forward directivity place large demands on structures and earth systems. Consequently, these motions cause damage that is more severe than that experienced by structures not subject to the same type of loading. As a result, in recent years, emphasis has been placed on the importance of including the directivity effect when predicting intensity of ground shaking at a site.

While it is crucial to include forward directivity effects for long period structures that are near highly active faults, available ground motion prediction equations (e.g. the NGA relations) do not explicitly account for directivity effects. However records with directivity effects have been included in the databases used to derive these empirical relations. Such effects are observed by Kuehn et al. (2009) as described in section 3.2.1. As a result, ground motion prediction equations tend to under-predict ground motion intensity in the near-fault region when pulses occur and over predict when pulses do not occur. To correct these relations, adjustment factors have been proposed (Somerville et al. 1997; Abrahamson 2000; Tothong et al. 2007; Shahi and Baker 2010). Abrahamson (2000) and Shahi & Baker (2010) provide factors to modify the spectral acceleration obtained from conventional ground motion prediction equations to account for the directivity effect. Bray & Rodriguez-Marek (2004) offer a modified attenuation relationship for predicting the peak ground velocity at a site when the directivity pulse is known to have occurred. BN formulations for the directivity models contained in Abrahamson (2000), Shahi & Baker (2010), and Bray & Rodriguez-Marek (2004) are developed in the following subsections.

#### 3.2.2.3.1 Directivity BN based on Abrahamson (2000)

Abrahamson (2000), adapting an earlier relation by Somerville et al. (1997), accounts for the directivity effect by defining a factor that amplifies or de-amplifies the ground motion intensity at a site that is in close proximity to a fault rupture. The model is a broad-band formulation, which monotonically increases/decreases spectral acceleration over a range of periods (Shahi and Baker 2010). The site-specific directivity amplification factor is added to the natural logarithm of the spectral acceleration in the ground motion prediction equation. Using Abrahamson (2000), directivity effects are included in the BN through the addition of a node,  $f_{d,i}$ , representing the amplification factor on ground motion intensity at site *i*.

 $f_{d,i}$  is defined as a function of regression coefficients, the fraction of the rupture length that propagates toward the site, and the angle between the fault and the ray-path from the earthquake epicenter/hypocenter to the site. For the average horizontal component of ground motion, the Somerville et al. (1997) factor (later modified by Abrahamson (2000)) is given by

$$f_{d,i}(F_{F,mod}, \theta, T) = c_1(T) + 1.88c_2(T)F_{F,mod}\cos\theta$$
(3.20)

where  $\theta$  is the angle between the fault strike and the ray path to the site (the azimuth angle for strike-slip faults; the zenith angle for dip-slip faults),  $c_i(T)$ , i = 1,2, are regression coefficients expressed as functions of the fundamental period *T*, and

$$F_{F,mod} = \begin{cases} F_F \text{ for } F_F \le 0.4\\ 0.4 \text{ for } F_F > 0.4 \end{cases}$$
(3.21)

where  $F_F = D_i/R_L$  is the fraction of the rupture that propagates towards the site.  $D_i$  is the length of the rupture that falls between the epi/hypocenter and the site. The parameters of the directivity model are shown in Figure 3.19 for (a) a strike-slip fault and (b) a dip-slip fault. In these illustrations,  $R_L$  is the rupture length,  $D_i$  is the length of the rupture propagating towards the site and, for a strike-slip fault,  $\theta$  is the azimuth angle between the fault rupture and the ray path to the site. For a dip-slip fault,  $\theta$  is the zenith angle between the fault rupture and the ray path to the site.

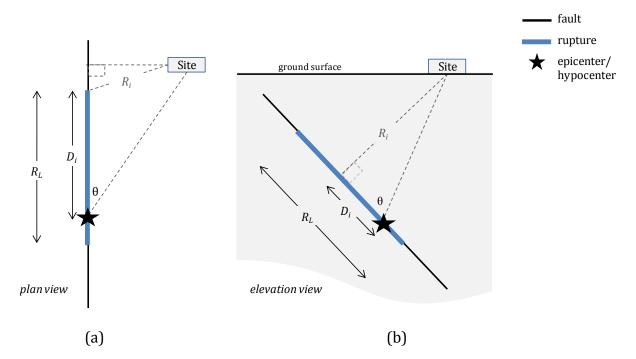


Figure 3.19: Directivity parameters for (a) strike-slip fault; and (b) dip-slip fault

Abrahamson (2000) added two modifications to the Somerville et al. (1997) model in (3.20). The first modification is added to ensure that the directivity effect reduces to zero for large source-to-site distances (i.e. greater than 60 km) and is given by

$$T_{r}(R_{i}) = 1 \quad \text{for } R_{i} \leq 30km$$
  
=  $1 - \frac{R_{i} - 30}{30} \quad \text{for } 30km < R_{i} < 60km$  (3.22)  
=  $0 \quad \text{for } R_{i} \geq 60km$ 

The second modification factor reduces directivity to zero for magnitudes less than 6.0. It takes the form

$$T_m(M) = 1 \text{ for } M \ge 6.5$$
 (3.23)

$$= 1 - \frac{m - 6.5}{0.5} \text{ for } 6.0 < M < 6.5$$
$$= 0 \text{ for } M \le 6.0$$

Two geometric expressions are necessary for computing the directivity amplification factor  $f_{d,i}(F_{F,mod}, \theta, T)$ : the fraction of the rupture propagating toward the site and the angle between the strike of the rupture and the ray path to the site. Four geometric cases (combinations of site and rupture locations) are considered for deriving the expressions needed for constructing the BN that includes the directivity effect:

- Case 1: the site is located to the right of the entire rupture:  $x_{s,i} > (X_r + R_L)$
- Case 2: the site is located to the right of the epicenter and to the left of the right end of the rupture:  $X_e < x_{s,i} < (X_r + R_L)$
- Case 3: the site is located to the left of the entire rupture:  $x_{s,i} < X_r$
- Case 4: the site is located to the left of the epicenter and to the right of the left end of the rupture:  $X_r < x_{s,i} < X_{epi}$

These cases are demonstrated in Figure 3.20 for a strike-skip fault. The derivations included here are based on the geometry of strike-slip faults, but similar expressions can be derived for dip-slip faults.

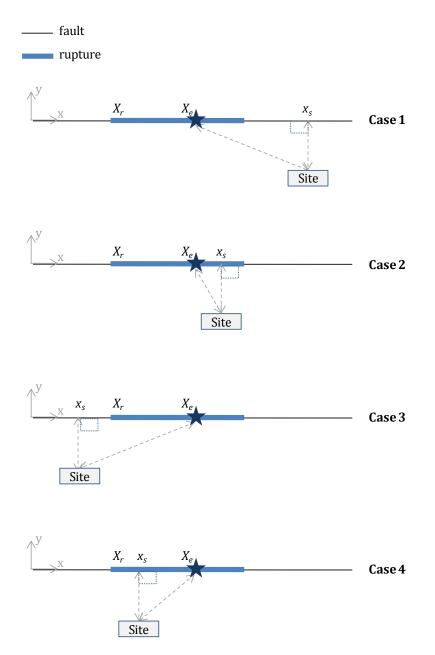


Figure 3.20: Geometric cases considered for computing the directivity amplification factor

For all cases in Figure 3.20, the cosine term in (3.20) is defined as

$$\cos(\theta) = \frac{|X_e - x_{s,i}|}{\sqrt{y_{s,i}^2 + z_{s,i}^2 + (X_e - x_{s,i})^2}}$$
(3.24)

The fraction of the rupture that propagates toward the site,  $F_F$ , is defined for each of the geometric cases as follows:

- Case 1:  $F_F = \frac{A_R}{R_L} = \frac{(X_r + R_L) X_e}{R_L}$
- Case 2:  $F_F = \frac{x_{s,i} X_e}{R_L}$

• Case 3: 
$$F_F = \frac{X_e - X_P}{R_L}$$

• Case 4: 
$$F_F = \frac{X_e - x_{s,i}}{R_L}$$

Cases 1 and 2 correspond to the event that  $X_e < x_{s,i}$  and cases 3 and 4 correspond to the event that  $X_e > x_{s,i}$ . Using this observation  $F_F$  is defined as

$$F_{F} = \frac{\min[(X_{r} + R_{L}) - X_{e}, x_{s,i} - X_{e}]}{R_{L}} \quad \text{if } X_{e} < x_{s,i}$$

$$= \frac{\min[X_{e} - X_{r}, X_{e} - x_{s,i}]}{R_{L}} \quad \text{if } X_{e} \ge x_{s,i}$$
(3.25)

Thus, the spectral acceleration at site *i*, when accounting for directivity effects, is computed as

$$\ln(S_{a,i}) = \ln(\overline{S_{a,i}}) + f_{d,i} * T_m(M) * T_r(R_i) + \varepsilon_m + \varepsilon_{r,i}$$
(3.26)

The BN incorporating the directivity effect is shown in Figure 3.21, where node  $S_i$  is taken to represent the spectral acceleration. Node  $f_{d,i}$  is a child of nodes defining the site and source geometry. Due to the modifications introduced by Abrahamson (2000),  $f_{d,i}$  is affected by the earthquake magnitude and the source-to-site distance due to the factors  $T_r(R_i)$  and  $T_m(M)$ . Thus, it is a child of nodes M and  $R_i$  as well. The CPT of node  $S_i$  is computed using (3.26) and Monte Carlo simulation.

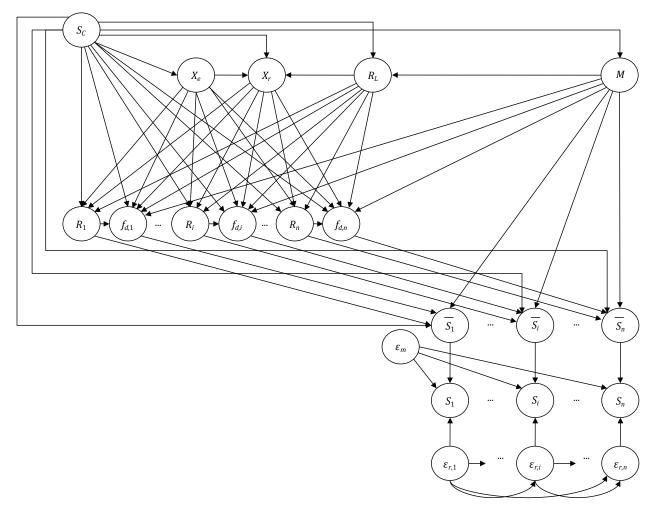


Figure 3.21: BN modeling of ground motion intensity at multiple sites accounting for directivity effects according to Abrahamson (2000) model

It should be noted that the formulations derived by Somerville et al. (1997) and Abrahamson (2000) use geometric quantities that assume an idealization of a fault as a single line segment. It is not clear how to extrapolate quantities such as  $\cos \theta$  in the context of faults with bends. Therefore, it is assumed that the above directivity model and BN formulation are only applicable to faults that can be reasonably idealized as straight lines. For multi-segment faults, the Shahi & Baker (2010) model, described next, should be used.

### 3.2.2.3.2 Directivity BN based on Shahi & Baker (2010)

Shahi & Baker (2010) developed an extension to the framework in Tothong et al. (2007) for including the directivity effect in seismic hazard analysis. Unlike Somerville et al. (1997), Shahi & Baker (2010) develop a narrow-band model, which only amplifies intensities in the range of periods near the predominant period of the directivity pulse. Using a database of

179 pulse-like ground motions, empirical models are developed to estimate: (1) the probability that a pulse-like ground motion will occur at a site; (2) the orientation of the expected pulse; (3) the pulse period; and (4) the amplification (or de-amplification) of the spectral acceleration resulting from the occurrence (or non-occurrence) of a pulse-like motion.

Shahi and Baker use logistic regression to develop a model for the probability of observing a pulse-like ground motion as a function of parameters representing site/source geometry (particularly  $D_i$ ,  $R_i$  and, in the case of dip-slip faults,  $\theta$ ). The resulting model for strike-slip faults is

Pr(pulse at site 
$$i|R_i, D_i$$
) =  $\frac{1}{(1 + \exp(0.642 + 0.167R_i - 0.075D_i))}$  (3.27)

and for dip-slip faults is

Pr(pulse at site 
$$i|R_i, D_i, \theta$$
) =  $\frac{1}{(1 + \exp(0.128 + 0.055R_i - 0.061D_i + 0.036\theta))}$  (3.28)

The amplification of spectral acceleration depends on the predominant period,  $T_p$ , of the pulse. Empirical results suggest that the pulse period is dependent on the earthquake magnitude. Shahi and Baker find that  $\ln T_p$  is normally distributed with mean  $\mu_{\ln T_p} = -5.73 + 0.99M$  and standard deviation  $\sigma_{\ln T_p} = 0.56$ .

Next, the spectral acceleration predicted by a ground motion prediction equation (e.g. a NGA model) is modified based on whether or not a pulse is observed. If a pulse is observed, then the spectral acceleration at site  $i (S_{a,i}^{pulse})$  is computed as

$$\ln S_{a,i}^{pulse} = \ln A_f + \ln S_{a,i}^{GMPE}$$
(3.29)

where  $S_{a,i}^{GMPE}$  is the spectral acceleration predicted by a ground motion prediction equation and  $A_f$  is the amplification factor due to the occurrence of a pulse. The ground motion intensity at site *i* accounting for amplification due to directivity ( $\ln S_{a,i}^{pulse}$ ) has the normal distribution with mean  $\mu_{\ln A_f} + \mu_{\ln S_{a,i}^{GMPE}}$  and standard deviation  $\sigma_{\ln S_{a,i}^{pulse}} = R_f \sigma_{\ln S_{a,i}^{GMPE}}$ . The mathematical form of  $\mu_{\ln A_f}$  is given in Shahi and Baker (2010) and is a function of  $T_p$ and the period of interest, *T*.  $\sigma_{\ln S_{a,i}^{GMPE}}$  is the standard deviation of the ground motion prediction equation and  $R_f$  is a reduction factor expressed as a function of *T* and  $T_p$ . The reduction factor accounts for the additional refinement resulting from inclusion of a directivity factor by decreasing the uncertainty associated with the ground motion prediction equation. Shahi and Baker (2010) provide a similar expression for the de-amplification of spectral acceleration in the absence of directivity:

$$\ln S_{a,i}^{no \ pulse} = \ln D_f + \ln S_{a,i}^{GMPE}$$
(3.30)

where  $S_{a,i}^{no\ pulse}$  is the spectral acceleration at the site when a directivity pulse is not observed and  $D_f$  is a de-amplification factor.  $\ln S_{a,i}^{no\ pulse}$  is normally distributed with mean  $\mu_{\ln D_f} + \mu_{\ln S_{a,i}^{GMPE}}$  and standard deviation  $\sigma_{\ln S_{a,i}^{GMPE}}$ . It is found that the standard deviation need not be altered in the absence of directivity.  $\mu_{\ln D_f}$  is given in Shahi and Baker (2010) as a function of the source-to-site distance and *T*.

The BN seismic demand model for spectral acceleration accounting for the directivity effect according to the Shahi and Baker (2010) model is shown in Figure 3.22. This figure uses two common representation conventions used in BNs. A rectangle with rounded corners is an object behind which a BN is hidden to reduce graphical clutter. The object in Figure 3.22 is hiding the random field model to be described in Chapter 4. The shaded rectangle with dotted borders represents a plate. Nodes within a plate repeat for each GMPP in the system and outside nodes are common to all plates.

In the BN in Figure 3.22, a two-state node  $P_i$  is introduced as a child of nodes  $X_r$ ,  $X_e$ ,  $R_L$  and  $S_c$ . Node  $P_i$  represents the existence of a pulse-like motion at site *i*. As described above, the Shahi & Baker (2010) model applies adjustment factors to the median and variance of spectral acceleration to account for the presence/absence of the directivity pulse. These factors are modeled in the BN by defining the median spectral acceleration a child of node  $P_i$  and by introducing a variance reduction factor node  $R_{f,i}$  as a child of  $P_i$  and a parent of the total error term associated with the spectral acceleration. The CPTs of nodes  $\bar{S}_i$  are defined conditioned on  $P_i$  using (3.29) and (3.30) as well as the conventional attenuation relationship.

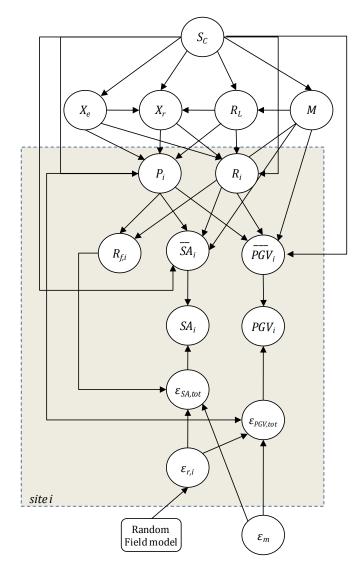


Figure 3.22: BN model of ground motion intensity at multiple sites accounting for directivity effects using the model by Shahi & Baker (2010)

### 3.2.2.3.3 Directivity BN based on Bray & Rodriguez-Marek (2004)

For peak ground velocity, Bray & Rodriguez-Marek (2004) provide alternate prediction equations that are used when directivity effects are present, rather than modifying existing ground motion prediction equations. They do not provide prediction equations in the absence of directivity nor provide estimates of the probability of experiencing directivity effects at a site for a particular earthquake. As a result, in the absence of directivity, the conventional NGA prediction equations for PGV must be employed. In the present study, we use the Bray & Rodriguez (2004) model for predicting PGV when the directivity pulse is present. We use the conventional NGA equation for PGV when the pulse is absent, though

the latter is somewhat conservative because of the "averaging" effect described earlier. The nodes required for the Bray & Rodriguez-Marek (2004) model are included in Figure 3.24 in conjunction with the pulse prediction models of Shahi & Baker (2010).

# 3.3 Hazard due to liquefaction

Liquefaction describes a phenomenon in which soil transitions from a solid to a liquid state (with consistency of a heavy liquid) as a result of rapid loading. Liquefaction is associated with a variety of situations in which deformations of saturated cohesionless soils (i.e. loose to moderately granular soil with poor drainage, such as silty sands and sands) are caused by transient, monotonic, and repeated disturbances. A common source of liquefaction-inducing loads is earthquakes. Ground shaking increases pore water pressure resulting in reduced effective stress and consequently decreased shear strength (Kramer 1996).

There are two types of liquefaction: (1) flow liquefaction; and (2) cyclic mobility. Flow liquefaction is characterized by flow failures caused by a static shear stress that is greater than the shear strength of the soil. Cyclic stresses induced by ground motion bring the soil to an unstable condition at which the strength drops sufficiently to allow flow failure. Flow liquefaction occurs suddenly and the liquefied material can travel long distances. Flow liquefaction occurs less frequently than cyclic mobility, but results in more severe consequences (Kramer 1996).

Cyclic mobility, driven by both cyclic and static stresses, occurs when the static stress is *less* than the shear strength of the soil. Unlike flow liquefaction, cyclic mobility develops incrementally throughout the duration of ground shaking and can occur even after shaking has stopped. Cyclic-mobility induced deformation is called *lateral spreading* and can occur on gently sloping and flat ground. Cyclic mobility can produce surface effects known as sand boils (Kramer 1996).

Liquefaction can cause significant damage to infrastructure systems in several ways (Rauch 1997): (1) foundations that are supported on liquefied soils may experience sudden loss of load-bearing capacity resulting in irregular settlement of the supported structure; (2) differential settlement of pipelines may result from sudden loss of capacity of the supporting soil at different points along the pipe; (3) buried components (e.g. pipelines) situated in liquefied soil may float to the surface causing breakage; and (4) slope failures may cause damage to infrastructure components or structures located on the incline.

The assessment of liquefaction potential and consequences involves the assessment of: (1) the likelihood of initiation of liquefaction at a site; (2) the post-liquefaction strength and stability of the soil at the site; and (3) the deformations and displacements resulting from liquefaction (Seed et al. 2003). There is a great deal of uncertainty associated with each of these three steps; however, only recently have probabilistic methods been applied to the assessment of liquefaction potential and consequences. As described above, certain soil types are more susceptible to liquefaction and significant research has strived to specify

vulnerable soil classes. The reader is referred to Seed et al. (2003) and Cetin et al. (2004) for additional information.

If potentially liquefiable soils are found at a site, quantitative assessment indicates the likelihood of liquefaction initiation. Research in this area is based on both laboratory and, more commonly, empirical relations derived from in situ observations. These empirical relations are typically expressed as a function of measurements from common tools such as the Standard Penetration Test (SPT), the Cone Penetration Test (CPT; this acronym should not be confused with the use of CPT to abbreviate conditional probability table), measurement of shear wave velocity, and the Becker Penetration Test (BPT).

Less mature is the assessment of potential consequences when liquefaction does occur. Seed et al. (2003) note that large deformations (> 1m) can typically be estimated within a factor of 2, but such accuracy is not encountered when predicting small to moderate displacements. Small to moderate displacements result from more complicated phenomena and are of more significant engineering interest because sites prone to large deformations are typically mitigated, while sites prone to smaller amounts may not be mitigated. The reader is referred to Seed et al. (2003) for more details related to the phenomenon of liquefaction.

# 3.3.1 Existing Bayesian network-based liquefaction models

Bayraktarli (2006) developed a BN for liquefaction assessment that is shown in Figure 3.23. Descriptions of the earthquake magnitude, source-to-site distance and damage nodes are contained in section 3.2.1 in connection with Figure 3.23. Node Soil Type takes on states rock, gravel, sand, silt, and clay. The spectral acceleration is computed as a function of the magnitude, source-to-site distance, soil type, as well as soil response (to account for the effect of liquefaction on ground motion experienced at the site). The node representing soil profile corresponds to the different layers of a bore profile at a site. Peak Ground Acceleration is defined as a function of nodes that represent earthquake magnitude, earthquake distance, soil type, and soil profile. The node representing liquefaction susceptibility is logically related to nodes corresponding to fines content and liquid limit using the Modified Chinese Criteria (Liam Finn et al. 1994). The probabilities required to populate the CPT for node *Liquefaction* are obtained via simulation using the limit state function:  $g(x) = CRR_{7.5} * MSF * K_{\sigma} - CSR = 0$ , where  $CRR_{7.5}$  is a function of SPT blow count  $((N_1)_{60})$ , *MSF* is a function of earthquake magnitude,  $K_{\sigma}$  is a correction factor, and *CSR* is a function of  $a_{max}$  (maximum acceleration in a soil layer, computed as a function of the PGA on rock using one-dimensional wave propagation analysis), effective stresses, and a depth reduction coefficient. Node Soil Response has two states: amplification and liquefaction.

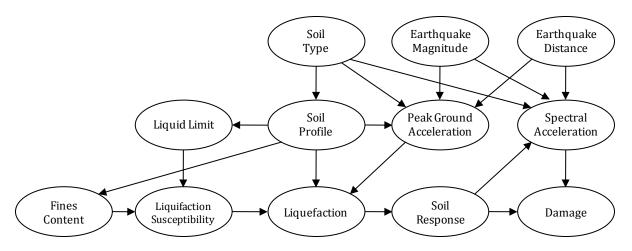


Figure 3.23: Liquefaction BN developed by Bayraktarli (2006)

## 3.3.2 Proposed formulation

We next present a preliminary BN for modeling seismic demands placed on the components of an infrastructure system due to liquefaction. Development of a more sophisticated and refined model is left as an area warranting further study. The proposed BN formulation for modeling liquefaction hazard developed in this section differs from the formulation in Bayraktarli (2006) in several ways. First, the liquefaction BN developed here assumes that all variables related to site characteristics are either deterministic or are stochastic but unobservable. For this reason, they are not explicitly modeled as nodes in the BN, as is the case in Bayraktarli (2006). Though not explicitly included, they influence the CPTs of nodes that appear in the BN. Nodes corresponding to these variables can be added, but at the expense of additional computational demand and increased discretization error. Second, differing from the model in Bayraktarli (2006), all layers are considered simultaneously rather than layer-by-layer. Furthermore, it is assumed that the output quantities of interest are the seismic demands placed on structural and geotechnical components of the infrastructure system and therefore nodes representing total seismic settlement are explicitly included. Finally, while Bayraktarli (2006) use Monte Carlo simulation to estimate the probability of liquefaction using a defined limit state function, we utilize published empirical relationships to estimate this probability as a function of ground motion intensity estimated using the previously described ground shaking models. This permits inclusion of spatial correlation in estimating likelihood of liquefaction at sites in close proximity.

The preliminary BN that models liquefaction-induced demands and the associated seismic settlements is presented in Figure 3.24. The construction of CPTs required by this formulation is described next. The soil profile at the site must be discretized into layers corresponding to changes in soil characteristics, e.g. soil class (silt, clay, sand, gravel, etc.), SPT blow count, fines content (*FC*), plasticity index (*PI*), depth to water table, unit weight ( $\gamma$ ), and relative density ( $D_R$ ). A layer of soil will experience liquefaction if the cyclic stress

ratio, CSR, (the demand) placed on it exceeds, the cyclic resistance ratio, CRR. CRR is computed using empirical relations as a function of SPT, CPT, BPT, and/or shear wave velocity measurements. The CSR for a soil layer is expressed as a function of the peak surface acceleration  $(a_{max})$ , the total and effective stresses in the layer (expressed as a function of unit weight and layer thickness), and the depth reduction coefficient  $r_d$ (available from published correlations, typically as a function of depth).  $a_{max}$  can be computed as a function of base rock acceleration using a variety of methods, including simplified procedures (Seed and Bray 1997), code-based methods (International Code Council 2006), and 1-dimensional site response wave propagation analysis (Ordonez 2009). Thus, in Figure 3.24,  $a_{max,i}$  is a child of  $PGA_{rock,i}$ . Cetin et al. (2002) and Moss et al. (2006) provide expressions for the probability that a layer will liquefy as a function of the CSR, earthquake magnitude, and layer properties characterizing the liquefaction resistance. These expressions are used to populate the CPTs associated with the binary (liquefaction/no liquefaction) nodes  $Liq_{i,k}$ , k = 1, ..., N where N is the number of discrete layers at site *i* that are susceptible to liquefaction, e.g. as defined by the modified Chinese criteria (Liam Finn et al. 1994). The expression for probability of liquefaction in Cetin et al. (2002) is applicable when the resistance is measured in terms of  $(N_1)_{60}$  (i.e. it uses a SPTbased correlation), while the expression in Moss et al. (2006) uses a CPT-based expression. For each state of *PGA*<sub>rock,i</sub>, whose value comes from the ground-shaking seismic demand model, the CPT for the node  $a_{max,i}$  is computed using one of the aforementioned procedures.

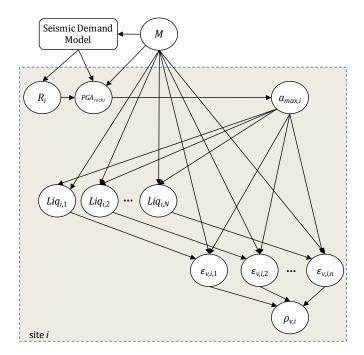


Figure 3.24: BN modeling seismic demands resulting from liquefaction

As described above, the goal of the liquefaction BN in Figure 3.24 is to model the effects of liquefaction, e.g. soil displacement and settlement. If the  $k^{\text{th}}$  layer liquefies, the induced

volumetric strain  $(\epsilon_{v,i,k})$  is computed as a function of layer properties using existing relations (Tokimatsu and Seed 1987; Ishihara and Yoshimine 1992). The total settlement at site *i* is then computed as  $\rho_{v,i} = \sum_{k=1}^{N} h_k * \epsilon_{v,i,k}$ , where  $h_k$  is the height of layer *k*. These relations are used to produce the CPTs for nodes  $\epsilon_{v,i,k}$  and  $\rho_{v,i}$  in Figure 3.24. The converging structure linking  $\rho_{v,i}$  with its parent nodes is computationally inefficient. The converging structure is replaced with a chain like structure as shown in Figure 3.25. The CPT associated with node  $\rho_{v,i,1}$ is based on the relation  $\rho_{v,i,1} = \mathbb{I}[\text{liquefaction occurs in layer 1}] * h_1 * \epsilon_{v,i,1}$ , where  $\epsilon_{v,i,1}$  is computed for layer 1 using relations available in the aforementioned references as a function of earthquake magnitude, *CSR*, and layer properties. For all other layers (k = 2, ..., N)

$$\rho_{\nu,i,k} = \mathbb{I}[\text{liquefaction occurs in layer } k] * h_k * \epsilon_{\nu,i,k} + \rho_{\nu,i,k-1}$$
(3.31)

Setting  $\rho_{v,i} = \rho_{v,i,N}$  gives the total liquefaction-induced seismic settlement at the site. The liquefaction model shown in Figure 3.25 should only be added to the BN for sites vulnerable to liquefaction; adding the model to other sites results in unnecessary increases in computational demands. Furthermore in Figure 3.25, it may be preferable to reduce computational demands by eliminating nodes  $Liq_{i,j}$  and  $\rho_{v,i,j}$  and computing the CPT  $\rho_{v,i}$  directly as a function of  $a_{max,i}$ .

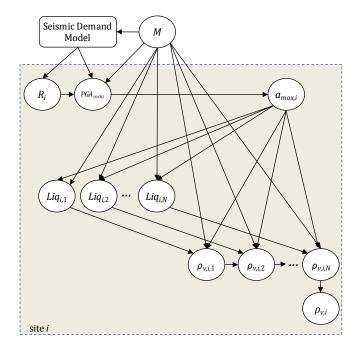


Figure 3.25: Liquefaction modeling seismic demands resulting from liquefaction with converging structure in Figure 3.24 at node  $\rho_{v,i}$  replaced by a chain structure

Similar models yield estimates of lateral spreading of sloping ground and ground with a vertical discontinuity (i.e. a free-face) as a function of site and earthquake characteristics

(see Youd et al. 2002). Finite element models have also been used to assess ground stability, see Seed et al. (2003) for a summary.

# 3.4 Hazard due to fault rupture

Hazard due to fault rupture is an important vulnerability of distributed infrastructure systems, which may have long components (e.g. pipelines, bridges, embankments) that cross a fault. A generic and preliminary BN model framework is described here to address this hazard. Development of a refined model is left as future work. It is assumed that fault rupture effects need only be included when a distributed component crosses a known fault that has potential to cause permanent differential displacement of engineering significance at the surface. The effects of incoherence of ground motions on either side of the fault are not considered, though such effects may be critical for certain distributed components. Development of models to consider such effects is left as an area warranting future study.

For sites crossing faults with potential to produce permanent surface displacement, several nodes are added to the BN to account for fault rupture effects. First, a binary node is added indicating whether the rupture passes the point at which the component crosses the fault. This node, labeled  $RC_i$ , is a child of nodes representing the geometry of the rupture, i.e.  $X_r, R_L, S_c$ . Additionally, a node  $SS_i$  is included to represent the maximum amount of surface slip associated with the fault rupture, modeled a function of earthquake magnitude and source characteristics. The performance of the distributed component crossing the fault is computed as a function of the induced ground displacement and the deformation capacity of the component. The associated BN is shown in Figure 3.26. For the sake of clarity, nodes not applicable to the fault-rupture model are not shown in this figure.

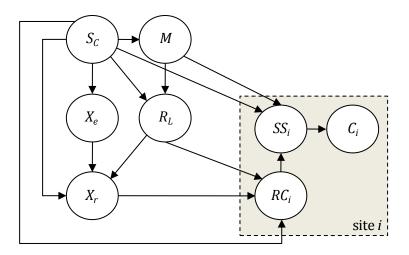


Figure 3.26: Preliminary fault rupture BN

## 3.5 Summary

This chapter presents BN-based formulations for modeling ground motion intensity and ground deformation at multiple sites distributed across a geographic region in the vicinity of one or more faults. Formulations in which an earthquake is idealized as both a point-source and a finite rupture are presented. The chapter focuses primarily on the hazard due to ground shaking, but preliminary models for liquefaction and fault rupture are also developed.

### Appendix 3.1: Earth-Centered Earth-Fixed coordinate system

The seismic demand model requires the definition of fault-specific local coordinate systems. Typically, information relating to coordinates of sites and faults are specified using latitude, longitude, and altitude. To facilitate the conversion to a local rectangular (Cartesian) coordinate system, the latitude/longitude/altitude geographic coordinates of a site or fault are first converted to Earth-center Earth-fixed (ECEF) coordinates (Clynch 2006).

The ECEF coordinate system is a rectangular coordinate system with its origin at the mass center of the Earth (0,0,0). In this system, the *z*-axis goes through the true North Pole. The *x*-axis extends through the intersection of the Prime Meridian and the Equator. The *y*-axis, consistent with a "right-hand-rule" coordinate system, passes through the intersection of the Equator and the 90° longitude meridian. Using assumptions about the shape of the Earth based on WGS-84 (DoD 2000), which approximates the Earth as an ellipsoid revolving about its minor axis and provides a standard coordinate frame for the Earth, latitude/longitude/altitude data are converted to ECEF data as described in Farrell & Barth (1999).

In this study, the coordinates of all sites are defined with respect to a source-specific local coordinate system. To do this, the ECEF coordinate system must be shifted from the origin at the center of the Earth to one centered at one end of the fault. Then, the coordinate system must be rotated so that the first segment of the fault coincides with the x-axis. This is accomplished by rotating once about the *z*-axis and once about the *y*-axis. Thus, the coordinate transformation consists of three steps:

1. Shifting of the origin from the center of the Earth to the end of the fault, which is assumed to have at ECEF coordinates  $(x_{f,0}, y_{f,0}, z_{f,0})$ :

$$\begin{aligned} x' &= x - x_{f,0} \\ y' &= y - y_{f,0} \\ z' &= z - z_{f,0} \end{aligned}$$
 (3.32)

2. Rotation about the *z*-axis:

$$x'' = x' \cos(\phi_1) + y' \sin(\phi_1) 
 y'' = -x' \sin(\phi_1) + y' \cos(\phi_1) 
 z'' = z'$$
(3.33)

where:

$$\begin{split} \phi_{1} &= \tan^{-1} \left( \frac{y_{f,1}'}{x_{f,1}'} \right) \quad \text{if } x_{f,1}' > 0 \\ &= \tan^{-1} \left( \frac{y_{f,1}'}{x_{f,1}'} \right) + \pi \quad \text{if } x_{f,1}' < 0 \text{ and } y_{f,1}' \ge 0 \\ &= \tan^{-1} \left( \frac{y_{f,1}'}{x_{f,1}'} \right) - \pi \quad \text{if } x_{f,1}' < 0 \text{ and } y_{f,1}' < 0 \\ &= \frac{\pi}{2} \quad \text{if } x_{f,1}' = 0 \text{ and } y_{f,1}' > 0 \\ &= -\frac{\pi}{2} \quad \text{if } x_{f,1}' = 0 \text{ and } y_{f,1}' < 0 \\ &= 0 \quad \text{if } x_{f,1}' = 0 \text{ and } y_{f,1}' = 0 \end{split}$$

$$(3.34)$$

3. Rotation about the *y*-axis:

$$x''' = x'' \cos(\phi_2) + z'' \sin(\phi_2) y''' = y'' z''' = -x'' \sin(\phi_2) + z'' \cos(\phi_2)$$

where  $\phi_2$  is defined in a manner similar to (3.34) but using  $x'_{f,1}$  and  $z'_{f,1}$  instead of  $x'_{f,1}$  and  $y'_{f,1}$ , respectively.

# Chapter 4: Bayesian Network Modeling of Gaussian Random Fields

### 4.1 Introduction

In civil engineering applications, it is often necessary to model vectors of random variables drawn from a random field. In the current application, the earthquake-induced ground motion intensities at the locations of the system components constitute a vector of random variables drawn from the ground motion random field. However, the modeling of random fields is necessary in many applications. For example, factors determining the progress of deterioration in elements of concrete surfaces are random variables drawn from environmental and material property random fields. Proper modeling of the dependence structure of such vectors of random variables is essential for accurate probabilistic analysis. In the special case when the field is Gaussian, or derived from a Gaussian field, the spatial dependence structure of the field is completely defined by the autocorrelation function and the correlation matrix fully defines the dependence structure of the random vector drawn from the field. Typically, this correlation matrix is fully populated. Although this chapter only deals with Gaussian random fields, the methods developed are equally applicable to non-Gaussian fields that are derived from Gaussian fields(e.g. Grigoriu 2000).

In some applications, it is of interest to update a probabilistic model in light of available or assumed observations of the random field. For the current application, one might be interested in updating the reliability of the system when ground motion intensities at one or more locations are observed. In the case of a concrete surface subject to deterioration, the reliability of the system can be updated when cracking (or no cracking) of the concrete in some of the elements is observed. As emphasized in previous chapters, the Bayesian network (BN) methodology is a powerful tool for such updating purposes, particularly when the available information evolves in time and the updating must be done in (near) real time (e.g. Straub 2009). However, there is a limiting characteristic of the BN that poses a challenge when modeling random variables drawn from a random field: due to the full correlation structure of the random variables, the BN becomes densely connected. When combining these random variables with system models that involve additional random variables, the computational and memory demands rapidly grow with the number of points drawn from the random field. In this chapter, we develop approximate methods to overcome this difficulty. Specifically, we present methods for reducing the density of the BN model of the random field by selectively eliminating nodes and links. The aim is to minimize the number of links in the BN while limiting the error in the representation of the correlation structure of the random variables drawn from the Gaussian random field.

The chapter begins with BN models of random variables drawn from a Gaussian random field. Approximation methods are then developed to achieve computationally tractable BN models. Using several generic and systematic spatial configuration models, numerical investigations are performed to compare the relative effectiveness of the proposed approximation methods. Finally, the effects of the random field approximation on estimated reliabilities of example spatially distributed systems are investigated. The chapter ends with a set of recommendations for BN modeling of random variables drawn from a random field.

#### 4.2 Bayesian network modeling of random fields

Let  $Y(\mathbf{x})$ ,  $\mathbf{x} \in \Omega$ , be a multi-dimensional Gaussian random field defined within the domain  $\Omega$  with mean function  $\mu_Y(\mathbf{x})$ , standard deviation function  $\sigma_Y(\mathbf{x})$ , and auto-correlation coefficient function  $\rho_{YY}(\mathbf{x}_1, \mathbf{x}_2)$ ,  $(\mathbf{x}_1, \mathbf{x}_2) \in \Omega$ . Without loss of generality, hereafter we work with the normalized random field

$$Z(\mathbf{x}) = \frac{Y(\mathbf{x}) - \mu_Y(\mathbf{x})}{\sigma_Y(\mathbf{x})}$$
(1)

 $Z(\mathbf{x})$  is a stationary Gaussian random field with zero mean, unit variance and, because of the linearity of the transformation, auto-correlation coefficient function  $\rho_{ZZ}(\mathbf{x}_1, \mathbf{x}_2) = \rho_{YY}(\mathbf{x}_1, \mathbf{x}_2)$ .

Consider random variables  $Z_i = Z(\mathbf{x}_i)$ , i = 1, ..., n, drawn from  $Z(\mathbf{x})$  at selected points  $\mathbf{x}_i$ within  $\Omega$ . The Gaussian vector  $\mathbf{Z} = [Z_1, Z_2, ..., Z_n]^T$  has zero means, unit standard deviations and correlation matrix  $\mathbf{R} = [\rho_{ij}]$ , where  $\rho_{ij} = \rho_{YY}(\mathbf{x}_i, \mathbf{x}_j)$ , i, j = 1, ..., n. We consider the general case where the correlation matrix is fully populated. Figure 4.1 shows a BN model of vector  $\mathbf{Z}$ . The correlation structure implies links between all pairs of *Z*-nodes. The particular formulation in Figure 4.1 requires specification of the conditional distribution of each  $Z_i$  given its parent nodes  $Z_1, ..., Z_{i-1}$ . That is, the conditional probability  $\Pr(Z_i =$   $z_i|Z_1 = z_1, ..., Z_{i-1} = z_{i-1}$ ) must be specified for each combination of the mutually exclusive states of  $Z_1, ..., Z_{i-1}, Z_i$ . It should be clear that the CPT of node  $Z_n$  can become extremely large as n increases.

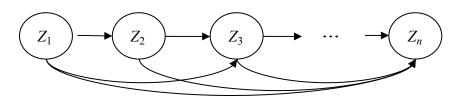


Figure 4.1: BN model of vector Z drawn from Gaussian random field Z(x)

Vector **Z** may be decomposed as a product of an  $n \times n$  transformation matrix **T** and an  $n \times 1$  vector of statistically independent, standard normal random variables **U**:

$$\mathbf{Z} = \mathbf{T}\mathbf{U} = \begin{bmatrix} t_{11} & \cdots & t_{1n} \\ \vdots & \ddots & \vdots \\ t_{n1} & \cdots & t_{nn} \end{bmatrix} \begin{bmatrix} U_1 \\ \vdots \\ U_n \end{bmatrix}$$
(4.2)

The transformation matrix **T** may be determined using an Eigenvalue (Karhunen–Loève) expansion, Cholesky factorization, or other decomposition methods that diagonalize the covariance matrix (Trefethen and Bau 1997). Alternatively, the transformation matrix may be determined approximately via numerical optimization (Song and Kang 2009; Song and Ok 2009), as described later in this chapter. The BN corresponding to the above transformation is shown in Figure 4.2 where the latent *U*-nodes are introduced as parents of the *Z*-nodes. Here, an element of the transformation matrix,  $t_{ij}$ , is interpreted as a factor on the link between  $U_j$  and  $Z_i$ . A value of  $t_{ij} = 0$  corresponds to no link between  $U_j$  and  $Z_i$ . Due to the unit covariance matrix of **U**, we have  $\mathbf{R} = \mathbf{TT}^{T}$ . The CPTs required by the BN in Figure 4.2 are easier to specify than those required for the BN in Figure 4.1, because each  $Z_i$  is a deterministic function of the its parent *U*-nodes.

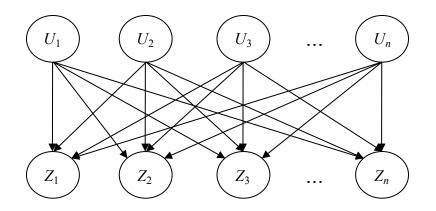


Figure 4.2: BN model of the decomposition of vector Z drawn from Gaussian random field Z(x)

The BNs in Figure 4.1 and Figure 4.2 are densely connected and, therefore, exact inference with these BNs becomes computationally intractable as the number n of random variables increases. To achieve computational tractability, the number of links in the BN must be reduced. However, removal of links introduces error. Hence, in developing a procedure for link elimination, the goal is to balance computational efficiency and model accuracy by removing as many links as possible without causing significant loss of accuracy. A procedure must be defined to identify and eliminate links that are least critical for accurately modeling the vector **Z**.

Define  $\widehat{\mathbf{T}} = [\widehat{t}_{ij}]$  as the approximate transformation matrix with some of its elements set to zero. Setting  $\widehat{t}_{ij} = 0$  implies removal of the link connecting  $U_j$  and  $Z_i$ . If column j of  $\widehat{\mathbf{T}}$  has all zero entries, then node  $U_j$  has no children and can be eliminated from the BN. The removal of links and nodes in the BN results in an approximation of the covariance matrix of  $\mathbf{Z}$ , including the on-diagonal variance terms. The errors in the variances are corrected by introducing an additional  $n \times 1$  vector of statistically independent standard normal random variables  $\mathbf{V}$  and a diagonal transformation matrix  $\mathbf{S}$ ,

$$\widehat{\mathbf{Z}} = \mathbf{S}\mathbf{V} + \widehat{\mathbf{T}}\mathbf{U} = \begin{bmatrix} s_1 & \cdots & 0\\ \vdots & \ddots & \vdots\\ 0 & \cdots & s_n \end{bmatrix} \begin{cases} V_1\\ \vdots\\ V_n \end{cases} + \begin{bmatrix} \widehat{t}_{11} & \cdots & \widehat{t}_{1m}\\ \vdots & \ddots & \vdots\\ \widehat{t}_{n1} & \cdots & \widehat{t}_{nm} \end{bmatrix} \begin{cases} U_1\\ \vdots\\ U_m \end{cases}$$
(4.3)

where  $\hat{\mathbf{Z}}$  denotes the approximated vector. Note that, after elimination of barren *U*-nodes,  $\hat{\mathbf{T}}$  is an  $n \times m$  matrix and **U** is and  $m \times 1$  vector, where  $m \leq n$ . To achieve unit variances for  $\hat{\mathbf{Z}}$ , we set the condition

$$s_i = \sqrt{1 - \sum_{k=1}^{m} \hat{t}_{ik}^2}, \quad i = 1, \dots, n$$
(4.4)

This correction does not affect the off-diagonal terms of the covariance matrix. Furthermore, the approximated correlation coefficients are given by

$$\hat{\rho}_{ij} = \sum_{k=1}^{m} \hat{t}_{ik} \, \hat{t}_{jk}, \quad i, j = 1, \dots, n \tag{4.5}$$

The BN corresponding to the formulation in (4.3) is shown in Figure 4.3. We note that the addition of the *V*-nodes does not significantly increase the computational complexity of the BN.

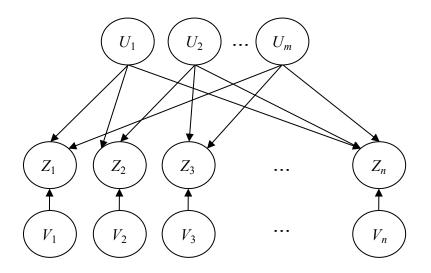


Figure 4.3: BN model of the approximate decomposition  $\hat{Z} = SV + \hat{T}U$ 

As suggested by Song and Kang (2009) and Song and Ok (2009), the transformation in (4.3) may be regarded as a generalization of the Dunnett-Sobel (DS) class of Gaussian random variables (Dunnett and Sobel 1955). This class of random variables is defined by

$$Z_i = s_i * V_i + t_i * U, \quad i = 1 \dots n$$
(4.6)

where  $V_i$  and U are independent standard normal random variables and  $s_i$  and  $t_i$  are variable-specific coefficients satisfying the conditions  $s_i = (1 - t_i^2)^{1/2}$  and  $-1 \le t_i \le 1$ . Note that U is common to all  $Z_i$  and, therefore, is the source of the correlation among them. For this model, one can easily show that  $Z_i$  are standard normal random variables having the correlation coefficients  $\rho_{ij} = t_i * t_j$  for  $i \neq j$  and  $\rho_{ij} = 1$  for i = j. As a special case, the  $Z_i$  are equi-correlated when all  $t_i$  are identical. The transformation in (4.6) corresponds to that in (4.3) with  $\hat{\mathbf{T}}$  being an  $n \times 1$  vector. The associated BN has a single common *U*-node as shown in Figure 4.4.

