UCSF

UC San Francisco Previously Published Works

Title

Association of β -Blockers With Functional Outcomes, Death, and Rehospitalization in Older Nursing Home Residents After Acute Myocardial Infarction

Permalink

https://escholarship.org/uc/item/8zr4f4gr

Journal

JAMA Internal Medicine, 177(2)

ISSN

2168-6106

Authors

Steinman, Michael A Zullo, Andrew R Lee, Yoojin et al.

Publication Date

2017-02-01

DOI

10.1001/jamainternmed.2016.7701

Peer reviewed



HHS Public Access

JAMA Intern Med. Author manuscript; available in PMC 2018 February 01.

Published in final edited form as:

JAMA Intern Med. 2017 February 01; 177(2): 254–262. doi:10.1001/jamainternmed.2016.7701.

Impact of beta blockers on functional outcomes, death, and rehospitalization in older nursing home residents following acute myocardial infarction

Michael A. Steinman, MD, Andrew R. Zullo, PharmD, ScM, Yoojin Lee, MS, MPH, Lori A. Daiello, PharmD, ScM, W. John Boscardin, PhD, David D. Dore, PharmD, PhD, Siqi Gan, MPH, Kathy Fung, MS, Sei J. Lee, MD, MAS, Kiya D.R. Komaiko, BA, and Vincent Mor, PhD Division of Geriatrics (MAS, WJB, SG, KF, KDRK) and Department of Epidemiology and Biostatistics (WJB), University of California, San Francisco and San Francisco VA Health Care System; Department of Health Services, Policy, and Practice, Brown University School of Public Health (ARZ, YL, LAD, DDD, VM); Providence VA Medical Center (VM); and Optum Epidemiology, Boston, MA (DDD)

Abstract

Importance—Beta blockers are a mainstay of treatment after acute myocardial infarction (AMI). Yet, these medications are commonly not prescribed for older nursing home residents after AMI, in part owing to concerns about potential functional harms and uncertainty of benefit.

Author contact information: Michael Steinman, MD, 4150 Clement St, VA Box 181G, San Francisco, CA 94121, T: 415.221.4810 x23677, F: 415.750.6641, mike.steinman@ucsf.edu.

Data Access and Responsibility: Drs. Steinman and Mor had full access to all the data in the study and take responsibility for the integrity of the data and the accuracy of the data analysis.

Financial Disclosures:

- Dr. Steinman: MAS is a paid consultant for iodine.com.
- Dr. Mor: VM's research is in a related area to that of several different paid activities. VM also periodically serves as a paid speaker at national conferences where he discusses trends and research findings in long term and post-acute care. VM holds stock of unknown value in PointRight, Inc. an information services company providing advice and consultation to various components of the long term care and post-acute care industry, including suppliers and insurers. PointRight sells information on the measurement of nursing home quality to nursing homes and liability insurers. VM was a founder of the company but has subsequently divested much of his equity in the company and relinquished his seat on board. In addition, VM Chairs the Independent Quality Committee for HRC Manor Care, Inc., a nursing home chain, for which he receives compensation in the \$20,000-\$40,000 range. VM also serves as chair of a Scientific Advisory Committee for NaviHealth, a post-acute care service organization, for which he also receives compensation in the \$20,000-40,000 per year range. VM serves as a Technical Expert Panel member on several Centers for Medicare/ Medicaid quality measurement panels. VM is a member of the board of directors of Tufts Health Plan Foundation; Hospice Care of Rhode Island; and The Jewish Alliance of Rhode Island.
- Dr. Dore: D.D.D. is an employee of Optum and stockholder in UnitedHealth Group, Optum's parent company.
- Other authors: The other authors report no conflicts of interest.

Author contributions: Study concept and design (Steinman, Boscardin, Mor), acquisition of data (Steinman, Zullo, Daiello, Dore, Fung, Komaiko, Mor), analysis of data (Zullo, Y. Lee, Gan, Fung), interpretation of results (Steinman, Zullo, Y. Lee, Daiello, Boscardin, Dore, Gan, Fung, S. Lee, Covinsky, Komaiko, Mor), preparation of the initial draft of manuscript (Steinman), and critical review of the manuscript (all authors).

Prior presentation: Presented at the national meetings of the American Geriatrics Society (May 2016) and Academy Health (June 2016)

Conflict of interest disclosures: Please see acknowledgments page for full disclosures

Objective—We studied the impact of beta blockers after AMI on functional decline, mortality, and rehospitalization among long-stay nursing home residents age 65 and older.