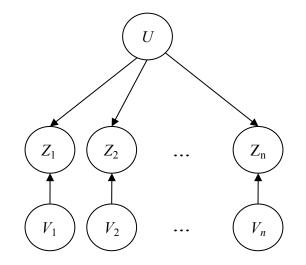


Figure 4.4: BN model of DS class of random variables

#### 4.3 Construction of Approximate Transformation Matrix $\hat{T}$

As described above, setting elements of the transformation matrix  $\hat{\mathbf{T}}$  to zero corresponds to removing links in the BN. Link removal may be accomplished by three means: (1) selectively removing *U*-nodes and all associated links by setting entire columns of the transformation matrix to zero; (2) selectively removing links by setting the corresponding elements of the transformation matrix to zero; or (3) combination of the first two, i.e., first reducing the number of *U*-nodes and then selectively removing links from the remaining *U*-nodes.

To select nodes and links for elimination, we introduce node and link importance measures. For a transformation matrix **T** (or  $\hat{\mathbf{T}}$ ), the node importance measure (NIM) for node  $U_i$  is defined as

$$M_{i} = \sum_{k=1}^{n} \sum_{l=k}^{n} |\Delta_{i}(k, l)|$$
(4.7)

where  $\Delta_i(k, l)$  is the (k, l)<sup>th</sup> element of the matrix

$$\boldsymbol{\Delta}_i = \mathbf{t}_i \mathbf{t}_i^{\mathrm{T}} \tag{4.8}$$

in which  $\mathbf{t}_i$  is the *i*<sup>th</sup> column of  $\mathbf{T}$  or  $\widehat{\mathbf{T}}$ .  $\Delta_i$  quantifies the contribution of  $U_i$  to the correlation matrix of  $\mathbf{Z}$ . Thus,  $M_i$  is a measure of the information contained in the correlation matrix  $\mathbf{R}$  that is lost by removing  $U_i$ . Clearly, eliminating the *U*-node associated with the smallest NIM will result in the least loss of accuracy. Nodes may be eliminated based on their NIM values until a pre-selected number of nodes remain, or until a pre-defined threshold on NIM is exceeded.

Similarly, a link importance measure (LIM) associated with element  $t_{ij}$  of the transformation matrix **T** or **T** is defined as

$$m_{ij} = \sum_{k=1}^{n} \sum_{l=k}^{n} |\delta_{ij}(k,l)|$$
(4.9)

where  $\delta_{ij}(k, l)$  is the  $(k, l)^{th}$  element of the matrix

$$\boldsymbol{\delta}_{ij} = \mathbf{t}_i \mathbf{t}_i^{\mathrm{T}} - \mathbf{t}_{i,j} \mathbf{t}_{i,j}^{\mathrm{T}}$$
(4.10)

in which  $\mathbf{t}_i$  is the *i*<sup>th</sup> column of  $\mathbf{T}$  or  $\hat{\mathbf{T}}$  and  $\mathbf{t}_{i,j}$  is equal to  $\mathbf{t}_i$  but with its *j*<sup>th</sup> element set equal to zero. Thus,  $m_{ij}$  is a measure of the information contained in the correlation matrix that is lost by eliminating the link from node  $U_j$  to node  $Z_i$ . It follows that eliminating the link associated with the smallest LIM will result in the least loss of accuracy. Links are eliminated based on their LIM until a pre-selected number remain, or until a pre-defined threshold on the LIM is exceeded. The node- and link-based approaches can be combined: U-nodes are first eliminated based on their NIMs, followed by selective elimination of links associated with the remaining nodes based on their LIMs. Below, we describe three methods for constructing the transformation matrix  $\mathbf{T}$  and its approximation  $\hat{\mathbf{T}}$  by use of the above measures.

#### 4.3.1 Decomposition using Eigenvalue expansion

Using an eigenvalue expansion, the transformation matrix can be obtained in the form (Jolliffe 2002)

$$\mathbf{T} = \mathbf{\Phi} \mathbf{\Lambda}^{1/2} \tag{4.11}$$

where  $\Phi = [\Phi_1, ..., \Phi_n]$  is the matrix of eigenvectors and  $\Lambda = \text{diag}[\lambda_i]$  is the diagonal matrix of eigenvalues obtained by solving the eigenvalue problem

$$\mathbf{R}\boldsymbol{\phi}_i = \boldsymbol{\phi}_i \lambda_i, \qquad i = 1, \dots, n \tag{4.12}$$

In this case the transformation matrix is generally full and the resulting BN takes the form in Figure 4.2. The eigenvalues  $\lambda_i$  may be interpreted as factors associated with the *U*-nodes, while the elements in the eigenvectors are interpreted as factors on the links. It is well known that the order of contribution of the eigenvectors to the covariance matrix is consistent with the order of magnitudes of the corresponding eigenvalues (This property is exploited by Principal Component Analysis (Jolliffe 2002)). Our numerical investigations reveal that the order of the NIMs generally agrees with the eigenvalue ordering. In the remainder of this chapter, the approach in which the transformation is defined via eigenvalue decomposition and U-nodes are eliminated based on their NIMs is referred to as the node-based eigenvalue approach (NEA). When the elimination is performed on individual links based on their LIMs, the approach is referred to as a *link-based eigenvalue* approach (LEA). It is noted that the use of a link-based elimination procedure in conjunction with eigenvalue decomposition was first suggested in Straub et al. (2008). However, the importance measure defined there is less general than the one introduced above, as the former is only applicable to transformation matrices obtained via eigenvalue decomposition.

### 4.3.2 Decomposition using Cholesky factorization

The decomposition approach using Cholesky factorization results in a transformation matrix **T** that is triangular (Trefethen and Bau 1997). Assuming **T** is a lower triangular matrix, the corresponding BN takes the form shown in Figure 4.5, where node  $Z_i$  has the parents  $U_1, ..., U_i$ . By virtue of the triangular form of the transformation matrix, the BN in Figure 4.5 is less densely connected than the BN is Figure 3. However, the sizes of the largest CPT and the largest clique,  $\{U_1, ..., U_n, Z_n\}$ , remain unchanged. Consequently, the order of computational complexity of exact inference associated with the formulations in Figure 4.2 and Figure 4.5 are the same. Using a node-elimination approach, the nodes corresponding to the rightmost columns of the transformation matrix are typically associated with the smallest NIMs and are eliminated first. Hereafter, the elimination of nodes based on a transformation matrix determined by Cholesky decomposition is referred to as the *node-based Cholesky approach* (NCA). When individual links are eliminated by zeroing the elements in the Cholesky decomposition matrix based on their LIMs, the approach is referred to as the *link-based Cholesky approach* (LCA).

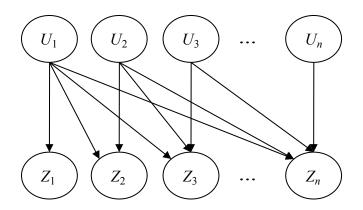


Figure 4.5: BN corresponding a transformation matrix obtained by Cholesky decomposition

#### 4.3.3 Node and link elimination using optimization

Numerical optimization offers an alternative approach to defining the approximate transformation matrix  $\hat{\mathbf{T}}$ . We propose an optimization-based node-elimination approach consisting of two steps: (1) specification of the number  $m \leq n$  of *U*-nodes to include in the BN; and (2) solution of the nonlinear constrained optimization problem

$$\min \sum_{i=1}^{n-1} \sum_{j=i+1}^{n} \left[ \rho_{ij} - \sum_{k=1}^{m} \hat{t}_{ik} * \hat{t}_{jk} \right]^{2}$$
(4.13)
subject to: 
$$\sum_{k=1}^{m} \hat{t}_{ik}^{2} \le 1, \qquad i = 1, ..., n$$

The objective of the above problem is to minimize the sum of squared deviations between the actual and approximated coefficients in the upper triangle of the symmetric correlation matrix (excluding the diagonal terms, which are later corrected using (4.3) and (4.4)). The constraint functions are based on (4.4). The use of nonlinear constrained optimization to determine the terms of an approximate transformation matrix was earlier suggested by Song and Ok (2009) and Song and Kang (2009). Hereafter, construction of the approximate transformation matrix by use of the optimization scheme in (4.13) is referred to as the *node-based optimization approach* (NOA).

Analogously, a link-based optimization problem may be formulated, in which the number of links rather than nodes is specified. The result is a mixed-integer, non-linear constrained optimization problem, which is difficult to solve in practice. Problems with convergence and over-sensitivity to initial values were encountered. Instead, an iterative procedure for link elimination is developed. First, the optimization problem in (4.13) is solved for a pre-

specified *m*. Then, the term in the resulting transformation matrix with the lowest LIM is set to zero and the remaining terms in the matrix are re-optimized. The link elimination is repeated until a pre-set number of links remain. The iterative algorithm is summarized in Figure 4.6. Hereafter, this procedure is referred to as the *iterative link-based optimization approach* (ILOA).

INITIALIZE
Specify: $m =  ext{maximum number of } U ext{-nodes to include}$ $L =  ext{number of links to eliminate}$ ELIMINATE NODES
1. Determine $n \times m$ matrix $\hat{\mathbf{T}}_0$ with elements obtained by solving the optimization problem in (4.13)
ELIMINATE LINKS
for $p = 1,, L$
1. Calculate LIMs for all elements of $\hat{\mathbf{T}}_{p-1}$
<ol> <li>Identify <i>p</i> elements in Î<sub>p-1</sub> with the smallest LIM values. (These include elements previously set equal to zero.)</li> <li>Determine the <i>n</i> × <i>m</i> matrix Î<sub>p</sub> by setting all elements identified in step 2 equal to zero and solving for the remaining elements according to (4.13)</li> </ol>
end

#### Figure 4.6: Iterative algorithm for determining $\widehat{T}$ through node and link elimination

An alternative to the above iterative scheme is as follows: In each iteration step, all elements in the transformation matrix that have LIM values below a threshold are set to zero. The remaining elements in the matrix are then re-optimized according to (4.13). The procedure is repeated until no link has an LIM below the specified threshold. Hereafter, this procedure is referred to as the *alternative iterative link-based optimization approach* (AILOA).

For the numerical analysis reported in this chapter, the function *fmincon* in Matlab, which finds the minimum of a constrained non-linear multi-variable function, is used to solve the optimization problem (4.13).

## 4.3.4 Qualitative comparison of methods for constructing $\widehat{T}$

It is important to recognize a significant distinction between the decomposition (eigenexpansion and Cholesky factorization) and optimization methods described above. In the former methods, after setting selected elements to zero, the remaining elements in the transformation matrix remain unchanged. In contrast, in the optimization methods, the matrix  $\hat{\mathbf{T}}$  is re-optimized after elimination of each set of nodes or links. In this sense, the optimization approaches are "dynamic," while the decomposition methods are "static." Because of these characteristics, one can expect that the optimization approaches will generally outperform the decomposition methods.

From a practical standpoint, a node-elimination approach is preferable because the effect on the computational complexity of the BN is systematic. The user can easily predict the memory demands of the resulting BN prior to performing inference. The same cannot be said when links are individually eliminated. However, the link-elimination approach is often better able to capture the correlation structure of a given configuration of points drawn from a random field because unimportant links are eliminated selectively rather than in large groups, as is the case with the node-elimination approach. These observations are borne out by numerical investigations that follow.

# 4.4 Numerical investigation of approximation methods

In this section, a numerical investigation is performed to determine the relative efficiencies of the proposed methods for approximating the correlation structure of random variables drawn from a Gaussian random field in the context of BN analysis. Four generic configurations for the locations of points drawn from the random field are considered:

- (1) Points arranged along a line (Figure 4.7a)
- (2) Points arranged concentrically on a circle (Figure 4.7b)
- (3) Points arranged in a rectangular grid (Figure 4.7c)
- (4) Points arranged in clusters (Figure 4.7d)

Different sizes of the above configurations are considered by changing the number of points, *n*, and the distance between the points. The latter is controlled through a distance measure, *d*, as defined in Figure 4.7. For the cluster configuration, the points in each cluster are distributed uniformly on the circumference of a circle of radius *d*, which is centered at a pre-defined coordinate. For the sake of brevity, only results for representative cases of the above configurations are shown and general trends are described.

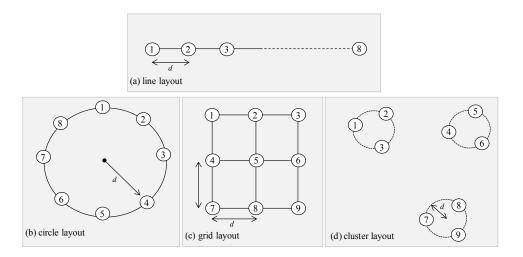


Figure 4.7: Generic configurations considered in numerical investigation

To measure the relative efficiencies of the aforementioned approximation methods (NEA, LEA, NCA, LCA, NOA, ILOA, and AILOA), the error measure

$$e = \frac{\sum_{i=1}^{n-1} \sum_{j=i+1}^{n} |\hat{\rho}_{ij} - \rho_{ij}|}{\sum_{i=1}^{n-1} \sum_{j=i+1}^{n} |\rho_{ij}|}$$
(4.14)

is used, where  $\rho_{ij}$  is the correlation coefficient between  $Z_i$  and  $Z_j$  and  $\hat{\rho}_{ij}$  is the approximated correlation coefficient as defined in (4.5). Note that the sums are over the upper triangle of the correlation matrix, excluding the diagonal terms. This error measure was selected because the normalization by the sum of absolute values of the correlation coefficients permits comparison of cases with varying sizes of vector **Z** and a wide range of correlation values. (Later in the chapter, we consider another error measure that is directly relevant to reliability assessment of an infrastructure system.) For each layout, this error is compared against a measure of the computational complexity of the resulting BN. For simplicity and intuitive appeal, we use the number of links in the BN as a proxy for computational complexity. There is not a perfect one-to-one relation between the number of links in a BN and direct measures of computational complexity, such as the tree width, which is equal to one less than the size of the largest clique (Wainwright and Jordan 2008). However, the number of links is sufficiently indicative to serve our purpose of comparing alternative approximation methods.

For the numerical investigation, the Gaussian random field is assumed to be homogenous and isotropic with the autocorrelation coefficient function

$$\rho_{VV}(\mathbf{x}_i, \mathbf{x}_j) = \exp\left(-\frac{\Delta \mathbf{x}_{ij}}{a}\right)$$
(4.15)

where  $\Delta \mathbf{x}_{ij} = |\mathbf{x}_i - \mathbf{x}_j|$  is the distance between sites  $\mathbf{x}_i$  and  $\mathbf{x}_j$  and a is a measure of the correlation length of the random field with value a = 6 in units of distance. This correlation model is taken from (Park et al. 2007) and is motivated by our interest in applying the BN methodology to seismic risk assessment of infrastructure systems.

## 4.4.1 Results

First consider a line layout system (Figure 4.7a) with 10 sites. For this system, an exact BN with the eigenvalue expansion method has  $10^2 = 100$  links (see Figure 4.3) and with the Cholesky decomposition method it has  $1 + 2 + \dots + 10 = 55$  links (see Figure 4.5). Figure 4.8 shows the error measure versus the number of links included in the BN for d = 1, 5 and 10 (corresponding to  $d = \frac{1}{6}a, \frac{5}{6}a$  and  $\frac{5}{3}a$ ) using the decomposition and optimization methods described above. All methods exhibit similar trends with increasing d, i.e., the error measure becomes larger and converges to zero at a slower rate. This is partly due to the definition of the error measure: For large *d*, the correlations are small and, therefore, the error measure is magnified. For all *d* values and for both node- and link-elimination approaches, the optimization methods consistently achieve smaller errors than the decomposition methods. Within the optimization methods, ILOA and AILOA outperform NOA, particularly with regard to the speed of convergence to zero. For small d, the performances of NEA and LEA are close to those of the corresponding optimization approaches, but the performance of the eigen-expansion approaches rapidly degrade as d increases. This is because, for large d, the correlation coefficients are small and the eigenvalues of the resulting correlation matrix are of similar magnitude. Furthermore, LEA exhibits non-monotonic convergence. Conversely, the approaches based on Cholesky decomposition perform poorly for small distances, but their performance approaches those of the optimization methods as *d* increases. The above investigation was repeated for 5and 15-site line layouts and similar trends were observed.

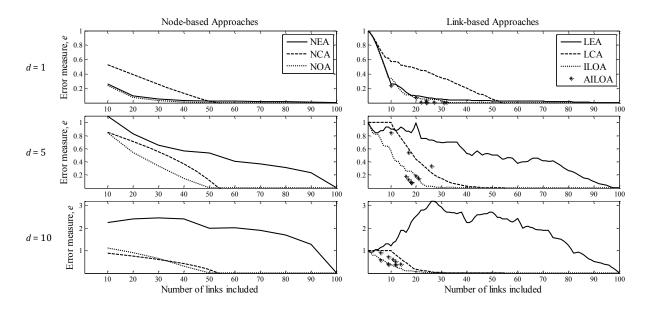


Figure 4.8: Error in approximating correlation matrix for 10-site line layouts

Figure 4.9 and Figure 4.10 respectively show the results for 10-site circle and 9-site grid layouts. As with the line layout, the optimization methods outperform the decomposition methods for all values of *d*. For small *d*, all the points in these layouts are closely spaced and, hence, the correlation coefficients are large and relatively uniform in magnitude. As a result, the correlation structure is well captured by a few *U*-nodes, or even a single *U*-node. As *d* increases, the elements of the correlation matrix become less uniform. A larger number of nodes is then required to capture the correlation structure when using node-based approaches. In this situation, the link-based approaches exhibit better performance because they offer versatility in distributing links among different *U*-nodes.

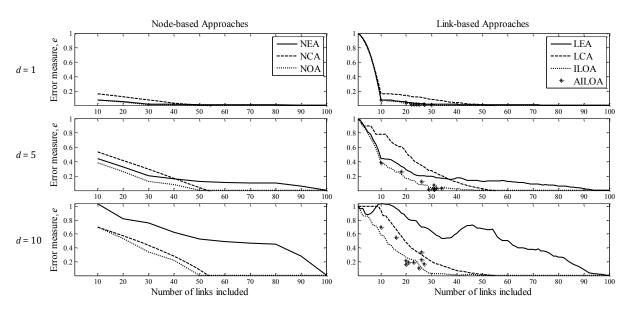


Figure 4.9: Error in approximating correlation matrix for 10-site circle layouts

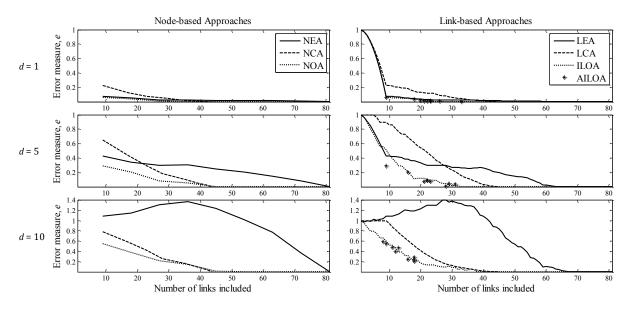


Figure 4.10: Error in approximating correlation matrix for 9-site grid layouts

Finally, the layout in which points are arranged in clusters is considered. Figure 4.11 shows examples of 9-site layouts arranged into three three-site clusters centered at coordinates (0,0), (50,50) and (25,75). For each cluster, points are equally distributed around a circle of radius *d*. Thus, *d* is a measure of how tightly the nodes in each cluster are arranged.

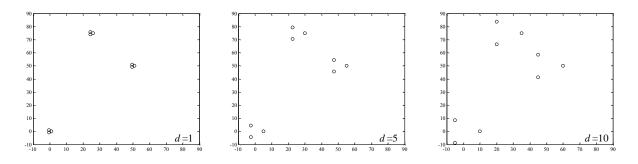


Figure 4.11: Example cluster layouts with *d* = 1, 5, 10

Figure 4.12 shows the errors for the 9-site cluster layouts with d = 1, 5 and 10. Among the node-based approaches, NCA and NOA perform similarly while NEA performs poorly, particularly for large d. Overall, the link-based approaches, excluding LEA, offer much better performance for the cluster layouts. This is because link-based approaches are able to distribute links among different U-nodes consistent with the geometry of the clusters. To illustrate this concept, consider the two BNs shown in Figure 4.13, which represent alternative approximations of the 9-node cluster system with 9 links. The top BN uses the ILOA, which distributes the 9-links such that the points in each cluster are connected to a common U-node and points in different clusters are uncorrelated. The bottom BN uses a node-based approach, in which 9 links connect a single U-node to all the Z-nodes. While both BNs in Figure 4.13 have the same number of links, ILOA distributes them more efficiently, consistent with the geometry of the layout. If clusters are sufficiently far apart, it is not necessary to include information paths (i.e. common U-nodes) between sites located in different clusters. This is why the link-based methods outperform node-based methods.

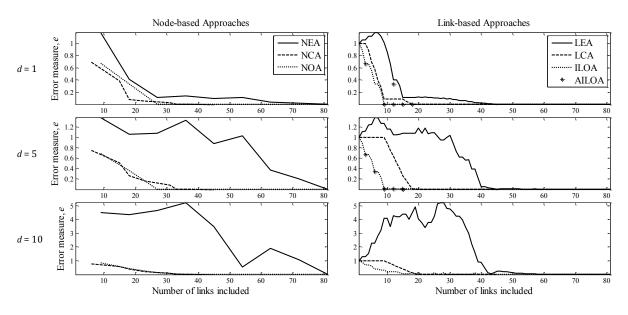


Figure 4.12: Error in approximating correlation matrix for 9-site cluster layouts

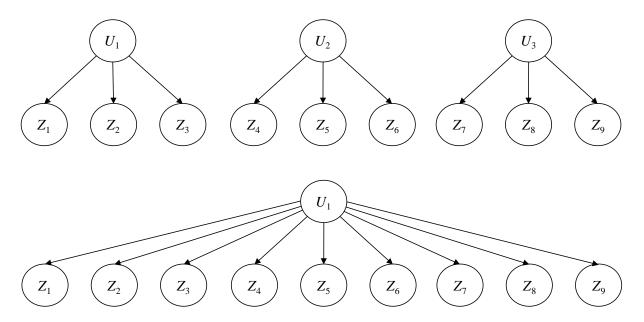
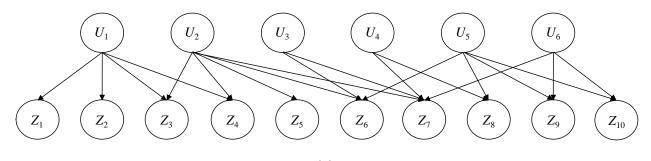


Figure 4.13: 9-link BN approximations for 9-site cluster layouts using ILOA (top) and NOA (bottom)

For comparison, select BNs obtained using the ILOA for the n = 10-site line and circle and n = 9-site grid layouts for d = 5 are shown next. For the line and circle layouts, each BN contains 2n = 20 links; the BN corresponding to the grid layout contains 2n = 18 links. Note that the ILOA is an iterative procedure for identifying important links. The iterative procedure is used in lieu of considering the full link-based optimization problem, which is difficult to solve in practice. Therefore, the BNs obtained from the procedure may not be

globally optimal. Geometric interpretations of the resulting BNs for the line, circle, and grid layouts are not as clear as that for the cluster configuration, however trends do exist.

For the line layout the BN using 20 links is shown in Figure 4.14a. In this BN, links are distributed so that there are over-lapping information paths between nodes in close proximity. Sites located farther apart do not share a common U-node. Figure 4.14b graphically illustrates the sites that are linked by common U-nodes. Shaded circles in this diagram indicate sites that share a common U-node. The BN for the 10-site circle layout with 20 links is shown in Figure 4.15a and the corresponding diagram of sites sharing common U-nodes is shown in Figure 4.15b. Because points are spaced concentrically in this layout, the sites are linked by a single common U-node when as few as 10 links are considered (see the second diagram from the left in Figure 4.15b). When more than 10 links are included, overlapping information paths around the circle are added, similar to the trend seen for the line-layout. For the grid layout, the BN containing the 18 most important links defined by ILOA is shown in Figure 4.16b. For this layout with 18 links, the links are distributed so that adjacent sites share at least one U-node.



(a)

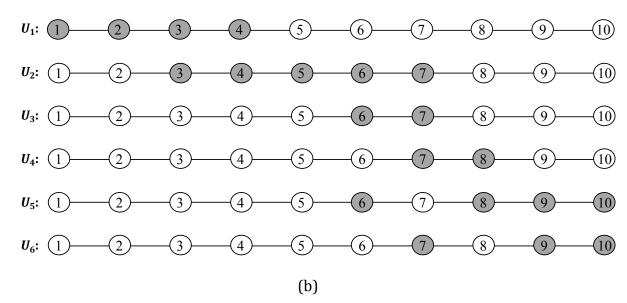
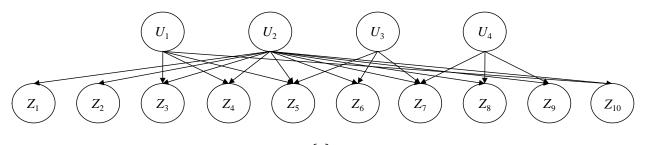


Figure 4.14: (a) BN approximation using ILOA with 20 links; (b) illustration of sites linked by common *U*-node for 10-site line layout with *d*=5



(a)

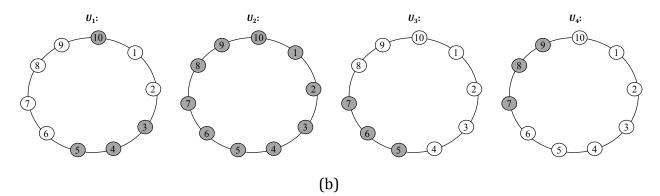
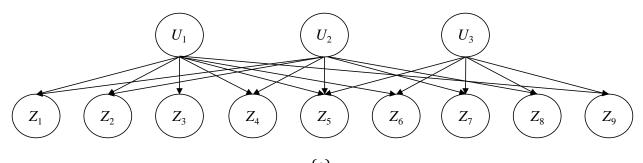


Figure 4.15: (a) BN approximation using ILOA with 20 links; (b) illustration of sites linked by common U-node for 10-site circle layout with d=5



(a)

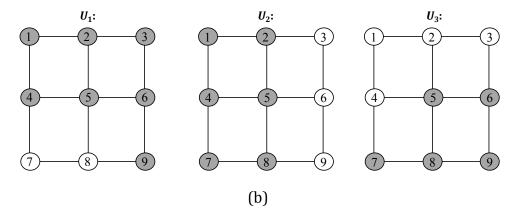


Figure 4.16: (a) BN approximation using ILOA with 18 links; (b) illustration of sites linked by common *U*-node for 9-site grid layout with *d*=5

#### 4.5 Effect of correlation approximations on system reliability

The objective in BN analysis is usually some sort of probabilistic assessment involving risk or reliability evaluation, life cycle cost analysis, statistical inference, or expected-utility decision-making. Thus, the effect of approximating the correlation structure of random variables drawn from a Gaussian random field should be evaluated within such a context. Towards that end, we investigate the effect of the correlation approximation on the estimates of reliability of example systems. The points in the layouts in Figure 4.7 are assumed to represent the locations of components of infrastructure systems subjected to an earthquake demand. Two system performance criteria are considered: (1) all components must survive for the system to survive (series system); (2) only one component needs to survive to ensure survival of the system (parallel system).

A limit-state function is assigned to each component of the system to model its performance. For component i, i = 1, ..., n, the limit-state function has the form

$$g_i = \ln(R_{c,i}) - \ln(S_i)$$
(4.16)

where  $R_{c,i}$  is the capacity of component *i* and  $S_i$  is the seismic demand on the component. The natural logarithms  $\ln(R_{c,i})$  are assumed to be statistically independent normal random variables with common means  $\lambda_R$  and common standard deviations  $\zeta_R$ . Consistent with the descriptions in Chapter 3, the natural logarithm of the seismic demand on component *i* is expressed as

$$\ln(S_i) = \ln(\overline{S}_i) + \varepsilon_m + \varepsilon_{r,i} \tag{4.17}$$

where  $\overline{S}_i$  is the median ground motion intensity at site *i*,  $\varepsilon_m$  is a normally distributed interevent error term with zero mean and variance  $\sigma_m^2$  that is common to all components in the system, and  $\varepsilon_{r,i}$  is a site-specific intra-event error term drawn from a homogeneous Gaussian random field with zero mean, variance  $\sigma_r^2$  and autocorrelation function  $\rho_{\epsilon\epsilon}(|\mathbf{x}_i - \mathbf{x}_j|)$  distributed over the spatial domain within which the system is located. It is assumed that  $\overline{S}_i$  is the same for all sites in the system. This assumption, as well as the earlier assumption of identical distributions for component capacities, ensures that all components have equal importance with regard to the performance of the system. This allows us to focus solely on the effect of the approximation of the correlation structure.

For a system of *n* components, there exist *n* limit state functions. Component *i* is in the fail state if  $g_i \le 0$ . Thus, for a series system, the failure probability is given by

$$\Pr(F_{series}) = \Pr[(g_1 \le 0) \cup (g_2 \le 0) \cup ... \cup (g_n \le 0)] = \Phi_n(-\mathbf{M}_g, \mathbf{\Sigma}_{gg})$$
(4.18)

while for a parallel system the failure probability is given by

$$\Pr(F_{parallel}) = \Pr[(g_1 \le 0) \cap (g_2 \le 0) \cap ... \cap (g_n \le 0)] = 1 - \Phi_n(\mathbf{M}_g, \mathbf{\Sigma}_{gg}) \quad (4.19)$$

The right-hand sides of the above equations are exact solutions of the system failure probabilities expressed in terms of the *n*-variate multinormal cumulative probability function  $\Phi_n(\mathbf{M}_g, \Sigma_{gg})$  with the mean vector  $\mathbf{M}_g$  having common elements  $\lambda_R - \ln \bar{S}$  and the covariance matrix  $\Sigma_{gg}$  having variances  $\zeta_R^2 + \sigma_m^2 + \sigma_r^2$  and covariances  $\sigma_m^2 + \rho_{ij}\sigma_r^2$ , where  $\rho_{ij} = \rho_{\epsilon\epsilon}(|\mathbf{x}_i - \mathbf{x}_j|)$  is the autocorrelation coefficient function of the random field from which  $\varepsilon_{S,i}$  are drawn. For the latter function, the form in (4.15) is used. Furthermore,  $\lambda_R = -0.9$ ,  $ln\bar{S} = -1.8$ ,  $\sigma_m = 0.2$  and  $\sigma_r = 0.5$  are used, while for  $\zeta_R$  two values as described below are selected. The multinormal probabilities are computed by an algorithm available in the general-purpose reliability code CalREL (Der Kiureghian et al. 2006).

Figure 4.17 shows the BN model of the system performance. Each node  $g_i$  is binary, indicating failure or survival of component *i*. It is a child of nodes representing the component capacity  $(R_i)$  and demand  $(S_i)$ . Node  $S_i$  is a child of nodes representing the common median demand  $\overline{S}$ , the common error term  $(\varepsilon_m)$  and the site-specific error term  $(\varepsilon_{r,i})$ . The correlations among the site-specific error terms are accounted for through the

latent *U*- and *V*-nodes according to the formulation in (4.3), as exemplified in Figure 4.3. The performance of the system is represented by a single node  $S_{sys}$ , which is a child of all the component limit-state nodes  $g_i$ . Note that the converging structure of links going into node  $S_{sys}$ , as shown in Figure 4.17, is computationally inefficient. Methods for efficiently constructing the system performance portion of the BN are contained in Chapter 6.

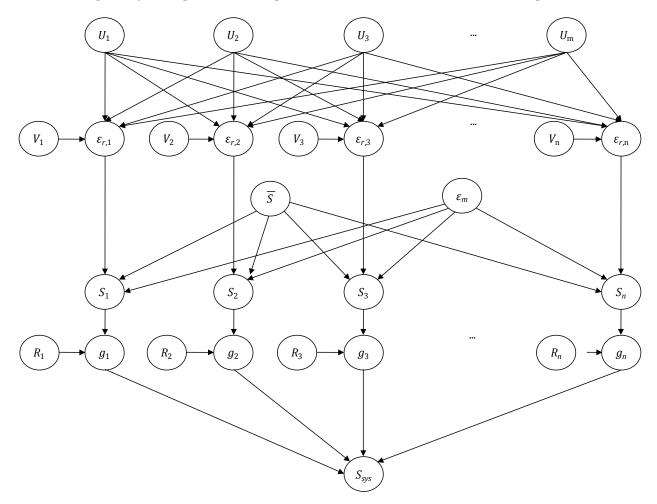


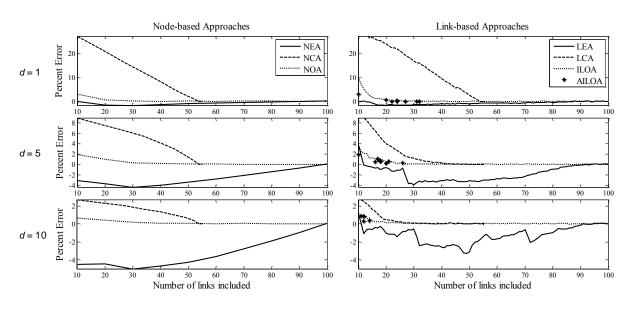
Figure 4.17: BN model of system performance

When using the proposed approximation methods, the correlation coefficients  $\rho_{ij}$  are replaced by their approximations  $\hat{\rho}_{ij}$ . In order to avoid mixing the discretization error of the BN with the error due to elimination of nodes and links, failure probabilities are computed by the same formulas as in (4.18) and (4.19) but using the approximate covariance matrix. In the following analysis, the error in computing the system failure probability is measured as

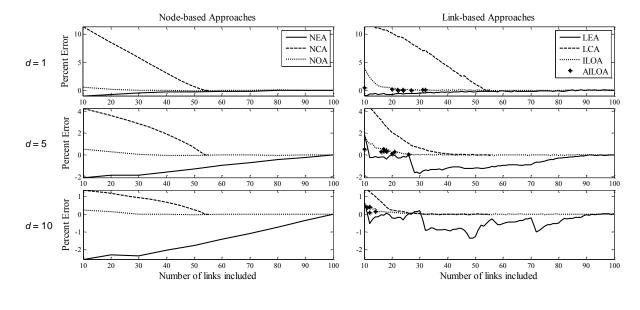
Percent Error = 
$$\frac{\hat{P}_{f,sys} - P_{f,sys}}{P_{f,sys}} * 100$$
(4.20)

where  $P_{f,sys}$  is the true system failure probability and  $\hat{P}_{f,sys}$  is the approximate failure probability computed based on the approximate correlation matrix. A negative (positive) error implies underestimation (overestimation) of the system failure probability.

For a 10-site line layout system, Figure 4.18 and Figure 4.19 plot the percent errors in estimating the failure probabilities for series and parallel systems, respectively, versus the number of links included for each of the approximation methods. These are shown for two values of  $\zeta_R$ , which approximately represents the coefficient of variation of the component capacities. Consistent with the results shown in the previous section, the optimization approaches outperform the decomposition approaches for both the parallel and series configurations. Among the decomposition approaches, the eigen-expansion methods are unconservative for the series system and the Cholesky factorization methods are unconservative for the parallel system. Overall, the errors associated with the series system are significantly smaller than those associated with the parallel system. This is because the series system failure probability is less sensitive to the correlation between component demands than that of a parallel system with the same components – a fact that has also been observed by other investigators (Grigoriu and Turkstra 1979). Comparing the graphs in parts (a) and (b) of Figure 4.18 and Figure 4.19, it is evident that the error due to the approximation in the correlation matrix of the component demands becomes less critical when the uncertainty in the component capacities is large ( $\zeta_R = 0.6$  versus  $\zeta_R$  = 0.3). Thus, accurate modeling of the correlation structure of the random field is less critical when important sources of uncertainty other than the random field are present.







(b)

Figure 4.18: Percent error in estimating failure probability for a 10-site line series system when (a)  $\zeta_R = 0.3$  and (b)  $\zeta_R = 0.6$ 

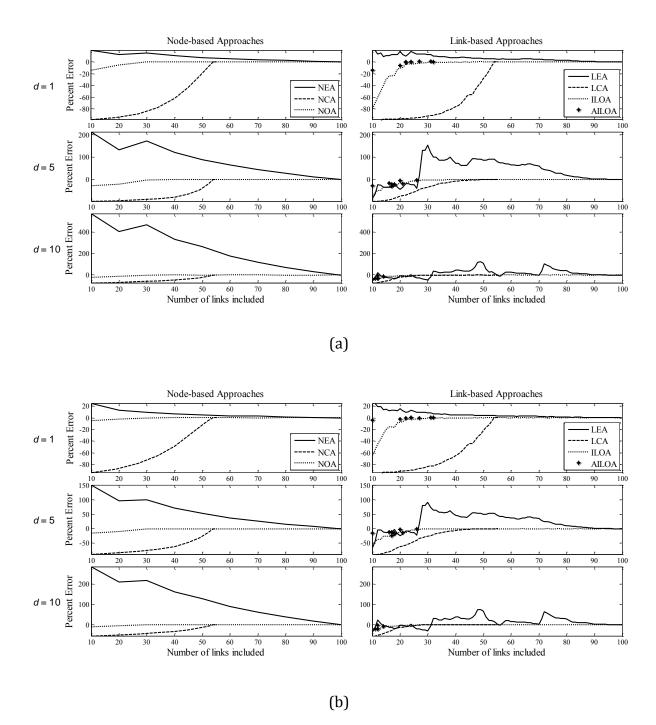


Figure 4.19: Percent error in estimating failure probability for a 10-site line parallel system when (a)  $\zeta_R = 0.3$  and (b)  $\zeta_R = 0.6$ 

To aggregate and compare the results of the numerical investigation for the various layouts, we consider the minimum number of links required in the BN model of each layout

to achieve an error less than 10% in the estimate of the system failure probability. Table 4.1 and Table 4.2 present the required minimum number of links for series and parallel systems, respectively, for each of the layouts and each of the seven approximation methods considered. Tables on the left list the results for  $\zeta_R = 0.3$ , while those on the right list the results for  $\zeta_R = 0.6$ . Results for d = 10 are not shown for the series system because, for this relatively long distance, the target accuracy threshold of 10% is achieved without including random field effects. For shorter correlation lengths, similar accuracy is obtainable when neglecting random field effects for some layouts, particularly for  $\zeta = 0.6$ , as indicated by a symbol (\*) in Table 4.1.

# Table 4.1: Number of links required to achieve less than 10% error in estimate offailure probability for series system when $\zeta_R = 0.3$ (left) and $\zeta_R = 0.6$ (right) (Bold numbers indicate theminimum number of links obtained with any approximation method)

<i>d</i> =1									<i>d</i> =1									
Layout	N	NEA	LEA	NCA	LCA	NOA	ILOA	AILOA	Layout	N	NEA	LEA	NCA	LCA	NOA	ILOA	AILOA	
Line	5	5	5	9	9	5	5	5	Line	5	5	4	5	4	5	4	5	
Line	10	10	9	40	40	10	9	10	Line	10	10	7	19	21	10	8	10	
Circle	5	5	5	5	5	5	5	5	Circle	5	5	4	5	4	5	4	5	
Circle	10	10	10	27	22	10	10	10	Circle	10	10	9	10	9	10	9	10	
Cluster	9	9	6	15	7	18	6	9	Cluster	9	*	*	*	*	44	201	*	
Grid	9	9	9	17	18	9	9	9	Grid	9	9	8	9	9	9	8	9	
									-									
				d = 5									d = 5					
Layout	N	NEA	LEA	NCA	LCA	NOA	ILOA	AILOA	Layout	N	NEA	LEA	NCA	LCA	NOA	ILOA	AILOA	
Line	5	5	2	5	6	5	2	5	Line	5	*	*	*	*	*	*	*	
Line	10	10	2	10	11	10	2	10	Line	10	*	*	*	*	44	201	*	
Circle	5	5	3	5	6	5	3	5	Circle	5	*	*	*	*	*	*	*	
Circle	10	10	8	27	24	10	8	10	Circle	10	10	5	10	8	10	4	10	
Cluster	9	9	1	6	1	9	1	3	Cluster	9	*	*	*	*	*	*	*	
					20				Grid									

\* Percent error below 10% acheiveable when neglecting random field effects

<i>d</i> =1									<i>d</i> =1										
Layout	N	NEA	LEA	NCA	LCA	NOA	ILOA	AILOA	Layout	N	NEA	LEA	NCA	LCA	NOA	ILOA	AILOA		
Line	5	10	9	14	14	5	5	5	Line	5	10	7	12	13	5	5	5		
Line	10	50	14	52	53	20	21	20	Line	10	30	22	52	53	10	19	10		
Circle	5	15	9	12	13	5	5	5	Circle	5	10	8	12	13	5	5	5		
Circle	10	10	10	45	45	10	10	10	Circle	10	10	10	45	44	10	10	10		
Cluster	9	72	35	33	18	27	9	9	Cluster	9	63	33	31	17	27	9	9		
Grid	9	27	13	39	39	9	9	9	Grid	9	18	13	35	37	9	9	9		
	<i>d</i> =5									<i>d</i> =5									
Layout	N	NEA	LEA	NCA	LCA	NOA	ILOA	AILOA	Layout	N	NEA	LEA	NCA	LCA	NOA	ILOA	AILOA		
Line	5	25	20	14	14	5	7	5	Line	5	25	4	14	13	5	7	5		
Line	10	100	85	54	39	30	24	20	Line	10	90	12	54	38	30	20	20		
Circle	5	25	19	14	14	5	5	5	Circle	5	25	17	14	14	5	5	5		
Circle	10	90	71	52	50	20	13	18	Circle	10	80	61	52	49	10	10	10		
Cluster	9	81	5	33	18	27	9	9	Cluster	9	81	5	31	17	27	9	9		
Grid	9	81	56	42	41	9	19	9	Grid	9	72	50	42	40	9	18	9		
	<i>d</i> = 10												d = 10						
Layout	N	NEA	LEA	NCA	LCA	NOA	ILOA	AILOA	Layout	N	NEA	LEA	NCA	LCA	NOA	ILOA	AILOA		
Line	5	25	20	14	10	5	7	5	Line	5	25	3	14	9	5	7	5		
Line	10	100	13	54	25	30	19	12	Line	10	100	13	52	23	20	15	12		
Circle	5	25	21	14	13	5	9	5	Circle	5	25	3	12	11	5	8	5		
Circle	10	100	81	54	47	10	23	10	Circle	10	100	77	52	44	10	21	10		
Cluster	9	81	4	23	17	27	10		Cluster	9	54	4	23	16	9	8			
Grid	9	81	55	42	37	9	32	9	Grid	9	81	51	42	34	9	31	9		
										• i	ndicates me	thod was no	t able to achi	eve a percer	nt error belo	wspecified	threshold		

Table 4.2: Minimum number of links required to achieve less than 10% error in estimate of failure probability for parallel system when  $\zeta_R = 0.3$  (left) and  $\zeta_R = 0.6$  (right) (Bold numbers indicate the minimum number of links obtained with any approximation method)

It is known that series systems are not strongly sensitive to neglecting correlation in demands. This observation is reflected in Table 4.1, which demonstrates that the accuracy threshold is met when considering very few links, if any. Note that, for the cases in which the threshold is met without inclusion of random field effects, an asterisk (\*) is included in the table. These cases essentially require zero links. Furthermore, for the series system, the optimization-based methods are generally not associated with significant gains in efficiency relative to the decomposition approaches, particularly LEA. Because the optimization approaches are computationally more expensive than the decomposition approaches, it may not be of value to compute transformation matrices using optimization techniques when working with series systems.

Conversely, it is well established that parallel systems are sensitive to inclusion of correlation in demands. This finding is reflected in Table 4.2 by the relatively large number of links that are needed in each case to achieve the required threshold of accuracy in the estimate of the system failure probability. With only a few exceptions, the optimization approaches are more efficient than the decomposition approaches. For small values of *d* (high correlations), the node- and link-based optimization approaches offer similar performance, except in the case of the cluster layout system. For large *d* values (low correlations), NOA and AILOA are more efficient than ILOA. In most cases, more links are required to achieve a percent error below the specified threshold when  $\zeta_R = 0.3$  than when  $\zeta_R = 0.6$ , particularly for larger systems.

#### 4.6 Summary

Methods for efficient Bayesian network (BN) modeling of correlated random variables drawn from a Gaussian random field, such as those arising in seismic risk assessment of spatially distributed infrastructure systems, are investigated. The modeling of broadly dependent random variables results in a BN that is densely connected. Because exact inference algorithms in densely connected BNs are demanding of computer memory, approximate methods are necessary to make the BN computationally tractable for large systems. This chapter develops methods for reducing the density of connections in a BN by eliminating nodes and/or links, while minimizing the error in the representation of the correlation structure. Methods based on classical decomposition techniques as well as numerical optimization are developed and compared. It is found that optimization methods are able to achieve the best trade-off of accuracy versus computational efficiency. The effects of the approximation methods on estimates of failure probability for idealized infrastructure systems are also considered. It is found that the optimization-based approaches offer significant increases in efficiency when modeling the performance of parallel systems. For series systems, which are known to be less sensitive to the correlation structure of component demands, classical decomposition approaches may suffice. While the work done in this chapter has been performed as part of a broader effort to develop a BN-based framework for seismic infrastructure risk assessment, it is believed that the findings are useful for more general applications.

# Chapter 5: Modeling Component Performance via BN

### 5.1 Introduction

This chapter presents BN formulations for modeling the performance of individual components of an infrastructure system as a function of the seismic demands placed upon them. Two types of components are considered: (1) point-site components, such as relatively short span highway bridges, buildings, culverts, pumps in a water distribution system, or transformers in an electrical grid; and (2) distributed components, such as pipelines, embankments, and rail lines.

Components of an infrastructure system are modeled as having multiple states, but in many cases binary states (e.g., fail or survive, open or closed, damaged or undamaged) are sufficient to describe their performance under seismic demands. For example, a single-lane tunnel in a transportation system will either be open or closed after an earthquake. Components of an electrical system will also typically have binary states. Usually more than two states are considered when a component has a "flow" characteristic associated with its performance. Water distribution systems have components with flow characteristics, e.g. a leaking pump in a water distribution system may exhibit decreased flow capacity, though it is not entirely inoperable. Multi-state components are also useful for distinguishing operating levels without consideration of flow. For example, in transportation systems, a bridge may be open for full capacity traffic, open only to light-weight (non-truck) vehicles due to load restrictions associated with moderate damage, or it may be closed entirely to traffic. Such a bridge has three states.

To perform infrastructure seismic risk assessment, it is necessary to define a mapping between the seismic demands placed on a component (modeled using the seismic demand BN formulation described in Chapter 3) and the associated component damage states. In the context of developing a BN framework for seismic risk assessment and decision support, it is assumed that only the demands and component damage states or operating level are observable. Nodes that are not observable and not required for modeling dependences between random quantities are not explicitly included in the BN. They are instead handled when constructing CPTs of nodes that are explicitly included.

The performance of a component can be predicted as a function of the ground motion intensity at its site using a variety of methods such as structural reliability analysis, stochastic finite elements, and statistical analysis of experimental or field data. For a point-site component, the probabilistic mapping between site-specific ground motion intensity and the performance of the component is expressed using a set of *fragility functions*. Each fragility function provides the probability that the component experiences a specified damage state or greater for a given level of ground motion intensity. For distributed components, similar functions give mean rates of occurrence of damages of a particular state or greater for a given level of ground motion intensity.

Fragility functions are widely used in probabilistic seismic hazard assessment. However, components in an infrastructure system are often complex systems themselves and fragility functions may represent an oversimplification of their response to seismic demands. For example, a highway bridge subjected to ground shaking may have a complex response and multiple correlated failure modes. Thus, proper computation of its reliability may require use of non-linear structural analysis as well as systems reliability methods. However, because our goal is to develop a framework that operates at the infrastructure system level, we are interested in aggregated results. Therefore, idealization of component performance using fragility functions is both reasonable and computationally necessary. Furthermore, sophisticated approaches are becoming increasingly available for developing fragility functions that are better able to predict complex post-earthquake performance of components in an infrastructure system, e.g. Straub and Der Kiureghian (2008). However, in defining fragility functions, care should be exercised to properly account for uncertainties that arise from idealization of structural performance as well as possible sources of correlation between estimated performances of similar components.

In recent years, BNs have been applied in a limited way for modeling component performance. For example, Mahadevan et al. (2001) use BNs for modeling simple structural systems with multiple failure modes. The BN consists of a node representing the overall structural performance that is a child of nodes corresponding to individual limit-state functions. However, the BN topology used in Mahadevan et al. (2001) will cause prohibitively large memory demands if used to model a large number of components in an infrastructure system. Shumuta et al. (2009) develop a BN-based damage estimation framework for electric power distribution equipment. The framework provides time-evolving estimates of post-earthquake damage by integrating information obtained relating to the hazard and performance of components, e.g. observed power outages, using BN.

They do not account for spatial correlation in demands nor system effects. Baykartali (2006) develops vulnerability curves for original and retrofitted structures (low-rise, bare frame reinforced concrete structures) and uses them to populate the CPTs for the seismic vulnerability BN model shown in Figure 3.22. More recently, Straub and Der Kiureghian (2010) developed a novel and sophisticated computational framework for integrating structural reliability methods with BNs to facilitate the risk and reliability assessment of structural systems that function as components of larger infrastructure systems. Additional applications of BNs for modeling the performance of components of infrastructure systems include embankment dams (Li et al. 2009) and marine structures (Friis-Hansen 2000).

### 5.2 Seismic fragility of point-site components

Seismic fragility is the conditional probability of failure of a component (or a point-site system) for a given measure(s) of ground motion intensity. Here, failure is defined generically as an event in which a particular level of damage, or other undesirable consequence, is met or exceeded. Using structural reliability methods, the failure of a component is defined in terms of a limit state function, which mathematically defines the boundary between failure and survival events in the outcome space of random variables influencing the state of the component (Ditlevsen and Madsen 1996). The reader is referred to other sources (Ditlevsen and Madsen 1996; Rackwitz 2001; Der Kiureghian 2005) for additional details on structural reliability methods.

Let  $g(\mathbf{X}, \mathbf{\theta})$  define the limit state function for a given component, where **X** is a set of random variables influencing the performance of the component and  $\mathbf{\theta}$  is a set of model parameters estimated from data or from theory of structural analysis. Conventionally, { $g(\mathbf{X}, \mathbf{\theta}) \leq 0$ } represents the failure event. The fragility of the component is then expressed as

$$F(\mathbf{s}, \mathbf{\theta}) = \Pr(\{g(\mathbf{X}, \mathbf{\theta}) \le 0\} | \mathbf{S} = \mathbf{s})$$
(5.1)

where **s** is a vector of variables defining the ground motion intensity. A plot of  $F(\mathbf{s}, \mathbf{\theta})$  versus **s** yields a fragility surface. In the special case when **s** is a scalar intensity measure, a plot of  $F(\mathbf{s}, \mathbf{\theta})$  versus *s* yields the conventional fragility curve (Der Kiureghian 2002). Hereafter, for the sake of simplicity, models are developed for a scalar intensity measure; however, the concepts that are developed can be easily extended to vector-valued intensity measures. Furthermore, the limit-state function is assumed to be comprised of a capacity term and a demand term, though each of these in general could be functions of other random variables.

As a general case, we consider an infrastructure system that is comprised of multiple components in different classes, e.g. bridges of different age and construction type. The capacity of components within each class to withstand damage are thought of as having been randomly drawn from a distribution that characterizes its inherent randomness arising from, for example, natural variability in material properties, construction and manufacturing processes, and changes due to deterioration. The components in different classes are assumed to have statistically independent capacities with different distributions. Let  $R_{c,i,j}$  denote the uncertain capacity of a component, where  $i \in (1, ..., n)$  is an index set defining the component numbers and  $j \in (1, ..., J)$  is an index set defining the class to which the component belongs. Capacities are measured in the same unit as the sitespecific seismic demand, which is denoted as  $S_i$  for component *i*.

In risk analysis of infrastructure systems, it is common to assume the lognormal distribution for component capacities. This distribution has several characteristics that make it well suited for modeling capacities and its use in a number of cases is supported by empirical observations. Let  $\lambda_i$  and  $\zeta_i$  denote the parameters of the lognormal capacity distribution for class *j*. For each class of components, the parameters  $\lambda_i$  and  $\zeta_i$  can be estimated using experimental or post-earthquake field observations and numerical or analytical structural models. As a result of having a limited number of observations, there is often statistical uncertainty associated with the estimates of the parameters. This implies that a distribution on the values of the parameters should be used rather than taking them to be deterministic quantities. This statistical uncertainty is common to all components within a class and, hence, gives rise to correlations among the perceived component capacities (Der Kiureghian 2002). In the context of BN modeling, this implies nodes corresponding to the uncertain parameters  $\lambda_i$  and  $\zeta_i$  must be included as common parents to nodes representing capacities of all components within each class. For illustration, Figure 5.1 shows BNs of capacities of components in several classes that are correlated by common uncertain parameters.

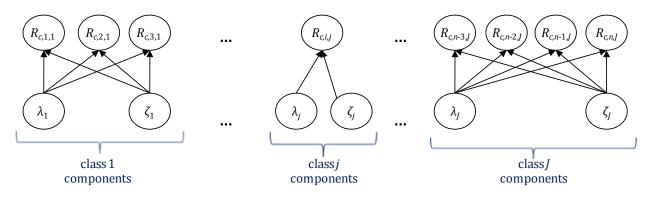


Figure 5.1: BN model of component capacities with uncertain distribution parameters

In accordance with the models described in Chapters 3 and 4, the ground motion intensity  $S_i$  is also modeled as a lognormal random variable. In that case, assuming the fragility model is exact, the limit-state function is conveniently formulated as  $g_i = \ln R_{c,i,j} - \ln S_i$ . This function defines the state of component *i* of class *j* for each value of  $R_{c,i,j}$  and  $S_i$ . The BN corresponding to such a model is shown in Figure 5.2, where the details of the ground motion intensity model are hidden in a BN object. In this BN, each node  $C_i$  is binary indicating the state of component *i*. The node indicates failure if  $g_i < 0$ , i.e.  $R_{c,i,j} < S_i$ . Note that, in the formulation in Figure 5.2, the ground motion prediction points (GMPPs)

selected in the random field seismic demand model correspond to the locations of the point-site components in the infrastructure system.

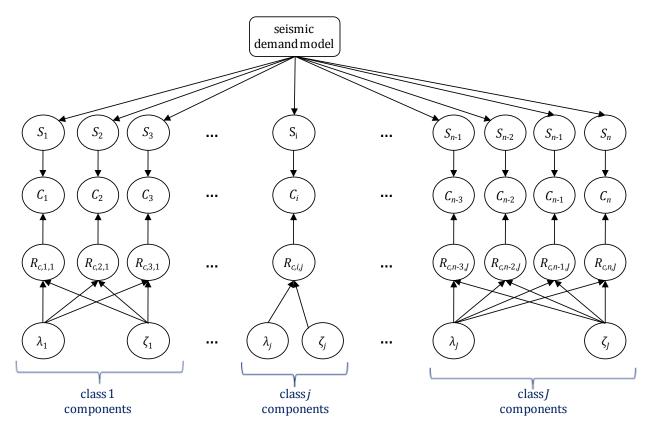


Figure 5.2: BN model of component performances assuming limit-state functions are exact

Limit-state functions and fragility models are idealizations of component performance. In constructing such models based on theory and observed data, it is usually necessary to account for the error resulting from this idealization, i.e. the model error. Following an earthquake, the observed states of components will differ from those predicted using a limit state function. In fact, given a set of identical components subjected to ground motions with exactly the same magnitude of ground motion intensity *S*, the observed performances will vary from component to component because the intensity measure *S* is not sufficient to fully characterize the ground motion at a site (Der Kiureghian 2002).

To account for the uncertainties that arise from incomplete characterization of the ground motion and other idealizations, a random correction term,  $\varepsilon_{i,j}$ , is added to the limit state function characterizing the performance of component *i* of class *j*:

$$g_i = \ln R_{c,i,j} - \ln S_i + \varepsilon_{i,j} \tag{5.2}$$

 $\varepsilon_{i,j}$  is zero mean and normally distributed with standard deviation  $\sigma_{i,j}$ . As described, the correction term captures uncertainty arising from the idealization of structural performance and missing variables, including ground motion characteristics not captured by the intensity measure and structural properties not explicitly included in the model.

At the component level, these uncertainties can be reasonably incorporated into the fragility model. As a result, a component level fragility model should not be interpreted as a distribution on the true capacity of the component because it accounts for the aforementioned sources of uncertainty. In fact, the uncertainty in the fragility model may be larger if a different, and less informative, measure of intensity is selected. The assumptions that  $R_{c,i,j}$  and  $S_i$  are lognormal random variables and  $\varepsilon_{i,j}$  is normally distributed implies  $P(g_i \leq 0)$  can be easily computed using the standard normal CDF,  $\Phi(\cdot)$ . This is the approach taken when developing many fragility models.

When modeling the performance of systems of components, a problem arises that is not an issue when considering individual components. A portion of the model error,  $\varepsilon_{i,j}$  is random from component to component. However, some of the missing variables and a portion of the uncertainty in the form of the mathematical model are common to different components of the same class. Therefore, the estimated fragilities of these components become statistically dependent. To separately account for the portions the model error that are independent versus correlated from component to component, it is useful to break the model error down into two components:  $\varepsilon_{i,j}^{(1)}$  and  $\varepsilon_j^{(2)}$ .  $\varepsilon_{i,j}^{(1)}$  captures the uncertainty that is random from component to component of class j.  $\varepsilon_j^{(2)}$  captures the portion of model error that is common to all components of the same class. The BN corresponding to such a model is shown in Figure 5.3. Consistent with an assumption of component exchangeability, it is assumed that the performances of all components of the same class are equi-correlated through  $\varepsilon_j^{(2)}$ . Thus, nodes  $\varepsilon_{i,j}^{(1)}$  are component specific and nodes  $\varepsilon_j^{(2)}$  are common to all nodes  $c_i$  of the same class j.

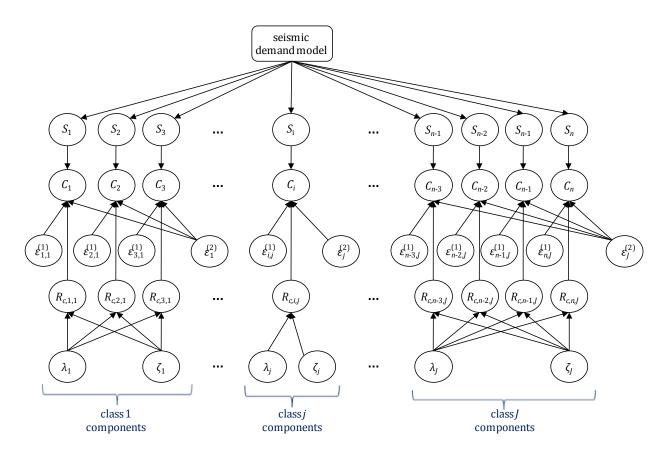


Figure 5.3: BN model of component performance including uncertainty in the functional form of the limit state functions

Because  $\varepsilon_{i,j}^{(1)}$  and  $R_{c,i,j}$  are component-specific, their uncertainties can be incorporated into the CPT of the node representing the uncertain component damage state,  $C_i$ . It is therefore not necessary to explicitly include nodes representing these quantities in the BN. This reduces computational demands and discretization error. The resulting BN is shown in Figure 5.4.

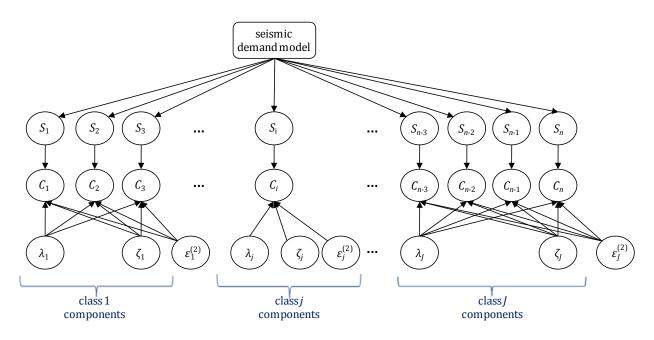


Figure 5.4: BN model of component performance including uncertainty in the functional form of the limit state functions, after eliminating nodes representing  $R_{c,ij}$  and  $\varepsilon_{i,i}^{(1)}$ 

Recall that the concept of "failure" described above is general and includes any event in which an unacceptable level of damage is experienced or exceeded. Define a damage index set  $k = \{0, 1, ..., K\}$  such that k = 0 corresponds to the intact (no damage) state and k = K corresponds to the most severe damage state considered. Node  $C_i$  has mutually exclusive states corresponding to each element of the damage index set. It follows from the above description that the fragility of component *i* for state *k* is

$$F_{k,i}(s_i) = \Pr(\text{component } i \text{ meeting or exceeding damage state } k | S_i = s_i)$$
 (5.3)

Note that  $F_{0,i}(s_i) = 1$ . The probability that component *i* is in damage state *k* is then computed as

$$P_{k,i}(s_i) = F_{k,i}(s_i) - F_{k+1,i}(s_i)$$
(5.4)

For illustration, Figure 5.5 shows example fragility curves and state probabilities for a five-state component (Der Kiureghian 2009).

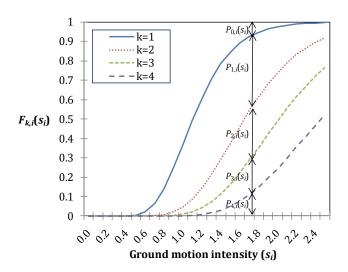


Figure 5.5: Example fragility curves and state probabilities for a five-state component

The BN in Figure 5.4 requires that the probability of each damage state be defined for each combination of the states of the parent nodes of  $C_i$ . Let

$$g_{i,k} = \ln R_{c,i,j,k} - \ln S_i + \varepsilon_{i,i,k}^{(1)} + \varepsilon_{i,k}^{(2)}$$
(5.5)

be the true limit-state function defining damage of state *k* or greater, where  $R_{c,i,j,k}$  represents the corresponding capacity and  $\varepsilon_{i,j,k}^{(1)}$  and  $\varepsilon_{j,k}^{(2)}$  are the associated model errors. Therefore the probability that component *i* will experience damage of state *k* or greater then is

$$P(g_{i,k} \le 0) = P(\ln R_{c,i,j,k} + \varepsilon_{i,j,k}^{(1)} \le \ln S_i - \varepsilon_{j,k}^{(2)})$$
(5.6)

As before,  $\ln R_{c,i,j,k}$  is assumed to be normally distributed with parameters  $\lambda_{j,k}$  and  $\zeta_{j,k}$  and  $\varepsilon_{i,j,k}^{(1)}$  and  $\varepsilon_{j,k}^{(2)}$  are normally distributed with zero mean and standard deviations  $\sigma_{i,j,k}^{(1)}$  and  $\sigma_{j,k}^{(2)}$ , respectively. In (5.6) the error terms have been split on either side of the inequality to facilitate the separate account of: (1) the portion of the model error that is independent from component to component and that need not be modeled explicitly as a node, as described above; and (2) the portion of the model error that is common to all components of the same class and thus must be modeled explicitly as a node in the BN. The conditional probability that a component is in the  $k^{\text{th}}$  damage state for a particular combination of the parent nodes of  $C_i$  is expressed as

$$P_{k,i}(s_i) = F_{k,i}\left(s_i, \varepsilon_{j,k}^{(2)}, \lambda_{j,k}, \zeta_{j,k}\right) - F_{k+1,i}\left(s_i, \varepsilon_{j,k+1}^{(2)}, \lambda_{j,k+1}, \zeta_{j,k+1}\right)$$
(5.7)

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$$= \Phi\left(\frac{\ln s_{i} - \varepsilon_{j,k}^{(2)} - \lambda_{j,k}}{\sqrt{(\zeta_{j,k})^{2} + (\sigma_{i,k}^{(1)})^{2}}}\right) - \Phi\left(\frac{\ln s_{i} - \varepsilon_{j,k+1}^{(2)} - \lambda_{j,k+1}}{\sqrt{(\zeta_{j,k+1})^{2} + (\sigma_{i,k+1}^{(1)})^{2}}}\right)$$

where  $\Phi(\cdot)$  is the standard normal cumulative distribution function. The above formulation imposes certain conditions on the true joint distribution of the parameters ( $\lambda_{j,k}, \zeta_{j,k}$ ), and model error variances  $\sigma_{i,j,k}, k = 1, ..., K$ ; namely they should be such that the probabilities computed by (5.7) are non-negative. That is, the fragility curves in Figure 5.5 must be non-intersecting.

While a few studies (e.g. Straub and Der Kiureghian (2008) for two-state components) provide information sufficient to populate the CPTs required for the above comprehensive formulation, sufficient information to assess parameter uncertainties and components of the model error are not available for many infrastructure components. As a result, it is necessary to either utilize engineering judgment to define quantities that are not available or to make simplifying assumptions about the BN formulation. Specifically, in the absence of information or assumptions about the statistical uncertainty of parameters or the statistical dependence of estimated fragility models, the corresponding simplified BN is shown in Figure 5.6. This BN disregards uncertainty in the fragility function parameters and assumes, given ground motion intensities, the component states are statistically independent. The probability that  $C_i$  is in each damage state k given  $S_i = s_i$  is defined using the expression

$$P_{k,i}(s_i) = F_{k,i}(s_i) - F_{k+1,i}(s_i) = \Phi\left(\frac{\ln s_i - \lambda_{j,k}}{\sqrt{\zeta_{j,k}^2 + \sigma_{i,j,k}^2}}\right) - \Phi\left(\frac{\ln s_i - \lambda_{j,k+1}}{\sqrt{\zeta_{j,k+1}^2 + \sigma_{i,j,k+1}^2}}\right)$$
(5.8)

where  $\sigma_{i,i,k}^2$  is the variance of the total model error.

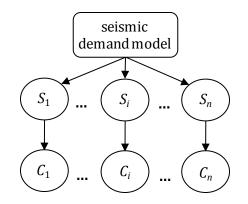


Figure 5.6: BN modeling performance of point-site components neglecting statistical uncertainty and correlation in component response other than through demands

Because the dependence among the estimated states of the components of the system is not included in the BN in Figure 5.6, the result is a simpler formulation with reduced computational demand. If the uncertainty in modeling seismic demands dominates the uncertainty in capacity, if the system is not highly redundant, or if the system is heterogeneous with few components of the same class, then this omission should not significantly degrade the accuracy of probability estimates obtained from the simplified BN.

## 5.3 Seismic fragility of distributed components

For distributed components, such as pipelines and rail segments, it is not adequate to model performance at a single point. Instead, the seismic fragility of these linear components is expressed as a mean number of damages of state k or greater per unit length, given a level of ground motion intensity. The failures along a distributed component are often assumed to follow a Poisson process. One or more failures along the length constitute failure of the entire component (Adachi and Ellingwood 2008).

While a distributed component is continuous, the BN requires discretization of the random field and thus estimates of ground motion intensity are only available at the discrete GMPPs represented by nodes  $S_i$ , i = 1, ..., n. As a result, it is necessary to discretize a distributed component into segments. GMPPs modeled in the BN are selected to correspond with the ends of each segment. Each discretized segment is considered a "component" in the system. Define  $C_{i,i+1}$  as the segment of a distributed component between GMPPs represented by nodes  $S_i$  and  $S_{i+1}$ .

Let  $\eta_i^k(s_i)$  be the mean rate of damage points of state k or greater along the component when ground motion intensity  $S_i = s_i$ . Associated with the component segment  $C_{i,i+1}$  are two mean damage rates:  $\eta_i^k(s_i)$  and  $\eta_{i+1}^k(s_{i+1})$ . Interpolating linearly between these two values gives the mean damage rate as a function of the coordinate along the component length. The mean number of damages of state k or greater for the entire component  $C_{i,i+1}$  is then

$$\mu_{i,i+1}^{k}(s_{i}, s_{i+1}) = \frac{L_{i,i+1}}{2} \Big( \eta_{i}^{k}(s_{i}) + \eta_{i+1}^{k}(s_{i+1}) \Big)$$
(5.9)

where  $L_{i,i+1}$  is the length of  $C_{i,i+1}$ . Assuming damages occur randomly, the probability that the component experiences damage of state k or greater (i.e. the probability of at least one incident of damage of state k or greater) is computed by treating the occurrence of such events as a non-homogenous Poisson process. The probability that  $C_{i,i+1}$  is in damage state k given ground motion intensities  $s_i$  and  $s_{i,i+1}$ , expressed  $p_{i,i+1}^k(s_i, s_{i,i+1})$ , is then computed as

$$P_{i,i+1}^{k}(s_{i}, s_{i+1}) = \exp\left(-\mu_{i,i+1}^{1}(s_{i}, s_{i+1})\right) \text{ for } k = 0$$
  
=  $\exp\left(-\mu_{i,i+1}^{k+1}(s_{i}, s_{i+1})\right) - \exp\left(-\mu_{i,i+1}^{k}(s_{i}, s_{i+1})\right), \text{ for } k = 1, \dots, K-1$  (5.10)  
=  $1 - \exp\left(-\mu_{i,i+1}^{K}(s_{i}, s_{i+1})\right), \text{ for } k = K$ 

Once again neglecting parameter uncertainties and correlation due to common components of model error, the BN corresponding to the above description is shown in Figure 5.7. The node representing the damage state of component segment  $C_{i,i+1}$  is modeled as a child of ground motion intensity nodes at either end of the segment:  $S_i$  and  $S_{i+1}$ . Note that this formulation preserves the correlation between the states of components sharing a common GMPP, even if random field effects were neglected.

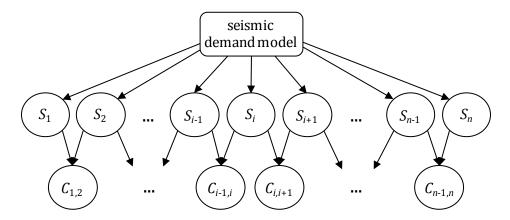


Figure 5.7: BN modeling performance of distributed components neglecting statistical uncertainty and correlation in component states other than through demands

#### 5.4 Examples of available fragility functions

Fragility curves have been developed for a vast range of infrastructure component types using varying degrees of sophistication relating to the treatment of uncertainties. Fragility models have been developed based on: (1) expert opinion or engineering judgment; (2) statistical analysis using data collected on observed performances of infrastructure system

components during past earthquakes; (3) simulation and modeling of component performance under earthquake loading, e.g. via finite element methods; and (4) a combination of physical and computer models as well as empirical information. Examples of these functions are described below. However, this is not an exhaustive list and updated models continually become available.

One of the most widely used sources for fragility curves is the HAZUS<sup>MH</sup> Technical Manual (DHS 2003). The document provides fragility functions, typically in the form of parameters  $\lambda_{j,k}$  and  $\zeta_{j,k}$  for the formulation in (5.8), for a large number of infrastructure types, including transportation systems (highway, air, and rail), utility systems (potable water, waste water, oil, natural gas, electric and communications), as well as building stocks. Information relating to statistical uncertainty and uncertainty in the functional form of the fragility models is generally not explicitly described. Because of the large number of infrastructures included in that document, the fragility functions are based primarily on engineering judgment. As a result, they are associated with a high degree of uncertainty. While the HAZUS<sup>MH</sup> Technical Manual is a good source of fragility functions for generic systems, the models for specific classes of infrastructure components that have been developed by researchers in discipline specific settings tend to be more accurate.

Researchers in bridge engineering have been the most prolific producers of fragility curves. As a result, the field has a wide range of models available including: older models developed using expert knowledge (e.g. Rojahn and Sharpe 1985); empirically based formulations (e.g. Basoz et al. 1999; Shinozuka et al. 2000); models based on structural analysis (e.g. Karim and Yamazaki 2003; Mackie and Stojadinovic 2004; Mackie and Stojadinovic 2006); and hybrid models that combine empirical and analysis-based methods (e.g. Gardoni et al. 2002; Kim and Shinozuka 2004). Highly specific models are available that provide, for example, the fragility of retrofitted structures (Padgett and DesRoches 2009), time-dependent fragility curves capturing the behavior of deteriorating structures (Gardoni and Rosowsky 2009; Sung and Su 2009), seismic fragility curves specifically for bridges in the central and southeastern United States (Nielson and DesRoches 2007), and bridges under ground motion with spatial variation (Kim and Feng 2003). While the majority of bridge fragility curves consider ground shaking as the hazard, models are also available that consider liquefaction and lateral spreading (Bowers 2007; Zhang et al. 2008).

Fragility curves for other components of transportation infrastructure are available, but in a more limited capacity, e.g. functions are available to estimate the mean number of postearthquake damages per kilometer of highway embankment (Mizuno et al. 2008) and for predicting the performance of rail viaducts (Kurian et al. 2006; Yoshikawa et al. 2007). However, fragility curves for transportation infrastructures, other than bridges, tend to be based on observations from a single earthquake or analyses of a single structure.

The American Lifelines Alliance provides fragility functions for components of water distribution systems including buried pipelines, water tanks, tunnels, and canals (Eidinger et al. 2001). The referenced document describes a variety of seismic hazards including ground shaking, liquefaction, landslide, and fault rupture. Additional studies are available

that model the performance of underground pipelines to ground shaking and/or fault rupture (Jacobson and Grigoriu 2005; Shih and Chang 2006; Toprak and Taskin 2007). The majority of studies related to seismic performance of pipelines are based on post-earthquake observations.

Fragility functions for components in electrical distribution systems are available from several authors including Vanzi (1996), who models the reliability of electrical networks by assessing the fragility of microcomponents and macrocomponents. Ang et al. (1996) develop fragility functions of critical components and substations. Straub & Der Kiureghian (2008) develop empirical fragility models of components in an electrical substation, while accounting for parameter uncertainties and statistical dependence among observations. Fragility models are also available for components of hydropower systems (Lin and Adams 2007) and nuclear power facilities (Kennedy and Ravindra 1984; Nakamura et al. 2010).

Fragility functions have also been developed for other types of infrastructure components such as oil storage facilities (Fabbrocino et al. 2005) or more general on-grade steel storage tanks used for storing a variety of liquids (O'Rourke and So 2000; Berahman and Behnamfar 2007). Furthermore, a vast literature exists relating to fragility curves for many classes of buildings (e.g. Ellingwood 2001; Ryu et al. 2008; Park et al. 2009; Rota et al. 2010).

# 5.5 Summary

This chapter presents BN formulations for modeling the performance of infrastructure components as a function of seismic demands. First, a generic model is presented for modeling component performance while accounting for potential sources of correlation. However, for many infrastructure types, seismic fragility research is not sufficiently mature, nor in possession of adequate data, to facilitate construction of the fragility models required to populate the CPTs of the generic formulation. As a result, simplified BN-based models for point-site and distributed components are developed that reflect the current-state of practice. This simplified model disregards statistical uncertainty in parameters of the fragility model and the correlation among component states that arises due to common modeling errors. The chapter ends with a brief description of fragility functions available for modeling the performance of components in a variety of infrastructure systems.

# Chapter 6: Modeling System Performance via BN

## 6.1 Introduction

This chapter presents BN formulations for modeling the performance of an engineered system as a function of the states of its constituent components. The chapter begins with a brief review of conventional methods for modeling system performance, including reliability block diagrams, fault trees, event trees, and minimal link and cut sets. These methods are then compared with the use of BNs for modeling system performance. Relative advantages are described. The chapter continues with the description of several methods for modeling system performance within the context of BNs. Emphasis is placed on the computational efficiency of these methods and algorithms are developed for minimizing computational demands. The chapter concludes with example applications.

### 6.2 Conventional methods of modeling system performance

A system is a collection of components, each having one or more states. The state of the system is defined as a function of the states of the individual components that comprise it. Critical infrastructure systems (e.g. transportation networks, water and power distribution systems) as well as structural systems (e.g. buildings and bridges) are all engineered systems for which it is important to know the system reliability. As described in previous chapters, these systems have statistically dependent component states due to several factors, including dependence in loading and in capacities.

A variety of methods have been developed to facilitate the performance modeling of engineered systems, all of which represent the system performance as a function of events

that affect its overall state. Several established methods of representing a system and its performance, both qualitatively and quantitatively, are briefly described below.

# 6.2.1 Qualitative methods

The qualitative systems analysis methods described below are typically used in organizational settings as means for the identification of failure modes, consequences, and remedial actions.

# 6.2.1.1 Failure modes and effects analysis

Failure modes and effects analysis is a technique employed for enumerating system failure modes. It is primarily an organizing tool in which a table is created with columns corresponding to, for example, system components that can fail, the corresponding component failure modes, the probability of occurrence for each mode, and the possible ways the governing organization can respond (Lewis 1987). The results of a failure models and effects analysis are often used to construct graphical models, such as fault trees (described in a subsequent section).

# 6.2.1.2 Reliability block diagrams

The reliability block diagram (RBD) is a graphical way of representing system topology. In a RBD, blocks represent system components or events and links represent component relationships. The blocks are arranged in series and parallel configurations to show logical interactions between components (Pagès and Gondran 1986). An example of a RBD is shown in Figure 6.1. In this RBD, a source and a sink node are connected by eight components. Components 1, 2, and 3 are shown in parallel, while components 4, 5, and 6 are constructed in series. Each of the components 1, 2, and 3 exists in series with components 7 and 8. Furthermore, components 4, 5, and 6 collectively exist in series with components 7 and 8. This example system is used throughout this chapter for illustration purposes.

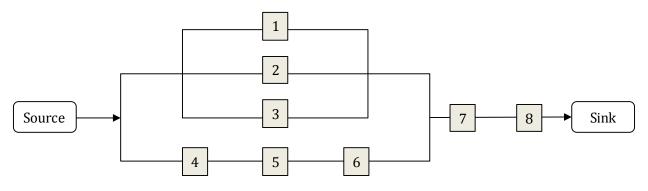


Figure 6.1: Example RBD

#### 6.2.1.3 Fault trees

A fault tree is a graphical model of events, which, when occurring individually (i.e. events in series) or in conjunction with other events (i.e. events in parallel), give rise to higher level events that eventually lead to the occurrence of system failure. A fault tree is constructed specifically for a particular undesired event or system failure mode. Therefore, a single fault tree cannot model all possible modes or causes of system failure. Fault trees are graphical models. They are first and foremost qualitative models, though quantitative analyses can be performed on them.

A fault tree consists of a *top event* corresponding to the undesired system event. *Gates* in a fault tree permit or inhibit the passage of fault logic up the tree. We limit ourselves here to consideration of two types of gates: *and gates* and *or gates*. An *and gate* indicates that all input failure events must occur in order for the output event to indicate failure. Conversely, an *or gate* signifies that the output event is in failure state if any of the input events is in the failure state (NRC 1981).

Fault trees are typically used to identify possible sequences of events leading to the occurrence of the undesired top event. They are used most often in operational settings, e.g. to enumerate sequences of events leading to core meltdown in a nuclear reactor. For example, Figure 6.2 shows a fault tree corresponding to the undesired event that a traffic signal at a dangerous intersection is not operational. It is assumed that there are two possible causes for the signal not to operate: loss of power *or* damage of the signal control box due to vehicle collision. Loss of power only occurs if off-site power supply *and* emergency back-up power supply are both lost. The emergency power supply system will fail if either the trigger which initiates the switch to back-up power fails *or* the emergency batteries are dead.

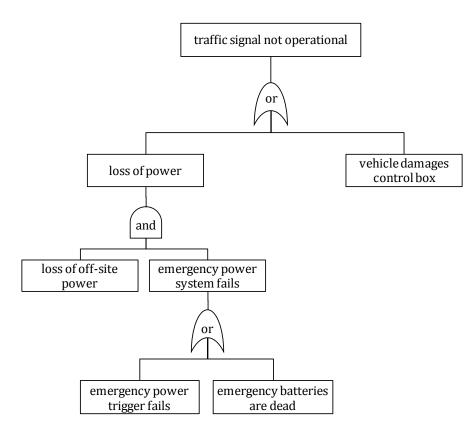


Figure 6.2: Example operational fault tree corresponding to traffic signal failure problem

It is uncommon to use fault trees to analyze topologically defined networks that are often found in civil engineering applications, e.g. transportation or water distribution systems. Such networks are often defined by a RBD. However, fault trees can be constructed from simple RBDs. A fault tree associated with the RBD in Figure 6.1 is shown in Figure 6.3. The circle nodes in this figure represent basic events which correspond to the failure of components  $C_1, ..., C_8$ . A bar over the component name denotes a failure event. For complex topological networks, direct construction of a fault tree from a RBD is difficult.

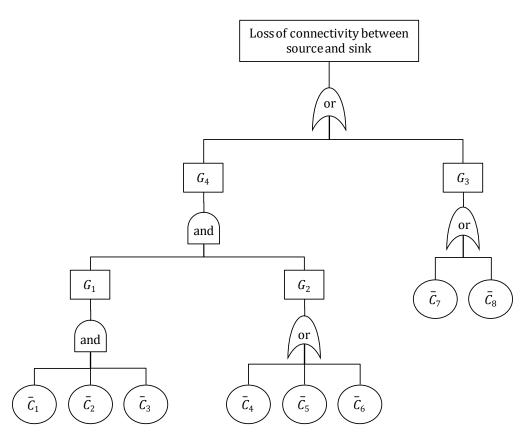


Figure 6.3: Fault tree corresponding to RBD in Figure 6.1

### 6.2.1.4 Event trees

An event tree is a graphical model that is used for inductive analyses. An event tree begins with an initiating event and then branches outward following possible progressions of subsequent events, e.g. failure or non-failure of other components (Pagès and Gondran 1986). Like fault trees, event trees are primarily qualitative tools useful for identification of failure sequences in operational settings. In the above example involving a traffic signal, one possible initiating event is the loss of off-site power. The corresponding event tree is shown in Figure 6.4.

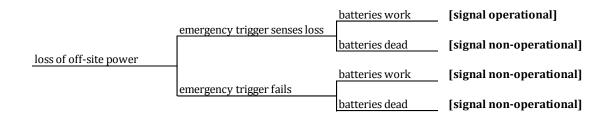


Figure 6.4: Example event tree corresponding to signal failure problem

# 6.2.2 Quantitative methods

Consider a system of *n* components. Each component  $i \in [1, n]$  has  $d_i$  discrete component states. Each system configuration represents a distinct combination of component states. Therefore the number of distinct configurations of the system is  $N = \prod_{i=1}^{n} d_i$ . Clearly, for systems with either a large number of components or components with many states, the number of distinct system configurations can be large. Define  $s_i$  as an indicator variable corresponding to the state of component *i*, with outcomes in the set  $\{1, 2, ..., d_i\}$ . Also define S as an indicator variable for the state of the system with outcomes in the set  $\{1, 2, ..., d_i\}$ , where *D* indicates the number of system states. The *N* distinct system configurations fall into the *D* distinct system states in accordance to the structure of the system. In general  $D \leq N$  and usually  $D \ll N$  (Der Kiureghian 2006).

Quantitative system analysis methods define a mapping  $S = f(s_1, \dots s_n)$ , known as the *system function*, that quantifies the relationship between the component and system states. Often the formulation of the system function is best done by individuals familiar with the disciplinary nature of a system. When the component states are random, and thus the system state is stochastic, the resulting quantitative analysis solves a *system reliability* problem (Der Kiureghian 2006). In such problems, the objective of quantitative analysis is to determine a system failure probability or, complementarily, the system reliability. When both the components and the system have binary states (e.g. on/off, open/close), the result is a special case of the system reliability problem in which  $N = 2^n$  and  $s_i$  and S are Boolean random variables defined such that  $S, s_i \in \{0,1\}$ . We use the convention that 0 and 1 respectively denote the failure (false) and survival (true) states.

# 6.2.2.1 Minimal link sets and cut sets

The analysis of systems with binary component and system states is often quantified through the use of *link sets* and *cut sets*. A link set is a set of components whose joint survival ensures survival of the system. Conversely, a cut set is a set of components whose joint failure constitutes system failure. A *minimal (or minimum) link set* (also referred to as a *minimal path set*) is a link set that ceases to be a link set if any component is removed from it. Similarly, a *minimal (or minimum) cut set* is a cut set containing no superfluous components.

The system failure probability is specified in terms of the minimal cut sets (MCSs) or the minimal link sets (MLSs). Using MLSs, the system failure probability is written as

$$P_f = 1 - \Pr\left(\bigcup_{1 \le i \le N_{MLS}} \left(\bigcap_{j \in MLS_i} \{s_j = 1\}\right)\right)$$
(6.1)

where  $N_{MLS}$  is the number of MLSs of the system and  $MLS_i$ ,  $i = 1, ..., N_{MLS}$ , denotes the set of indices corresponding to the components in the *i*-th MLS. It follows from (6.1) that each MLS is a series system of its components, i.e. all components must survive for the MLS to survive. The system state is a parallel system of its MLSs, i.e. the system survives if any MLS survives.

In terms of the MCSs, the system failure probability is written as

$$P_f = \Pr\left(\bigcup_{1 \le i \le N_{MCS}} \left(\bigcap_{j \in MCS_i} \{s_j = 0\}\right)\right)$$
(6.2)

where  $N_{MCS}$  is the number of MCSs of the system and  $MCS_i$ ,  $i = 1, ..., N_{MCS}$ , denotes the component indices of the *i*-th MCS. Each MCS is a parallel system of its components and the system state is a series system of the MCSs. The enumeration of MCSs (or MLSs) is an essential part of conventional quantitative systems analysis. Qualitatively, MCSs may be interpreted as weak points in the system and sources of false redundancy (Pagès and Gondran 1986). A variety of algorithms are available for determining system MCSs/MLSs. These include methods ranging from inductive analysis to approaches utilizing simulation (Shin and Koh 1998; Suh and Chang 2000; Yeh 2006). MCSs can also be determined directly from fault trees (Pagès and Gondran 1986; Fard 1997).

### 6.2.2.2 Additional methods

While the use of MLSs and MCSs is one of the most common methods for quantifying system performance, other approaches do exist. Such methods include definition of Boolean *operating* and *failure functions* as well as construction of *state transition (Markov) diagrams*. Additional details on these less widely employed methods can be found in Pagès and Gondran (1986).

### 6.3 Existing works using BNs to model system performance

Work related to the use of BNs for modeling system performance is limited. The majority of work in this area relates to the conversion of conventional qualitative system modeling techniques into BNs. For example, Torres-Toledano and Sucar (1998) address converting RBDs into BNs for simple system configurations. However, they only consider simple RBDs and do not address issues related to computational efficiency or tractability, nor do they

provide specific algorithms for automating the construction of BNs from RBDs. Doguc et al. (2009) provide a generic algorithm for learning the structure of a system performance BN automatically from raw system behavior data using an algorithm that is efficient for complete systems and large data sets. The authors seek to eliminate the need for expert interference by constructing BNs automatically. However, for many civil engineering systems, data is sparse and the BNs resulting from such an algorithm may thus be inefficient, intractable, or inaccurate.

A fault tree may be considered a deterministic special case of a BN (Lampis and Andrews 2009). Because of the advantages of using BNs over fault trees, several authors have described algorithms for generating BNs from fault trees (Bobbio et al. 2001; Liu et al. 2008; Lampis and Andrews 2009). While the algorithms differ slightly from author to author, they generally consist of the following steps: (1) converting basic (bottom) events in the fault tree into root (top) nodes in the BN; (2) expressing logical gates in the fault tree as nodes in the BN; (3) assigning links in the BN consistent with the input-output relationships in the fault tree; and (4) assigning CPTs consistent with the logic gates in the fault tree 6.5.

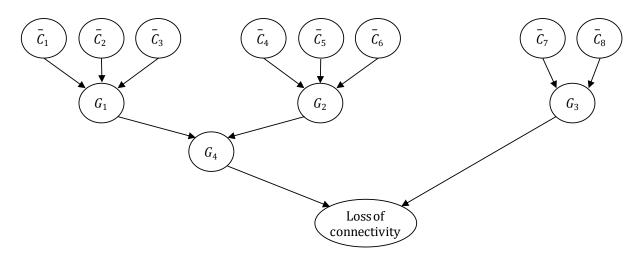


Figure 6.5: BN corresponding to fault tree in Figure 6.3

The CPTs of the nodes  $G_i$  are constructed using the logical relations

$$\bar{G}_1 = \bigcap_{i=1}^3 \bar{C}_i$$

$$\bar{G}_2 = \bigcup_{i=4}^6 \bar{C}_i$$

$$\bar{G}_3 = \bar{C}_7 \cup \bar{C}_8$$

$$\bar{G}_4 = \bar{G}_1 \cap \bar{G}_2$$

$$LC = \bar{G}_4 \cup \bar{G}_3$$
(6.3)

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where  $G_i$  is the event that the output event from the gate is in the survival state and  $\overline{G}_i$  indicates its complement, the failure state. *LC* represents the event of loss of connectivity between the source and sink.

While both BNs and fault trees are graphical models and are useful for representing systems, BNs have capabilities that fault trees do not have:

- (1) In BNs, the dependence relations between components are not restricted to be deterministic and BNs are not limited to including only binary components, as is the case when working with fault trees (Bobbio et al. 2001).
- (2) The BN framework enables the modeling of complex dependences among the performance of individual components.
- (3) BNs permit information to be entered at any node and this evidence propagates throughout the BN.
- (4) BNs avoid repeating nodes associated with common causes (Mahadevan et al. 2001).

# 6.4 BN formulations for modeling system performance

In this section, five formulations are presented for modeling system performance using BNs. For brevity and simplicity, we focus first on systems with binary component and system states. The components of such a system have two states and the system itself also has two states, e.g. either connectivity between the source and sink nodes exists or does not exist. The analysis for this class of systems is obviously simpler than when the components or the system have multiple states. The BN methodology is in not restricted to systems with binary component states and consideration of the multi-state problem is presented in subsequent sections. We initially introduce five methods for modeling system performance: (1) naïve formulation; (2) explicit connectivity formulation; (3) explicit disconnectivity formulation; (4) MLS formulation; and (5) MCS formulation. Later, we adapt formulations (4) and (5) with the goal of optimizing computational efficiency.

### 6.4.1 Naïve formulation

Consider a system of *n* components. As described previously, each component  $i \in [1, n]$  has  $d_i$  discrete component states. Therefore, the number of distinct configurations of the system is  $\prod_{i=1}^{n} d_i$ . We refer to the BN formulation that corresponds to the joint consideration of all combinations of component and system states as the *naïve BN* formulation. In this formulation, a single node  $S_{sys}$ , representing the state of the system, is defined as a child of all nodes representing the states of the components forming a converging structure, as shown in Figure 6.6. For a system with *n* binary components and binary system state, node  $S_{sys}$  has a CPT of size  $2^{n+1}$ . For systems with a large number of components, the size of the CPT of node  $S_{sys}$  quickly causes the BN to become the number of components is small, e.g. in the reliability assessment of simple structural systems. Mahadevan et al. (2001) utilize a formulation as in Figure 6.6 with components

representing limit-state functions corresponding to the failure modes of a structural system. The number of limit-state functions in their example is sufficiently small so that the size of the CPT associated with the system node is not prohibitively large. However, for realistic infrastructural systems, such an approach to modeling is impractical. Hence, while easy to formulate, this is not a pragmatic approach for many real-world applications, where the number of system components is often large.

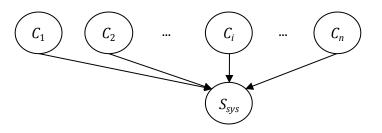


Figure 6.6: Naïve BN system formulation

Consider the example system represented by the RBD in Figure 6.1. In this system, eight components (squares) connect the system source and sink. It is assumed that only the square components in the system can fail. The required system performance is connectivity between the source and sink nodes. The naïve BN formulation for this system is shown in Figure 6.7. For binary components, the state of the system must be specified for each combination of the states of the 8 components. Thus, the size of the CPT associated with node  $S_{sys}$  is  $2^9 = 512$ . Assuming each member of the CPT takes up 8 bytes of memory, the CPT requires 4 kilobytes of computer memory to store the table. While this is not a substantial amount of memory, if the number of states associated with each node is increased from 2 to 10, then the size of the CPT is 10<sup>9</sup> and amount of memory required to store the CPT in the computer increases to over 7600 megabytes (7.6 gigabytes). Adding just one more component node to the BN increases the memory requirement to  $10^{10} * 8$ bytes (nearly 75 gigabytes) simply to hold the CPT in the computer memory. It is clear that the naïve formulation is computationally impractical when either the number of states associated with each node is large or the system has many components. The total clique table size associated with performing inference on the BN in Figure 6.7 is 512 for binary components. When 10 states are considered per node, the total clique table size is  $10^{10}$  and it is then not possible to perform inference using standard exact inference algorithms on a workstation with 32 GB of RAM.