Design, setting, participants, exposure—Observational study of nursing home residents with AMI in 2007–2010, using national data from the Minimum Data Set and Medicare Parts A and D. Subjects with beta blocker use prior to AMI were excluded. We used propensity score-based methods to compare outcomes in people who did vs. did not initiate a beta blocker after AMI hospitalization.

Main outcomes—Functional decline, death, and rehospitalization in the first 90 days after AMI. Functional status was measured using a validated 28-point scale that evaluates independence in activities of daily living.

Results—The study cohort included 5,496 new beta blocker users and an equal number of non-users. Mean age was 84 years. Beta blocker users were more likely than non-users to experience functional decline (OR 1.14, 95% CI 1.02–1.28), with a number need to harm of 52. Conversely, beta blocker users were less likely than non-users to die (HR 0.74, 95% CI 0.67–0.83) and had similar rates of re-hospitalization (HR 1.06, 95% CI 0.98–1.14). Nursing home residents with moderate or severe cognitive impairment or severe functional dependency were particularly likely to experience functional decline from beta blockers (OR 1.34, 95% CI 1.11–1.61 and OR 1.32, 95% CI 1.10–1.59, respectively). In contrast, there was little evidence of functional decline from beta blockers in subjects with better cognitive and functional status (ORs 0.99 to 1.05; P value for effect modification 0.03 and 0.06, respectively). Mortality benefits of beta blockers were similar across all subgroups.

Conclusions/relevance—Use of beta blockers after AMI is associated with functional decline in older nursing home residents with substantial cognitive or functional impairment, but not in those with relatively preserved mental and functional abilities. Beta blockers yielded considerable mortality benefit in all groups.

Beta blockers are a mainstay of guideline-recommended care for adults following acute myocardial infarction (AMI).^{1,2} Randomized trials in middle-aged and "young-old" adults show that treatment with beta blockers after AMI reduces mortality by 25–30%.^{3–5} Multiple observational studies have found a similar level of mortality reduction in adults 85 years and older and in those with functional impairment.^{6–9}

Despite the benefits of beta blockers across the age span, these medications are less often prescribed to older adults, especially those with functional impairment or multimorbidity. 6,7,10,11 Although studies have suggested that beta blockers are generally well-tolerated in older adults, 12–14 there are little data on their adverse event profile in frail and highly vulnerable elders, including potential harms such as orthostasis, fatigue, and depression, which can negatively impact daily functioning and quality of life. This dilemma, where potential mortality benefits are weighed against an unclear level of harms, is common in the care of vulnerable older adults. 15–18 It is particularly important for the 1.4 million Americans who reside in nursing homes, who are at high risk of functional decline and often strongly value preserving whatever remaining functional independence they have. 19,20

In this study, we evaluated the impact of beta blockers on functional outcomes in older nursing home residents with myocardial infarction, and compared these functional outcomes with the impact of beta blockers on death and re-hospitalization in this population.

Methods

Data Sources and Subjects

Data came from Medicare Part A and Part D (prescription drug benefit) claims; the Online Survey Certification and Reporting System (OSCAR), which provides facility-level information on nursing home characteristics, staffing, and quality indicators; and the Minimum Data Set (MDS) version 2.0, which comprises assessments made on nearly all nursing home residents in the U.S. MDS assessments occur a minimum of every 3 months, and more often for patients with a major recent change in clinical status and those receiving care under the Medicare Skilled Nursing Facility (SNF) benefit.

Our study population comprised U.S. nursing home residents age 65 and older who were hospitalized for AMI between May 1, 2007 and March 31, 2010, had resided in a nursing home for at least 30 days prior to the AMI hospitalization, were not using a beta blocker for at least four months prior to hospitalization, and returned to a nursing home after hospital discharge (see Appendix 1 and Zullo et al²¹ for additional details). We defined hospitalization with AMI based on a hospital admission or discharge claim with ICD9 code 410.XX or 411.1 as a primary or secondary diagnosis. We excluded patients who died, were rehospitalized, or otherwise left the nursing home within 14 days of hospital discharge, because in such short-stay situations it is difficult to reliably ascertain beta blocker use. We also excluded patients with very poor prognosis at baseline (Changes in Health, End-Stage Disease, and Signs and Symptoms [CHESS] score of 5 or hospice), ²² patients who were not continuously enrolled in Medicare Part D during the study period or had no Part D claims following hospitalization, and patients who were enrolled in a Medicare Advantage plan at any point during this period. Finally, we excluded subjects with extremely poor functional status prior to hospitalization (Morris ADL score 24/28) since they had little room for further functional decline.²³