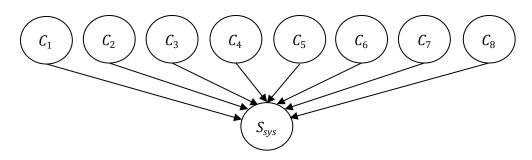


Figure 6.7: Naïve formulation for example system

# 6.4.2 Explicit formulations

Two intuitive approaches for modeling system performance via BN are next developed, which are referred to as the *explicit connectivity* and *explicit disconnectivity* formulations. These explicit formulations use a casual approach, in which connectivity paths or disconnectivity sequences are constructed manually based on the topology of the system. The explicit formulations, while not systematic in their construction, are designed to facilitate interaction with third-party system owners/operators; their intuitive nature allows for validation by interested parties familiar with a particular system but not with quantitative systems analysis techniques or BNs.

## 6.4.2.1 Explicit Connectivity formulation

Friis-Hansen (2004) proposed a BN approach to modeling system performance, in which system connectivity is modeled by exploiting causal relationships between the system components necessary for its survival. The *explicit connectivity* (EC) *BN formulation* defined in this study is a formalization of an approach to modeling system performance advocated by Friis-Hansen (2004). In promoting an intuitive approach to modeling system performance, Friis-Hansen notes that, when modeling complex systems, it is of utmost importance that the modeling can be validated by third parties. Therefore, the modeling approach should focus on transparency. Causal modeling helps assure this is possible. As a result, the EC BN formulation expresses system connectivity using a causal interpretation of the connectivity paths. We think of it as the optimist's formulation (in contrast to the explicit disconnectivity formulation, which is viewed as a pessimist's formulation), because one models the system by directly defining paths that ensure *survival* of the system. This approach does not formally require identification of MLSs, though the causal logic indirectly employs them. Due to the encoded causal relationships, the resulting BN is likely to produce CPTs that are smaller than those of the naïve formulation. As a result, the formulation is typically superior to the naïve formulation in terms of its computational efficiency.

The EC formulated BN must be constructed specifically for each system of interest. Because the construction of the explicit formulation is not systematic, accuracy is not guaranteed,

particularly when working with for large and/or complex systems. Furthermore, there is no formal control on the computational demands associated with performing inference on the resulting BN. To illustrate the construction of a BN using the EC formulation, Figure 6.8 shows two possible BNs that may be constructed using this approach for the RBD shown in Figure 6.1. The BN formulation in Figure 6.8a distinguishes two paths through the system: an upper path and a lower path, as shown in Figure 6.9.

Let  $C_i$  be the event that component *i* is in the survival state. The upper path is open only if components 7 and 8 are functioning and at least one of components 1, 2, and 3 is also working. The Boolean logic corresponding to this causal interpretation is used to construct the CPT of the node labeled upper path in Figure 6.8a. The event in which the upper path is operational (denoted *UP*) is expressed as

$$UP = C_7 \cap C_8 \cap (\bigcup_{i=1}^3 C_i)$$
(6.4)

Similarly, the CPT attached to the node labeled *lower path* is constructed using the Boolean logic:

$$LP = C_7 \cap C_8 \cap (\bigcap_{i=4}^6 C_i)$$
(6.5)

where *LP* is the event in which the lower path is functioning. Connectivity between the source and sink is maintained if either path is functioning. Letting *Sink* be the event in which there is connectivity to the sink, the CPT attached to the corresponding node is expressed using the logic

$$Sink = UP \cup LP \tag{6.6}$$

Alternatively, one can model the system bottleneck explicitly while distinguishing between the upper and lower paths up to the bottleneck. The associated BN is shown in Figure 6.8b. The CPT for node *upper path* is constructed using the logic:  $UP = \bigcup_{i=1}^{3} C_i$ . The CPT for node *lower path* is similarly constructed:  $LP = \bigcap_{i=4}^{6} C_i$ . The node *bottleneck* is constructed using a causal interpretation of the connectivity paths: one is only able to traverse the bottleneck only if one is first able to reach the bottleneck through either the lower or upper path *and* both components 7 and 8 are functioning. The CPT for node *bottleneck* is thus constructed:

$$BT = (UP \cup LP) \cap C_7 \cap C_8 \tag{6.7}$$

where *BT* corresponds to the event in which one is able to traverse the system bottleneck. Relative to the naïve formulation, the EC formulation is advantageous because it facilitates understanding of system topology, whereas the naïve formulation aggregates all system states into a single node. Furthermore, the EC formulation is typically associated with significantly smaller CPTs, and thus lower memory demands, than the naïve formulation.

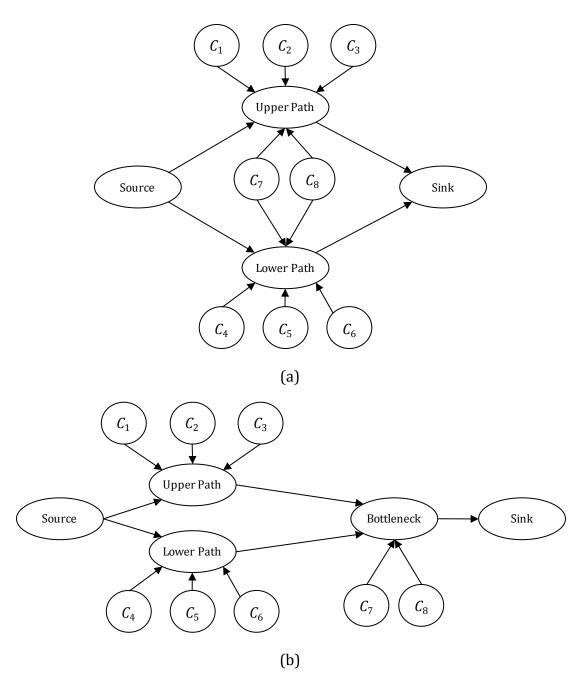


Figure 6.8: Examples of possible EC formulations for the system in Figure 6.1

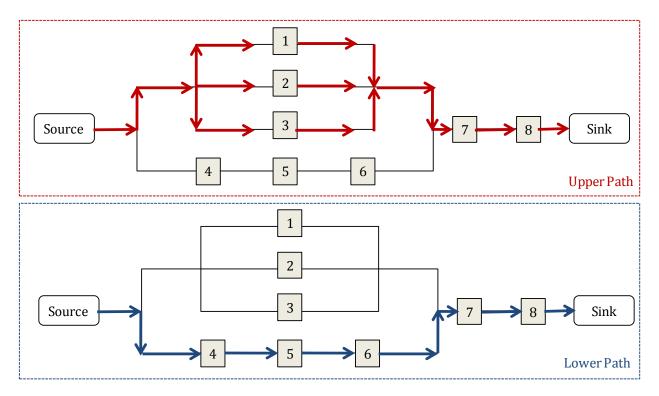


Figure 6.9: Illustration of upper and lower paths through the example system, used for construction of EC BN formulations in Figure 6.1

## 6.4.2.2 Explicit Disconnectivity formulation

The dual of the EC formulation is defined as the *explicit disconnectivity* (EDC) *BN formulation*. Rather than tracing paths that ensure survival of the system, one pursues causal event paths that ensure *failure* of the system. This is a less intuitive approach than the EC formulation as it follows a pessimist's perspective. Similar to the EC formulation, using the EDC formulation, one can often improve upon the naïve formulation, particularly when the number of failure event paths (or MCSs) are small relative to the number of components. The EDC formulation does not explicitly require the MCSs, though the causal logic indirectly employs them. Though less intuitive for many applications, the EDC formulation may be instinctual for situations in which the construction of fault and event trees is common, because these models are constructed by enumerating failure sequences.

Once again, consider the example system in Figure 6.1. Working backward from the sink, it is observed that disconnectivity will occur if the bottleneck is blocked, or if connectivity is blocked further "upstream." The upstream connectivity is lost only if both the upper and lower paths through the system up to the bottleneck are closed. The lower path is not operational if any component along it fails. The upper path losses connectivity only if all components fail. The above description leads to the BN in Figure 6.10. This formulation has

the same topology and conditional relationships as the BN in Figure 6.5, which was created directly from a fault tree.

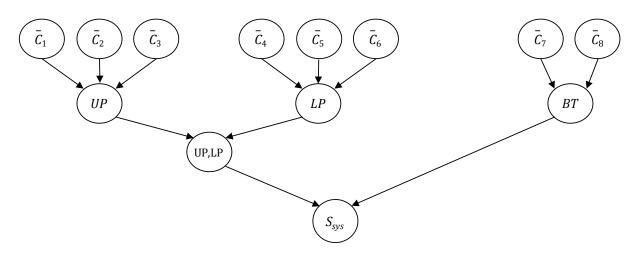


Figure 6.10: Example EDC BN formulation for system in Figure 6.1

Like the EC formulation, the EDC formulation is more computationally efficient than the naïve formulation due to the generally smaller CPT sizes. Furthermore, it likewise lends itself to easier interpretation than the naïve formulation.

# 6.4.3 Minimal link set and minimal cut set formulations

As described previously, a common approach to making systems analysis more methodical is through the use of minimal link sets and minimal cut sets. Recall that a minimal link set (MLS) is a minimum set of components whose joint survival constitutes survival of the system. The *minimal link set BN formulation* introduces intermediate nodes between the component and system nodes that correspond to the MLSs. Analogously, the *minimal cut set BN formulation* introduces intermediate nodes between the system. Torres-Toledano and Sucar (1998) use a MLS-based BN formulation for modeling system performance, though with less formality and generality than is described here. In the next two sections, the MLS and MCS BN formulations are described.

# 6.4.3.1 Minimal link set formulation

As described above, the MLS BN formulation introduces intermediate nodes corresponding to the system MLSs between the nodes representing component and system states. Figure 6.11 shows an example of the MLS BN formulation. The binary states of the MLS nodes are defined such that each MLS node is in the survival state only if all its constituent components have survived; otherwise it is in the failure state. The system node is in the survival state if any MLS node is in the survival state. The MLS BN formulation takes advantage of the fact that each MLS node is a series system of its components, and that the system node is a parallel system of its MLS parents. Clearly, this formulation is advantageous to the naïve formulation described above, particularly when the system has fewer MLSs than components. With binary component and system states, the size of the CPT for each MLS node is 2 to the power of the number of its constituent components plus one, and the size of the system node CPT is 2 to the power of the number of MLSs plus one. As a result, when the number of MLSs is large, the size of the CPT associated with the system node  $S_{sys}$  becomes large. A similar problem occurs for a MLS node when the number of its constituent components is large. To address the exponential growth in computational demands, the MLS formulation described here will be adapted in later sections to address issues related to computational efficiency.

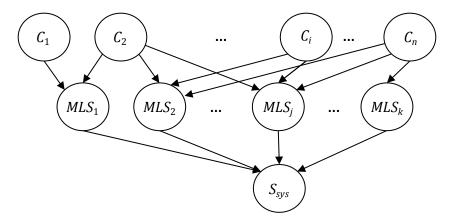


Figure 6.11: System performance BN using the MLS BN formulation

The example system in Figure 6.1 has four MLSs:  $MLS_1 = \{1,7,8\}$ ;  $MLS_2 = \{2,7,8\}$ ;  $MLS_3 = \{3,7,8\}$ , and  $MLS_4 = \{4,5,6,7,8\}$ . The MLS formulated BN for the example system is shown in Figure 6.12. For this example, the largest CPT occurs at node  $MLS_4$  and has size  $2^6$ . The total memory requirement associated with performing inference in this BN, measured as the total clique table size, is 208, assuming independent component states.

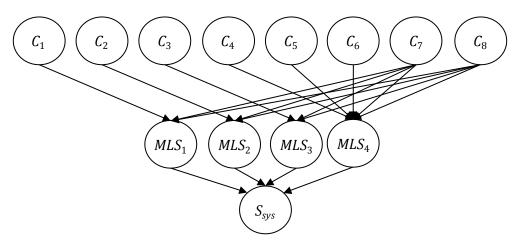


Figure 6.12: MLS formulation for example system in Figure 6.1

A reduction in the CPT size can be achieved through introduction of additional intermediate nodes between the component and MLS nodes and/or between the MLS and the system nodes. A common "rule of thumb" when working with BNs suggests that, wherever possible, the number of parents to a node should be no more than three. For the BN in Figure 6.12, nodes  $MLS_4$  and  $S_{sys}$  have more than three parents. Intermediate nodes are introduced into this BN to reduce the number of parents associated with any node to below three, as shown in Figure 6.13.

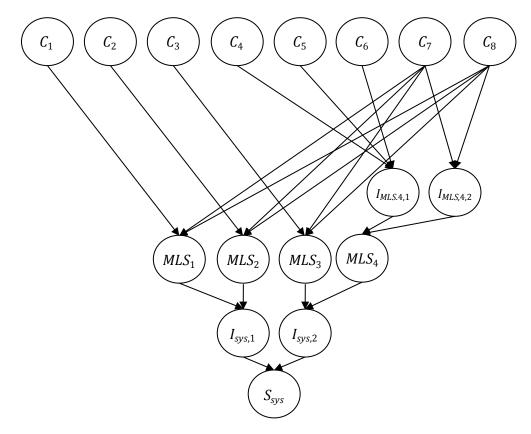


Figure 6.13: MLS BN formulation for example system in Figure 6.1, with intermediate nodes introduced to keep the number of parents to a node to no more than three

The CPTs of the intermediate nodes  $I_{MLS,4,i}$ , i = 1,2, are constructed such that they are in survival state only if all their parent nodes are in the survival state. The CPT of the child to the intermediate nodes is constructed similarly. The CPTs of nodes  $I_{sys,i}$ , i = 1,2, are defined such that they assume survival state if any of their parent MLS nodes is in the survival state. Node  $S_{sys}$  is in survival state if any node  $I_{sys,i}$  is in survival state. The introduction of additional intermediate nodes in the BN decreases the size of the largest CPT to  $2^4$  and the total clique table size is reduced to 176. While a small computational advantage results from the introduction of intermediate nodes, the increase in efficiency is not significant. In later sections, this MLS formulation, which will be referred to as the standard MLS (MCS) BN formulation, is adapted to increase computational efficiency.

### 6.4.3.2 Minimal cut set formulation

We define the dual of the MLS formulation as the *minimal cut set BN formulation*. In this formulation, the system node is a child of parents corresponding to MCSs, and each MCS node is a child of nodes representing its constituent components. Figure 6.14 shows a conceptual BN employing the MCS formulation. The system node is a series system of all the MCS nodes, whereas each MCS is a parallel system of its parent nodes.

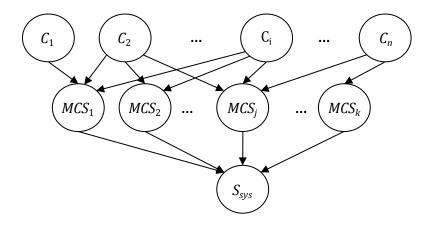


Figure 6.14: System performance using the MCS BN formulation

The example system in Figure 6.1 has 5 MCSs:  $MCS_1 = \{1,2,3,4\}; MCS_2 = \{1,2,3,5\}; MCS_3 = \{1,2,3,6\}; MCS_4 = \{7\}, and MCS_5 = \{8\}$ . The MCS BN for the example system is shown in Figure 6.15. For this BN formulation, the largest CPT is associated with the system node and it has size  $2^6$ . The total clique table size for this BN, assuming independent binary components, is 232. Introducing intermediate nodes between the MCS and system node so no node has more than three parents (as was done for the MLS formulation shown in Figure 6.13) reduces the total clique table size to 200.

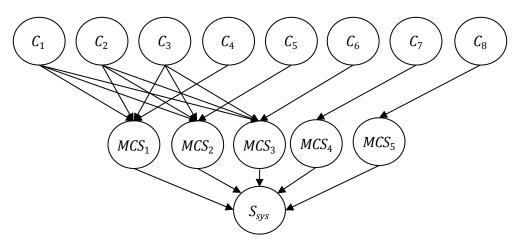


Figure 6.15: MCS formulation for example system in Figure 6.1

Generally, relative to the MLS formulation, the MCS formulation is advantageous when the number of MCSs is smaller than the number of MLSs. However, their relative advantage also depends on the number of components within the individual MLSs and MCSs. As was encountered with the MLS formulation, the CPTs in the MCS formulation become large as the number of MCSs increases and/or when the number of components in individual MCSs grows large.

## 6.4.4 Extension to multi-state problems

Thus far, the BN formulations described have only considered binary-state components and systems. Multi-state problems are easily handled within the above MCS BN formulation through application of the Max-Flow Min-Cut theorem (Elias et al. 1956; Ford and Fulkerson 1956). We are not aware of any similar theories that allow a direct adaptation of the MLS formulation to multi-state problems. The application of the Max-Flow Min-Cut theorem to the MCS BN formulation is described next.

Let  $Cap_i$  denote the operating capacity of component *i*. Assign to each MCS a *value* by taking the sum of the capacities of the components in the MCS. For distributed components, one should only count the capacities going from the source side to the sink side of the cut. The Max-Flow Min-Cut theorem states that the maximum possible flow from source to sink is equal to the minimum value among all MCSs, i.e. the bottleneck in the system. The theorem allows adaptation of the MCS formulation to multi-state problems, without changing the topology of the BN created when assuming components are binary. It is only necessary to: (1) increase the number of states associated with each node to correspond to the component operating levels; and (2) use arithmetic expressions rather than Boolean logic to define the relationships between the nodes.

Once again, consider the RBD in Figure 6.1 with the MCS BN formulation in Figure 6.15. Nodes  $C_i$  are modified to contain states corresponding to operating levels of the components. If the component has flow characteristics, then the range of possible flow levels must be discretized; in such a case, node  $C_i$  will be an interval node. Similarly, nodes  $MCS_i$  must have multiple states corresponding to the values of each MCS with CPTs defined using the relation

$$Value(MCS_i) = \sum_{C_j \in MCS_i} Cap_j$$
(6.8)

Note that, because the value of  $MCS_i$  is an additive function of the capacities, the states of the system node must be modified to contain all possible system operating levels ( $Cap_{sys}$ ) with CPTs defined using the relation

$$Cap_{sys} = \min_{\text{all } MCS_i} Value(MCS_i)$$
(6.9)

### 6.5 Efficient minimal link set and minimal cut set formulations

Due to the potentially large computational demands associated with the MLS and MCS BN formulations when either the number of components belonging to a MLS/MCS or the number of MLS/MCSs is large, alternative MLS and MCS formulations are next proposed. The alternative formulations take advantage of the observation that BNs with nodes arranged in chain structures are generally significantly more efficient than the converging structures characterizing the naïve and standard MLS/MCS formulations. Consider the two BN topological structures shown in Figure 6.16, in which the figure on the left shows a converging structure and the one on the right illustrates a chain structure. Both BNs model systems whose components are correlated by a common demand, *D*. A formal description of the construction of the BN in Figure 6.16b and its CPTs is reserved for later sections.

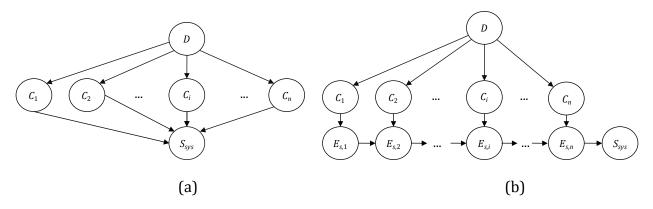


Figure 6.16: Illustration of BNs with (a) converging structure and (b) chain structure

Figure 6.17 shows a comparison of the computational demands, measured in terms of the total clique table sizes, associated with binary and five-state components when the system analysis BNs are arranged in converging and chain structures as in Figure 6.16. It is observed in Figure 6.17 that the converging structure is associated with exponentially increasing memory demands, while the complexity of the chain structure grows linearly. However, the converging structure is more efficient than the chain structure when less than four components are considered. Thus, when the node modeling system performance in Figure 6.16a has more than 3 parents, it is advantageous to model the system as in Figure 6.16b.

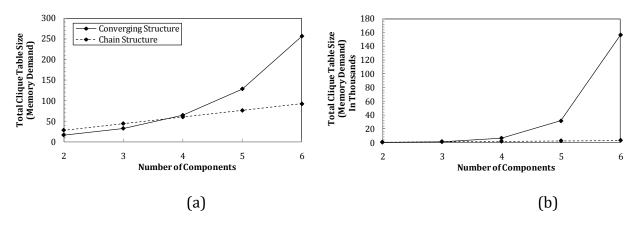


Figure 6.17: Comparison of computational demands associated with converging and chain BN topologies when components have (a) two-states and (b) five states

The computational demands associated with inference are influenced not only by the system configuration, but also by the number of common parents to the component nodes in Figure 6.16. Figure 6.19 shows a comparison of computational demands associated with the converging and chain topologies for binary component nodes with 1, 2, and 3 common parent demand nodes. It is observed that increasing the number of common parent nodes increases the computational demands. However, the chain structure remains advantageous to the converging structure as the number of components increases. Note that, as the number of common demand nodes increases, the "cut-off" point at which the converging structure is better than the chain structure moves upward slightly. For three common demand nodes, the converging structure is more efficient than the chain structure when considering up to 4 components. For one common demand node, the converging structure is more efficient for up to three components.

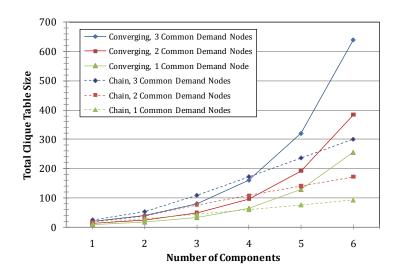


Figure 6.18: Comparison of computational demands associated with converging and chain BN topologies for binary nodes with one or more common parent demand nodes

### 6.5.1 Construction of efficient BN formulations for series and parallel systems

In this section efficient chain-structure BN formulations are constructed for series and parallel systems with binary states. In subsequent sections, the approach is extended to more general system topologies. Beginning with these simple system types serves to illustrate the construction of the BN formulation without the increased conceptual complexity associated with more general systems.

Define a *survival path sequence* (SPS) as a chain of events, corresponding to a MLS, in which the terminal event in the sequence indicates whether or not all the components in the MLS are in the survival state. Note that the term "sequence" does not have any time-based implications. A series system has one MLS and a parallel system has *n* MLSs. It follows that a series system has one SPS and a parallel system has *n* SPSs. A SPS is comprised of a chain of *survival path events* (SPEs), each of which describe the state of the sequence up to that event. SPEs are represented in the BN by nodes labeled  $E_{s,i}$ ; the subscript *i* indicates that the particular SPE is associated with component *i*. The state of  $E_{s,i}$  is defined as

$$E_{s,i} = 1 \text{ if } E_{s,Pa(i)} = 1 \cap C_i = 1$$
  
= 0 otherwise (6.10)

where  $E_{s,Pa(i)}$  defines the state of the SPE node that is parent to  $E_{s,i}$ ;  $E_{s,i} = 1$  indicates that the node is in the survival state and  $E_{s,i} = 0$  indicates its failure.  $C_i$  denotes the state of component *i* with  $C_i = 1$  ( $C_i = 0$ ) indicating the survival (failure) state. Thus, for a series system, the BN formulation takes the form shown in Figure 6.19. The state of node  $E_{s,1}$  is equal to the state of node  $C_1$ .  $E_{s,2}$  is in the survival state only if  $E_{s,1}$  is in the survival state and  $C_2$  is in the survival state. This pattern continues such that  $E_{s,n}$  is in the survival state only if both  $E_{s,n-1}$  and  $C_n$  are in the survival state. Consequently, the state of  $E_{s,n}$  describes the state of the entire SPS (i.e. it indicates whether all components in the MLS have survived) and, therefore, that of the system.

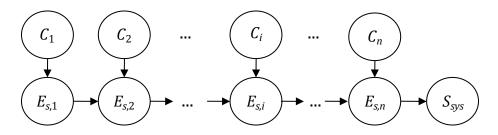


Figure 6.19: BN using SPEs to define performance of a series system

A parallel system has a SPS corresponding to each component. The resulting BN formulation is shown in Figure 6.20. The system node indicates system survival if any node  $E_{s,i}$  is in the survival state. Like the naïve formulation, the exponential growth in the size of the CPT associated with node  $S_{sys}$  renders this BN intractable when the number of components in the system is large.

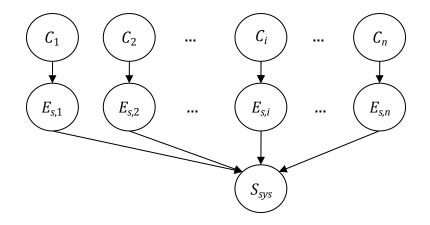


Figure 6.20: BN using SPEs to define performance of a parallel system

Define a *failure path sequence* (FPS) as a chain of events, corresponding to a MCS, in which the terminal event in the sequence indicates whether or not all the components in the MCS are in the failure state. For a parallel system, there is only one MCS and thus one FPS. For a series system with *n* components, there are *n* FPSs, one corresponding to each component. A FPS is comprised of a chain of *failure path events* (FPEs), each of which gives the state of the sequence up to that event. Let  $E_{f,i}$  be the FPE associated with component *i*. The state of  $E_{f,i}$  is expressed as

$$E_{f,i} = 0 \text{ if } E_{f,Pa(i)} = 0 \ \cap \ C_i = 0 \tag{6.11}$$

#### = 1 otherwise

where  $E_{f,Pa(i)}$  defines the state of the FPE node that is parent to  $E_{f,i}$ ;  $E_{f,i} = 0$  indicates that the node is in the failure state and  $E_{f,i} = 1$  indicates its survival. Thus, for a parallel system, the BN formulation takes the chain form shown in Figure 6.21. The Boolean logic used to construct the CPTs in this BN is the dual of that used for SPSs, i.e.  $E_{f,i}$  is in the failure state only if the parent FPE is in the failure state and  $C_i$  is also in the failure state. For a series system, the BN formulation using FPSs is shown in Figure 6.22. The size of the CPT associated with  $S_{sys}$  is  $2^{n+1}$  and there is no computational advantage to this approach over the naïve formulation for series systems. These findings suggest that a combination of SPS and FPS formulations can be used to efficiently model general systems. This approach is described in the next section.

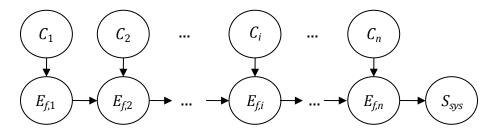


Figure 6.21: BN using FPEs to define a parallel system

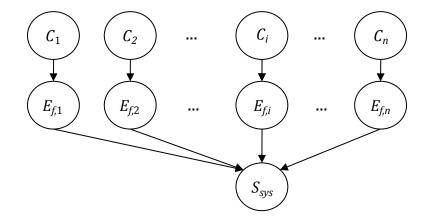


Figure 6.22: BN using FPEs to define a series system

#### 6.5.2 General systems

Next, efficient MLS and MCS formulations are developed for general system topologies.

### 6.5.2.1 Efficient MLS formulation

As described earlier, a MLS is a series system of its constituent components. Therefore, based on the above discussion, one can construct a SPS to describe each MLS. Again, consider the example system in Figure 6.1, which has four MLSs:  $MLS_1 = \{1, 7, 8\}$ ,  $MLS_2 = \{2,7,8\}, MLS_3 = \{3,7,8\}$  and  $MLS_4 = \{4,5,6,7,8\}$ . In Figure 6.23 each MLS is modeled as an individual SPS. The SPEs,  $E_{s,i}^{j}$ , in each SPS are indexed by a subscript corresponding to the associated component *i* and a superscript corresponding to SPS/MLS number *j*. The dependence between the SPEs corresponding to the same component is modeled through a common parent node. The system node is in the survival state if the terminal node of any SPS is in the survival state. For reference, the BN formulation in which the MLSs are arranged in chain structures is named the efficient MLS BN formulation. The total clique table size associated with the BN in Figure 6.23 is 224. This is higher than the demands required by the standard MLS formulation (208). This increase arises because MLSs 1-3 only have three components, and therefore it is less efficient to model them using a chain structure than a converging structure, as indicated by the graphs in Figure 6.17. Furthermore, the dependence between SPEs sharing a component increases the computational demand when performing inference in the BN. By coalescing common SPEs that appear in multiple SPSs, the number of nodes and links in the BN, and hence the computational demand, are reduced. In the example system, components 7 and 8 appear in all SPSs. We take advantage of this observation and introduce only one "instance" of the SPEs associated with these components. The resulting BN is shown in Figure 6.24. The states of SPE nodes having multiple SPEs as parents (e.g. node  $E_{s,7}$  in Figure 6.24) are specified using the Boolean relation

$$E_{s,i} = 1 \text{ if } \left[ \cup \{ E_{s, \text{Pa}(i)} = 1 \} \right] \cap C_i = 1$$
  
= 0 otherwise (6.12)

A notational change has been introduced in Figure 6.24: the superscript associated with each SPE node, which previously indicated the MLS/SPS number, now represents the *instance* of the SPE, i.e. if multiple SPEs are associated with the same component, then they are recognized as different instances of the SPE and are distinguished through the superscript. Because, for this system, each component is associated with only one SPE, all superscripts in Figure 6.24 are 1. The total clique table size associated with the BN in Figure 6.24 is 108, which is smaller than the demand associated with any other formulation.

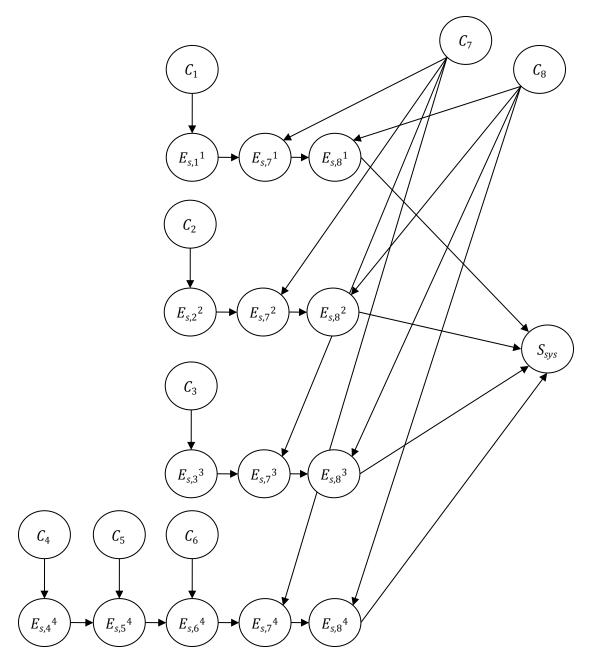


Figure 6.23: Efficient MLS BN formulation for system in Figure 6.1 with distinct SPSs

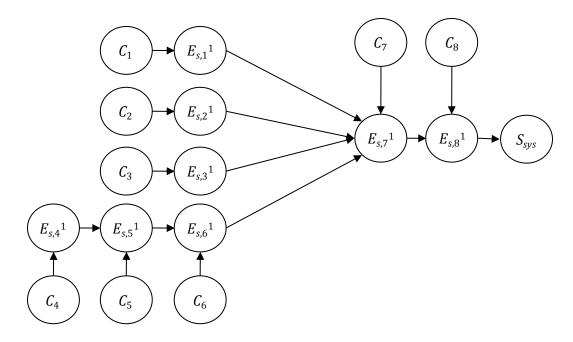


Figure 6.24: Efficient MLS BN formulations for the example system with coalesced SPEs associated with components 7 and 8 using a converging structure

It is noted that node  $E_{s,7}$  in Figure 6.24 has more than 3 parents. Earlier, it was indicated that chain structures are more efficient than converging structures when the number of parents is greater than 3. Thus, the BN in Figure 6.24 is further modified by replacing the parallel SPE nodes associated with components 1, 2, 3, and 6 with nodes arranged in a chain, resulting in the BN in Figure 6.25, with CPTs defined using the relation

$$E_{s,i} = 1 \text{ if } \left[ (C_i = 1) \cap \left( \cup \{ E_{s,Pa(i)'} = 1 \} \right) \right] \cup \{ E_{s,Pa(i)''} = 1 \}$$
  
= 0 otherwise (6.13)

where  $E_{s,Pa(i)'}$  are the SPE nodes that are parent to  $E_{s,i}$  before the addition of the chain modification and which remain parents after it;  $E_{s,Pa(i)''}$  are the SPE nodes that become parents to  $E_{s,i}$  after the chain structure is added (identified by dashed links in Figure 6.25). The total clique table size associated with this BN is 64.

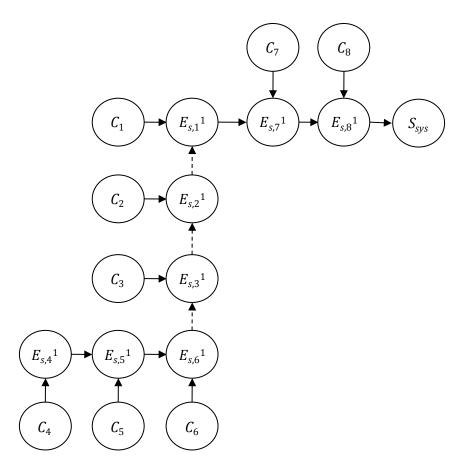


Figure 6.25: Efficient MLS BN formulations for the example system with coalesced SPEs associated with components 7 and 8 using a chain structure

## 6.5.2.2 Efficient MCS formulation

The efficient MCS formulation is constructed in a manner similar to the efficient MLS formulation. In the efficient MCS formulation, strings of FPSs are constructed corresponding to each MCS. To illustrate the construction of the efficient MCS formulation, return to the example system with RBD in Figure 6.1. Figure 6.26 shows a BN formulation with a FPS corresponding to each MCS and with the superscripts on the FPE nodes indicating MCS numbers. The total clique table size associated with this BN is 320. In Figure 6.27, FPSs associated with components 1, 2, and 3 are coalesced, reducing the total clique table size to 132. Once again, a notational change has been made: the superscript on each FPE node, which previously indicated the MCS number, now represents the instance of the FPE. Because, for this system, each component is associated with only one FPE, all superscripts are 1. If the component and system states are binary, the CPTs of the FPE nodes are constructed using the Boolean logic:

$$E_{f,i} = 0 \text{ if } \left[ \cup \left\{ E_{f, \text{Pa}(i)} = 0 \right\} \right] \cap \left[ C_i = 0 \right]$$
(6.14)

#### = 1 otherwise

where  $E_{f,Pa(i)}$  indicate the FPE nodes that are parent to  $E_{f,i}$ . The node representing the system state indicates failure if any of its parent FPE nodes is in the failure state.

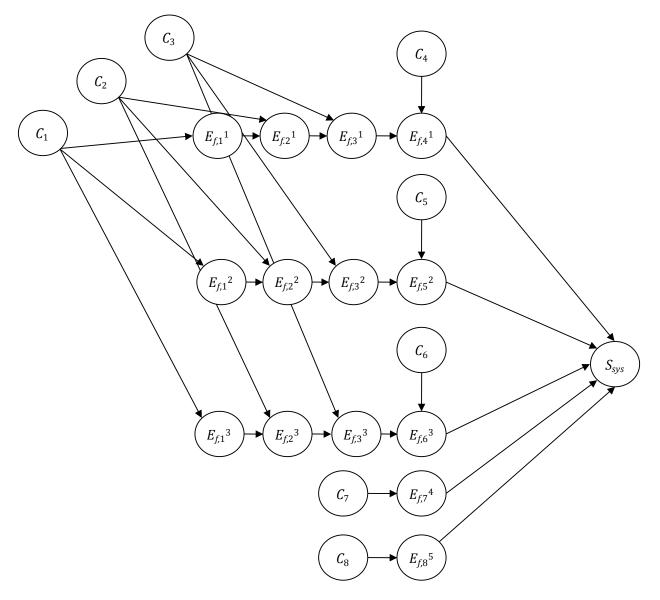


Figure 6.26: Efficient MCS BN formulation for the system in Figure 6.1 with distinct FPSs

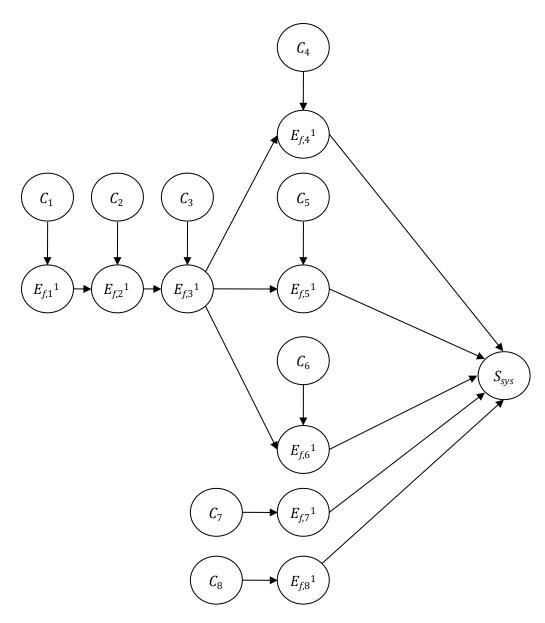


Figure 6.27: Efficient MCS BN formulation for the example system with coalesced FPEs associated with components 1, 2 and 3 using a converging structure

Note that node  $S_{sys}$  in Figure 6.27 has more than three parents. To increase computational efficiency, the converging connection is replaced by a chain structure as was done previously with the efficient MLS formulation. The resulting BN is shown in Figure 6.28.

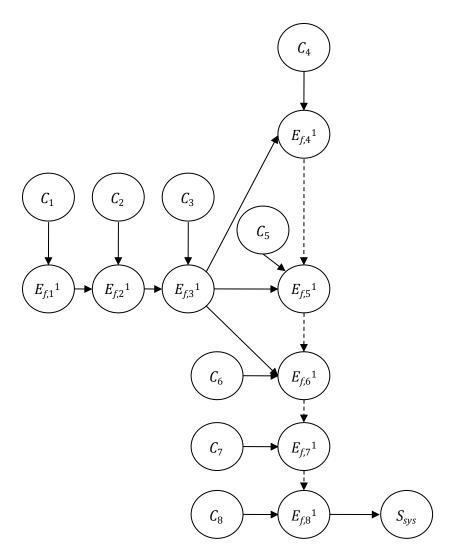


Figure 6.28: Efficient MCS formulations for the example system with coalesced SPEs associated with components 1, 2, and 3 using a chain structure

Returning to the special case of binary state components and system, the CPTs of the FPE nodes are defined using the relation

$$E_{f,i} = 0 \text{ if } \left[ (C_i = 0) \cap \left( \cup \{ E_{f, Pa(i)'} = 0 \} \right) \right] \cup \{ E_{f, Pa(i)''} = 0 \}$$
  
= 1 otherwise (6.15)

where  $E_{f,Pa(i)'}$  are the FPE nodes that are parents to  $E_{f,i}$  before addition of the chain modification and which remain parents after it;  $E_{f,Pa(i)''}$  are the FPE nodes that become parents to  $E_{f,i}$  after the chain structure is added (identified by dashed links in Figure 6.28).

The total clique table size associated with this BN is 80, assuming independent binary components.

Like the standard MCS formulation, the efficient MCS formulation can be adapted to handle multi-state problems through the application of the Max-Flow Min-Cut Theorem. To consider multi-state problems, the topology of the BN need not differ from the topology used for the binary state problem. It is only necessary to increase the number of states associated with all nodes and use arithmetic expressions instead of Boolean relations to define the CPTs. In the multi-state problem, the states associated with the FPE nodes correspond to *values* rather than logical outcomes, similar to the values assigned to the MCS nodes when adapting the standard MCS formulation to the multi-state problem. The values of the FPE nodes for the multi-state problem are defined as

$$Value(E_{f,i}) = \left[\min_{E_{f,j} \in \operatorname{Pa}(E_{f,i})} Value(E_{f,j})\right] + Cap_i$$
(6.16)

That is, the value assigned to node  $E_{f,i}$  is equal to the value of any incoming node plus the capacity of the associated component. Thus, each FPE node can be thought of as representing a "running total" of the capacities of the MCSs that the current FPE is a part of. Node  $S_{sys}$  now represents the maximum operating level of the entire system rather than a Boolean outcome. Its value corresponds to the minimum value among all MCS and thus is defined using the expression

$$Cap_{sys} = \min_{E_{f,j} \in \operatorname{Pa}(Sys)} Value(E_{f,j})$$
(6.17)

Considering the formulation modified with the addition of the chain structure to replace a converging structure (e.g. Figure 6.28), the value of a node  $E_{f,i}$  is expressed as:

$$Value(E_{f,i}) = \min_{E_{f,j} \in \operatorname{Pa}(E_{f,i})^{\prime\prime}} \left[ Value(E_{f,j}), \left( \left( \min_{E_{f,k} \in \operatorname{Pa}(E_{f,i})^{\prime}} Value(E_{f,k}) \right) + Cap_i \right) \right]$$
(6.18)

Table 6.1 shows a summary of the total clique table sizes associated with the above formulations, considering binary and statistically independent component states. It is observed that all MLS/MCS-based formulations are more efficient than the naïve formulation. For this small example system, the efficient MLS and MCS formulations only offer an advantage over the standard MLS/MCS formulations if care is used with regard to the structure of chains. However, when properly constructed, the efficient formulations are associated with computational demands of a lower magnitude than all other methods. The next section describes an optimization-based algorithm for determining the optimal order of the SPE or FPE nodes so as to minimize the number of links in the BN (as a proxy for minimizing total computational demands).

BN	Total clique table size
Figure 6.7: naïve formulation	512
Figure 6.12: standard MLS formulation	208
Figure 6.13: standard MLS formulation with intermediate nodes	176
Figure 6.15: standard MCS formulation	232
standard MCS formulation with intermediate nodes (no figure)	200
Figure 6.23: efficient MLS formulation with distinct SPEs	224
Figure 6.24: efficient MLS formulation with coalesced SPEs	108
Figure 6.25: efficient MLS formulation with chain structure for system	64
Figure 6.26: efficient MCS formulation with distinct FPEs	320
Figure 6.27: efficient MCS formulation with coalesced FPEs	132
Figure 6.28: efficient MCS formulation with chain structure for system	80

### Table 6.1: Comparison of computational complexity

## 6.6 Optimal ordering of efficient MLS and MCS formulations

Thus far, the SPEs in a SPS (FPEs in FPSs) corresponding to a particular MLS (MCS) have been arranged in an arbitrary order. However, for complex systems, the arrangement of the SPEs in the SPSs may strongly influence our ability to coalesce SPEs in multiple SPSs (and analogously for FPEs in FPSs). As was shown, increases in computational efficiency are achieved only if nodes in different SPSs (FPSs) are coalesced. The order in which SPEs (FPEs) appear can be optimized such that SPEs (FPEs) in as many SPSs (FPSs) as possible are coalesced. This reduces the number of nodes and links in the BN. This optimization problem is described next. For brevity, only the formulation employing SPSs is presented; a dual formulation applies to FPSs. To obtain the dual formulation, one simply needs to replace references to MLSs with MCSs when specifying the optimization problem.

Let  $L(i^m, j^n) = 1$  indicate the existence of a directed link from  $E_{s,i}^m$  to  $E_{s,j}^n$  in the efficient MLS BN formulation and  $L(i^m, j^n) = 0$  indicate its absence, where *i* and *j* are component indices and *m* and *n* are indices denoting the instances of these SPE nodes in the BN. Similarly, let  $K_i^m = 1$  indicate a directed link between the node representing component *i* and node  $E_{s,i}^m$  and  $S_i^m = 1$  indicate a directed link between  $E_{s,i}^m$  and the system node (with  $K_i^m = 0$  and  $S_i^m = 0$  respectively denoting their absences).  $L(i^m, j^n)$ ,  $K_i^m$  and  $S_i^m$  are the decision variables in the optimization problem. Formulation of the optimization problem assumes the use of only SPE nodes and a converging structure at the system node. To further increase computational efficiency of the resulting BN, the converging structure at any node with more than 3 SPE nodes as parents is replaced by a chain structure in the manner described above.

The objective of the optimization problem is to minimize the number of links in the BN, i.e. the objective function is

$$\min\left[\sum_{i=1}^{N_c}\sum_{j=1}^{N_c}\sum_{m=1}^{N_I}\sum_{n=1}^{N_I}L(i^m, j^n) + \sum_{i=1}^{N_c}\sum_{m=1}^{N_I}K_i^m + \sum_{i=1}^{N_c}\sum_{m=1}^{N_I}S_i^m\right]$$
(6.19)

where  $N_C$  is the number of components in the system and  $N_I$  is the maximum number of instances of any SPE. It is desirable that  $N_I$  be as small as possible, but its value is not known prior to solving the optimization problem. Thus, an iterative procedure must be pursued to find the smallest  $N_I$  value for which the optimization problem is feasible.

The existence of links between the component and SPE nodes as well as between the SPE nodes and the system node are controlled by the arrangement of SPE nodes in the BN. Specifically,  $K_i^m = 1$  if node  $E_{s,i}^m$  exists in the BN, which occurs if the decision variables  $L(j^n, i^m)$  or  $L(i^m, j^n)$  indicate a link going into or out of node  $E_{s,i}^m$ , respectively. (A node without links going into or out of it can be removed from the BN.) Mathematically, this constraint is written as

$$\left[\sum_{j=1}^{N_C} \sum_{n=1}^{N_I} \{L(i^m, j^n) + L(j^n, i^m)\} \ge 1\right] \Rightarrow K_i^m = 1$$
(6.20)

Techniques are available for modeling "if-then" and "*k*-out-of-*n*" (which are needed later) constraints in numerical optimization (e.g. see Sarker and Newton 2008). Appendix 6.1 provides ans explanation of these techniques. The decision variable  $S_i^m = 1$  if node  $E_{s,i}^m$  is a terminal node in a SPS, i.e.  $E_{s,i}^m$  exists and has no other SPE node as a child. Mathematically, this is written as

$$\left[\sum_{j=1}^{N_C} \sum_{n=1}^{N_I} L(j^n, i^m) \ge 1\right] \cap \left[\sum_{j=1}^{N_C} \sum_{n=1}^{N_I} L(i^m, j^n) = 0\right] \Rightarrow S_i^m = 1$$
(6.21)

There are two additional constraints governing the arrangement of the SPE nodes in the BN: (1) each MLS must be represented by a SPS; and (2) no SPS may exist that is not strictly a MLS. If the first constraint is violated, then one or more MLSs are excluded resulting in overestimation of the system failure probability. If the second constraint is violated, then the BN includes one or more fictitious MLSs and thus underestimates the system failure probability.

The first constraint requires that each MLS be represented as a SPS, i.e., at least one *permutation* of the SPEs associated with the components in each MLS must be connected as a chain. Define  $MLS_i$  to be the set of components contained in the *i*<sup>th</sup> MLS and let  $N_{MLS,i}$  be the number of components in  $MLS_i$ . For the system in Figure 6.1,  $N_{MLS,1} = N_{MLS,2} =$ 

 $N_{MLS,3} = 3$  and  $N_{MLS,4} = 5$ . Let  $P_i$  be the set of permutations, without replacement, of the component indices in  $MLS_i$  and define  $p_i^{\alpha} = \{p_{i,1}^{\alpha}, p_{i,2}^{\alpha}, \dots, p_{i,N_{MCS,i}}^{\alpha}\}$  as the  $\alpha$ <sup>th</sup> permutation contained in the set  $P_i$ . As an example, for the system in Figure 6.1,  $P_1 = [p_1^1 = \{8,7,1\}, p_1^2 = \{8,1,7\}, p_1^3 = \{7,8,1\}, p_1^4 = \{7,1,8\}, p_1^5 = \{1,7,8\}, p_1^6 = \{1,8,7\}].$ 

Next, let  $Q_i$  be the set of permutations with replacement of  $N_{MCS,i}$  draws from the instance index set  $\{1, ..., N_I\}$ . Define  $q_i^{\beta} = \{q_{i,1}^{\beta}, q_{i,2}^{\beta}, ..., q_{i,N_{MLS,i}}^{\beta}\}$  as the set of instance indices ordered according to the  $\beta^{th}$  member of  $Q_i$ . Using the same example and assuming  $N_I = 2$ , we have  $Q_1 = \{q_1^1 = (1,1,1), q_1^2 = (1,1,2), q_1^3 = (1,2,1), q_1^4 = (1,2,2), q_1^5 = (2,1,1), q_1^6 = (2,1,2), q_1^7 = (2,2,1), q_1^8 = (2,2,2)\}$ . Note that  $P_i$  has  $N_{MLS,i}$ ! members, while  $Q_i$  has  $N_I^{N_{MLS,i}}$  members.

Define a set  $r_i^{\alpha,\beta} = [r_{i,1}^{(\alpha,\beta)}, r_{i,2}^{(\alpha,\beta)}, \dots, r_{i,N_{MLS,i}}^{(\alpha,\beta)}]$  which combines the elements of  $p_i^{\alpha}$  and  $q_i^{\beta}$ . Specifically,  $r_i^{\alpha,\beta}$  includes the set  $p_i^{\alpha}$  with superscripts given by the set  $q_i^{\beta}$ . For the example system,  $r_1^{1,1} = \{8^1, 7^1, 1^1\}, r_1^{1,2} = \{8^1, 7^1, 1^2\}, r_1^{2,4} = \{8^1, 1^2, 7^2\}$ , etc. Overall, for this specific MLS, there are  $3! * 2^3 = 48$  possible ways to arrange the component indices given by  $p_i^{\alpha}$  and the instance superscripts given by  $q_i^{\beta}$ .

For convenience, define the sum  $X_i^{(\alpha,\beta)} = \sum_{l=1}^{N_{MLS,i}-1} L[r_{i,l}^{(\alpha,\beta)}, r_{i,l+1}^{(\alpha,\beta)}]$ , where  $r_{i,l}^{(\alpha,\beta)}$  is the  $l^{\text{th}}$  element of  $r_i^{(\alpha,\beta)}$ .  $X_i^{(\alpha,\beta)} = N_{MLS,i} - 1$  only if the SPEs corresponding to the components in  $MLS_i$  form a SPS in the order specified by  $p_i^{\alpha}$  and instance indices according to  $q_i^{\beta}$ . For a required SPS to exist in the BN,  $X_i^{(\alpha,\beta)} = N_{MLS,i} - 1$  for *at least one* component/instance index ordering from the set  $r_i^{(\alpha,\beta)}$ . The constraint is written as

$$\max_{\alpha,\beta} X_{i}^{(\alpha,\beta)} = N_{MLS,i} - 1 \qquad \alpha = 1, ..., N_{MLS,i}!, \qquad \beta = 1, ..., N_{I}^{N_{MCS,i}}, \quad \forall i$$
(6.22)

The second constraint requires that no SPS exist in the BN which does not correspond to a MLS. Consider the BN shown in Figure 6.29a. Let the shaded nodes  $(E_{s,1}^1 \rightarrow E_{s,2}^1 \rightarrow E_{s,3}^1 \rightarrow E_{s,4}^1)$  represent a particular permutation of component/instance indices  $r_i^{(\alpha,\beta)} = \{1^1, 2^1, 3^1, 4^1\}$  resulting in a valid SPS. Constraint (2) prohibits a SPE  $E_{s,j}^n$ , for any n, from "branching-off" the SPS at any point (i.e. being a child of any node in the chain), unless component j exists in a MLS with all components of the preceding SPEs in the sequence. For example, in Figure 6.29a,  $E_{s,j}^n$  cannot exist as a child of  $E_{s,3}^1$  unless components 1,2,3, and j all exist together in a MLS. If components 1,2,3, and j do not exist in a MLS, then the false survival path shown by nodes with dashed edges is introduced into the BN. The associated constraint is written as

$$\left[X_{i}^{(\alpha,\beta)} = N_{MLS,i} - 1, \forall i, \alpha, \beta\right] \Rightarrow$$
(6.23)

$$L\left[r_{i,l}^{(\alpha,\beta)}, j^n\right] = 0, \qquad \forall j: \left\{p_{i,1}^{(\alpha)}, \dots, p_{i,l}^{(\alpha)}, j\right\} \notin MLS_m, \forall n, \forall m, \forall l$$

$$\underbrace{\left(E_{s,1}^{(\alpha,\beta)}, j^n\right)}_{\left(E_{s,1}^{(\alpha)}, \dots, p_{i,l}^{(\alpha)}, j\right)} \underbrace{\left(E_{s,1}^{(\alpha)}, \dots, p_{i,l}^{(\alpha)}, j\right)}_{\left(E_{s,1}^{(\alpha)}, \dots, p_{i,l}^{(\alpha)}, j\right)} \underbrace{\left(E_{s,1}^{(\alpha)}, \dots, E_{s,l}^{(\alpha)}, j\right)}_{\left(E_{s,1}^{(\alpha)}, \dots, E_{s,l}^{(\alpha)}, j\right)} \underbrace{\left(E_{s,1}^{(\alpha)}, \dots, E_{s,l}^{(\alpha)}, j\right)}_{\left(E_$$

Figure 6.29: BNs used to illustrate constraint (2)

Furthermore, the constraint must prohibit SPE  $E_{s,j}^n$ , for any n, from being a parent to any node in a valid SPS, unless component i exists in a MLS with all components of the subsequent SPEs in the sequence. For example, in Figure 6.29b,  $E_{s,j}^n$  cannot be a parent of  $E_{2,1}$  unless components 2,3,4, and j all exist together in a MLS. The second constraint takes the form

$$\begin{bmatrix} X_i^{(\alpha,\beta)} = N_{MLS,i} - 1, \forall i, \alpha, \beta \end{bmatrix} \Rightarrow$$

$$L\left[j^n, r_{i,l}^{(\alpha,\beta)}\right] = 0 \quad \forall j \colon \left\{j, p_{i,l}^{(\alpha)}, \dots, p_{i,N_{MLS,i}}^{(\alpha)}\right\} \notin MLS_m, \forall n, \forall m, \forall l$$
(6.24)

The combination of these two requirements along with the objective function, the minimization of which ensures links that are not necessary for constructing the required SPSs are not in the BN, prohibits invalid SPSs in the BN. Combining (6.23) and (6.24) results in the constraint

$$\begin{bmatrix} X_{i}^{(\alpha,\beta)} = N_{MLS,i} - 1, \forall i, \alpha, \beta \end{bmatrix} \Rightarrow$$

$$\sum_{m=1}^{N_{MLS}} \sum_{n=1}^{N_{I}} \sum_{l=1}^{N_{MLS,i}} \sum_{\forall j: \{p_{i,1}^{(\alpha)}, \dots, p_{i,l}^{(\alpha)}, j\} \notin MLS_{m}} L\left[r_{i,l}^{(\alpha,\beta)}, j^{n}\right]$$

$$+ \sum_{m=1}^{N_{MLS}} \sum_{n=1}^{N_{I}} \sum_{l=1}^{N_{MLS,i}} \sum_{\forall j: \{j, p_{i,l}^{(\alpha)}, \dots, p_{i,N_{MLS,i}}^{(\alpha)}\} \notin MLS_{m}} L\left[j^{n}, r_{i,l}^{(\alpha,\beta)}\right] = 0$$

$$(6.25)$$

The binary optimization problem described above requires consideration of permutations of components. Consequently, it becomes difficult to solve this problem in practice for large systems. To overcome this problem, several heuristics have been developed to reduce the size of the optimization problem that must be considered. Specifically, groups of components are considered as single "super components," thus reducing the number of components in each MLS or MCS. Furthermore, measures are taken to reduce the number permutations of component indices considered in the optimization problem, without significantly impacting the optimality of the final solution.

Before describing the heurtistics, note that significant differences in performance were encountered depending in the optimization algorithm/software used to solve the above problem in conjunction with the heuristics described in the next section. For example, the Matlab-based Tomlab optimization environment (toolbox) (Holmström 2008), which is not produced by Mathworks (producer of Matlab), was able to solve example applications described later in this chapter in under 30 seconds on a HP xw8600 workstation with 3.00GHz Xeon processor. However, using the same computer, the native binary integer program solver in the Matlab optimization toolbox often "timed-out" at over 1500 seconds without finding a solution.

# 6.7 Heuristic augmentation

For systems with MLSs or MCSs with many components, the optimization problem may become computationally infeasible because of the large number of permutations to be considered. One way to reduce the size of the optimization problem is through the use of super components, a process used in multi-scale modeling. The first heuristic discussed in this section is based on an algorithm for identifying super components. The second heuristic identifies sets of components that appear in multiple MLSs (MCSs) and uses them to reduce the number of permutations that must be considered in solving the optimization problem.

# 6.7.1 Heuristic based on super components

Multi-scale modeling is an approach by which elementary components in a system are grouped into "super components" (Der Kiureghian and Song 2008). Analysis is performed individually for each super component and then results are aggregated on the system-level. The super components typically are comprised of simple sub-systems, such as components that exist in series or parallel along a system link. In this study, only super components made up of series and parallel components are considered. The introduction of super components reduces the size and number of MCSs and MLSs. Consequently, the number of permutations that must be considered in the optimization problem is reduced.

Once again, consider the simple system shown in Figure 6.1. Components  $C_4$ ,  $C_5$ , and  $C_6$  exist in series as do components  $C_7$  and  $C_8$ . We replace these individual components by two super components  $SC_1$  and  $SC_2$ , as shown in Figure 6.30. The system still has 4 MLSs, but the number of components in them is reduced: {1,  $SC_2$ }, {2,  $SC_2$ }, {3,  $SC_2$ }, { $SC_1$ ,  $SC_2$ }.

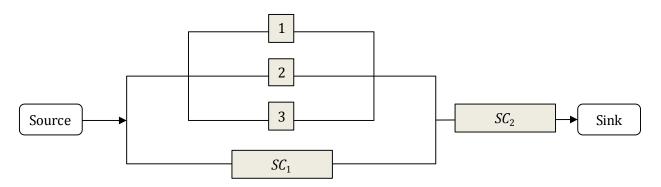


Figure 6.30: System in Figure 6.1 with components  $C_4$ ,  $C_5$  and  $C_6$  a well as components  $C_7$  and  $C_8$  replaced by super components

Examination of Figure 6.30 reveals that components  $C_1$ ,  $C_2$ ,  $C_3$  and  $SC_1$  exist in parallel. These components are next replaced by a single super component resulting in the RBD shown in Figure 6.30. Now, components  $SC_2$  and  $SC_3$  exist in series and can be replaced by another super component. The BN resulting from this incremental procedure for identifying components that may be grouped and replaced by a super component is shown in Figure 6.32. For super components containing less than 4 constituent components, a converging structure is used. For super components with 4 or more constituent components, a chain structure is utilized.

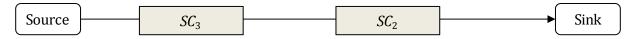


Figure 6.31: System from Figure 6.30 with components  $C_1$ ,  $C_2$ ,  $C_3$  and  $SC_1$  replaced by a super component

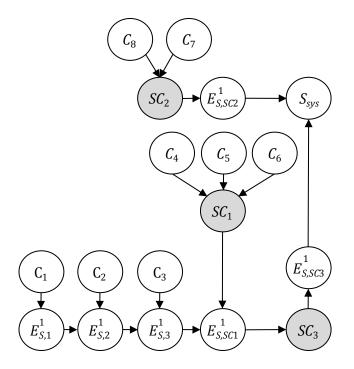


Figure 6.32: BN constructed for system in Figure 6.1 using super component heuristic

Note that components in a super component need not be contiguous. For example, consider the simple system in Figure 6.33. From an analysis perspective, components 1 and 4 can be combined because, with regard to formation of MLSs and MCSs, they have the same effect as if they physically existed in series.

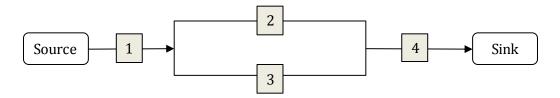


Figure 6.33: Example system illustrating non-contiguous components that can be combined in a super component.

An algorithm has been developed to automate the incremental identification and replacement of elementary components by super components (or sets of super components by other super components). The first step in the algorithm is to construct an initial matrix  $\mathbf{M}^0$  that contains a row corresponding to each MLS (MCS) and a column corresponding to each component. The elements of the initial matrix  $\mathbf{M}^0$ ,  $M_{ij}^0$ , are defined such that

$$M_{ij}^0 = 1$$
 if component *j* is a member of MLS (MCS) *i* (6.26)

#### = 0 otherwise

For the example system in Figure 6.1, **M**<sup>0</sup> has the following elements:

	Comp 1	Comp 2	Comp 3	Comp 4	Comp 5	Comp 6	Comp 7	Comp 8
MLS 1	1	0	0	0	0	0	1	1
MLS 2	0	1	0	0	0	0	1	1
MLS 3	0	0	1	0	0	0	1	1
MLS 4	0	0	0	1	1	1	1	1

Two types of super components are considered in the algorithm. Class A super components are made up of groups of components that always appear together in a MLS (MCS) and never appear separately. In a MLS-based formulation, Class A super components correspond to components that exist in series. For a MCS-based formulation, class A super components represent components that exist in parallel. Define a quantity  $m_j^{(A)}$  that is assigned to each component *j* at the *p*<sup>th</sup> iteration of the algorithm:

$$m_j^{(A)} = \sum_{i=1}^{N_{MLS}} 2^i M_{ij}^p \tag{6.27}$$

where  $M_{ij}^p$  is the element of the  $\mathbf{M}^p$  at step p of the algorithm; for the initial step, p = 0. This quantity is identical only for components that always appear together in MLSs (MCSs) and do not appear separately in different MLSs (MCSs). Therefore, components having identical values of this quantity can be grouped into a super component. Returning to the example system:  $m_1^{(A)} = 2$ ,  $m_2^{(A)} = 4$ ,  $m_3^{(A)} = 8$ ,  $m_4^{(A)} = m_5^{(A)} = m_6^{(A)} = 16$ , and  $m_7^{(A)} = m_8^{(A)} = 30$  at the first iteration. Thus, components 4, 5, and 6 can be grouped into one super component. Components 7 and 8 can be grouped to form another super component. This was shown graphically in Figure 6.30. Matrix  $\mathbf{M}^0$  must now be updated to reflect the new super components by removing columns corresponding to components. To illustrate the updating process, first consider  $SC_1$  containing components 4, 5, and 6. Adapt the matrix  $\mathbf{M}^0$  to produce a matrix  $\mathbf{M}^1$  that accounts for the new super component. This is done by removing columns from  $\mathbf{M}^0$  corresponding to components 4, 5, and 6 and adding a column that corresponds to  $SC_1$ . Let  $M^{1'}$  denote this column. Each row k of this column is defined such that

$$M_{\operatorname{row} k}^{p}{}' = 1 \text{ if } \sum_{i \in C_{grp}^{p-1}} M_{ki}^{p-1} > 0$$
  
= 0 otherwise (6.28)

where  $C_{grp}^{p-1}$  is the set of components grouped into a super component at stage (p-1), i.e. components 4, 5, and 6 for (p-1=0) in the example. For the example system, **M**<sup>1</sup> is

	Comp 1	Comp 2	Comp 3	Comp 7	Comp 8	$SC_1$
MLS 1	1	0	0	1	1	0
MLS 2	0	1	0	1	1	0
MLS 3	0	0	1	1	1	0
MLS 4	0	0	0	1	1	1

The above procedure is repeated to construct  $SC_2$  consisting of components 7 and 8. The resulting matrix  $\mathbf{M}^2$  is

	Comp 1	Comp 2	Comp 3	$SC_1$	SC <sub>2</sub>
MLS 1	1	0	0	0	1
MLS 2	0	1	0	0	1
MLS 3	0	0	1	0	1
MLS 4	0	0	0	1	1

The second class of super components, class B, contain components that appear in separate MLSs (MCSs), but with the same set of other components. For a MLS formulation, this corresponds to components in parallel; for a MCS formulation it corresponds to components in series. For the example in Figure 6.30, components 1, 2, 3, and  $SC_1$  appear in separate MLSs, but with the same set of other components (7 and 8) in those separate MLSs. Define the quantity for each component *j* 

$$m_{j}^{(B)} = \sum_{i=1}^{N_{MLS}} \left\{ M_{ij}^{p} * \left( \sum_{\substack{k=1\\k \neq j}}^{N_{c}} 2^{k} M_{ik}^{p} \right) \right\}$$
(6.29)

At iteration *p* any set of components (or super components) for which the value of  $m_j^{(B)}$  is the same, can be grouped into a super component. For the example system in Figure 6.1,  $m_1^{(B)} = m_2^{(B)} = m_3^{(B)} = m_{SC1}^{(B)} = 32$  and  $m_{SC2}^{(B)} = 30$ . As expected, components 1, 2, 3, and super component  $SC_1$  have the same value of  $m_j^{(B)}$  and can be grouped into a super component. The matrix  $\mathbf{M}^2$  must be adapted by removing columns corresponding to the grouped components and adding columns that correspond to the new super component. The rows of the column that is added to the matrix  $\mathbf{M}^2$  are defined using (6.28). Redundant rows of the updated  $\mathbf{M}^p$ -matrix can be removed. It follows that  $\mathbf{M}^3$  is

$$\begin{array}{c|c} SC_2 & SC_3 \\ MLS 1 & 1 & 1 \\ \end{array}$$

The above iterative procedure for finding and replacing components with super components of classes A and B is repeated until no super components remain. The matrix

 $\mathbf{M}^p$  that corresponds to the last iteration is then used to specify the components and MLSs or MCSs required to define the optimization problem.

# 6.7.2 Second heuristic for reducing the number of permutations

The second heuristic identifies components that appear in many (but not all) MLSs (MCSs) and uses this observation to reduce the number of permutations that must be considered when solving the optimization problem. This heuristic can be used after first reducing the size of the problem through identification of super components. This heuristic is likely to result in a solution to the optimization problem that is suboptimal.

To facilitate the explanation of the heuristic, an example system is used to illustrate each step of the procedure. Consider an arbitrary system with seven MLSs:

$$MLS_1 = \{1,2,3,4\}$$
$$MLS_2 = \{1,4,5,6,8\}$$
$$MLS_3 = \{1,4,7\}$$
$$MLS_4 = \{2,3,5\}$$
$$MLS_5 = \{1,5,7\}$$
$$MLS_6 = \{1,2,9\}$$
$$MLS_7 = \{7,9\}$$

**Step 1:** The first step of the heuristic is to create an ordered list of components based on the number of times a component appears in a MLS or MCS. Call this list *O*. For the example system, the number of occurrences of each component within a MLS is shown in the following table:

Component	number of appearances in a MLS
1	5
2	3
3	2
4	3
5	3
6	1
7	3
8	1
9	2

The above table leads to the ordered list:  $O = \{1,2,4,5,7,3,9,6,8\}$ . In the case of ties in the above table, the order is arbitrarily based on component index value.

**Step 2:** Re-order the components within the MLSs or MCSs based on the order *O*. For the example system, the new MLS component orders are:

 $MLS_{1} = \{1,2,4,3\}$  $MLS_{2} = \{1,4,5,6,8\}$  $MLS_{3} = \{1,4,7\}$  $MLS_{4} = \{2,5,3\}$  $MLS_{5} = \{1,5,7\}$  $MLS_{6} = \{1,2,9\}$  $MLS_{7} = \{7,9\}$ 

**Step 3a:** Determine all pair-wise intersecting sets between MLSs or MCSs. Define the set of components in the intersection of  $MLS_i$  (or  $MCS_i$ ) and  $MLS_i$  (or  $MCS_i$ ) as  $G_{ij}$ :

$$G_{ij} = MLS_i \cap MLS_j, i = 1, \dots, (N_{MLS} - 1), j = (i + 1), \dots, N_{MLS}$$
(6.30)

For the example problem:

 $\begin{array}{l} G_{12}=G_{13}=G_{23}=\{1\;4\}\\ G_{16}=\{1\;2\}\\ G_{35}=\{1\;7\}\\ G_{25}=\{1\;5\}\\ G_{14}=\{2\;3\}\\ G_{15}=G_{24}=G_{26}=G_{36}=G_{37}=G_{45}=G_{46}=G_{56}=G_{57}=G_{67}=\text{sets of cardinality 1}\\ G_{17}=G_{27}=G_{34}=G_{47}=\emptyset \end{array}$ 

**Step 3b:** Define *G* as a set containing the unique sets  $G_{ij}$ ,  $\forall i, j$  with cardinality greater than 1.

For the example,  $G = \{\{1, 4\}, \{1, 2\}, \{1, 5\}, \{1, 7\}, \{2, 3\}\}$ .

**Step 4:** Sequentially assign to each MLS (or MCS) a set from within *G* that corresponds to the set whose intersection with the MLS (or MCS) has the largest cardinality. Let  $g_{k,i}$  be the intersection of set *i* within *G* and the  $k^{\text{th}}$ MLS (or MCS), i.e.

$$g_{k,i} = G_i \cap MLS_k \tag{6.31}$$

where  $G_i$  is the *i*<sup>th</sup> set contained within *G*. Define  $g'_k$  as the set  $g_{k,i}$  which has the longest length:

$$g_{k}' = \{g_{k,max} : |g_{k,max}| = \max_{i} |g_{k,i}|\}$$
(6.32)

where  $|\cdot|$  denotes the cardinality of a set. In the case of ties in the length of the intersection, the set within *G* that appears first is assigned to the MLS (MCS). Once a set from within *G* has been assigned to a MLS (or MCS), place it in a new set *G'*, if it has not already been placed there. Define  $g'_{c,i}$  as the number of times component *i* appears within

the sets in G'. When a component has appeared in  $N_I$  (the parameter representing number of instances in the optimization problem described previously) sets within G', all remaining sets containing that component should be removed from G, unless the set is already a member of G'. This step is necessary because, if  $g'_c > N_I$  for any component, the optimization problem may become infeasible.

Return to the example system and assume  $N_I = 2$ . Begin with the first MCS. The sets within G with the longest intersection with  $MCS_k$  are {1,4}, {1,2}, and {2,3}. Because set {1,4} appears first in G, it is assigned to the first MCS. Therefore,  $g'_{k=1} = \{1,4\}$  and this set is added to the set of sets G'. It follows that  $g'_{c,1} = g'_{c,4} = 1$  and  $g'_{ci} = 0$  for i = 2,3,5,...,9. These results for  $MLS_1$  as well as the results for  $MLS_k$ , k = 2, ..., 7, are shown below:

For 
$$k = 1, MLS_1 = \{1,2,4,3\}$$
:  
 $g'_1 = \{1,4\}; G' = \{\{1,4\}\}$   
 $g'_{c,1} = 1; g'_{c,4} = 1, g'_{c,i} = 0, i = 2,3,5, ...,9$   
 $G = \{\{1,4\}, \{1,2\}, \{1,5\}, \{1,7\}, \{2,3\}\}$ 

For 
$$k = 2$$
,  $MLS_2 = \{1,4,5,6,8\}$ :

$$g'_{2} = \{1,4\}; G' = \{\{1,4\}\}$$
  

$$g'_{c,1} = 1; g'_{c,4} = 1, g'_{c,i} = 0, i = 2,3,5, \dots,9$$
  

$$G = \{\{1,4\}, \{1,2\}, \{1,5\}, \{1,7\}, \{2,3\}\}$$

For 
$$k = 3$$
,  $MLS_3 = \{1,4,7\}$ :  
 $g'_{c,1} =$ 

$$g'_{3} = \{1,4\}; \ G' = \{\{1,4\}\} \\ g'_{c,1} = 1; g'_{c,4} = 1, g'_{c,i} = 0, i = 2,3,5, ...,9 \\ G = \{\{1,4\}, \{1,2\}, \{1,5\}, \{1,7\}, \{2,3\}\}$$

For 
$$k = 4$$
,  $MLS_4 = \{2,5,3\}$ :  
 $g'_4 = \{2,3\}; G' = \{\{1,4\}, \{2,3\}\}$   
 $g'_{c,1} = 1; g'_{c,4} = 1, g'_{c,2} = 1; g'_{c,3} = 1, g'_{c,i} = 0, i = 5, ..., 9$   
 $G = \{\{1,4\}, \{1,2\}, \{1,5\}, \{1,7\}, \{2,3\}\}$ 

For 
$$k = 5$$
,  $MLS_5 = \{1,5,7\}$ :  
 $g'_5 = \{1,5\}$ ;  $G' = \{\{1,4\}, \{2,3\}, \{1,5\}\}$   
 $g'_{c,1} = 2$ ;  $g'_{c,4} = 1$ ,  $g'_{c,2} = 1$ ;  $g'_{c,3} = 1$ ,  $g'_{c,5} = 1$ ,  $g'_{c,i} = 0$ ,  $i = 6$ , ..., 9  
 $G = \{\{1,4\}, \{1,5\}, \{2,3\}\}$ 

For 
$$k = 6$$
,  $MLS_6 = \{1,2,9\}$ :  
 $g'_6 = \emptyset$ ;  $G' = \{\{1,4\},\{2,3\},\{1,5\}\}$   
 $g'_{c,1} = 2$ ;  $g'_{c,4} = 1$ ,  $g'_{c,2} = 1$ ;  $g'_{c,3} = 1$ ,  $g'_{c,5} = 1$ ,  $g'_{c,i} = 0$ ,  $i = 6$ , ..., 9  
 $G = \{\{1,4\},\{1,5\},\{2,3\}\}$ 

For 
$$k = 7, MLS_7 = \{7,9\}$$
:  
 $g'_7 = \emptyset; G' = \{\{1,4\}, \{2,3\}, \{1,5\}\}$   
 $g'_{c,1} = 2; g'_{c,4} = 1, g'_{c,2} = 1; g'_{c,3} = 1, g'_{c,5} = 1, g'_{c,i} = 0, i = 6, \dots, 9$   
 $G = \{\{1,4\}, \{1,5\}, \{2,3\}\}$ 

Note that, after the fifth MCS (k = 5) is considered, component 1 has appeared in the set G' two times. Because component 1 has appeared  $N_I = 2$  times, all sets containing component 1 are removed from the set G unless they already appear in G'.

**Step 5:** The last step of the heuristic is to modify the optimization problem such that the permutations on components contained within the set  $g'_k$  assigned to  $MLS_k$  are not considered. Recall that  $r_k^{\alpha,\beta} = \left[p_{k,1}^{(\alpha,\beta)}, p_{k,2}^{(\alpha,\beta)}, \dots, p_{k,n_k}^{(\alpha,\beta)}\right]$  combines the elements of  $p_k^{\alpha}$  and  $q_k^{\beta}$ . Specifically,  $r_k^{\alpha,\beta}$  includes the set  $p_k^{\alpha}$  with superscripts given by the set  $q_k^{\beta}$ . To modify the optimization problem, remove from the set of constraints all  $r_k^{\alpha,\beta}$  for which the component indices do not appear in the order specified by  $g'_k$ . Furthermore, remove all  $r_k^{\alpha,\beta}$  for which the superscript on component *i* in the set  $g'_k$  exceeds the value  $g'_{c,i}$ .

## 6.8 Example Applications

Two example systems are considered to illustrate the application of the above optimization scheme, with and without the heuristics, and compare the computational complexities associated with the efficient MLS/MCS BN formulations versus the standard MLS/MCS formulations.

## 6.8.1 Example application #1

Consider the system shown in Figure 6.34 (e.g. a utility distribution system), in which a source and sink are connected by a series of numbered links that can fail. The system has 5 MCSs: {1,2}, {1,4,6}, {2,3,5}, {3,4,6}, {2,5,7}, and {6,7}. The efficient MCS formulation obtained via the optimization algorithm, without use of heuristics, is shown in Figure 6.35. The first heuristic is not useful for this problem because no component appears in series or parallel and thus no components are eligible to be grouped as a super component. Furthermore, it is not necessary to apply the second heuristic in this example because the size of the problem is small. Figure 6.36 illustrates the FPSs corresponding to each MCS, as shown in grey. The total clique table size associated with this BN is 164. This is compared with a total clique table size of 1024 when the standard MCS formulation is used without the addition of intermediate nodes (recall Figure 6.13). The total clique table size is 680 when intermediate nodes are added to the standard MCS BN formulation to ensure no node has more than 3 parents.

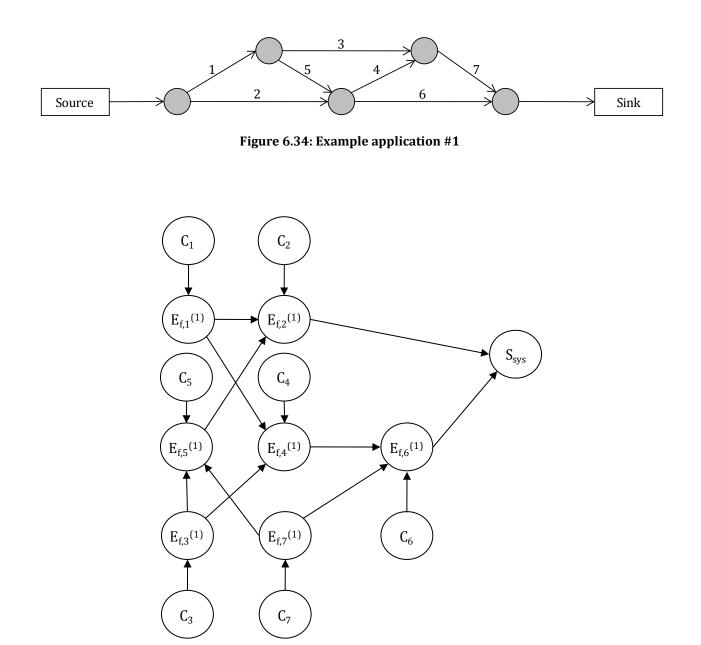


Figure 6.35: Efficient MCS formulation without heuristics for example application #1

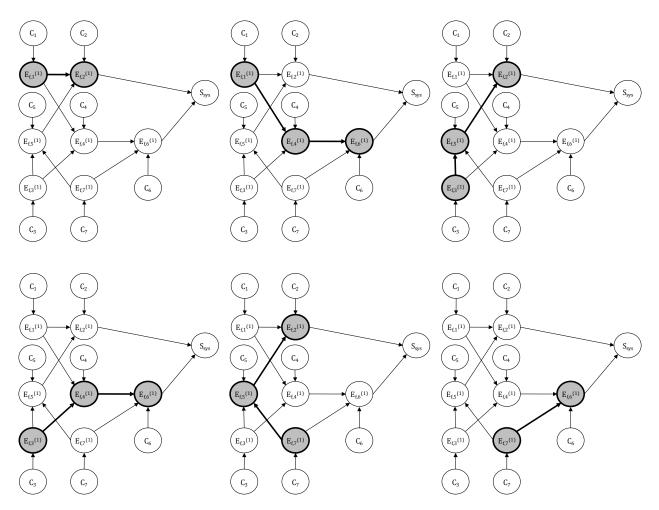


Figure 6.36: Illustration of FPSs corresponding to MCSs for example application #1

## 6.8.2 Example application #2

Next consider the structural system in Figure 6.37 consisting of 10 labeled components fail. 11 MCSs: {1,2} , {3,4} , {1,3,10} , {1,4,10}, that can The system has {2,3,10}, {2,4,10}, {5}, {6}, {7}, {8}, and {9}. The BN obtained using the optimization algorithm is shown in Figure 6.38. Figure 6.39 explicitly shows the FPSs corresponding to each of the multiple component MCSs. The super component heuristic is employed to obtain this solution. The MCSs containing a single elementary component (components 5-9) are grouped into a super component in which the relationship between the nodes is defined as a series system, i.e. if any component 5-9 fails, the system will be in the failure state. The total clique table size associated with this BN is 184. Note that the system node in Figure 6.39 has four parents. Replacing the converging structure with a chain structure, as shown by dotted lines in Figure 6.40, reduces the total clique table size to 164. The total clique table size of the standard MLS BN formulation with the converging structure is 5,140 without use of intermediate nodes. When intermediate nodes are introduced to ensure no node in the standard MCS BN formulation has more than three parents, the total clique table size reduces to 804, but remains substantially higher than when the optimization algorithm is used to obtain the efficient MCS BN topology. Thus, the efficient MCS BN obtained via optimization is significantly more efficient.

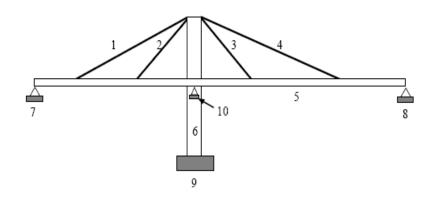


Figure 6.37: Example application #2

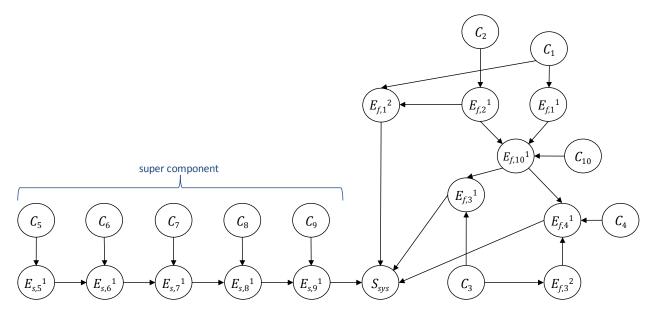


Figure 6.38: Efficient MCS formulation obtained using optimization algorithm with use of super components

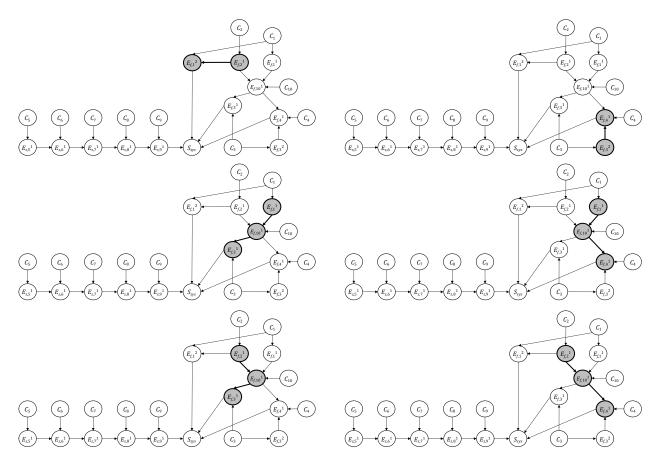


Figure 6.39: Illustration of FPSs corresponding to multi-component MCSs for example application #2

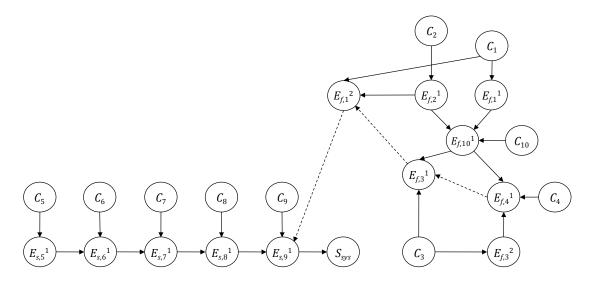


Figure 6.40: Efficient MCS formulation obtained using optimization algorithm with chain structure

The above formulations were obtained without using the heuristic described in section 6.7.2. This idealized structural system is sufficiently simple that it is possible to obtain a solution without requiring the heuristic, but for larger systems this may not be case. The BN obtained for this system using the optimization algorithm augmented by the second heuristic is shown in Figure 6.41. The solution obtained is suboptimal. It contains a total of 29 links, whereas the optimal solution contains 26 links. The total clique table size for this BN is 252. However, using the Tomlab optimization environment (Holmström 2008), it took over 12 times longer to obtain a solution without the heuristic than when the heuristic was employed. Thus, in using the heuristic, there is a trade-off between the among of time required to solve the optimization problem and the optimality of the resulting BN topology. However, even with the heuristic, the efficient MCS formulation is substantially more efficient than the standard MCS formulation.

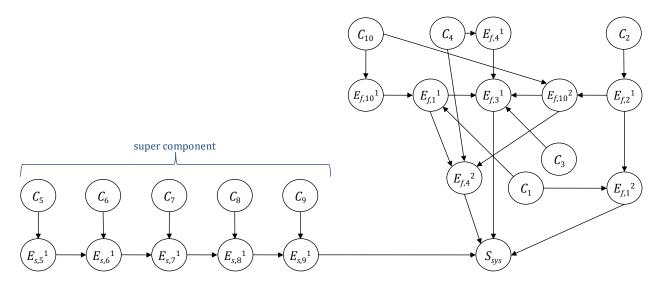


Figure 6.41: Efficient MCS formulation obtained using optimization algorithm with both heuristics

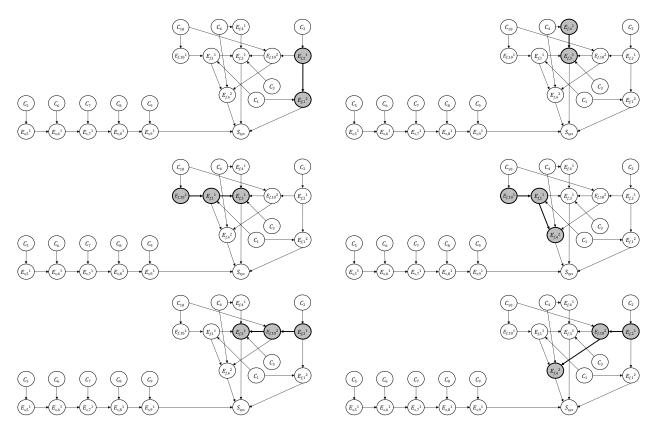


Figure 6.42: Illustration of FPSs corresponding to multi-component MCSs for example application #2 using both heuristics

#### 6.9 Summary

This chapter describes BN formulations for modeling the performance of engineered systems. The chapter begins with an overview of conventional methods of modeling system performance and compares them with using BNs for this purpose. Next, five methods for modeling system performance are described including a naïve approach, two approaches based on intuition, and two approaches that utilize minimal link and cut sets. The latter two formulations are then adapted with the goal of minimizing computational demands. A heuristically augmented optimization-based procedure is developed to automate the construction of the efficient BN formulations. The chapter concludes with several example applications.

#### Appendix:6.1: Modeling "either-or" & "k-out-of n" constraints

Methods are available for translating "either-or" and "*k*-out-of-*n*" constraints into a format standard to linear programming (Sarker and Newton 2008). First consider the "either-or" constraint. Let  $e_1$  denote the first constraint and  $e_2$  denote the second constraint. By way of an example, assume they are defined as

$$e_1: a_1x_1 + a_2x_2 \le a_3$$
  
 $e_2: b_1x_1 + b_2x_2 \le b_3$ 

We require that *either* constraint  $e_1$  or constraint  $e_2$  must hold. To achieve this condition, add an arbitrarily large number, M, to the right hand side of constraint  $e_1$  to create a modified constraint  $e'_1$ :

$$e_1': a_1x_1 + a_2x_2 \le a_3 + M$$

Given a sufficiently large M, any values of  $x_1$  and  $x_2$  will satisfy the modified constraint  $e_1'$ . It is thus an inactive constraint. Similarly, when a large number is added to constraint  $e_2$ , it becomes inactive.

Next, we introduce a binary design variable, y, in conjunction with the arbitrarily large number M to achieve two new modified constrains,  $e_1''$  and  $e_2''$ , which achieve the desired behavor:

$$e_1'': a_1x_1 + a_2x_2 \le a_3 + M * y$$
  
 $e_2'': b_1x_1 + b_2x_2 \le b_3 + M * (1 - y)$ 

When y = 0, constraint  $e_1''$  must hold and  $e_2''$  is inactive. Conversely, when y = 1, constraint  $e_2''$  must hold and  $e_1''$  is inactive. Thus the "either-or" behavior is achieved.

A similar procedure can be used for the case in which "k-out-of-n" constraints must hold. Consider n possible constraints:

$$e_1: f_1(x) \le c_1$$
$$\vdots$$
$$e_n: f_N(x) \le c_n$$

Introduce *n* binary design/selection variables,  $y_i$ , in conjunction with an arbitrarily large number *M* to each constraint:

$$e_1'': f_1(x) \le c_1 + M * y_1$$
  
$$\vdots$$
  
$$e_n'': f_N(x) \le c_N + M * y_N$$

Constraint  $e''_i$  is active when  $y_i = 0$  and it is inactive when  $y_i = 1$ . To achieve the desired "*k*-out-of-*n*" behavior, we must satisfy the constraint on the integer variables

$$\sum_{i=1}^{n} y_i = n - k$$

# **Chapter 7: Decision Support**

#### 7.1 Introduction

Decision theory provides a rational basis for solving a diverse range of practical decision problems encountered in civil engineering. Examples of these decision problems include: structural system selection (e.g. the decision to construct a steel versus concrete bridge); choice and timing of maintenance, rehabilitation, and retrofit actions; and inspection prioritization. Associated with these decisions are costs and benefits. For example, there are costs associated with performing a retrofit to a component of an infrastructure system, e.g. materials and labor. However, there are also benefits associated with the retrofit, including increased reliability and service life. Decision theory provides a rational basis for weighing the costs and benefits associated with a set of alternative choices and for selecting the best among them in accordance with the preferences of the decision-maker (Kübler 2007). However, the costs and benefits associated with decision alternatives are not deterministic. For example, while a seismic retrofit may be expected to improve the reliability of a component, there is no guarantee that it will prevent damage to the component in a future earthquake. While several theories exist to provide solutions to uncertain decision problems, we utilize the paradigm in which the optimal decision is the one associated with the maximum expected utility. The concept of expected utility is described below.

Consider an arbitrary decision problem in which a decision-maker must select among alternative actions  $a_i$ , i = 1, ..., m. Each alternative action can lead to one of multiple potential outcomes  $b_j$ , j = 1, ..., n. Associated with decision alternative  $a_i$  and outcome  $b_j$  is a utility  $u_{ij} = u(a_i, b_j)$ , which maps the relative desirability of this pair on the decision-maker's value scale. Usually, a positive utility indicates a benefit and a negative utility

indicates a cost. Let  $p_{ij} = \Pr(b_j | a_i)$  be the probability that outcome  $b_j$  will be observed given alternative  $a_i$  is chosen. The expected utility associated with decision alternative  $a_i$  is

$$E(u|a_i) = \sum_{j=1}^{n} p_{ij} u_{ij}, \quad i = 1, ..., m$$
(7.1)

The "best" decision alternative is the one that achieves the maximum expected utility,  $u^*$ , (Kübler 2007) i.e.,

$$u^* = \max_{i=1,...,m} E(u|a_i)$$
(7.2)

A variety of methods exist for solving decision problems under uncertainty, including conventional methods such as decision trees and the more compact representation resulting from influence diagrams.

The goal of this chapter is to demonstrate how BNs can be extended by decision and utility nodes to solve decision problems. Once decision and utility nodes have been added to the BN, the resulting graphical model is known as an influence diagram. To achieve the goal of the chapter, we focus on developing preliminary influence diagrams to solve a specific problem involving the post-earthquake inspection and shutdown of components. Consider a post-earthquake scenario in which an earthquake has occurred and placed seismic demands on the components of an infrastructure system. Immediately following the earthquake, the owner/decision-maker must quickly decide, for each component of the system, whether to keep it open or shut it down, or to conduct an inspection before deciding on the fate of the component. Because of finite resources, the owner cannot simultaneously inspect all components. Therefore, for components for which inspections are required, the owner must choose the order in which to perform the inspections.

This chapter begins with a brief overview of decision trees. Then an introduction to influence diagrams, including perfect recall and limited memory influence diagrams, is presented. For brevity we utilize the aforementioned example application to demonstrate the value of using influence diagrams for solving decision problems. However, the framework described here can be applied to more general decision problems. We focus first on developing an influence diagram for the inspection-shutdown decision at the component level. Then we consider the decision at the system level and develop a prioritization heuristic based on a value of information criterion. Finally, we provide illustrative results for an example system.

#### 7.2 Introduction to decision trees and influence diagrams

A classical method for solving decision problems under uncertainty involves the use of decision trees. In a decision tree, all possible sequences of alternatives, observations, and outcomes are expanded in a tree-like structure containing decision, chance, and utility

nodes. Each path from the root of the tree to a leaf represents a possible complete decision alternative/observation/outcome *sequence*. An optimal decision strategy is found by finding the sequence(s) resulting in the highest expected utility. Because of the need to list all possible combinations of decision alternatives, observations, and outcomes, decision trees quickly become excessively and impractically large for decision problems with many variables. This problem is further exacerbated in decision scenarios in which the order of decisions is not known a priori. For example, in the inspection prioritization problem, the order in which inspections are to be performed is not known when the decision problem is formulated. In fact, determination of the optimal order is an outcome of the decision problem. Therefore, to solve the decision problem via a decision tree, it is necessary to consider all permutations of the orders in which components can be inspected, in addition to considering whether or not to shutdown components with or without first making an inspection. Thus, the number of decision alternatives is large and using a decision tree is not practical. Additional details on decision trees can be found in a number of sources (Benjamin and Cornell 1970; Raiffa 1997; Jordaan 2005).

An alternative to the decision tree is the *influence diagram* (ID). An ID can generally be viewed as a compact representation of the same information contained in the decision tree. The following brief description of IDs is taken from Kjaerulff & Madsen (2008) and Jensen & Nielson (2007). The reader is referred to these sources for details beyond the overview presented here.