Measures

Our exposure of interest was use of a beta blocker in the immediate post-hospital period. We defined this as a Part D claim for an oral beta blocker within 30 days of resuming Part D coverage after hospital discharge. Part D covers at least 81% of nursing home residents and in most cases is the sole source of prescription drug coverage for these patients.²⁴ For the subset of patients who return to the nursing home under the Medicare Skilled Nursing Facility (SNF) benefit, resumption of Part D claims is temporarily delayed. Therefore, we conducted a companion validation study to evaluate the performance of our beta blocker exposure measure in this subset. This study confirmed the validity of our measure (see Appendix 1).

Our primary outcome was functional decline. We defined this as a loss of 3 points on a validated 28-point scale of independence in activities of daily living (ADLs) between the

pre-hospital baseline and the first available assessment following hospitalization, up to 3 months after discharge. ²³ A 3-point drop corresponds to a major loss of independence in 1 ADL or incremental losses in 2 or more ADLs. In a sensitivity analysis, we evaluated the outcome as a 4-point (more substantial) decline in function. We chose a 90-day outcome period because it is long enough to be clinically meaningful, but short enough that many of these highly vulnerable patients have not yet died, a competing outcome which complicates interpretation of longer-term functional outcomes.

Other key outcome measures included death and re-hospitalization within 90 days of the index hospital discharge. We used data from Medicare Part A and Medicare enrollment files to identify hospital admissions and date of death. We also explored two composite outcomes: time to hospitalization or death, and time to hospitalization, death, or functional decline.

Information on chronic conditions and characteristics of the index hospitalization were obtained from Medicare Part A data. Overall, this data source is more accurate for identifying chronic conditions than MDS 2.0.^{25–28} MDS 2.0 provided data on other patient characteristics including functional and cognitive status, geriatric syndromes, and symptoms, including validated scales such as the Cognitive Performance Score (CPS) and CHESS score.^{22,29}

We used the OSCAR dataset to evaluate a variety of nursing home facility characteristics such as staffing, resident mix and quality indicators.

Analyses

We used propensity score-based methods to evaluate the relationship between beta blocker exposure and our outcomes of interest. Following an intention-to-treat framework, we defined subjects as beta blocker users or non-users throughout the study period based on their exposure in the immediate post-AMI period.

We estimated the propensity score via a logistic regression model that used 93 variables to predict beta blocker use. Variables included sociodemographic characteristics, chronic medical conditions, baseline medication use, prior hospitalization history, baseline functional and cognitive status, geriatric syndromes, symptoms, characteristics of the AMI hospitalization, and nursing home characteristics (Appendix 2). To evaluate whether vital signs, laboratory test results, and measures of cardiac function could result in unmeasured confounding, we conducted a companion validation study using national VA data, which unlike Medicare claims data contains information on these parameters. We found no evidence that the absence of these factors would substantially alter our results (Appendix 3).

To match beta blocker users with non-users who had similar propensity scores, we first discarded subjects in the top and bottom 1% of the propensity score distribution so as to exclude areas of non-overlap. We then applied a 1:1 greedy 5-to-1 digit matching algorithm without replacement.³⁰ We evaluated the quality of resulting matches by comparing standardized differences between groups for each covariate in our model, and by using t-tests to assess differences in the distribution of propensity scores.^{31,32}

Our propensity matching yielded excellent covariate balance, so we did not further adjust for baseline covariates in our models. Because we excluded people who died or were rehospitalized during the first 14 days after hospital discharge, we did not consider outcomes that occurred during this period, thus effectively beginning our outcome analyses at day 14 after hospitalization.

We used Cox proportional hazards models to determine the impact of beta blocker use on time to death. We used the method of Fine and Gray (similar to Cox regression) to evaluate the impact of beta blocker use on time to rehospitalization while accounting for the competing outcome of death. Finally, we used multinomial logit models to evaluate the impact of beta blocker use on functional decline. At the end of the 90 day followup period, subjects were classified as alive without functional decline, having had functional decline documented in the first MDS assessment of that period, or having died without evidence of functional decline on the first MDS assessment.