An ID is a probabilistic network used to aid decision-making under uncertainty. It encodes both the probability model and the decision problem structure, which represents a sequence of observations, outcomes, and decision alternatives. Solving a decision problem via ID amounts to (1) calculation of the expected utility associated with each decision alternative, and (2) selecting the optimal decision strategy, i.e. the strategy which maximizes expected utility from the perspective of the decision-maker. In decision problems there are two types of decisions: (1) action decisions, which proactively "change the state of the world" through some activity, and (2) test decisions, i.e. decisions to look for more evidence that can be included in the model before an action decision is made (Raiffa and Schlaifer 1970; Friis-Hansen 2000). The inspection decision described above is an example of a test decision. The shutdown decision represents an action decision.

Like a BN, an ID consists of nodes representing discrete random variables (hereafter called chance nodes) that are connected by directed links and are associated with CPTs. An ID differs from a BN because it is extended by utility and decision nodes. A utility node represents the value associated with a specific decision(s) and/or the outcome of a chance node(s). A utility node is commonly represented by a diamond, though some authors and software applications use other shapes. A utility node has no states; instead it is assigned a utility value (e.g. monetary units) as a function of the states of its parent variables. A utility node cannot have children

A decision node encodes alternative actions available to the decision maker. It is represented by a rectangle and is defined by a finite set of states corresponding to decision alternatives. To define an order on decisions and observations in the ID, precedence and information links are added to the graphical model. A link going from a chance node to a decision node is referred to as an information link and indicates that the value of the chance node is observed before the decision is made. A precedence link connects decision nodes and represents an order on decisions, i.e. the parent decision is made before the child decision. It is important to emphasize that the links going into a decision node are not associated with quantitative requirements; the links strictly specify the information that is available to the decision maker at the time the decision is made. By specifying a required order on decisions and observations, precedence and information links reduce the solution space that must be searched to find an optimal decision (Jensen 2001a). The perfect recall *ID* is the conventional form of an ID. It is associated with a structural requirement in which the ID describes a temporal sequence on all decisions. In other words, there must be a directed path through the ID that contains all decision nodes. Several challenges arise when constructing IDs with a directed path between all decision nodes: (1) not all decision problems can be reasonably formulated with a known ordering on decisions; and (2) considering all preceding decisions may result in an intractably large number of decision sequences. Perfect recall IDs are based on a "no forgetting" assumption (Jensen and Nielson 2007). This assumption means that there is only one order in which decisions can be made and that, when making each decision, the decision-maker remembers all preceding observations, outcomes, and decisions. A decision problem with these properties is called symmetric because it is possible to formulate it as a symmetric decision tree. In symmetric decision trees, all paths from a root node to a leaf node include all variables in the problem. A decision problem is asymmetric if a decision tree representation of the problem exists in which not all paths from a root node to a leaf node include all variables in the problem (Shenoy 1996).

Three types of asymmetry exist in decision problems: (1) functional asymmetry, in which outcomes and decision options may vary as a function of the past, (2) structural asymmetry, in which the occurrence of an observation or decision in the network is a function of the past, and (3) order asymmetry, in which the ordering of decisions and observations in the network does not follow a predefined sequence, i.e. the order is not known when the ID is constructed. The inspection-shutdown problem described previously contains decisions that can be performed in any order, e.g. the owner can decide to inspect bridge *i* before bridge *j* or vice-versa. The order in which the inspections are performed is a desired outcome of the decision formulation. The inspection prioritization problem as formulated in this report is order asymmetric.

Order asymmetric problems cannot be modeled using perfect recall IDs. However, two more general classes of IDs called *limited memory influence diagrams* (LIMID) and *unconstrained influence diagrams* (UID) relax the requirement that there be an explicit order on decisions. A LIMID drops the no-forgetting assumption and, instead, assumes that only nodes that are explicitly represented as parents to a decision node are known at the time the decision is made. Thus, the LIMID solves decision problems with a smaller domain and, therefore, the solution of a LIMID is an approximation of the solution obtained for a perfect recall ID (Jensen 2001b). UIDs permit decision problems in which it is not just the

optimal choice for each decision that is of concern but also the best ordering of those decisions. The result is an exponential growth in complexity when using exact solution algorithms. Additional details on UIDs can be found in a variety of sources (Jensen and Vomlelova 2002; Jensen and Nielson 2007; Luque et al. 2008). In this study, we utilize LIMIDs, which use an iterative approximate algorithm for solution and are supported in popular commercial BN software, e.g. Hugin (Hugin Expert A/S 2008). LIMIDs are described below in greater detail.

## 7.2.1 Overview of Limited Memory Influence Diagrams

Recall that a LIMID differs from a perfect recall ID due to the relaxation of the no-forgetting assumption. Rather than considering that all preceding decisions are known at the time a decision is made, only the states of nodes that are direct parents to the decision node are assumed to be known. The types of links found in a LIMID are shown in Figure 7.1. These links have the following meanings. Diagram (a) in Figure 7.1 shows a link between two chance nodes which, as in previous chapters, indicates that the distribution of  $X_2$  is defined conditionally on  $X_1$ . Diagram (b) has a chance node X that is dependent on a decision node *D*, indicating that the distribution of *X* depends on the alternative of *D* that is selected. The link between decision nodes in Figure 7.1c indicates temporal precedence, which is necessary for the solution algorithm. It does not define a conditional distribution as with the links in Figure 7.1a-b. Figure 7.1d shows an information link going from chance node X to decision node *D*. This information link implies that the state of *X* is known before making decision *D* and may influence its outcome. Last, Figure 7.1e shows a utility node *U* that is a child of chance node X and decision node D. This relationship indicates that the decision-makers' utility is dependent on the outcome of random variable X and the selected decision alternative. Recall that a utility node cannot have children. These connection types will be referenced again later to provide specific examples of each type.

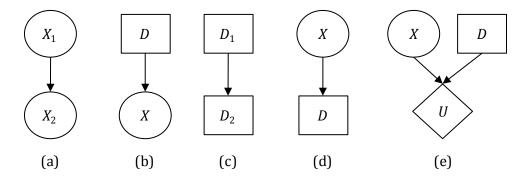


Figure 7.1: Types of links in limited memory influence diagrams

#### 7.2.1.1 Computations in Limited Memory Influence Diagrams

Computations in LIMIDs are explained here using a simplified adaptation of a more detailed explanation contained in Lauritzen & Nilsson (2001); details of mathematics and specifics of the algorithm are available in that source. When solving either a perfect recall

ID or a LIMID, the goal is to determine a *policy* for each decision node that maximizes the expected utility for any given configuration of its precedent nodes. In the case of a LIMID, this policy is defined for each combination of the states of the direct parents of the decision node, i.e. for any given combination of states of the parents of decision *D*, the policy specifies the action to take. Note that a policy may be deterministic or stochastic. The set of utility maximizing policies for all decision nodes in an ID or LIMID is referred to as a *strategy*. For a perfect recall ID, this strategy is a global optimum.

Because LIMIDs do not require a temporal ordering on decisions, the possible combinations of orders on decision alternatives can be extremely large. Rather than explicitly consider all combinations and possible orders of decision alternatives, (locally) optimal solutions are determined using an iterative procedure known as *single policy updating*. The algorithm begins with an initial strategy. This strategy can, and typically should, be random. A *cycle* of the algorithm includes the updating of policies for all decision nodes in the LIMID. Let  $q_i$  be the current strategy. Begin a cycle by updating the first decision  $D_1$ .  $q_{i+1}$  is computed by finding the local maximum policy for  $q_i$  for decision  $D_1$ . The algorithm moves on to compute a locally optimal policy for decision  $D_2$ . When all policies for all decisions have been updated, the cycle is complete. The algorithm converges when the expected utilities associated with successive cycles are the same (Lauritzen and Nilsson 2001). Because single policy updating works locally, it is not guaranteed to give a globally optimal solution.

In computing optimal strategies for the example inspection-shutdown decision considered in this chapter (described in detail later), problems with convergence to solutions that were not optimal at the system (global) level were occasionally encountered. Because the structure of the example problem is predictable, it was possible to recognize these situations and modify strategies accordingly.

# 7.3 Post-earthquake inspection and closure decision

Within the context of the aforementioned inspection-shutdown decision, we consider two decision problems. First we consider decision-making at the component-level without addressing system-level effects. Second we consider decision-making when accounting for the effects of decisions at the system level.

# 7.3.1 Component-level decision-making

The decision to shut down or reduce the capacity of a component is made under competing objectives: on the one hand the owner does not want to lose revenue by unnecessarily shutting down or reducing the operating level of a component that may not have experienced serious damage, while on the other hand the owner does not want to incur a liability by making an unsafe decision, i.e. keeping the component open when it may have sustained serious damage. To reduce the uncertainty associated with this decision, the owner may elect to conduct an inspection of the component, incurring a certain cost, which will yield information about the state of the component. The decision to close or not close

the component will then be made after receiving information gained from the inspection. To develop an ID for the component-level joint inspection-shutdown decision, we first individually consider three specific decision scenarios: (a) the owner must decide whether to shut down or continue operation of the component without the option to conduct an inspection; (b) the owner makes the shutdown/operation decision knowing the damage state of the component with certainty; and (c) the owner makes the shutdown/operation decision with imperfect information gained from an inspection of the component. Temporarily, we neglect costs associated with the inspection.

The ID corresponding to the component shutdown/operation decision without consideration of inspection is shown in Figure 7.2a. We do not distinguish between IDs and LIMIDs here because there is only one decision being made. In Figure 7.2a, the true damage state of component *i* is modeled by node  $C_i$ , which is defined conditionally on the ground motion intensity at the site, node  $S_i$ . Node  $S_i$  has mutually exclusive, collectively exhaustive interval states spanning the range of possible ground motion intensities at the site of the component. Node  $C_i$  has states corresponding to component damage levels, e.g. not damaged, slightly damaged, moderately damaged, severely damaged. The decision node Shutdown? corresponds to the decision whether or not to shut down or reduce the operating level of the component. For notational brevity, the node corresponding to this decision is labeled *Shutdown*?, however it should be understood that this node need not be binary and can include decision alternatives corresponding to actions that change the operating level of a component without completely shutting it down. Utility node  $L_I$  is a child of nodes C<sub>i</sub> and Shutdown?. This is an example of connection type (e) shown in Figure 7.1. The utility node  $L_i$  models potential loss associated with outcomes of the true component damage state and the shutdown decision. It assigns a value to every combination of the states of its parent nodes. A liability is associated with making an unsafe decision, i.e. keeping the component open when it is damaged. A revenue loss is incurred whenever the owner chooses to shut down or reduce the operating level of the component, regardless of its damage state. If the component is undamaged and remains in operation, no loss is incurred.

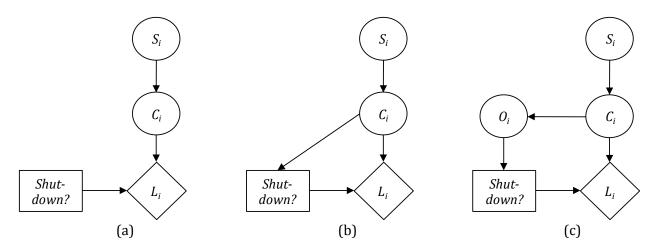


Figure 7.2: IDs modeling shutdown/operation decision based on (a) no information; (b) perfect information; and (c) imperfect information about the damage state of component *i* 

Next, consider the decision scenario in which the owner has perfect information about the damage state of the component. This situation corresponds to the case in which a high-quality and thorough inspection of the component is performed. The ID corresponding to this decision scenario is shown in Figure 7.2b. The information link between node  $C_i$  and decision node *Shutdown*? corresponds to link type (d) in Figure 7.1. The link indicates that the true damage state of the component is known prior to making the decision. Thus, the optimal policy will be defined for each true component damage state.

However, inspections do not always yield accurate results, particularly when a cursory or visual inspection is performed. In the absence of a perfect inspection, the true component damage state is a latent variable. The ID corresponding to the shutdown decision when an inspection provides imperfect information is shown in Figure 7.2c. The flow of information between  $C_i$  and the decision node is mediated by node  $O_i$  corresponding to the observed damage state. The conditional relationship between  $C_i$  and  $O_i$  reflects the accuracy of the inspection method. This is commonly described by the test likelihood, i.e., conditional probabilities of each observation given the true state of the component. The solution of this ID yields a decision policy for each observed damage state, the "correctness" of the shutdown decision is not guaranteed. The three formulations in Figure 7.2 correspond to common forms of influence diagrams used to model decisions made under no information, perfect information, and imperfect information.

The IDs in Figure 7.2 must be coalesced onto one ID to solve the inspection-shutdown decision at the component level. In the inspection-shutdown decision problem, the decision-maker first decides whether to inspect the component (and what inspection quality to use), and then decides whether to shut down the component based on the inspection results, if an inspection was performed. The coalesced ID must behave as Figure 7.2a when an inspection is not performed, as Figure 7.2b when perfect information is

available, and as Figure 7.2c when an imperfect inspection is performed. The ID in Figure 7.3 achieves the desired behavior by adding a node that defines the inspection type (no inspection, imperfect inspection, perfect inspection). It is assumed that the inspection decision choice, and the associated outcome if an inspection is made, is known before a decision is made regarding component shut down. The precedence link between *Inspect?* and *Shutdown*? is of the type in Figure 7.1c. It encodes an ordering on these decisions. Defining this order encodes the value of the information obtained by performing the inspection. The link from *Inspect*? to  $O_i$  is of the type in Figure 7.1b. This link differentiates the three decision scenarios in Figure 7.2. It also encodes the inspection quality relationship. If a decision is made not to inspect the component (i.e. *Inspect*? = no inspection), the distribution of  $O_i$  is uniform over all its states, regardless of the state of its other parent node  $C_i$ , indicating that the true component damage state remains a latent variable. If *Inspect*? = perfect inspection, the state of  $O_i$  is equal to the test likelihood.

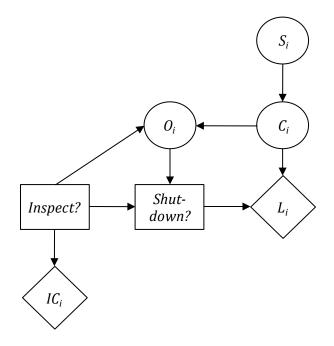


Figure 7.3: ID modeling inspection-shutdown decision at the component level

In the above, we neglected inspection costs. Associated with the inspection choice is a cost which depends on the inspection type. In Figure 7.3, the utility node  $IC_i$  is functionally dependent on the inspection decision node and represents the cost of performing an inspection, which normally increases as the inspection quality improves. Once again, as a function of the closure decision and the true component damage state, there is a utility node that represents the potential losses associated with the outcomes of the shutdown decision and the true component damage state, i.e. losses due to shutting down the component as well as liability. This perfect recall ID (designated perfect recall because it

contains a total order on decisions) can now be used to make decisions at the componentlevel. The formulation can be solved using exact procedures that yield a global optimum.

# 7.3.2 System-level decision-making

While the above description focuses on immediate post-event decisions made at the component level, decisions regarding post-event inspection and closure of components should ideally be made at the system level. This requires creation of an ID which contains a decision model, of the form in Figure 7.3, for each component. The result is a LIMID as shown in Figure 7.4. Because there is no longer an explicit order among all the decision nodes, the LIMID must be solved using the single policy updating procedure, which provides an approximate solution. To account for correlations among demands placed on the components, the seismic demand model is included. Consequences of shutting down or reducing the capacity of the component (e.g. lost revenue) should be calculated by considering the effect of the operating level of the component on the performance of the system. This is achieved by introducing the system performance model into the LIMID, as shown near the bottom of Figure 7.4. A utility node representing lost revenue at the system level is attached to node  $S_{sys}$ . If the system state is defined as a binary node, then the utility will be defined as a loss associated with system closure and no loss associated with system operation. If the system has multiple states, then each state below full capacity will be associated with a loss. Consideration of losses at the system level allows the decision framework to account for redundancies in the system.

To account for multiple operating levels, node  $Cap_i$  is introduced at the component level as a child of the node representing the true damage state and node *Shutdown comp i*?. Node  $Cap_i$  represents the component capacity or operating level, which is functionally dependent on the damage state of the component and the actions of the decision-maker. This allows for a multi-state mapping between the component damage state  $C_i$  and the operating state of the component,  $Cap_i$ . Accordingly, the component operates at reduced/zero capacity if it is damaged or if a decision is made to shut it down. The system node in Figure 7.4 is a child of nodes  $Cap_i$ , which represents the dependence of the system performance on the operating levels of the components.

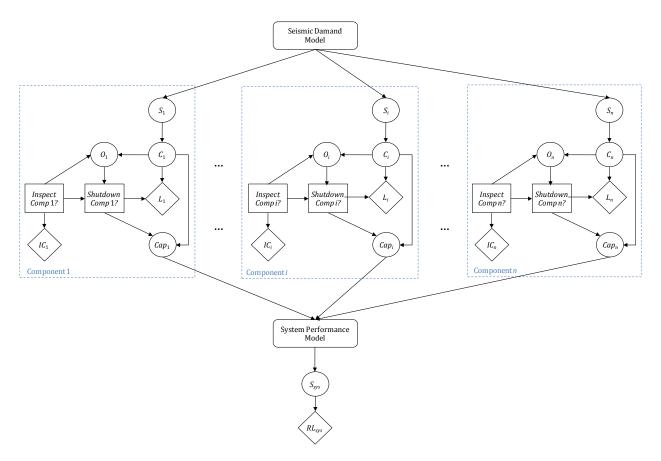


Figure 7.4: ID modeling joint inspection-shutdown decision at the system level

For a given post-earthquake scenario, the solution for a particular component obtained from the LIMID may indicate that the component should be kept in full operation or have its operating level reduced/shutdown without first making an inspection, while other components may require an inspection before the shutdown decision is made. The LIMID will indicate a component should be kept in operation without inspection if the expected loss due to liability is less than the expected loss at the system level due to the shutdown of the component, regardless of the information that can be gained through inspection. Conversely, the LIMID will indicate a component should be shutdown or have its operating level reduced, without inspection, if the expected loss due to liability exceeds the expected value of the component to the system, regardless of what information can be gained by inspection. For components for which there is value to be gained through inspection, it is necessary to have a method of prioritizing the order in which inspections are performed. In this study, inspection prioritization is based on a value of information-based heuristic, as described below. This heuristic represents a first attempt at solving post-earthquake decision problems involving component shutdown and inspection using the proposed BNbased methodology. Development of a more sophisticated decision framework based on the preferences of actual decision-makers remains an area warranting further study.

#### 7.4 Value of information heuristic

The *value of information* is a concept commonly employed in decision analysis to quantify the amount of money or resources a decision-maker is willing to expend to acquire more information before making a decision. In the context of the joint inspection-shutdown decision, illustrative calculations of the values of perfect and imperfection information at the component level are demonstrated by example in Appendix 7.1.

We first consider the component-level problem. Let EU|NI be the maximum expected utility associated with the available decision options given no information, i.e. the maximum expected utility associated with the ID in Figure 7.3 when decision node *Inspect*? is set equal to the no inspection option. Define EU|PI as the expected utility associated with making the optimal decision given perfect information. EU|PI is the expected utility associated with the optimal strategy in Figure 7.3 when decision node *Inspect*? is set equal to the perfect inspection option. Finally, define EU|II as the expected utility associated with making the optimal decision given imperfect information, i.e. the expected utility associated with Figure 7.3 when decision node *Inspect*? is set equal to the perfect inspected utility associated with Figure 7.3 when decision node *Inspect*? is set equal to the imperfect inspection option. The expected value of perfect information (excluding the cost to obtain the information) is

$$EVPI = EU|PI - EU|NI$$
(7.3)

and the expected value of imperfect information (excluding information acquisition costs) is

$$EVII = EU|II - EU|NI \tag{7.4}$$

To facilitate the prioritization of post-earthquake inspections of components at the infrastructure system level, we define a value of information-based heuristic. The heuristic assumes that, at each stage, the decision-maker is looking for the "next best component" to inspect. It is based on an assumption that only one inspector is available. It asks: if only one more component can be inspected, which component should it be? Furthermore, the heuristic does not account for issues related to the time it takes to inspect a component -- something that is important in post-earthquake applications and should be addressed in further studies. The LIMID in Figure 7.4 is used to determine the optimal inspection decision for each component. For components for which inspection is deemed unnecessary, the optimal shutdown decision is made according to the policy that maximizes the expected utility. For these components, the policy on *Shutdown Comp i*? will be a deterministic policy indicating whether, in the absence of any additional information, it is locally optimal to keep the component open or shut. (Note that, in contrast, when an inspection.)

For the components for which an inspection is recommended, the value of information is computed as follows. First, the total system-level expected utility, neglecting inspection

costs, is computed using the LIMID and assuming that all components requiring inspection will be inspected according to the optimal inspection type indicated. This is analogous to quantities *EU*|*PI* and *EU*|*II*, depending on the type of inspection, as described above. Then, for component *i* requiring inspection, the inspection decision node is set to the "no inspection" state. The remaining components are left at their optimal decisions. Then, the system level utility when the component inspection is disregarded (output from the LIMID) is recorded, not including inspection costs. This is analogous to  $EU^*|NI$ , which was described above for the component level problem. The value of information is computed as the difference between the two quantities. The value of information gained from the inspection is thus the difference between (1) the system level utility that is expected when inspecting all recommended components and (2) the expected system-level utility when the inspection of component *i* is disregarded. The component with the highest value of information, i.e. the component for which neglecting to perform an inspection has the most adverse impact on the system-level expected utility, is inspected first. The procedure is repeated until all components for which an inspection is recommended are inspected. In the event that multiple inspectors are able to visit *n* components simultaneously, the above procedure is modified to select the *n* components for which the value of information is highest at each stage.

Using the LIMID methodology, the inspection prioritization order evolves as new information becomes available. Sources of information may include measurements of ground motion intensity and information from structural health monitoring sensors. Furthermore, as inspections are performed and decisions are made regarding the shutdown of components, this information is also entered into the LIMID. The information propagates through the LIMID to provide an up-to-date probabilistic characterization of the system model and decision structure. At any stage (e.g. following the inspection of a component and subsequent entry in the model of inspection results), the recommendations regarding component inspections may differ from those made at a previous stage, e.g. a component which was previously deemed to require an inspection may no longer need one, and vice versa. Thus, the LIMID provides the decision-maker with guidance on optimal decisions relating to inspection and component closure, at any point in time, based on all available information up to that time. Note that, the above heuristic does not consider all possible orderings on the inspection and shutdown decisions explicitly -- a task which is computationally expensive and possibly intractable. Therefore the above procedure may arrive at a suboptimal solution.

#### 7.4.1 Example application

To illustrate the methodology described above, consider the simple infrastructure system shown in Figure 7.5. This system has a predictable topology that facilitates intuitive interpretation of results obtained from the LIMID analysis. The system contains a source and a sink node and eight components. It is assumed that, when operational, components 1-6 carry 25 units of capacity each, while components 7 and 8 each have a capacity of 100 units. Observe that components 7 and 8 form a system bottleneck. Thus, when all components are fully operational, the system carries a maximum of 100 units from the

source to the sink. Each component is assumed to have binary states, i.e. either the component carries the full capacity or its capacity is zero. The performance of the system is defined considering multiple states corresponding to the total capacity of the system. Let  $Cap_i$  indicate the operating capacity of component *i* at a given time. The system operating capacity is computed as a function of the individual component operating capacities using the expression:

$$Cap_{svs} = \min[\min(Cap_7, Cap_8), Cap_1 + Cap_2 + Cap_3 + \min(Cap_4, Cap_5, Cap_6)]$$
 (7.5)

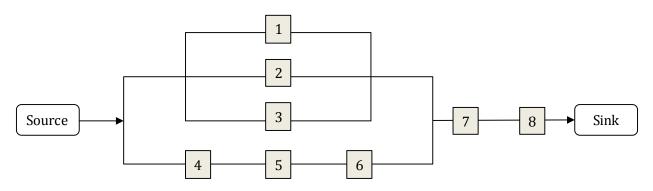


Figure 7.5: Example infrastructure system

LIMIDs are constructed for the system and the above prioritization procedure is executed considering varying conditions: (1) assuming all components have identical versus heterogeneous reliabilities; (2) accounting for correlation in the performance of components of the same class versus neglecting correlation; and (3) accounting for or neglecting the effect of system performance on the optimal decisions, i.e. ignoring the system object in Figure 7.4. The result is 8 combinations of cases under which the above methodology is carried out. These are described in more detail below.

In the cases in which system components are assumed to have identical reliabilities, all components have a probability 0.05 of failure (reliability = 0.95). When components are treated as heterogeneous, they are grouped into five classes with reliabilities given in Table 7.1.

Component	Class	Reliability
1	1	0.95
2	2	0.96
3	2	0.96
4	3	0.99
5	1	0.95
6	3	0.99
7	4	0.96
8	5	0.98

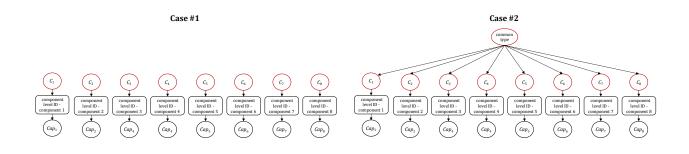
#### Table 7.1: Component classes

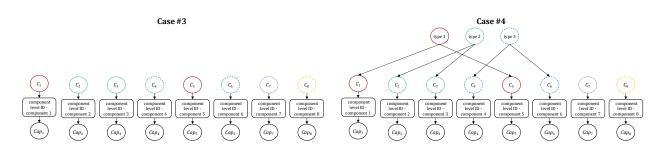
For simplicity, we do not consider a seismic demand model. Instead we consider a simple and intuitive correlation structure. In the cases in which correlation is considered, it is modeled through a common parent node that has binary states. Components of the same class are connected through this common parent node with marginal PMF defined according to the failure probabilities in the above table. The CPTs of the child nodes, which represent the states of the individual components, are defined such that there is a 0.8 probability that the component will have the same state as its parent node and a 0.2 probability that it will not.

The eight cases that arise from considering the combinations of identical/heterogeneous components, including/neglecting correlation, and including/neglecting system effects are summarized in Table 7.2. The cases are also shown graphically in Figure 7.6. In this figure, the component-level IDs are hidden behind objects. Furthermore, colors and line textures are used to distinguish components of different types.

Case	System effects	Components	Demand correlation
1	Neglected	Identical	Independent
2	Neglected	Identical	Correlated
3	Neglected	Heterogenous	Independent
4	Neglected	Heterogenous	Correlated
5	Included	Identical	Independent
6	Included	Identical	Correlated
7	Included	Heterogenous	Independent
8	Included	Heterogenous	Correlated

Table 7.2: System configurations for example inspection-shutdown decisions





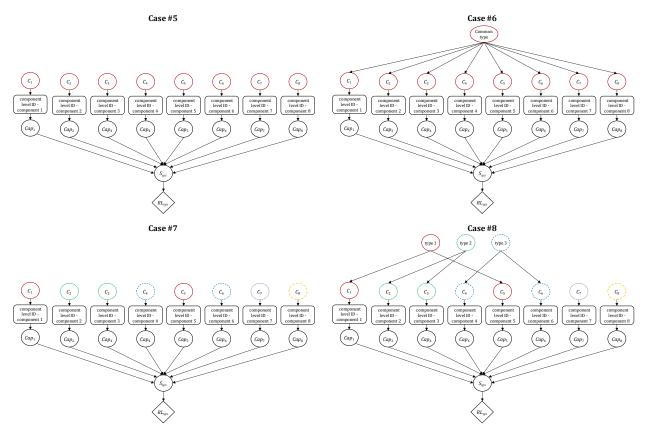


Figure 7.6: System configurations for example inspection-shutdown decisions

For all cases, the test likelihood associated with an imperfect inspection is:

	Component undamaged	Component damaged
Component observed to be undamaged	0.9	0.25
Component observed to be damaged	0.1	0.75

### 7.4.1.1 Assumed utilities

In addition to defining the CPTs of nodes, it is necessary to specify utility values. For cases 1-4, the assumed utility values are given in Table 7.3. In the cases that components have identical reliabilities (cases 1 & 2), all components are assumed to have the same utility values. For the cases in which system effects are neglected and components are heterogeneous (cases 3 & 4), the assumed utilities values depend on the component classes. Because components 7 & 8 carry a higher capacity, they are assumed to be associated with higher losses at the component level than components 1-6.

#### Table 7.3: Assumed utility values for cases 1-4

Utility type	Components 1-8 for cases 1 &2 Components 1-6 for cases 3 & 4	Components 7 & 8 for cases 3 & 4
Liability	-250 cost units/component	-1000 cost units/component
Loss due to component being non-operational	-25 cost units/component	-100 cost units/component
Perfect (imperfect) inspection cost	-5 (-2.5) cost units/component	-10 (-5) cost units/component

The utilities associated with the cases 5-8 are given in Table 7.4. Note that there are no explicit costs associated with an individual component being non-operational; instead costs are incurred at the system level, as a function of the system operating capacity.

#### Table 7.4: Assumed utility values for cases 5-8

Utility type	Components 1-8 for cases 5 &6 Components 1-6 for cases 7 & 8	Components 7 & 8 for cases 7 & 8	
Liability	-2500 cost units/component	-10000 cost units/component	
Perfect (imperfect) inspection cost	-5 (-2.5) cost units/component	–10 (–5) cost units/component	
System cost function	$Cap_{svs} = 50: -$ $Cap_{svs} = 75: -$	00 cost units 300 cost units 200 cost units 100 cost units : 0 cost units	

The utilities values in the above tables are arbitrary, but have been selected to produce "interesting" results, i.e. results such that the recommended decision actions are not always

the same. To achieve this effect it was necessary to modify the liability values for the cases in which system effects are included. Because the assumed utilities are significantly different when system effects are included, the computed values of information should not be compared between cases 1-4 and cases 5-8. However, the prioritization orders can be reasonably compared.

# 7.4.1.2 Numerical results

For cases 1-4, the value of information associated with making a perfect inspection of each component as well as the prioritization orders are shown in Table 7.5a, assuming no inspections have yet been performed. Table 7.5b presents the same results, but considers that component 1 has been perfectly inspected and found to be damaged. While the possibility of performing imperfect inspections was included in the model, given the model assumptions, it was never recommended as the optimal choice. Therefore, in the following section, if an inspection is recommended, it should be taken to mean a perfect inspection. For the outcomes in these tables for which the no inspection choice is recommended, abbreviated "No Ins," the shutdown decision that is prescribed, in lieu of inspection, is shown in parentheses with "SD" indicating component shutdown and "O" indicating the component is to remain operational.

# Table 7.5: Value of information and prioritization rankings for cases 1-4 given (a) no evidence and (b) component 1 has been perfectly inspected and found to be damaged (VoI = Value of information)

	Cas	se 1	Cas	se 2	Cas	se 3	Ca	se 4
Component	VoI	Ranking by VoI	VoI	Ranking by VoI	VoI	Ranking by Vol	VoI	Ranking by VoI
1	11.25	1	19.25	1	11.25	3	19.25	4
2	11.25	1	19.25	1	9	4	19.4	3
3	11.25	1	19.25	1	9	4	19.4	3
4	11.25	1	19.25	1	No Ins (O)	-	19.85	2
5	11.25	1	19.25	1	11.25	3	19.25	4
6	11.25	1	19.25	1	No Ins (O)	-	19.85	2
7	11.25	1	19.25	1	36	1	36	1
8	11.25	1	19.25	1	18	2	18	5

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#### (b)

	Cas	se 1	Cas	se 2	Cas	se 3	Cas	se 4
Component	VoI	Ranking by VoI	VoI	Ranking by VoI	VoI	Ranking by Vol	VoI	Ranking by VoI
1	-	-	-	-	-	-	-	-
2	11.25	1	17.4	1	9	4	19.4	3
3	11.25	1	17.4	1	9	4	19.4	3
4	11.25	1	17.4	1	No Ins (O)	-	19.85	2
5	11.25	1	17.4	1	11.25	3	17.4	5
6	11.25	1	17.4	1	No Ins (0)	-	19.85	2
7	11.25	1	17.4	1	36	1	36	1
8	11.25	1	17.4	1	18	2	18	4

For the unconditional scenario in which no inspections have yet been performed (Table 7.5a), cases 1 and 2 indicate that all components have equal importance. This is expected because all components are identical and system effects are not included. The difference between cases 1 and 2 is due to the inclusion of correlation among the components. This inclusion is only important when information becomes available about the state of a component, as shown in Table 7.5b. In the absence of information about the states of other components, the inclusion of correlation only has the effect of increasing the value of information. However, it does not alter the prioritization order. Note that, when correlation is included, inspecting one component will provide partial information about other components.

Cases 3 and 4 include heterogeneous components. For case 3, components 4 and 6, which have the highest reliability, do not require an inspection and can be kept in operation. An inspection is recommended for the remaining components. Among these components, component 7 is deemed the most critical, followed by component 8. This is expected

because these components are associated with the highest liability. Component 7 has a reliability of 0.96, the same as components 2 and 3, but because of the difference in potential liability, component 7 should be inspected before the other two. While components 7 and 8 have the same liability associated with making an unsafe decision, component 7 is recommended to be inspected before component 8 because component 8 is more reliable. When evidence is observed about the state of component 1, there is no effect on component values of information because correlation between components is not included.

Case 4 adds correlation to the model contained in case 3. The inclusion of correlation once again has the effect of increasing the value of information for any component that is correlated to another component. Note that, in the absence of correlation, it was not recommended that components 4 and 6 be inspected. However, when correlation is included, the recommendation changes and there is value in inspecting these components. Furthermore, the prioritization order of the components changes from the order recommended in case 3. When the observation has been made about the performance of component 1, all values of information remain unchanged, relative to the unconditional evidence scenario, except for component 5, which is correlated with component 1. The inspection of component 1 provides information not only about the state of component 1, but also about the likely state of component 5. Therefore, the value of information associated with an inspection of component 5 decreases. It is now lowest in the prioritization order.

Next, we consider the system-level problem (case 5-8). Case 5 assumes all components have identical reliabilities and liabilities. The results shown in Table 7.6a for case 5 reflect the importance of each component to the system. Components 7 and 8 are part of the system bottleneck and are therefore given first priority. Components 1-3 exist in parallel. Operation of any one of these components will allow some flow through the system. Furthermore, because they have the same reliabilities and utility values, and provide the same function to the system, there is no differentiation between them with regard to prioritization. Components 4-6 exist in series and have the same reliabilities. All these components must be operational in order for the link to be operational. Because of the relatively low probability that the link will be operable (0.95 \* 0.95 = 0.86), they are given lowest prioritization order because this case assumes no correlation between the component states. However, the expected losses associated with the remaining components increase because of the change in the system configuration and its reduced redundancy.

# Table 7.6: Value of information and prioritization rankings for cases 5-8 given (a) no evidence and (b) component 1 has been perfectly inspected and found to be damaged (VoI = Value of information)

	Cas	se 5	Cas	se 6	Cas	se 7	Ca	se 8
Component	VoI	Ranking by Vol	VoI	Ranking by Vol	VoI	Ranking by Vol	VoI	Ranking by VoI
1	86	2	49	2	89	4	72	4
2	86	2	49	2	90	3	73	3
3	86	2	49	2	90	3	73	3
4	77	3	31	3	25	6	46	5
5	77	3	31	3	88	5	46	5
6	77	3	31	3	25	6	46	5
7	125	1	177	1	358	1	264	1
8	125	1	177	1	200	2	200	2

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#### (b)

	Cas	se 5	Cas	se 6	Cas	se 7	Ca	se 8
Component	VoI	Ranking by VoI	VoI	Ranking by VoI	VoI	Ranking by Vol	VoI	Ranking by VoI
1	-	-	-	-	-	-	-	-
2	86	2	42	2	90	3	73	2
3	86	2	42	2	90	3	73	2
4	77	3	27	3	25	5	42	3
5	77	3	27	3	88	4	42	3
6	77	3	27	3	25	5	42	3
7	125	1	112	1	268	1	188	1
8	125	1	112	1	200	2	188	1

The inclusion of correlation in case 6 with identical component reliabilities has the effect of increasing the value of information associated with components 7 and 8 and decreasing that associated with components 1-3. This is because components 7 and 8 are in series and form the system bottleneck, while components 1-3 are in a parallel configuration. However, the prioritization order remains unchanged. The observation of component 1 being damaged reduces the value of information associated with all components for this case, though the prioritization order is unaltered. This is because the observation decreases the reliability of all components due to the common correlation structure. Therefore, when an inspection is made, it is less likely that the component will be observed to be undamaged and thus able to be kept in operation. In the absence of an inspection, the optimal decision is to shut the component down rather than risk a liability. Therefore, there is smaller difference between the maximum expected utility when the component is and is not inspected, due to revised reliability estimate of the components, given the observation of damage to component 1.

Cases 7 and 8 consider heterogeneous components as well as system effects. Case 7 neglects correlation between components. Thus component importance is governed jointly by the component reliabilities/liabilities as well as the importance of the component to the system. As expected, component 7 is given first priority followed by component 8 (recall that component 7 is less reliable than component 8). Note that the value of information for component 5 is higher than the value of information associated with components 4 and 6, even though the three components lie in series. This is because, in the absence of the inspection of component 4 or 6, the optimal decision is to keep these components in operation due to their relatively high reliabilities and low liability. However, when inspection of component 5 is disregarded, the optimal action is to shutdown that component (due to its low reliability), and thus the entire link becomes inoperable. Consequently, there is a relatively high value of information in inspecting component 5.

In cases 7 and 8, the values of information associated with components 2 and 3 are the same. This is expected, because they have the same reliabilities/liabilities and importance to the system. The value of information associated with component 1 is less than the value of information associated with components 2 and 3. In the absence of an inspection of any of the component 1-3, the optimal decision is to shut it down. Components 2 and 3 are more reliable than component 1. Therefore, since the three components are in parallel, it is preferable to shutdown component 1 than components 2 and 3. Thus component 1 is given lower priority. In comparing the results of case 7 before and after observing component 1 is damaged, it is found that the values of information of all but component 7 are unchanged.

Case 8 includes heterogeneous components with correlation and system effects included. It is thus the most comprehensive case considered. Components 7 and 8 are given the same priority due to the effects of the shutdown decision made in lieu of an inspection. In the absence of an inspection of component 7 (or similarly, component 8), the optimal action is to shutdown the component, and thus the entire system, rather than risk a liability. Thus, when the inspection of component 7 (or component 8) is not performed, the reliability of the component is inconsequential because the component will be shutdown and thus the total system-level loss is 400 units (recall Table 7.6). When component 7 or component 8 is inspected, the system level utility is the same because all components are inspected in accordance with the optimal component-level inspection policy. Therefore, the value of information is the same for components 7 and 8. Components 2 and 3 are given priority over component 1 for the reason described above. Components 4-6 are given lowest priority. Overall, the value of information associated with each component is lower when correlation is included (i.e. versus case 7) due to changes in system level reliability. The observation of component 1 being damaged has no effect on the value of information associated with components 2 and 3, given the number of significant digits reported. However, the observation reduces the value of information associated with components 7 and 8 as well as components 4-6. The reason for this is as follows. In the absence of an inspection of component 7 (or 8), the optimal decision is to shutdown the component and thus the system. This results in a revenue loss of 400 units, regardless of the operating levels of the remaining system components. Before the observation of damage to component 1 is made, the expected system level utility is higher than after the observation

is made. This is because damage to component 1 implies the system can, at best, operate at 75% its full operating level. Therefore, the maximum expected utility when component 7 (or 8) is inspected, for the case in which component 1 is not observed to be damaged, is larger than after the damage is observed. The value of information for component 7 (or 8) is computed as the difference between the maximum expected utility assuming the component is inspected and the maximum expected utility when it is not inspected. This difference turns out to be smaller when component 1 is observed to be damaged. A similar argument applies to components 4-6 and the shutdown of the link.

# 7.5 Summary

The goal of this chapter is to demonstrate how BNs are extended by utility and decision nodes to create influence diagrams (IDs). Limited memory influence diagrams (LIMIDs) are IDs that do not require an ordering on decisions; therefore all IDs are special cases of LIMIDs. LIMIDs are useful for decision-making under uncertainty when multiple decisions without predefined orders must be made. To demonstrate the usefulness of the BN/ID framework for solving decision problems, preliminary ID models are constructed to address immediate post-event decision making in which an infrastructure owner must decide whether or not to inspect and/or shutdown components. For components for which inspections are recommended, the order in which inspections are performed must be prioritized. A heuristic is developed that seeks to find the "next best component" to inspect considering that components may need to be shutdown if not inspected. Development of decision models for other problems, or alternative formulations and heuristics for the problems described here, remain topics for future study.

# Appendix 7.1: Example calculation of the value of perfect and imperfect information

The following provides illustrative calculations of the value of information for the joint inspection-shutdown decision problem at the component level. These are intended to demonstrate the concept of the value of information. With the help of a software such as Hugin (Hugin Expert A/S 2008), these calculations are easily performed once the ID is defined.

# Example: calculating expected value of perfect information

Consider the IDs in Figure 7.2a-b, which respectively model the cases that no information and perfect information are available to the decision-maker about the state of the component before a decision is made. We now demonstrate calculation of the value of perfect information. Assume the component state is binary (damaged/undamaged) and set the prior probability that the component is damaged to be 0.05. Assign liability and cost-of-closure utilities according to the following table:

Shut-down?	No (C	)pen)	Yes (Shut-down)		
Ci	undamaged	damaged	undamaged	damaged	
Utility	0	-1000	-100	-100	

In the absence of information about the component damage state (Figure 7.2a), the expected utility, *EU*, associated with each decision is:

$$EU(open) = 0.95(0) + 0.05(-1000) = -50$$
  
 $EU(closed) = 0.95(-100) + 0.05(-100) = -100$ 

It follows that, in absence of additional information, the optimal decision is to keep the component open. Define EU|NI as the maximum expected utility associated with the available decision options, given no information other than prior probabilities. Based on the above calculation, EU|NI = -50 units (e.g. dollars).

Having perfect information (Figure 7.2b) implies the optimal decision is made for any outcome of  $C_i$ . If  $C_i$  is in the damaged state, the optimal decision is to shut down the component for a loss of 100 units. If  $C_i$  is in the undamaged state, the optimal decision is to the keep the component open and thus incur no penalty. Thus, given perfect information (*PI*), the expected utility is

EU|PI = 0.05(-100) + 0.95(0) = -5

It follows that the value of obtaining perfect information, EVPI, is

$$EVPI = EU|PI - EU|NI = -5 - (-50) = 45$$

The value of performing an inspection capable of providing perfect information is 45 units of utility.

#### Example: calculating the expected value of imperfect information

Consider the prior damage probabilities and utility values defined above. Also, consider the following test likelihood table to define the mapping between nodes  $C_i$  and  $O_i$ :

	<i>C</i> <sub><i>i</i></sub>	undamaged	damaged
	observe no damage	0.9	0.2
<i>O</i> <sub><i>i</i></sub>	observe damage	0.1	0.8

Let  $c_i$  denote the event that component *i* is not damaged and  $\bar{c}_i$  denote the damaged state. Similarly, let  $o_i$  denote the event that component *i* is observed not to be damaged and  $\bar{o}_i$  indicate that damage is observed. The expected utility associated with imperfect information is computed as follows.

• Compute the probability that a component is damaged/undamaged given damage/no damage is observed:

$$\Pr(c_i|o_i) = \frac{\Pr(o_i|c_i)\Pr(c_i)}{\Pr(o_i)} = \frac{0.9(0.95)}{0.9(0.95) + 0.2(0.05)} = 0.988$$

$$\Pr(c_i|\bar{o}_i) = \frac{\Pr(\bar{o}_i|c_i)\Pr(c_i)}{\Pr(\bar{o}_i)} = \frac{0.1(0.95)}{0.1(0.95) + 0.8(0.05)} = 0.704$$
  
$$\Pr(\bar{c}_i|o_i) = \frac{0.2(0.05)}{0.2(0.05) + 0.9(0.95)} = 0.011$$
  
$$\Pr(\bar{c}_i|\bar{o}_i) = \frac{0.8(0.05)}{0.8(0.05) + 0.1(0.95)} = 0.296$$

• Compute the expected utility associated with each decision alternative given damage/no damage is observed:

 $EU(Open|o_i) = -1000 * \Pr(\bar{c}_i|o_i) + 0 * \Pr(c_i|o_i) = -11$   $EU(Closed|o_i) = -100 * \Pr(\bar{c}_i|o_i) - 100 * \Pr(c_i|o_i) = -100$   $EU(Open|\bar{o}_i) = -1000 * \Pr(\bar{c}_i|\bar{o}_i) + 0 * \Pr(c_i|\bar{o}_i) = -296$  $EU(Closed|\bar{o}_i) = -100 * \Pr(\bar{c}_i|\bar{o}_i) - 100 * \Pr(c_i|\bar{o}_i) = -100$ 

Given no damage is observed, the optimal decision is to keep the facility open with an expected loss of -11 units. Given damage is observed, the optimal decision is to close the facility with an associated expected loss of -100 units. These conclusions are intuitive, but their quantitative values lead to the more interesting result of the expected utility given imperfect information(EU|II):

 $EU|II = -11 * Pr(o_i) - 100 * Pr(\bar{o}_i) = -23$ 

Note that the quantity EU|II is the utility associated with the optimal policy corresponding to the ID in Figure 7.2c. It is therefore output directly from the ID, and the calculations demonstrated above need not be performed manually.

• Compute the expected value of imperfect information:

$$EVII = EU|II - EU^*|NI = -23 - (-50) = 27$$

As expected, the value associated with imperfect information is less than the value associated with perfect information.

# **Chapter 8: Example Applications**

## 8.1 Introduction

As emphasized in previous chapters, the capacity of Bayesian networks to facilitate nearreal time information updating renders them highly useful for post-earthquake decisionsupport. However, BNs are also useful for performing conventional probabilistic seismic hazard assessment (PSHA). The BN framework described in Chapter 3 naturally provides the output of conventional PSHA, i.e. the probability that the ground motion intensity at a site exceeds a given level in a specified time period. The output of the BN can thus be used to construct hazard curves. BNs can also be used to perform hazard deaggregation. However, BNs go beyond conventional PSHA by: (1) providing insight into the mostprobable earthquake scenarios to affect a site; (2) facilitating consideration of hazard for a multitude of post-earthquake situations; and (3) allowing for system hazard assessment by incorporating random field effects and measures of component and system performance.