We used both multiplicative and additive interaction terms to evaluate whether the impact of beta blockers on outcomes varied across subject characteristics. These characteristics included levels of baseline functional status, cognitive function, age, and presence or absence of an ICU or CCU stay during the AMI hospitalization. The distribution of propensity scores was very similar for beta blocker users and non-users within each subgroup, suggesting that stratifying patients into subgroups did not threaten covariate balance (Appendix 4).

The decision to exclude patients who died or were rehospitalized within 14 days after the AMI discharge has the potential to create selection bias. To evaluate this, we repeated our main analyses using inverse probability of selection weighting. ^{34,35} This approach weighted subjects according to their similarity to individuals who were excluded due to death (N=1,859) or re-hospitalization (N=2,444) in the first 14 days, thus estimating treatment effects as if these people had been included in the analysis. In another sensitivity analysis, we controlled for use of other cardiovascular medications post-AMI using multinomial logistic regression in our propensity-matched cohort.

We also evaluated several alternate approaches to determine if our results were stable across different analytic techniques. These included stratifying by propensity score quintile and deciles, controlling for propensity score as a covariate, using inverse probability of treatment weights, performing time-dependent analyses. In each case, results were similar to our main approach (Appendix 5).

Results

Our initial cohort included 8,953 new beta blocker users and 6,767 non-users. Before matching, beta blocker users were more likely to have been in an ICU or CCU during the hospital stay and to return to the nursing home on the Medicare SNF-benefit care pathway, and less likely to have a prior diagnosis of angina pectoris or unstable angina (Table 1 and Appendix 6).

Propensity score matching yielded a cohort of 5,496 new beta blocker users and an equal number of non-users (Table 1). Mean age was 84 years. The distribution of propensity scores was nearly identical between the matched groups (P=0.63), and all but 2 variables had standardized mean differences of 0.03 or less (Appendix 6). This is consistent with excellent covariate balance between groups. Beta blocker users and non-users had equal time between nursing home readmission and their first ADL assessment (median 22 days, IQR 11–29 days, P=0.97 for difference). New beta blockers users were more likely than non-users to be prescribed other cardiovascular medications in the post-AMI period, including statins (49% vs 32%, P<.0001) and ACE inhibitors (44% vs. 31%, P<.0001), but not angiotensin receptor blockers (8% vs 7%, P=.17).

Within 3 months after hospital discharge, 1,328 of 10,992 subjects (12%) experienced functional decline, 2,782 (25%) were rehospitalized, and 1,541 (14%) died. Some patients experienced more than one outcome; e.g. were rehospitalized and then died.

Beta blocker users had a higher rate of functional decline than non-users. In the first 90 days after AMI, the odds of functional decline were 1.14 (95% CI, 1.02–1.28) times greater in patients receiving beta blockers than in those not using beta blockers (Table 2). The number needed to treat to cause one patient to have functional decline was 52 (95% CI, 32–141). Results were similar using the more stringent threshold of a 4-point decline on the Morris ADL scale: using this definition, 1,165 subjects (11%) had functional decline, and beta blocker users were more likely to decline (OR 1.16, 95% CI, 1.02–1.31).

Beta blocker users were less likely than non-users to die within 90 days of hospital discharge (hazard ratio [HR] 0.74, 95% CI, 0.67–0.83; Figure 1 and Table 2). The number needed to treat to prevent one death was 26 (95% CI, 19–39). Beta blocker use had no impact on time to re-hospitalization (HR 1.06, 95% CI 0.98–1.14).

Beta blocker use had no significant effect on a composite outcome of time to death, hospitalization, or functional decline (HR 0.98, 95% CI 0.94–1.03). Beta blocker use showed a borderline small protective effect for a composite outcome that only included time to death or hospitalization (HR 0.94, 95% CI 0.88–1.00).