Two examples are presented in this chapter. The first example considers a simple transportation system and demonstrates how BNs can be applied to perform, and go beyond, conventional PSHA. The second example is based on an idealization of a portion of the proposed California High Speed Rail system. The emphasis of the second application is on post-earthquake information updating. Both examples are intended to demonstrate the capabilities and advantages of the BN framework for seismic hazard assessment and risk management of infrastructure systems.

#### 8.2 Example #1 – Connectivity of a transportation network

Consider the simple transportation network in Figure 8.1. Four cities, represented by circles labeled A-D, are connected to a hospital (circle labeled H) by roadways. The roadways connecting the cities and the hospital must cross a river in six places. Six

identical bridges, represented by numbered squares, are shown in Figure 8.1. The transportation network is located in close proximity to three active faults. The locations of the bridges relative to the faults are shown in Figure 8.2. This figure shows the fault-specific coordinate system associated with source 1. The fault lengths are 65, 83.5, and 105 km, for sources 1, 2, and 3, respectively. It is of interest to know the probability that any city will be disconnected from the hospital due to damage to bridges following an earthquake. It is assumed that the roadways are not prone to damage due to an earthquake.

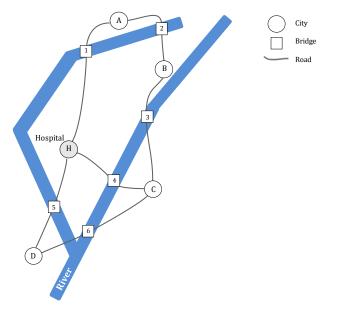


Figure 8.1: Example transportation network

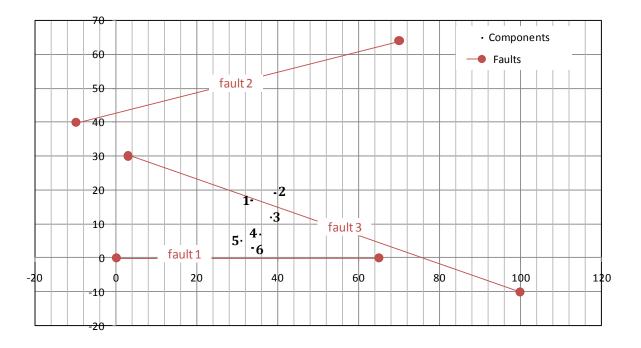


Figure 8.2: Location of transportation system relative to three active faults

#### 8.2.1 Using BN for conventional site-specific PSHA

The goal of PSHA is to quantify the probability of exceeding a given level of ground motion intensity. For a given earthquake occurrence, this probability is defined by:

$$\Pr(S_i > s_i) = \sum_{n=1}^{N_{sc}} \Pr(E_{sc,n}) \int_M \int_{\mathcal{E}} \int_{\mathcal{E}} \Pr(S_i > s_i | m, r, \varepsilon) f_{M,R}(m, r) f_{\varepsilon}(\varepsilon) dm dr d\varepsilon$$
(8.1)

where  $\Pr(S_i > s_i)$  is the probability that the ground motion intensity at site *i* exceeds level  $s_i$  in the event of the earthquake occurring on one of the faults.  $\Pr(E_{sc,n})$  is the probability that the earthquake will occur on source n;  $\Pr(S_i > s_i | m, r, \varepsilon)$  is the conditional probability that the ground motion intensity at site *i*, resulting from the earthquake on source n, exceeds level  $s_i$  for the given magnitude, source-to-site distance and  $\varepsilon$  value;  $f_{M,R}(m,r)$  is the joint distribution of the magnitude (M) and source-to-site distance (R); and  $f_{\varepsilon}(\varepsilon)$  is the distribution of the error on the attenuation model.  $N_{sc}$  (= 3 in the present example) is the number of sources in the vicinity of the site. The integral in (8.1) is a common form used in PSHA, though it does not explicitly show the effects of finite rupture length and directivity, which were discussed in Chapter 3. It also combines the inter- and intra-earthquake error terms in the attenuation model, which is reasonable when considering a single site.

Often PSHA expresses the hazard at a site as an annual rate at which the ground motion intensity level  $s_i$  is exceeded. Denoted  $v_i(s_i)$ , this rate is computed as

$$v_i(s_i) = \sum_{n=1}^{N_{sc}} v_n \iint_{M} \iint_{R} \iint_{\varepsilon} \Pr(S_i > s_i | m, r, \varepsilon) f_{M,R}(m, r) f_{\varepsilon}(\varepsilon) dm dr d\varepsilon$$
(8.2)

where  $v_n$  is the mean rate at which earthquakes occur on source *n*.

The BN in Figure 3.22 is used to model the seismic demand on a component in the example system. The modeling assumptions used to construct the CPTs of this BN, consistent with the methodology described in Chapter 3, are described below.

Figure 8.3a shows the assumed relative likelihoods of an earthquake occurring on each source. These probabilities are proportional to the mean occurrence rates  $v_n$  defined above. The distribution of the magnitude follows the truncated GR law (Gutenberg and Richter 1944) with source-specific parameters. The assumed prior discretized distribution of the magnitude for each source is shown in Figure 8.3b. The distribution of the location of the epicenter in each fault-specific coordinate system is assumed to be uniform along the fault length, as shown in a discretized form in Figure 8.3c. Note that, for source 2, the probability mass associated with the state 80-85 is slightly less that the mass associated with other states because source 2 is assumed to be 83.5 units in length, which falls near the middle of the state 80-85.

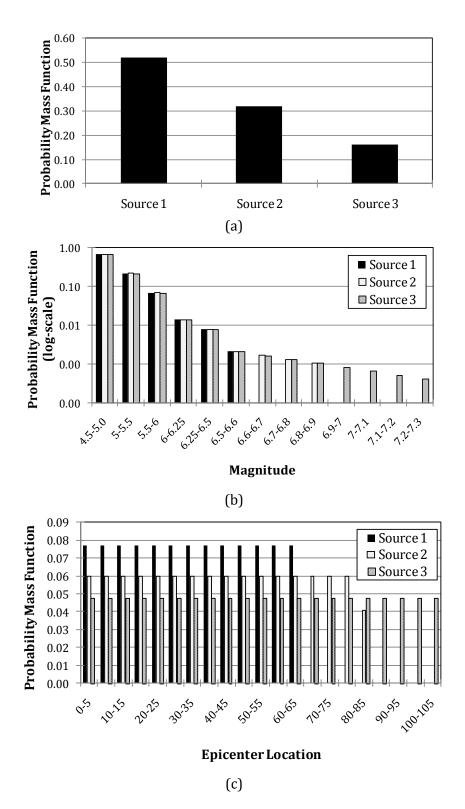


Figure 8.3: Prior probability mass functions of (a) source; (b) magnitude; and (c) epicenter location

The magnitude-rupture length relationship is given by (Wells and Coppersmith 1994)

$$\log R_L = a + b * M + \varepsilon_L \tag{8.3}$$

Statistical uncertainty on the values of the parameters *a* and *b* as well as the model error  $(\varepsilon_{I})$  are accounted for when defining the distribution of the rupture length for a given magnitude. For all faults, a and b are assumed to be normally distributed with means -3.55 and 0.74 and standard deviations 0.37 and 0.05, respectively.  $\varepsilon_L$  is normally distributed with zero mean and standard deviation of 0.23. The distributions of the rupture reference coordinate,  $X_r$ , and source-to-site distance, R, are defined using the expressions defined in Chapter 3. The Shahi & Baker (2010) directivity model is used to define the probability of experiencing a directivity pulse at a site as well as to determine the distribution of the variance reduction term,  $R_f$ . The ground motion intensity is measured in terms of spectral acceleration at 1 second. The Campbell and Bozorgnia (2006) NGA relationship is used to define the spectral acceleration at a site as a function of source and site characteristics. All faults are assumed to be strike-skip faults that crop-out at the ground surface and all sites have a soil shear wave velocity of 600m/s. For simplicity, all fault parameters required in the attenuation model (i.e.  $Z_{1.0} = 1$ ,  $F_{RV} = 0$ ,  $F_{NM} = 0$ ,  $Z_{tor} = 0$ , Dip = 90), other than fault length, are assumed to be identical for all sources. These simplifying assumptions are made to ensure that the illustrative results obtained from the BN are intuitive and predictable. This aids understanding of the BN model and verifiability of the obtained results.

The BN described in Chapter 3 yields the discretized distribution of ground motion intensity at a site (referred to as a ground motion prediction point, GMPP, in Chapter 3) as the output of node  $S_i$  in Figure 3.22. For the unconditional case, i.e. the case in which no evidence has been entered in the BN, summation of the probability masses above a given threshold yields the prior probability that the ground motion intensity at the site exceeds the given threshold in a future earthquake; this is the standard output of PHSA in (8.1). Conversely, to obtain this quantity, a binary node can be added into the BN that is in the true state if the ground motion intensity at the site exceeds the threshold and is the in the false state otherwise.

The remainder of this section describes site-specific hazard assessment for site 1. Unless noted otherwise, the results presented in this section are obtained using exact inference with the software application Hugin (Hugin Expert A/S 2008). The distribution of spectral acceleration for site 1, for a random earthquake occurring on one of the faults, is shown in Figure 8.4. This figure also demonstrates how this distribution changes when the earthquake is known to have occurred on each of the faults. The updated source-dependent distributions are obtained by entering evidence at node  $S_c$  in Figure 3.22 corresponding the occurrence of an earthquake on each fault. Because the faults are assumed to be identical with the exception of length, the longest fault (Source 3), which is capable of producing the largest earthquake magnitude, is associated with the distribution of ground motion intensity that is shifted farthest to the right.

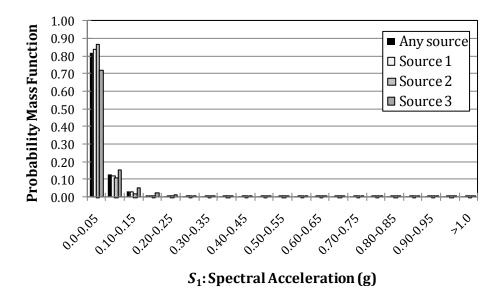


Figure 8.4: Distribution of ground motion intensity at site 1

The distribution in Figure 8.4 facilitates the construction of the site specific hazard curve. A hazard curve gives the annual frequency (*y*-axis) that the ground motion intensity at a site will exceed a given level (*x*-axis). Earthquakes occur on *each* fault with rate  $v_n$  and on *any* fault within the vicinity of the example system with a rate of  $v_1 + v_2 + v_3$ . The rate at which earthquakes occur on any fault with ground motion intensity greater than  $s_i$  is  $Pr(S_i > s_i) * (v_1 + v_2 + v_3)$ , where  $Pr(S_i > s_i)$  is obtained by summing the probability masses for all states greater than  $s_i$  in Figure 8.4, when considering all sources. Assuming earthquakes occur according to a homogenous Poisson process, the annual probability that the ground motion intensity at site *i* will exceed the level  $s_i$  due to earthquakes occurring on one of the  $N_{sc}$  faults is

$$Pr(at least one earthquake with S_i > s_i in one year) = 1 - exp (-Pr(S_i > s_i) * (v_1 + v_2 + v_3))$$
(8.4)

To produce a hazard curve for an individual fault, the total rate  $(v_1 + v_2 + v_3)$  in the above expression is replaced with the fault specific rate  $v_n$  and  $Pr(S_i > s_i)$  is computed using the corresponding distribution in Figure 8.4. The above description leads to the hazard curves for site 1 shown in Figure 8.5.

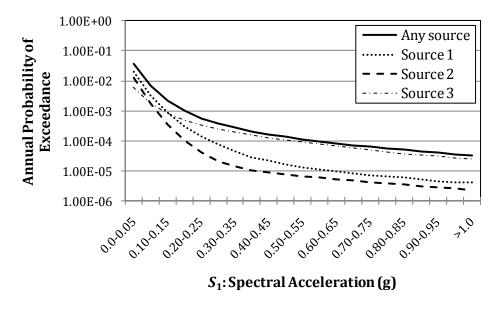


Figure 8.5: Example hazard curve for site 1

Source 3 produces the largest earthquakes but at the lowest rate. Conversely, source 1 produces the smallest earthquakes, but at the highest rate. Therefore, at low intensity thresholds (high levels of annual probability of exceedance; short return periods), source 1 provides the largest contribution to hazard. At high intensity thresholds (low annual probabilities of exceedance; long return periods), source 3 dominates the hazard.

The hazard curve gives the probability of exceeding a level of ground motion intensity when considering all possible combinations of earthquake magnitude and location (and thus source-to-site distance). When aggregated, it is difficult to gain insight into the combinations of magnitude and location (source-to-site distance) that govern the hazard. Often ground motion records required for structural analysis are selected based on the magnitude of the earthquake associated with the ground motion record and the distance between the earthquake source and the recording instrument. To gain insight into the combinations of magnitude and source-to-site distance that dominate the hazard, a process known as hazard *deaggregation* is used. Deaggregation breaks down the hazard curves into contributions from different earthquake scenarios. Most commonly, these scenarios are defined by combinations of magnitude and source-to-site distance. The hazard deaggregation gives the fractional contribution of a particular combination to the total hazard at a site and is computed according to the Bayes' rule:

$$P(M^{(1)} \le M \le M^{(2)}, R^{(1)} \le R \le R^{(2)} | S_i > S_i)$$

$$=\frac{P(S_i > s_i | M^{(1)} \le M \le M^{(2)}, R^{(1)} \le R \le R^{(2)}) * P(M^{(1)} \le M \le M^{(2)}, R^{(1)} \le R \le R^{(2)})}{P(S_i > s_i)}$$
(6.5)

(0 T)

where the superscripted values indicate lower and upper limits of intervals on the magnitude and source-to-site distance. It is seen that the conditional joint distribution of M and R given the hazard is proportional to the product of the conditional hazard and the unconditional joint distribution of M and R. Deaggregation is performed for a specific hazard level. In this example, the ground motion intensity in the hazard curve which corresponds to an annual probability of exceedance of  $10^{-4}$  is used. For the hazard curve in Figure 8.5, the discrete  $S_A$  state with ground motion in the range 0.55-0.6g at site 1 has a probability of exceedance of approximately  $10^{-4}$ .

The unconditional joint PMF of magnitude and source-to-site distances for site 1 is shown in Figure 8.6. In the three-dimensional space, the depth-wise axis has the lowest magnitudes located closest to the reader. As expected, low magnitude events at moderate source-to-site distances are most likely. The conditional probability that spectral acceleration at the site is in the state (0.55-0.6g) or higher given combinations of magnitude and source-to-site distances is shown in Figure 8.7. This figure shows that the exceedance event is most likely to result from a large magnitude event at a close distance. The distributions in Figure 8.6 and Figure 8.7 can be obtained directly from the BN. Finally, multiplying the conditional distribution times the unconditional joint distribution and dividing by the hazard level yields the deaggregation shown in Figure 8.8 according to (8.5). Conversely, the discretized joint distribution of *M* and *R* for a given evidence scenario (e.g.  $S_i > s_i$ ) can be obtained directly as output from the BN by entering the evidence as the appropriate node. Thus, the intermediate quantities in Figure 8.6 and Figure 8.7 need not be computed. Instead, Figure 8.8 can be obtained directly as natural output from the BN. As seen in Figure 8.8, the hazard at site 1 is dominated by low to moderate magnitude events located close to the site.

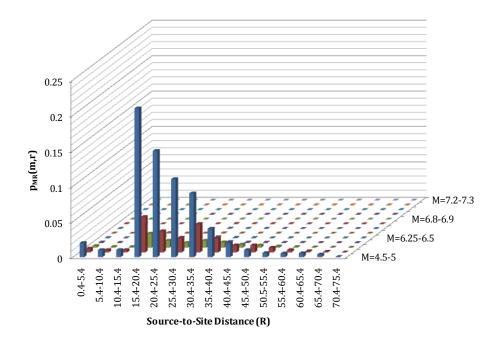


Figure 8.6: Joint distribution of magnitude and source-to-site distance

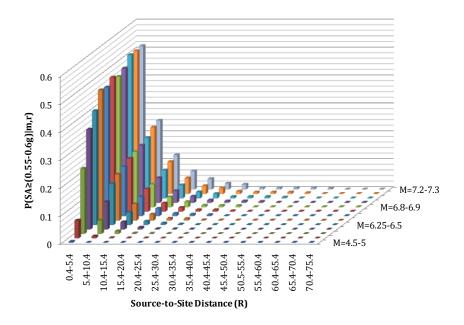


Figure 8.7: Conditional probability that spectral acceleration is equal to the state (0.55-0.6g) or greater at site 1 given magnitude and source-to-site distance

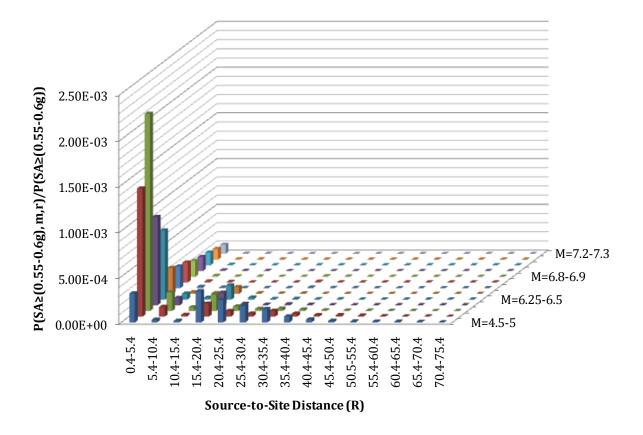


Figure 8.8: Hazard deaggregation for site 1

A variety of existing applications (e.g. see http://earthquake.usgs.gov/hazards/apps/) provide site-specific deaggregation. However, the BN framework is particularly useful for facilitating the understanding of hazard at a site *beyond* performing deaggregation. By entering evidence at any node in the BN, it is possible to identify the influnce on the hazard at the site. For example, Figure 8.9a shows the distribution of ground motion intensity at site 1 for several evidence cases (ECs): (1) it is known that an event has occurred on one of the faults with magnitude in the range 6.5-6.6; (2) it is known that the 6.5-6.6 magnitude event has occurred on source 2; and (3) it is known that the 6.5-6.6 magnitude event has occurred on source 2 with epicenter located at the center of the fault, but with unknown rupture length and direction. All evidence cases shift the distribution of ground motion intensity to the right relative to the unconditional case.

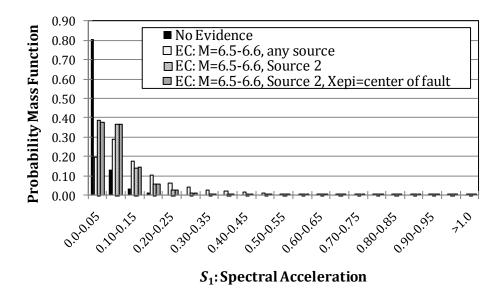


Figure 8.9: Distribution of spectral acceleration at site 1 given evidence cases

The ease with which BNs facilitate calculation of the updated distributions for all nodes in the BN is useful in understanding the hazard at a site. For example, the updated distributions of node  $S_c$  for a magnitude 6.5-6.6 event on any source as well as for events producing intensities of 0.3-0.35g and 0.55-0.60g at site 1 are shown in Figure 8.10a. The updated distributions of node *M* for the latter two evidence case are shown in Figure 8.10b.

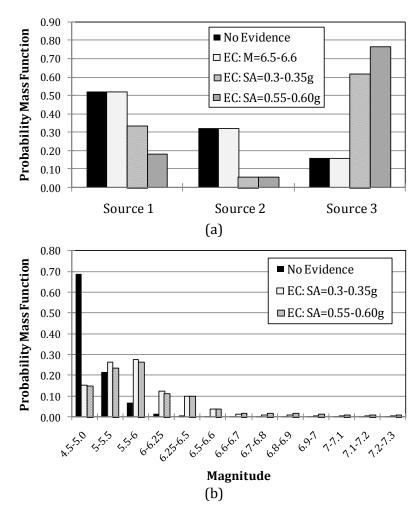


Figure 8.10: Updated distributions of (a) earthquake source and (b) magnitude, given evidence cases

Another useful application of BNs with regard to understanding the hazard at a site is the *max-propagation algorithm*, which provides the most-probable configuration of node states for a given evidence scenario. The algorithm can be applied to find the most likely configuration for any subset of nodes in the BN given evidence on a different subset. The exact version of the algorithm (in contrast to an approximate algorithm described later) is carried out using the inference procedures described in Chapter 2, but with summations over random variables replaced by maximization operators (Jensen and Nielson 2007). To differentiate the inference methods described in Chapter 2 from the max-propagation algorithm, the methods in Chapter 2 are generally referred to as *sum-propagation* algorithms.

To illustrate the value of using the max-propagation algorithm for understanding seismic hazard, consider that it is observed that ground motion intensity at site 1 is in the state 0.55-0.6g or greater, consistent with the hazard level considered previously. To facilitate

specification of this evidence case, a node is added to the BN as a child of node  $S_A$  corresponding to the binary event  $S_A \ge (0.55 \text{ to } 0.6\text{g})$ . This binary node is then set equal to the "true" state to reflect the evidence. The most probable configuration of source node states in the BN resulting from utilization of the exact max-propagation algorithm for this evidence scenario is shown in Table 8.1. The larger than expected ground motion intensity is most likely to result from a moderate magnitude event with epicenter located about one-third from the left end of source 3.

Node	State	
Source	Source 3	
Magnitude	5.0-5.5	
Rupture length	0-5 km	
Epicenter location	30-35 km	
Rupture reference coordinate	30-35 km	

Table 8.1: Most probable source node state configuration corresponding to evidence scenario in			
which spectral acceleration is in the state 0.55-0.6g or greater			

The configuration of the states of nodes of maximal probability corresponding to the evidence scenario described above is obtained using an approximate algorithm in GeNIe, which is developed by Yuan et al. (2004). It is important to note that the discretization intervals used in conjunction with continuous random variables can strongly influence results obtained via max propagation. Recall that the algorithm seeks to find the most probable joint combination of node states. When unequal bin sizes are used for the interval states of a continuous random variable, the algorithm may indentify certain states as being the most probable simply because they are associated with a large amount of probability mass, regardless of the physical meaning of the combination of states.

# 8.2.2 Using BN for system-level PSHA

The previous section focused on site-specific hazard. However, when working with infrastructure systems, it is necessary to consider the hazard at multiple sites. Analysis of seismic hazard at multiple sites requires consideration of ground motion intensities that vary in space and must be considered as a random field. For the subsequent analysis of this example, the size and complexity of the BN renders exact inference intractable on a workstation with 32 GB of RAM. Thus, an importance sampling-based algorithm known as EPIS-BN (Yuan and Druzdzel 2003) available in the GeNIe software application (DSL 2007) is used for sum propagation. The hazard at the site of each bridge in Figure 8.1 will differ depending on the distance to the earthquake source as well as due to site effects. For simplicity in the presentation, all sites are assumed to be identical; differences in hazard are thus due to source location and wave propagation effects only. Hazard curves for the different sites are shown in Figure 8.11. It can be seen that site 6 has the highest hazard level, while site 2 has the lowest hazard. These results can be explained by the relative

distances of these sites to the three faults and the mean rates of earthquake events on these faults.

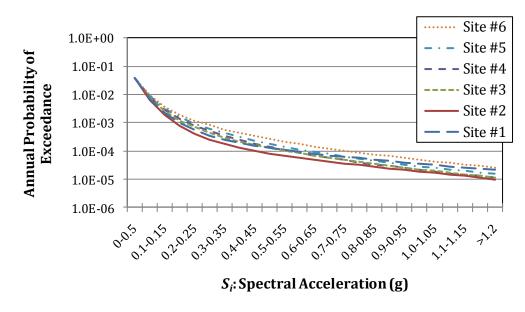


Figure 8.11: Hazard curves for all sites for unconditional case

Figure 8.12 shows how the hazard curves differ depending on the earthquake source. These curves are obtained by entering evidence at node  $S_c$  indicating the source for which the hazard curve is required. The resulting distribution of ground motion intensity at each site is then used to construct the hazard curve in the manner described in the previous section. The hazard curves associated with sources 1 and 3 are more dispersed from site to site. For source 1, site 6, which is located closest to the source, is associated with the highest hazard. However, this site is located farthest from source 3 and is, therefore, associated with the lowest hazard for the source 3 hazard curve. The hazards for source 2 are all of similar magnitude for all sites because all sites are located at similar distances from the fault. Unlike sources 1 and 3, no site is located sufficiently close to source 2 to result in a significant increase in the hazard over other sites. Figure 8.13 shows site-specific hazard curves given an earthquake on each source as well as for the unconditional evidence case, i.e. considering all earthquake sources. Consistent with previous observations, for all sites, hazard at low intensities is dominated by source 1. For sites that are close to source 3 (sites 1-3), the hazard at large intensity values is dominated by source 3. For sites 4-6, source 1 dominates the hazard at all intensities. For all sites, source 2 provides the smallest contribution to hazard.

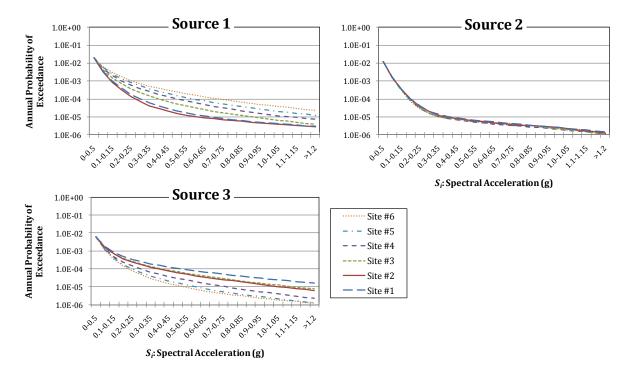


Figure 8.12: Hazard curves for all sites for an earthquake on each source

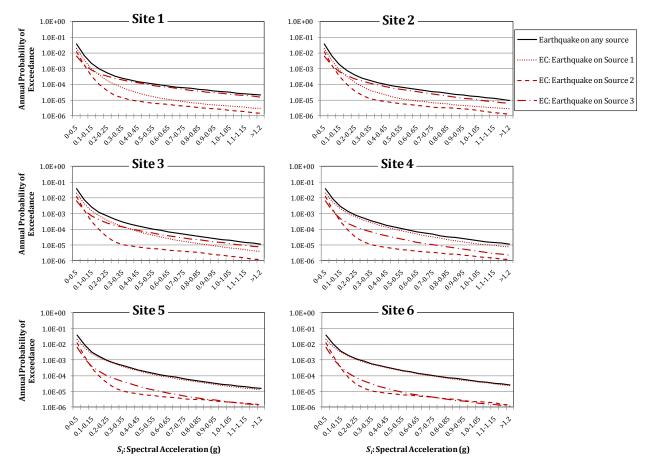


Figure 8.13: Site specific hazard curves for earthquake on each source

### 8.2.3 Incorporating component performance in assessing system seismic risk

The sites for which hazard curves were previously calculated correspond to the locations of bridges in the example transportation system. Consistent with Chapter 5, the performance of each bridge is defined as a function of the seismic demands placed upon it. For simplicity and predictability, it is assumed that all bridges are of the identical class. The fragility model used takes the form in (5.7), with  $\lambda = 0.6$  and  $\zeta = 0.4$ . Statistical uncertainty in the value of the parameters is neglected in this example. The component of model error that is common to all bridges is assumed to have a standard deviation of 0.1. It is assumed that the uncertainty associated with the site-specific component of model error is incorporated in  $\zeta$ . Figure 8.14 shows the fragility function for different realizations of the common component of model error.

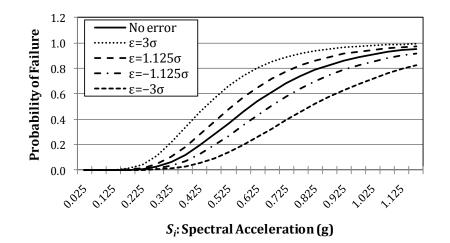


Figure 8.14: Fragility function for different realizations of model error

Component failure probabilities associated with the unconditional case in which an earthquake occurs on an unknown source, as well as the conditional cases in which the source is known, are shown in Figure 8.15. For the unconditional case, the failure probabilities are more homogenous than when considering earthquakes on sources 1 and 3. For the unconditional case, the highest failure probabilities are associated with components 5 and 6, which are located closest to fault 1 (the most active fault). Component 1, which is located extremely close (less than 1 km) to fault 3, which is capable of producing the largest earthquake, and component 4, which is located in between faults 1 and 3, have failure probabilities of similar magnitude. For the source-specific cases, the failure probabilities increase with proximity to the fault. Because source 3 is associated with the largest magnitude events, the associated failure probabilities are larger. The component failure probabilities associated with source 2 are the lowest and relatively homogenous because all sites are located a longer distance (at least 34 km) from the fault.

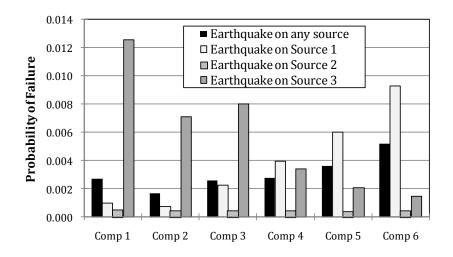


Figure 8.15: Probability of component failure given an earthquake

### 8.2.4 Incorporating system performance in assessing system seismic risk

The next step in assessing the seismic risk to the example system is the incorporation of a measure of system performance. The transportation system is classified in the failure state if any city is unable to reach the hospital following an earthquake as a result of bridge failures. A city will be disconnected from the hospital if a connectivity path does not exist between the city and the hospital. Treating a city as a source and the hospital as a sink, MLSs can be defined representing these connectivity paths. The MLSs corresponding to the source at City A and the sink at the hospital are: {1}, {2,3,4}, and {2,3,6,5}. For the source at City B the MLSs are: {2,1}, {3,4}, and {3,5,6}; for City C: {4}, {3,2,1}, {5,6}; and for City D: {5}, {4,6}, and {1,2,3,6}.

The portion of the BN utilized for modeling system performance is shown in Figure 8.16 using a MLS formulation. Node  $S_{sys}$  is defined as a child of nodes "*city*"  $\rightarrow$  *H*, which represent the event that a city is able to reach the hospital.  $S_{sys}$  will be in the survival state only if all nodes "*city*"  $\rightarrow$  *H* are in the survival state. Nodes "*city*"  $\rightarrow$  *H* have, as parents, nodes that represent the MLSs between each source corresponding to a city and the sink corresponding to the hospital. The MLS nodes are defined as children of their respective constitutive components. The BN in Figure 8.16 is not optimized for computational efficiency as converging connections have been used instead of a chain structure. Here, for illustrative proposes, the converging structure is shown because it is easier to interpret than the chain structure characterizing the more efficient formulation. For comparison, the efficient formulation is shown in Figure 8.17. In this BN, the formulations corresponding to each source-sink pair are hidden behind objects. The BNs behind the objects are included below the system BN. The total clique table sizes associated with the converging and efficient MLS formulations, neglecting correlation in demands, are 2624 and 1440 respectively.

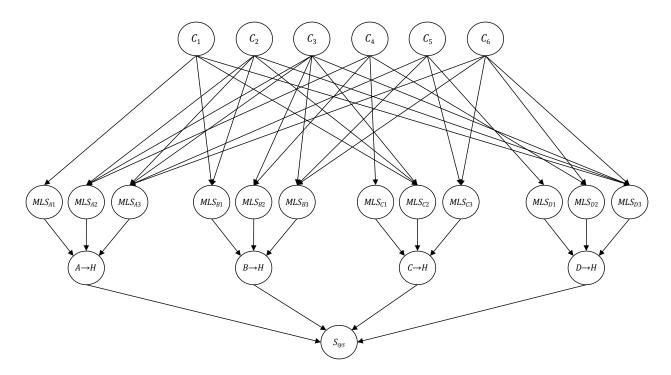


Figure 8.16: BN for modeling system performance in Example #1 using MLS formulation

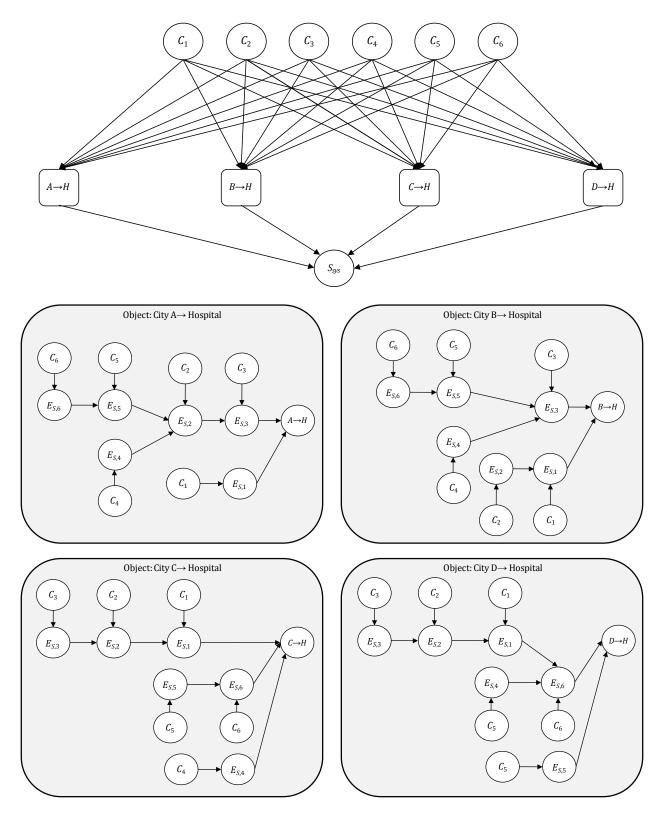


Figure 8.17: BN for modeling system performance in Example #1 using efficient MLS formulation

Figure 8.18 shows the system failure probability as well as the probability that each city is disconnected from the hospital in the event of an earthquake on one or any of the faults. For an earthquake on source 1, city D, which is located closest to the source and which will be disconnected if bridges 5 and 6 fail, is associated with the highest probability of disconnect. Conversely, for an earthquake on source 3, cities A and B, which are located closest to the source, are associated with the highest probabilities of disconnect. Source 3 is likely to produce the largest magnitude earthquakes and therefore the highest levels of ground motion intensity. As shown in Figure 8.15, the failure probabilities of bridges 1-3 are significantly higher than for earthquakes on other sources. These bridges are important for maintaining connectivity between cities A and B and the hospital.

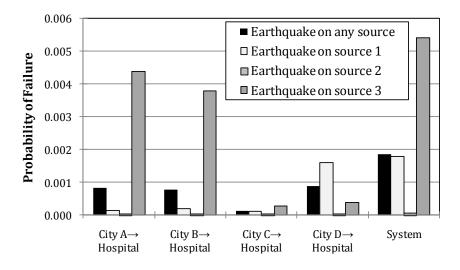


Figure 8.18: System failure probabilities for Example #1

Max-propagation is also useful for system identification. Table 8.2 shows the most probable configuration of component states corresponding to the case where node  $A \rightarrow H$  is observed to be in the failure state, i.e. city A is disconnected from the hospital. The results in the table imply that disconnect of city A from the hospital is mostly likely to result from the failure of bridges 1 and 3. This configuration is obtained using the approximate max propagation algorithm in GeNIe (Yuan et al. 2004). The most probable configuration is useful in indentifying the most vulnerable components or subsystems to cause failure of the infrastructure system in a future earthquake. Furthermore, Table 8.2b shows the most probable joint configuration of the nodes that describe the source. Based on Table 8.2b, the most probable event to cause disconnect of city A from the hospital is a 5.5-6.0 magnitude event on source 3 with epicenter located 30-35 km from the end of the fault and rupturing 0-5km to the right. Thus, the earthquake rupture is located in close proximity to the bridges that are the most likely to result in disconnect of city A from the hospital. Such information is useful in identifying system vulnerabilities and dangerous earthquake scenarios.

Table 8.2: Most probable node state configuration of (a) component states and (b) source characteristics, for the evidence scenario in which the city A is disconnected from the hospital

(b)

(°)			
Component	State	Node	State
1	Failure	Source	Source 3
2	Survival	Magnitude	5.5-6.0
3	Failure	Rupture length	0-5 km
4	Survival	Epicenter location	30-35 km
5	Survival	Rupture reference coordinate	30-35 km
6	Survival		

#### 8.2.5 Using BNs in post-event applications

(a)

The value of using the information updating capabilities of BNs for understanding the hazard at an individual site as well as at the system level was demonstrated above. Next, we briefly demonstrate how the information updating capabilities are useful in near-real time applications. More extensive consideration of post-earthquake decision-making is contained in the description of Example #2. To illustrate the information updating capabilities of BNs, seven post-earthquake evidence cases (ECs) are considered:

- EC #1: No evidence
- EC #2:  $S_{A,1} = 0.3 0.35g$
- EC #3:  $S_{A,1} = 0.55 \cdot 0.60$ g
- EC #4: Earthquake on Source 2, M = 6.5-6.6,  $X_{epi} =$  center of fault
- EC #5: Earthquake on Source 2, M = 6.5-6.6,  $X_{epi} = \text{center of fault}$ ,  $S_{A,1} = 0.3-0.35\text{g}$
- EC #6: Earthquake on Source 2, M = 6.5-6.6,  $X_{epi} = \text{center of fault}$ ,  $S_{A,1} = 0.55-0.60$ g

• EC #7: Earthquake on Source 2, M = 6.5-6.6,  $X_{epi} =$  center of fault,  $C_1$  in failure state

where  $S_{A,1}$  is the spectral acceleration at site 1. The first evidence case is the baseline unconditional scenario in which no information is available other than the occurrence of the earthquake. Evidence cases 2 and 3 correspond to the situation in which an observation has been made regarding the ground motion intensity as site 1 (e.g., from a recording instrument), but no source information is available about the earthquake. Evidence cases 4-6 include information about the earthquake source and, in cases 5 and 6, additional information about the ground motion intensity at site 1. Evidence case 7 replaces the information relating to a measurement of ground motion intensity with an observation of structural performance. Figure 8.19 shows the distributions of spectral accelerations at sites 2, 4, and 6 under evidence cases 1-3.

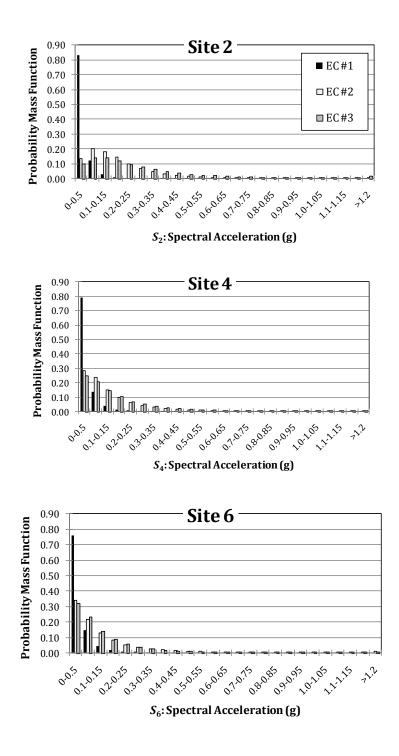


Figure 8.19: Distribution of spectral acceleration at sites 2, 4, and 6 for evidence cases 1-3

In Figure 8.19, the distributions of ground motion intensity for all three sites are shifted toward higher values when the intensity is observed at site 1. However, the observation

has the strongest influence on site 2, which is located in close proximity to site 1. The pronounced effect of this observation on the sites in close proximity, results from the way in which the observation updates the distributions of earthquake location parameters as well as the random field model. The observation of ground motion intensity at site 1 updates the BN in the backwards direction. The observed intensity at node  $S_1$  updates the probability distributions of the unobserved source terms ( $X_e$ ,  $X_r$ , and  $R_L$ ) as well as the random field model. Figure 8.20 shows the updated distribution of the location parameters given information about the ground motion intensity at site 1. Evidence case 2 shifts the distribution of node  $S_c$  to reflect the increased likelihood that the earthquake occurs on source 3 (the source in closest proximity to site 1). The shift is more pronounced for the third evidence case than the second evidence case. Furthermore, for evidence cases 2 and 3, the distributions of the epicenter location and the rupture reference coordinate reduce in entropy and become more peaked at values near the center of the fault. The distribution of the rupture length also shifts toward higher values.

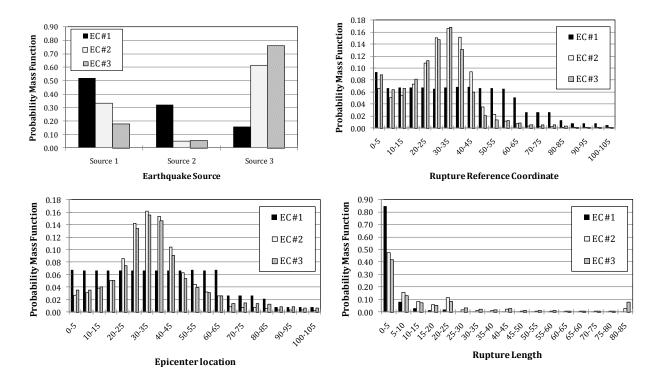


Figure 8.20: Distributions of earthquake location parameters for evidences cases 1-3

The effect of the observation of ground motion intensity at site 1 on the intra-event error terms of the random field for sites 1, 2, 4, and 6 is shown in Figure 8.21. The distribution of  $\varepsilon_{r,1}$  shifts strongly to the right as a result of the observed high level of ground motion intensity at site 1. A similar though less pronounced trend is observed for  $\varepsilon_{r,2}$ . The effect of the observation on the distribution of  $\varepsilon_{r,4}$  and  $\varepsilon_{r,6}$  is negligable. The approximate random field model is shown in Figure 8.22. Note that sites 1-3 are linked by a common *U*-node, as

are sites 5 and 6. All sites are linked by node  $U_2$ . Obviously, there is a stronger effect of an observation at site 1 on the intensity at sites with which it shares multiple *U*-nodes than on sites that are linked by only one information path.

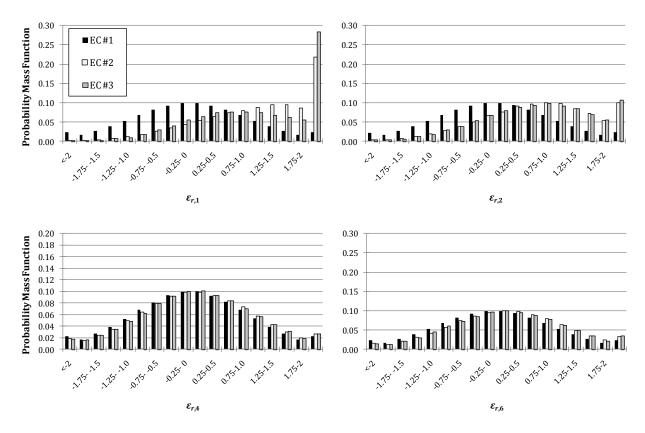


Figure 8.21: Distributions of random field intra-event error terms for evidence cases 1-3

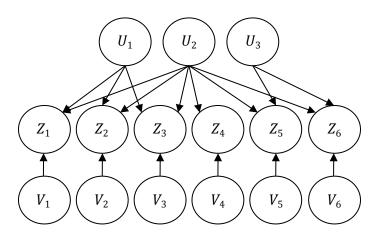


Figure 8.22: Approximate random field model

Figure 8.23 shows the updated distributions of ground motion intensities at sites 2, 4, and 6 for evidence cases 4-7. The distribution of ground motion intensity at all sites shifts to the right based on the observation of only source characteristics (EC 4). The distribution shifts further to the right when ground motion intensity is observed to be larger than expected at site 1 (ECs 5 and 6). The observation of component failure (EC 7) shifts the distribution to the right relative to evidence case 4; however it is not shifted as far as for evidence case 6. The updated distribution of ground motion intensity at site 1, given component 1 fails is shown in Figure 8.24. The peak of this distribution is at a value lower than the state 0.55-0.6g. Recall that the mapping between component performance and ground motion intensity is uncertain. Therefore, the observation of component performance is less informative to the seismic demand model than a direct observation of ground motion intensity.

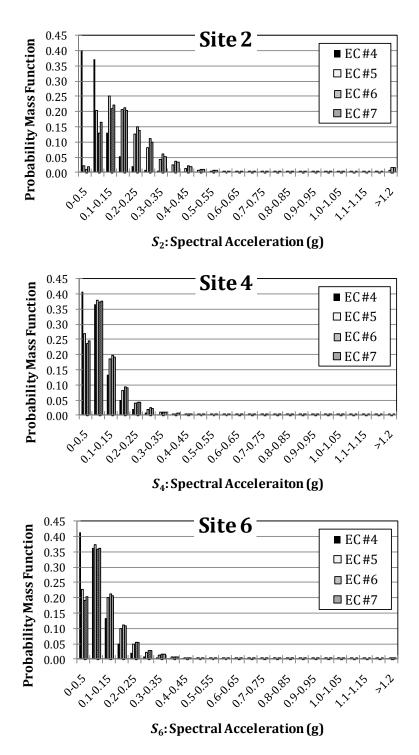


Figure 8.23: Distributions of spectral accelerations at sites 2, 4, and 6 for evidence cases 4-7

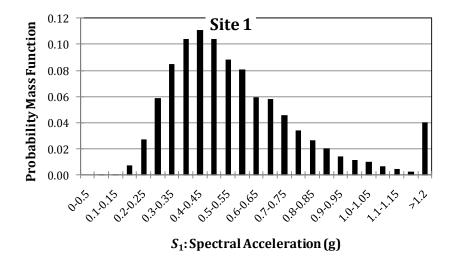


Figure 8.24: Distribution of spectral acceleration at site 1 for evidence case 7

# 8.3 Example #2 – Proposed California high speed rail system

The proposed California high speed rail (HSR) system is currently in the early stages of development under the guidance of the California High-Speed Rail Authority (CA HSR 2010). The proposed system will initially link San Francisco and Los Angeles in approximately 2.5 hours in addition to connecting intermediate cities along the route. It will later be expanded to link other major cities, such as San Diego and Sacramento. The system is situated in a highly seismic region and thus earthquakes pose a significant risk to the system. Figure 8.25 shows the proposed alignment of the northern portion of the HSR system (in blue) as well as known faults in the region (in red) (Bryant 2005). In this example, we focus on the portion of the system between San Francisco and Gilroy, California.

Through sound engineering and decision-making, it is possible to mitigate risks posed to the HSR system by seismic hazard. In this example we focus on hazard assessment and decision-making in the immediate aftermath of an earthquake. Often, at this stage, emphasis is placed on life safety (e.g. stopping the train to avoid derailment) and restoration of critical services (e.g. providing electricity to ensure trains can reach stations to aid evacuees). Post-event decisions include the dispatching of rescue, inspection, and repair crews as well as decisions to shut-down or continue component and system operations. These decisions are made under the competing demands to maintain operability (to prevent revenue loss), while not sacrificing safety (to avoid deaths and injury and consequent liability). The post-event decision-phase is characterized by an environment in which information is both uncertain and quickly evolving in time. Sources of information at this phase may include ground motion intensity measurements, structural health monitoring sensors, inspection results, and even observations reported by train operators, maintenance/repair crews, and passengers. This section provides an overview of the BN models developed for this purpose and their preliminary application to the proposed California HSR project.

Because the HSR system is in an early stage of design, data about the system (including specific component types, ground conditions, and finalized alignment) are not available. The analysis contained in this study is largely based on assumptions about the system and all results must be treated as hypothetical. Nonetheless, the study illustrates the power of the BN approach for infrastructure seismic risk assessment and provides a general framework that can be adapted when more specific information becomes available. Furthermore, this example draws attention to areas where further research and development are needed, if accurate seismic hazard assessment and risk-based decision support are to be provided to the management of the HSR in the design, operation, and post-event emergency response and recovery phases.



Figure 8.25: Google Earth Map of the Northern California portion of the proposed California High Speed Rail system route (blue) (CA HSR 2010) and nearby faults (red and pink) (Bryant 2005)

# 8.3.1 Model assumptions

There are three major faults in the vicinity of the portion of the HSR system under consideration in this study: the Hayward fault, the Calaveras fault, and a portion of the San Andreas fault. The faults are idealized as straight or broken line segments, as shown in Figure 8.26. If more detailed fault geometry is available and required, the framework can be extended to include it. In that case, modifications need only be made to the evaluation of the CPTs, which are generated external to the BN using Monte Carlo simulation, as described in Chapter 2. All calculations performed in this section are based on exact inference using the software application Hugin (Hugin Expert A/S 2008).

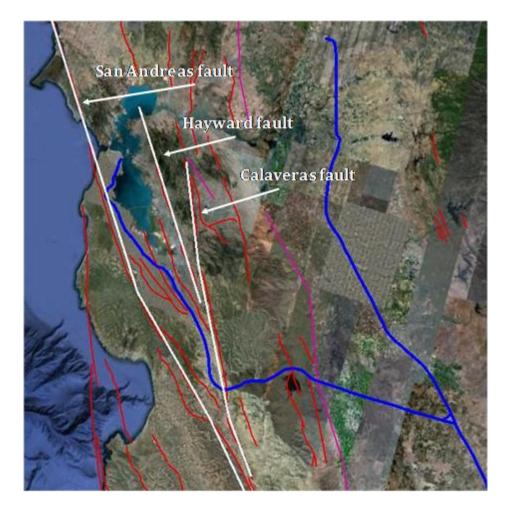


Figure 8.26: Google map overlaid with idealizations of Calaveras, Hayward and San Andreas (Peninsula, Santa Cruz Mountains, Creeping sections and a portion of North Coast section) faults

For each earthquake fault, source-dependent distributions of magnitude and location of earthquake epicenter are assumed. A characteristic magnitude distribution model (Youngs and Coppersmith 1985) is used with the fault parameters taken from the Caltrans fault parameter database (Caltrans 2007). The Wells & Coppersmith (1994) model is used to define the rupture length on the fault as a function of the earthquake magnitude and source characteristics. The Campbell & Bozorgnia (2006) NGA attenuation relationship is used to describe the ground motion intensity at each site as a function of source and site characteristics. Various components within the system are sensitive to different measures of ground motion intensity. The seismic demand BN model thus has nodes representing spectral acceleration, peak ground acceleration, and peak ground velocity. We use the model by Shahi and Baker (2010) to describe the directivity effects on spectral acceleration. For peak ground velocity, we use a prediction equation developed by Bray & Rodriguez-Marek (2004), which incorporates the effect of directivity. The site shear velocity for all sites is in the range 250-325 m/s.

For each fault, the geometry of the geographic region and of the infrastructure system is defined within a local coordinate system with the origin at one end of the fault and the *x*-axis oriented along its first linear segment. In this coordinate system, the epicenter is defined as a point along the fault. The rupture is assumed to occur anywhere along the fault with uniform likelihood, while containing the epicenter and not extending beyond the known ends of the fault. For a given rupture on a fault, the source-to-site distance  $R_i$  (the shortest distance between site *i* and the rupture) is obtained as described in Chapter 3. Site and fault coordinates are considered deterministic.

The approximated random field BN model used for the 18 GMPPs (ground motion prediction points, see Figure 8.27) in the HSR system is shown in Figure 8.28. Because of the computational demands of exact inference, an approximated model with few links is necessary to ensure computational tractability. Because the HSR is a series system, its reliability is relatively insensitive to approximations in the correlation matrix, so that we are able to eliminate many nodes and links (see Chapter 4). Thus, in the resulting BN, only relatively closely spaced GMPPs are linked thorough common *U*-nodes.



Figure 8.27: Ground motion prediction points along alignment

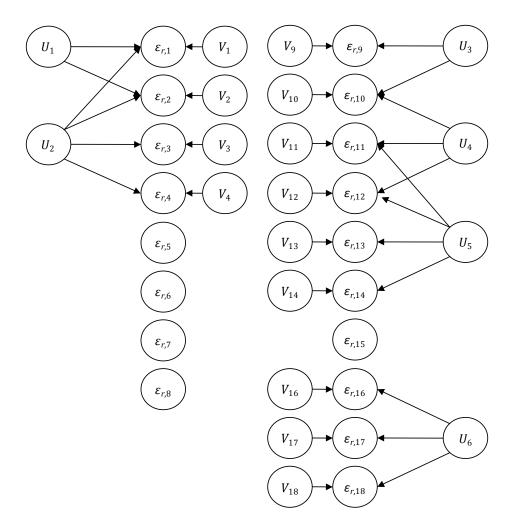


Figure 8.28: Approximated random field model for HSR

Aside from ground shaking, other earthquake-related hazards potentially affecting the HSR system include fault rupture for segments of the system that cross active faults, liquefaction and lateral spreading of saturated sandy soils, stability of slopes, embankments, retaining structures, tunnels that are affected by various aspects of the ground motion, and the spatial variability of the ground motion on extended components, such as the rail line or long-span bridges. BN models in this example focus on ground-shaking; models for some of the other hazards listed above are briefly discussed in Chapter 3, but their full development and application are left as areas warranting future study.

The HSR system consists of two general classes of components: point-site components, e.g. short span bridges over small waterways or roads, culverts, and station buildings; and distributed-site components, e.g. rail tracks, long embankments, aerial structures, and tunnels. At this stage, little is known about the locations of the point-site components in the system; the alignment only provides crude estimates of the locations of the distributed

components, as shown in Figure 8.29. The assumed discretized component types are listed in Table 8.3 along with their estimated length and associated GMPPs. The idealizations contained in this table are extremely crude. Significant modification and expansion is necessary when more detailed information becomes available.

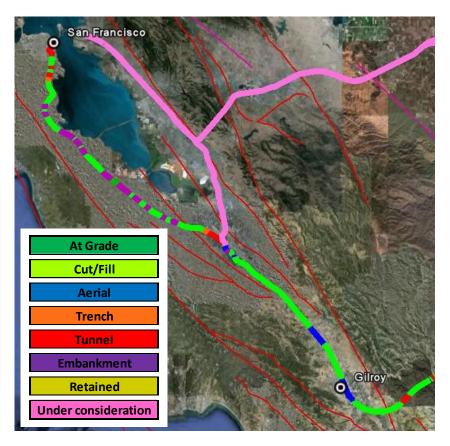


Figure 8.29: Google Map of HSR alignment showing locations of major distributed components (CA HSR 2010)

Componment	Туре	Length (km)	GMPP(i)	GMPP( <i>i</i> +1)
1	Tunnel	4.0	1	2
2	Embankment	2.7	2	3
3	Tunnel	3.5	3	4
4	Embankment	8.8	4	5
5	Embankment	8.7	5	6
6	Embankment	8.4	6	7
7	Embankment	8.8	7	8
8	Embankment	8.8	8	9
9	Embankment	7.7	9	10
10	Embankment	6.9	10	11
11	Tunnel	6.0	11	12
12	Aerial	point	12	-
13	Embankment	2.7	12	13
14	Aerial	point	13	-
15	Embankment	8.1	13	14
16	Embankment	9.0	14	15
17	Embankment	8.9	15	16
18	Aerial	5.5	16	17
19	Embankment	9.5	17	18

Table 8.3: Assumed component types, length, and associated GMPPs

As little is known about the final structural or geotechnical systems to be employed for the HSR system, broad assumptions have been made about the fragility/performance functions for the components listed in Table 8.3. Thus, all such functions used in this example should be regarded as "placeholders" that must be replaced when more specific component descriptions and corresponding fragility/performance models become available. Simple fragility models for the structural systems (tunnels and aerial structures) are adapted from HAZUS<sup>MH</sup> (DHS 2003). The HAZUS<sup>MH</sup> fragility functions are developed for standard highway systems; however the HSR system will be designed with more rigorous performance criteria. Furthermore, the HAZUS<sup>MH</sup> fragility functions include a large amount of uncertainty to account for the broad class of structural systems to which they are applied. As such, the associated fragility functions are not directly applicable to the HSR. To account for the more rigorous design standards and quality control of the HSR, HAZUS<sup>MH</sup> fragility functions are modified by increasing the median resistance and reducing the standard deviation.

The majority of fragility functions developed for bridges and tunnels are applicable only to point-site structures and do not account for spatial variability of the ground motion intensity along the structure. Analysis of structures subjected to differential ground motions is an area of developing research (Kim and Feng 2003; Konakli and Der Kiureghian 2010). Until fragility curves are available specifically for the structures employed by the

HSR system, we must make simplifying assumptions in order to proceed with development of the overall framework and produce illustrative (though hypothetical) results. For the present analysis, we estimate the performance of tunnels and aerial structures distributed between GMPPs i and i + 1 as

$$p_{i,i+1}^{k}(s_{i},s_{i+1}) = F_{k}^{i}(\max(s_{i},s_{i+1})) - F_{k+1}^{i}(\max(s_{i},s_{i+1}))$$
(8.6)

where  $p_{i,i+1}^k(s_i, s_{i+1})$  denotes the probability that the component between GMPPs *i* and i + 1 will experience damage of state *k* and  $F_k^i(.)$  is the fragility model associated with damage state *k* (see chapter 5). The embankment segments are PGV sensitive and are modeled using fragility functions adapted from the available literature (Mizuno et al. 2008). These models provide conditional rates of damage per unit length, from which fragility functions are developed by use of the non-homogeneous Poisson model as described in Chapter 5. In the absence of fragility functions are preliminarily applied to these components as well.

Neglecting infrastructure interdependencies (e.g. the dependence of the HSR system on the electrical grid) and assuming tracks going in either direction are associated with the same supporting structures, the HSR can be considered as a series system. The components of the system are provisionally assumed to have three states: not damaged, slightly damaged, and moderately to severely damaged. If the component is not damaged, it is assumed to be fully operational. Slight damage is associated with a reduction in performance (e.g. the trains must slow-down when traversing the component), and the moderate to severe damage state is associated with complete component closure.

In the present analysis, the system failure event is defined as the inability to reach Gilroy from San Francisco (either at full or reduced speed). However, we assume that, in the event a portion of the alignment is not passable, passengers can take an alternative mode of transportation to bypass the damaged segment and board the train further down the line to complete the journey to Southern California. With this in mind, four intermediate failure events are considered: Inability to travel from (1) San Jose to Gilroy; (2) Palo Alto to Gilroy; (3) Millbrae to Gilroy; and (4) San Francisco to Gilroy. The train is unable to travel from San Jose to Gilroy at normal speed if any component along this segment is damaged, and the connectivity is completely lost if any component along the segment experiences moderate or severe damage. Connectivity between Palo Alto and Gilroy is lost if any component between Palo Alto and San Jose is moderately or severely damaged or connectivity is lost between San Jose and Gilroy. The above pattern continues such that connectivity between San Francisco and Gilroy is possible if no component between San Francisco and Millbrae is moderately or severely damaged and connectivity is not lost somewhere farther down the line. The above description is used to create the BN for modeling the HSR system performance shown in Figure 8.30. Because this is a series system, the efficient formulation is constructed directly as described in section 6.5.1 without requiring use of the optimization algorithm.

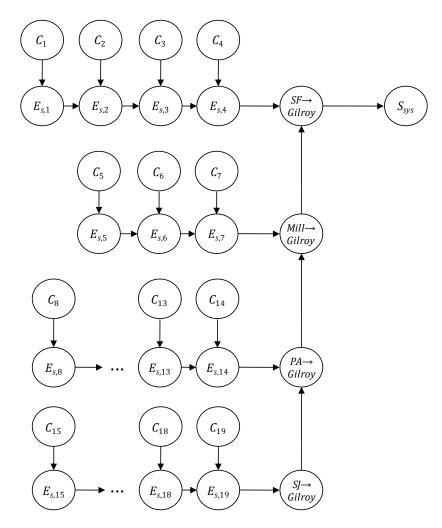


Figure 8.30: BN modeling system performance of HSR

In Figure 8.30, nodes labeled "*city-name* $\rightarrow$ *Gilroy*" indicate whether connectivity exists between the indicated city and Gilroy. Node *SJ* (*San Jose*)  $\rightarrow$  *Gilroy* is in an open (reduced/closed) state if node  $E_{s,19}$  is in open (reduced/closed) state. Node *PA* (*Palo Alto*)  $\rightarrow$  *Gilroy* indicates full capacity connectivity between Palo Alto and Gilroy only if node *SJ* (*San Jose*)  $\rightarrow$  *Gilroy* and  $E_{s,14}$  both are in the open (full capacity) state. Similarly node *PA* (*Palo Alto*)  $\rightarrow$  *Gilroy* indicates loss of connectivity if node *SJ* (*San Jose*)  $\rightarrow$  *Gilroy* or  $E_{s,14}$  are in the closed state. Otherwise, node *PA* (*Palo Alto*)  $\rightarrow$  *Gilroy* indicates the reduced capacity state. Similar patterns are used for constructing the CPTs of nodes *Mill*(*brae*)  $\rightarrow$  *Gilroy* and *SF*  $\rightarrow$  *Gilroy*.

Next, we extend the BN by decision and utility nodes to begin to address the postearthquake inspection-shutdown decision problem, as described in Chapter 7. When more information becomes available about the HSR system and the preferences of the system decision-makers, the decision framework described here must be modified and expanded. For now, we present a first-attempt at developing a post-earthquake decision-framework for the HSR system. Preliminarily, it is assumed that the shutdown decision will happen at the "link-level," where a link is the set of components located between stations. Link 1 connects San Francisco to Millbrae and contains components 1-4, link 2 goes from Millbrae to Palo Alto and includes components 5-7, link 3 contains components 8-14 and runs between Palo Alto and San Jose, and finally, link 4 is between San Jose and Gilroy and contains components 15-19. For reference in the following description, the LIMID corresponding to the decision model for link 1 is shown in Figure 8.31. LIMIDs for the other links are similar and are thus not shown here. Note that the decision problem described here considers each link as a separate system rather than considering the entire system simultaneously. Inspection decisions are made at the link as well as component levels. Decision making at the entire system level proved to be computationally impossible with the Hugin software application and our computing resources. In any case, decision making at the level of a link subsystem is reasonable because (a) a given earthquake is likely to affect only one link, (b) it is likely that separate inspection/repair crews will be available for the separate links and, hence, decisions can be made locally.

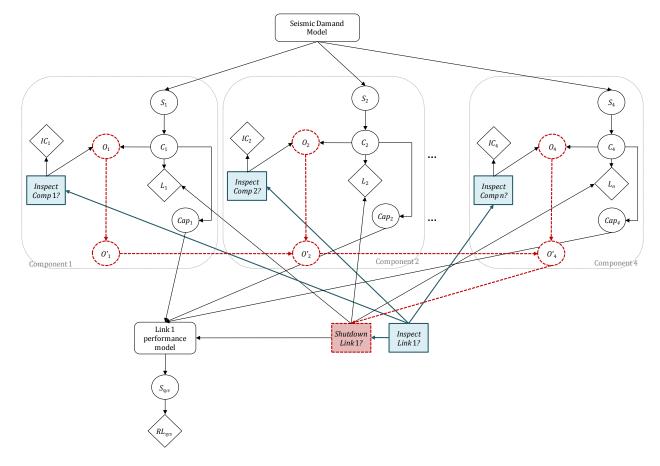


Figure 8.31:LIMID for solving inspection-shutdown decision for link #1 of HSR system

The decision framework described in Chapter 7 is applied to the LIMID for link #1 with a few modifications. In the modified decision framework, the component-level inspection-shutdown decision is replaced by a link-level decision, as shown by nodes *Shutdown Link* 1? and *Inspect Link* 1?. Node *Shutdown Link* 1? is a parent of the link performance object with conditional relationships defined such that link 1 is closed (resp. has reduced operating level) if any component is moderately or severely damaged (slightly damaged), or a decision is made to shutdown (reduce the operating level of) the link. The decision node is also a parent of nodes  $L_i$ , which model component liabilities. Node *Inspect Link* 1? is a parent of the node that represents the link-level shutdown decision to indicate decision precedence. It is also a parent of nodes *Inspect Comp i*?, i = 1, ..., 4, which represent component-level inspection decisions. Here, we assume that component-level inspections are prioritized *after* a link has been chosen for inspection. (Alternative heuristics for prioritizing component inspections can be easily considered.) Nodes  $O_i$  and  $IC_i$  are children of the component-level inspection decisions, consistent with the descriptions in Chapter 7.

To indicate information precedence, it is necessary to include links that signify that, if inspections are made, the inspection results (i.e. the outcomes of nodes  $O_i$ ) are known before the shutdown decision is made. Including all nodes  $O_i$  as parents of node *Shutdown Link* 1? results in an inefficient converging structure. Instead, the converging structure is replaced by a chain structure of nodes  $O'_i$  as shown (in thick dashed lines) in Figure 8.31. Each node of the chain indicates the worst damage state that is observed along the link up to that node, analogous to the logic used to construct the efficient MLS formulation in Chapter 6.

For this example, inspection costs and liabilities (in an arbitrary unit of utility) are defined as a function of component type and length and are summarized in Table 8.4. The penalties associated with closure of a link are defined as a function of the number of components along the link, as a proxy for revenue lost as a function of shutting down the component, e.g. due to lost passengers and the cost to run a bus bridge around the shutdown link. The link closure costs are defined as 10 times the number of components located along the link. The cost associated with reducing the capacity of the link is assumed to be 10 units, regardless of the number of components comprising the link.

Componment	Туре	Length (km)	Inspection Cost (imperfect)	Inspection Cost (perfect)	Liability (open when damage is slight)	Liability (open when damage is moderate or worse )	Liability (reduced when damage is moderate or worse )
1	Tunnel	4.0	0.080	0.160	6.00	600	60.0
2	Embankment	2.7	0.027	0.054	1.35	135	13.5
3	Tunnel	3.5	0.070	0.140	5.25	525	52.5
4	Embankment	8.8	0.088	0.176	4.40	440	44.0
5	Embankment	8.7	0.087	0.174	4.35	435	43.5
6	Embankment	8.4	0.084	0.168	4.20	420	42.0
7	Embankment	8.8	0.088	0.176	4.40	440	44.0
8	Embankment	8.8	0.084	0.168	4.40	440	44.0
9	Embankment	7.7	0.077	0.154	3.85	385	38.5
10	Embankment	6.9	0.069	0.138	3.45	345	34.5
11	Tunnel	6.0	0.120	0.240	9.00	900	90.0
12	Aerial	point	0.0005	0.001	2.50	250	25.0
13	Embankment	2.7	0.027	0.054	1.35	135	13.5
14	Aerial	point	0.0005	0.001	2.50	250	25.0
15	Embankment	8.1	0.081	0.162	4.05	405	40.5
16	Embankment	9.0	0.090	0.180	4.50	450	45.0
17	Embankment	8.9	0.089	0.178	4.45	445	44.5
18	Aerial	5.5	0.905	1.810	13.75	1375	137.5
19	Embankment	9.5	0.095	0.190	4.75	475	47.5

#### Table 8.4: Assumed component inspection costs and liabilities

# 8.3.2 Illustrative results

To illustrate the value of using BNs for immediate post-earthquake decision support, we consider the evolution of information following the occurrence of an earthquake. Suppose a 6.8 magnitude earthquake has occurred with the epicenter located 30 km from the northern end of the Hayward fault.

Following the earthquake, information becomes incrementally available to the decision maker. For illustration, we show how predictions by the BN model evolve as three pieces of evidence become available. For reference, we refer to each stage of information acquisition as an evidence case (EC):

- EC #1: M=6.8,  $X_e$ =30km (the base case)
- EC #2: EC #1 + spectral acceleration at GMPP 2 = 0.45-0.50g
- EC #3: EC #2 + pulse-like ground motion at GMPP 9
- EC #4: EC #3 + extensive damage of  $C_{17}$

Below, we consider each of these ECs individually.

Evidence case 1 corresponds to the information that will (likely) be available to the decision-maker from a source such as the United States Geological Survey almost immediately following the occurrence of an earthquake. The values reported initially are not certain, however in this example we treat the immediate post-earthquake estimate of the magnitude and earthquake epicenter as deterministic quantities. (Uncertainty in these estimates can be easily incorporated in the BN model by adding two additional nodes in Figure 3.22.) Under evidence case 2, the decision-maker receives information from a sensor located at GMPP 2 that indicates the ground motion intensity (spectral acceleration) at that site was in the range 0.45-0.50g (larger than expected for this event). This evidence is entered into the BN at the corresponding node and the information propagates throughout the BN. The observation updates the model in a forward direction, i.e. the known ground motion intensity provides a more accurate estimate of the failure probabilities of the associated components and, therefore, updates the system failure probabilities. This observation also provides information updating in the backward direction. It updates the error terms in the ground motion intensity model, which in turn update the ground motion intensity distributions at GMPPs for which observations are not available. The distributions of ground motion intensities at four GMPPs in the system (1, 6, 11, 16) are shown in Figure 8.32. Note that the distribution of the ground motion intensity at site 1 is shifted to the right relative to the distribution predicted for evidence case 1. This is due to the updating of the intra-event error random field ( $\epsilon_r$ ) and the common event error ( $\epsilon_m$ ). The updated distributions of these error terms in standardized form are shown in Figure 8.33. Note that the distribution of the intra-event deviation for GMPP 1 (which is near the point of observation) is significantly shifted, but the distributions for sites located at farther distances remain relatively unchanged due to loss of correlation of the random field with distance. Also, note that the distribution of the intra-event error for GMPP 6 will not change for any observation at other GMPPs because it is independent of other sites in the approximated random field model (see Figure 8.28).

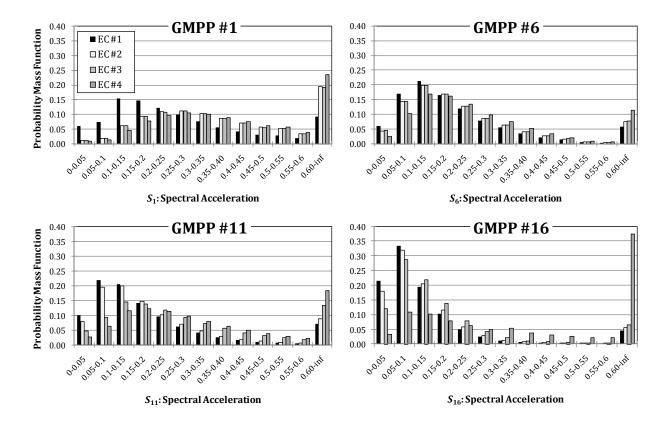


Figure 8.32: Distributions of ground motion intensities at GMPPs 1,6 11, 16 given evidence cases

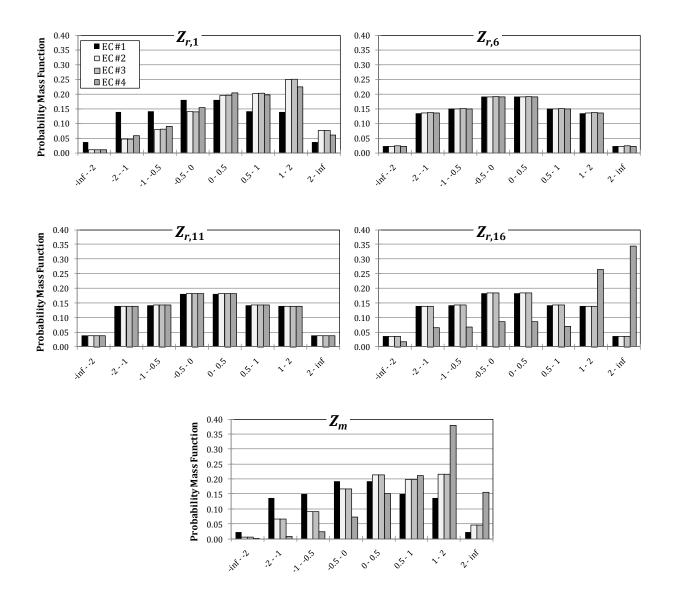


Figure 8.33: Distributions of normalized intra-event error terms ( $z_{r,i} = \varepsilon_{r,i}/\sigma_{r,i}$ ) for GMPPs 1,6,11,16 and the normalized inter-event error term ( $z_m = \varepsilon_m/\sigma_m$ )

The third evidence case corresponds to a situation where a partially processed ground motion record at GMPP 9 indicates the existence of a directivity pulse at that site, even though the intensity could not be determined. Note in Figure 8.32 that, in light of this information, the distributions of ground motion intensities at GMPPs 11 and 16 have shifted to larger values, while the distributions of the intra- and inter-event error terms have remained unchanged. The observation has also updated the distributions of the parameters defining the location of the rupture on the fault (see Figure 8.34). The shifted distributions suggest a rupture that is longer (though assumed bounded at ½ the fault

length) and that propagates more towards GMMP 9 than away from it. In turn, these updated distributions of source geometry increase the probability that the ground motion at the southern sites in the system will experience directivity effects.

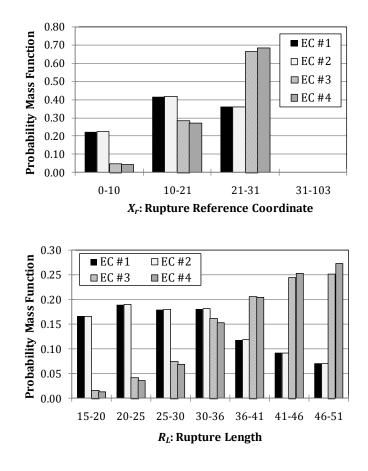


Figure 8.34: Distributions of parameters defining rupture location  $(X_r, R_L)$  given ECs

Evidence case 4 represents a situation in which, for example, a structural health monitoring sensor or an inspector reports that component 19 (the aerial structure located closest to Gilroy) is extensively damaged. This information is entered in the BN and updates the probabilistic model in several ways. First, it provides forward updating by decreasing the system reliability in light of evidence that a component has failed. For this particular evidence scenario, the system reliability is zero because the damaged component interrupts connectivity between all cities and Gilroy. The observation also updates the model in the backward direction by shifting the distributions of ground motion intensities of GMPPs at either end of the component towards larger values. This in turn updates the complete seismic demand model providing updated distributions of ground motion intensities at other sites in the system. In particular, the observation updates the intraevent error terms at sites near the failed component, consistent with the random field

model. The inter-event error distribution is also updated in a significant way. This updating of the common error term results from the accumulation of evidence, which all suggest higher than expected ground motion intensities at geographically dispersed locations. Consequently, the common inter-event error distribution shifts to the right to indicate that the ground motion prediction equations are systematically under-predicting intensities for all sites for this particular earthquake.

Table 8.5 shows the component state probabilities for the no damage and moderate or severe damage states. (The probability of slight damage is computed as 1 minus the sum of these two probabilities.) Table 8.6 shows the system failure probabilities for the four evidence cases. When the BN framework is extended with decision and utility nodes, these updated probabilities are important for accurately determining optimal decisions based on expected utilities.

	EC #1		EC #2		EC #3		EC #4	
Component	No damage	Moderate or severe damage						
1	0.94	0.06	0.89	0.10	0.89	0.10	0.86	0.13
2	0.95	0.05	0.91	0.08	0.91	0.08	0.89	0.11
3	0.95	0.05	0.90	0.10	0.90	0.10	0.86	0.13
4	0.95	0.05	0.91	0.08	0.91	0.08	0.88	0.12
5	0.96	0.04	0.95	0.05	0.95	0.05	0.92	0.08
6	0.96	0.04	0.95	0.05	0.94	0.05	0.91	0.09
7	0.96	0.04	0.94	0.05	0.94	0.06	0.91	0.09
8	0.96	0.04	0.95	0.05	0.93	0.06	0.90	0.10
9	0.95	0.05	0.94	0.06	0.92	0.07	0.88	0.11
10	0.95	0.05	0.93	0.07	0.90	0.09	0.85	0.14
11	0.96	0.04	0.95	0.05	0.93	0.07	0.89	0.10
12	0.95	0.03	0.94	0.04	0.92	0.06	0.88	0.08
13	0.96	0.04	0.95	0.05	0.93	0.06	0.90	0.10
14	0.97	0.02	0.96	0.03	0.94	0.04	0.91	0.07
15	0.96	0.04	0.95	0.05	0.94	0.06	0.91	0.09
16	0.97	0.03	0.96	0.04	0.95	0.05	0.51	0.49
17	0.97	0.03	0.96	0.04	0.95	0.05	0.00	1.00
18	0.95	0.03	0.93	0.04	0.92	0.06	0.67	0.24
19	0.98	0.02	0.97	0.03	0.96	0.04	0.90	0.09

### Table 8.5: Component state probabilities for states no damage and moderate or worse damage, under different evidence cases

	State	EC #1	EC #2	EC #3	EC #4
$SF \rightarrow Gilroy$	Open	0.64	0.50	0.46	0.00
SI-7 dill by	Reduced/Closed	0.36	0.50	0.54	1.00
Millbrae $\rightarrow$	Open	0.71	0.64	0.58	0.00
Gilroy	Reduced/Closed	0.29	0.36	0.42	1.00
Palo Alto $\rightarrow$	Open	0.75	0.69	0.63	0.00
Gilroy	Reduced/Closed	0.25	0.31	0.38	1.00
San Jose $\rightarrow$	Open	0.88	0.84	0.81	0.00
Gilroy	Reduced/Closed	0.12	0.16	0.19	1.00

Table 8.6: System state probabilities under different evidence cases

Next, we examine how the inspection prioritization changes under each evidence scenario. Preliminary (and hypothetical) prioritization orders for each evidence case are shown in Table 8.7. Under the first evidence case, link 3 is given highest priority, followed by links 1 and 2. Link 4 is located furthest from the epicenter and has components with generally the lowest probability of moderate or severe damage. Therefore, link 4 is not recommended for inspection. Instead, it is recommended that the link operate at reduced capacity. The prioritization order for the remaining links, under evidence case 1, is strongly influenced by the liabilities and system shutdown penalties. Link 3 has the largest number of components and thus the potential liability associated with making an unsafe decision is highest for this link. The prioritization of links 1 and 2 follows similar logic. The prioritization order changes when the ground motion intensity is observed at GMPP2. This GMPP lies along link 1 and, thus, the evidence of a high intensity there increases the probability that components along the link will experience damage. Therefore, even though link 3 is a longer link with more components and thus higher potential liability, it is second in the prioritization order. Evidence case 3 provides observation of a pulse-like ground motion at GMPP 9. This GMPP lies along link 3. Interestingly, this observation reduces the value of information associated with inspecting the components along link 3 and an inspection is no longer required. Instead, it is recommended to simply reduce the operating level of the link without inspection. That is, the recommendation to reduce the component operating level is not sufficiently likely to change, based on the results of an inspection, to warrant the cost of making inspections. Given knowledge that component 17 is extensively damaged, the prioritization order returns to the order suggested under evidence case 1, however the values of information associated with each link are higher. This reflects the accumulation of evidence that suggests ground motion intensities are higher than expected everywhere. Therefore, the prioritization order is once again controlled by the potential liabilities. Note however, that link 4 still does not require an inspection. This is because the observation of extensive damage of component 17 implies the link must be shutdown and thus there is no value in inspecting the other components along link 4.

Link	EC #1	EC #2	EC #3	EC #4
1	2	1	1	2
2	3	3	2	3
3	1	2	No Insp	1
4	No Insp	No Insp	No Insp	No Insp

#### Table 8.7: Prioritization order associated with each links under evidence cases

Once a link has been chosen for inspection, the individual components along the link are prioritized, based on the assumed heuristic described in Chapter 7. For illustration, Table 8.8 provides a preliminary prioritization order for link 1 for evidence case 2. The tunnels (components 1 and 3) require inspection first, followed by the embankment segments (components 2 and 4).

Table 8.8: Value of information and prioritization order associated with components along link 1
under evidence case 2

Comp	VoI	Order
1	27	1
2	14	3
3	18	2
4	18	2

#### 8.4 Summary

This chapter presents two example applications developed to demonstrate the value of using BNs for assessing and understanding seismic risk to infrastructure components and systems. These applications also serve to demonstrate the methodologies described in previous chapters. The first example in this chapter utilizes a simple example system in the vicinity of multiple faults to show how BNs can be utilized to perform and go beyond conventional probabilistic seismic hazard assessment. The second application utilizes an idealization of the proposed California high speed rail system to show the value in using BNs for post-earthquake applications, particularly in understanding system response and providing decision-support.

# **Chapter 9: Conclusions**

#### 9.1 Summary and major contributions

This study outlines a Bayesian network methodology for seismic infrastructure risk assessment and decision support. Bayesian networks have a variety of characteristics that make them well suited for the proposed application, including their graphical nature and, more importantly, the ease with which they facilitate modeling and information updating in complex problems. We primarily focus on the use of Bayesian networks to facilitate nearreal time updating of models of seismic hazard as well as component and system performance in the immediate aftermath of an earthquake. Sources of information that can be incorporated in the model include measurements of ground motion intensity from seismological recording instruments, data received from structural health monitoring sensors, and observations of component and system states.

A brief introduction to Bayesian networks is presented in Chapter 2. The goal of the chapter is to provide an overview of Bayesian networks sufficient to enable the reader to understand the contents in the remainder of the study. Wherever possible, the review avoids complex and/or theoretical descriptions in favor of intuitive explanations.

The Bayesian network methodology contained in this study is comprised of several components, which are described below and summarized in Figure 9.1.

- (1) A seismic demand model of ground motion intensity as a spatially distributed Gaussian random field accounting for multiple seismic sources with uncertain characteristics and including finite fault rupture and directivity effects.
- (2) A model of the performance of point-site and distributed components under seismic loading.
- (3) Models of system performance as a function of component performances.

(4) The extension of the Bayesian network to include decision and utility nodes to aid post-earthquake decision-making.

Chapter 3 and 4 address the first component of the framework. Components (2) and (3) are the subjects of Chapters 5 and 6, respectively. Chapter 7 describes component (4). A summary of the main contributions in each chapter is presented below.

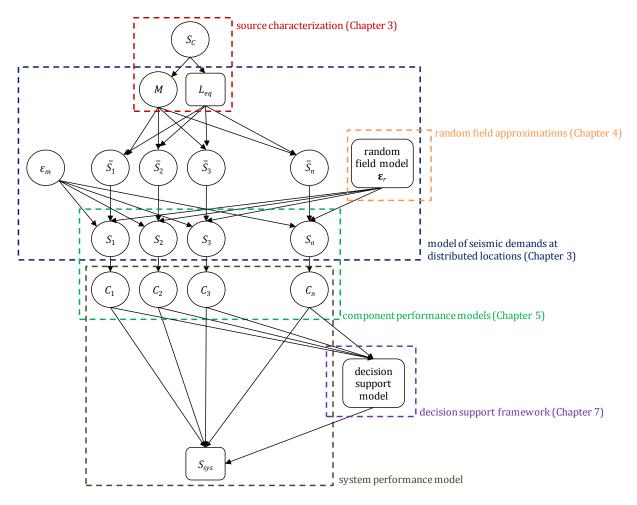


Figure 9.1: Summary of Bayesian network methodology for seismic infrastructure risk assessment and decision support

Chapter 3 presents a seismic demand model that provides distributions of ground motion intensity at discrete points in the geographic domain of a spatially distributed infrastructure system. Source characterization models are included in which a fault is treated as a single line segment or a contiguous set of line segments. Point-source and finite rupture idealizations are considered and accounting is made for directivity effects. The model provides a full random field characterization of the ground motion. Full probabilistic characterization of the ground motion as a spatially distributed random field enables observations of ground motion intensity made at one point in the system to be used to update distributions at other sites. Thus, the seismic demand model is useful not only for performing conventional probabilistic seismic hazard analysis, but also for use in near-real time post-earthquake hazard assessment and decision-making. While the presented Bayesian network model of seismic hazard can be further refined and extended, it is believed that it represents the most comprehensive model of its kind at the present time.

The modeling of random fields via Bayesian network results in a densely connected topology that can render the model computationally intractable when used for moderate or large size infrastructure systems. It is shown that neglecting to include correlation in seismic demands will result in erroneous estimates of system reliability, particularly in redundant systems. In view of the importance of including this correlation, Chapter 4 presents several approaches for approximating the random field model by removing links and nodes in the Bayesian network. The goal of these approaches is to balance computational complexity and accuracy by eliminating as many links and nodes as possible, without significantly sacrificing the accuracy with which the random field is modeled. Approximations based on Cholesky decomposition, eigenvalue expansion, and numerical optimization are considered. It is found that numerical optimization is able to achieve the best trade-off of accuracy versus efficiency. The methods in Chapter 4 are necessary to facilitate accurate and efficient updating of the seismic demand model in nearreal time applications, as well as to ensure accurate modeling of seismic hazard across the system. However, the proposed approximation methods are also useful within a broader context. The procedures outlined in Chapter 4 improve the feasibility of using Bayesian networks to model a variety of problems involving random fields that can benefit from the information updating capabilities of this approach.

Chapter 5 presents models of component performance under seismic demand using fragility functions. Fragility models are presented for point-site (e.g. short-span bridges, buildings) as well as distributed components (e.g. pipelines, railroads). A generic Bayesian network framework is outlined that accounts for statistical uncertainties and model errors, which can give rise to correlations between estimated performances of similar components during a single earthquake. Because few fragility models have been developed that properly account for statistical uncertainty and model error, simpler Bayesian network formulations are also developed that reflect the state-of-the-art.

Chapter 6 presents several methods for modeling system performance, as a function of component performances, within the context of Bayesian networks. Five formulations for modeling system performance are described, including a naïve formulation, two approaches based on an intuitive interpretation of system performance, and two approaches that utilize minimal link and cut sets. The last two formulations are adapted and refined with the goal of minimizing computational demands. Computational demands are reduced by utilizing a Bayesian network topology, in which nodes are arranged in chain structures rather than the converging structures that characterize most Bayesian network system performance models, including the aforementioned formulations. The result is an approach for defining system performance formulations that greatly improves computational efficiency relative to other methods. To automate the construction of efficient Bayesian network formulations, a novel heuristically augmented optimization

algorithm is presented. Both system connectivity and flow problems are considered. The optimization algorithm provides a method for constructing efficient Bayesian network models of the performance of topologically defined systems in a wide range of applications beyond seismic infrastructure risk assessment.

Enabling the inclusion of models of component and system performance, in addition to seismic hazard, allows the framework outlined in this study to go beyond many existing methodologies. The Bayesian network provides time-evolving estimates of component and system performance following an earthquake, while accounting for all available information such as measurements of ground motion intensity and observations of component and system performance. Thus, the framework goes beyond the abilities of any existing methodology for component and infrastructure seismic risk assessment that we have reviewed with regard to near-real time applications.

Chapter 7 extends the proposed Bayesian network methodology to solve decision problems. This is achieved by adding decision and utility nodes to the Bayesian network. Once these nodes have been added to the network, the new graphical construct is called an influence diagram. An influence diagram provides decision guidance based on principles of expected utility maximization. This guidance changes as evidence is entered into the network and as decisions are made. By considering an application involving postearthquake inspection and shutdown decisions, we demonstrate the extension of the proposed Bayesian network methodology to solve decision problems. Furthermore, we define a heuristic, based on value of information criteria, that aid post-earthquake inspection using the output of the influence diagram.

Finally, Chapter 8 presents two example applications that demonstrate the value of using Bayesian networks for the proposed application to seismic infrastructure risk assessment and decision support. The first application focuses on using the Bayesian network to perform and go beyond conventional probabilistic seismic hazard assessment. The second application focuses on post-earthquake information updating and decision-making. The subject of this example is a hypothetical model of a segment of the proposed California high speed rail system.

The Bayesian network provides a powerful tool for modeling risks posed to infrastructure systems both before and after a hazard event. While the use of Bayesian networks in civil engineering remains limited, this study shows the power in applying this framework to aid post-event decisions for emergency response and recovery of infrastructure systems. The study develops a methodology for using Bayesian networks to perform seismic infrastructure hazard assessment and provide decision support. This represents a novel approach to assessing and responding to seismic risk that draws upon and integrates knowledge from a range of disciplines, including computer science, hazard modeling, structural engineering, and system analysis methods. The result is a fully updateable probabilistic methodology that goes well beyond the capabilities of existing tools. In addition to demonstrating the value of using Bayesian networks for seismic infrastructure risk assessment and decision support, this study develops models necessary to construct

efficient Bayesian networks for this purpose. These models are developed with the goal of minimizing computational demands, while accurately reflecting the full probabilistic characterization of the problem. Although the hazard, component, system, and decision models in this report have been developed with the goal of defining a comprehensive methodology for the proposed earthquake engineering application, many of the models and algorithms are useful in a broader context. As a result, the methodology described in this study provides the foundation for a new and innovative computational tool for performing risk assessment and providing decision support for infrastructures subjected to variety of natural and man-made hazards.

#### 9.2 Future research and development

The methodology outlined in this report lays the groundwork for a comprehensive seismic infrastructure risk assessment and decision support system. The framework can be extended and refined in many ways. Examples for future extensions and refinements may include:

(1) Expanded models:

- The seismic demand model in Chapter 3 considers line-idealizations of faults. More complex finite rupture plane and seismologically-based formulations may improve the accuracy of the source characterization.
- The liquefaction and fault rupture seismic demand models in Chapter 3 are not fully developed. Furthermore, hazards such as landslide have not been addressed. To comprehensively model seismic hazard to infrastructure systems, it is necessary to develop models of these and other hazards with a level of detail that is consistent with the ground-shaking model.
- The component performance models developed in this study use fragility functions. For some infrastructure system components, the idealization of component performance through this mapping may be insufficient. Development of discipline-specific component performance formulations remains an area warranting further development.
- Chapter 6 presents a heuristically-augmented optimization algorithm to automate construction of efficient Bayesian network formulations for modeling system performance. Development of additional heuristics that further reduce problem size and increase efficiency is recommended.
- The methodology described in chapter 6 considers generic system connectivity and flow problems. The development of discipline-specific Bayesian network models is needed, e.g. transportation demand and travel time models, hydrological models for water distributions systems. Furthermore, it is known that system interdependency can strongly affect system reliability. Development of formulations to capture system interdependency is an area warranting future research.
- Chapter 7 demonstrates the value of extending the Bayesian network framework by utility and decision nodes to solve decision problems. However, only a single application to the post-earthquake inspection-

shutdown decision is considered. Gross assumptions and simplifications were made relating to economic consequences. Additional research is needed to accurately define utility values, particularly at the system level. Furthermore, the current formulation does not address the amount of time required to carry out actions, e.g. to inspect a component, nor such issues as the effect of downtime or loss of service. The decision-formulation needs to be expanded to consider time effects, which are critical in immediate post-earthquake applications.

- Significant work remains relating to the development of decision formulations for additional applications such as the dispatch of rescue personal in the immediate aftermath of an earthquake and allocation of resources for repair and reconstruction. The decision models should ideally be formulated in collaboration with actual decision-makers.
- (2) Computational issues:
  - When applying the framework developed in this study to large and complex infrastructure systems, a multi-scale modeling approach will need to be utilized. Guidance on multi-scale modeling is not contained in this study.
  - This study focuses on the use of exact and some sampling-based algorithms, e.g. (Yuan and Druzdzel 2006), for performing probabilistic inference. Future work is needed to examine the accuracy of the sampling-based methods as well as the viability, precision, and convergence (in time) of other methods, e.g. Markov-Chain Monte Carlo. Furthermore, sensitivities to discretization of continuous variables should be examined.
  - To ensure computational feasibility in the application of Bayesian networks to problems involving large infrastructure systems, it may be necessary to develop inference algorithms specifically designed to handle the challenges encountered by this specific class of problems. Variational and sampling-based approaches developed specifically for the topology of the Bayesian network developed for seismic infrastructure risk assessment may significantly increase computational efficiency (in time and space). Efficiency is crucial in near-real time applications.

(3) Integration with external information sources:

• To make the methodology described in this study viable for practical applications, it must be integrated directly with tools that process data from external sources, such as structural health monitoring sensors and ground motion recording instruments. Data should be fed directly into the Bayesian network, without external action by the operator. This will require the development of a computer application that links these data acquisition resources with the Bayesian network. Furthermore, the development of an interface which processes the probabilistic output of the Bayesian network to create charts and graphs that provide, for example, hazard curves, prioritization lists, and other decision guidance would be valuable.

#### (4) Additional applications:

- The proposed framework is useful for modeling hazard beyond earthquakes. Because the methodology developed in this study is modular, it can be expanded to consider other hazards (e.g. hurricanes, fire). Models for other hazards can be substituted in place of the seismic demand model. Modification of the mapping between the demands and component performance is also required. However, the system performance model likely need not be changed. The development of models for additional hazards will increase the usefulness of the proposed framework.
- It is envisioned that the Bayesian network methodology can be used to aid decision-making under normal operating conditions, e.g. maintenance and rehabilitation planning. As described in the introduction, decisions that improve system reliability under normal operating conditions and resilience to extreme loads should ideally be made in tandem. Special decision models are necessary to consider issues related to normal operating conditions.
- Seismic events provide information about seismic hazard and component and system performance. This information should be used to update the hazard models, component fragilities, and system characterizations that are used in the Bayesian network. The methodology described here should be extended to facilitate this updating.

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