The impact of beta blocker use on death was similar across a variety of patient characteristics (Figure 3). However, the impact of beta blocker use on functional decline varied according to patients' baseline cognitive and functional status (Figure 3 and Appendix 7). Among nursing home residents with moderate or severe cognitive deficits, beta blocker users were substantially more likely than non-users to experience functional decline (OR 1.34, 95% CI 1.11 – 1.61), with a number needed to harm of 36 (95% CI, 24–76). Similarly, among residents with severe functional dependence at baseline, beta blocker users had greater risk of functional decline than did non-users (OR 1.32, 95% CI 1.10–1.59), with a number needed to harm of 25 (95% CI, 16–55). In contrast, beta blocker use did not increase the risk of functional decline in people with intact cognition or mild dementia (OR 1.03, 95% CI 0.89–1.20) or in those with less impaired levels of functioning prior to their hospitalization for AMI (OR 1.05, 95% CI 0.86–1.27 and OR 0.99, 95% CI 0.77–1.26,

respectively). The P values for effect modification on the multiplicative scale were 0.03 for baseline cognitive status and 0.06 for baseline functional status.

The main results were similar after applying inverse probability of selection weights, although the point estimate for the impact of beta blockers on functional decline was slightly attenuated, with 95% confidence intervals crossing 1 (OR 1.09, 95% CI 0.96–1.24). Similar patterns held for results of subgroup analyses using selection weights (Appendix 8). Finally, results were similar after controlling for use of other cardiovascular medications in the post-AMI setting (Appendix 5).

Discussion

In this national study of older nursing home residents, using beta blockers after acute myocardial infarction resulted in a 26% relative reduction in 90-day mortality, with a number needed to treat of 26 to prevent one death. Similar levels of risk reduction were found across a wide variety of patient subgroups. However, beta blockers conferred a 14% relative increase in the odds of functional decline, with a number need to harm of 52 to cause one case of functional decline. This risk was particularly high for people with moderate or severe cognitive impairment or a high degree of functional dependence at baseline. In these groups, beta blockers increased the odds of functional decline by 32–34%, with a number needed to harm of 25 to 36. In contrast, nursing home residents with relatively preserved cognitive and functional abilities did not appear to suffer adverse functional consequences from receiving beta blockers.

Our findings of mortality benefit are consistent with the results of other observational studies of beta blocker use among the old-old, frail, and functionally impaired.^{6–9,36,37} Regarding harms, little is known about the impact of beta blockers on functional status. However, these agents increase risk of fatigue (particularly first-generation agents such as propranolol)¹² and have been associated with increased rates of dizziness^{38,39} and decreased subjective sense of well-being,^{40,41} although no consistent effect has been found on rates of depression¹² or falls.^{13,42}

Our results confirm the suspicion of many physicians that poor cognitive and functional status increase the risk of medication-induced harms in older adults. However, they call into question the more general practice whereby older adults are less likely to receive guideline-recommended medications after AMI regardless of their mental or physical abilities.^{7,10,11,43} For nursing home residents with intact cognition or mild dementia, and in those with non-severe levels of functional dependency, we found substantial mortality benefit and no functional harms. So, for most such patients treatment is appropriate. In contrast, for nursing home residents with extensive functional dependency or moderate to severe dementia (roughly corresponding to a Folstein Mini Mental State Exam score of 14/30 or lower),²⁹ resolving the tradeoff between reduced mortality and increased risk of functional decline will depend on patient preferences, as expressed directly or through surrogate decision-makers.^{44,45} For cognitively or functionally impaired nursing home residents who are more concerned about functional decline than death, avoiding treatment may be preferable. This is

a large population: more than half of nursing home residents have high levels of functional dependence, and two-thirds have moderate or severe cognitive impairment.⁴⁶

Because this is an observational study, we cannot rule out the possibility of confounding. However, several factors support the robustness of our findings. We obtained excellent balance of baseline covariates across treatment groups and consistent results using several alternate analytic approaches. Moreover, younger and healthier patients are more likely to receive secondary prevention medications after AMI.^{7,10,11,43,47} This would bias results toward better outcomes in beta blocker users. Instead, functional outcomes were in the opposite direction of this expected bias. Another important consideration is co-interventions. People who used beta blockers after AMI were also more likely to receive statins and ACE-inhibitors in the post-AMI period. Controlling for these differences slightly attenuated the observed associations between beta blocker use and our outcomes of interest, although the overall pattern remained.

To enable robust assessment of beta blocker exposure, we excluded subjects who died or were rehospitalized within the first 14 days of hospital discharge. This prevented us from evaluating the impact of beta blockers on outcomes during this period. Thus, our results should be interpreted as providing evidence about the impact of beta blocker use on outcomes starting 14 days after discharge, among people who had survived and remained in the nursing home until then. In addition, these exclusions could induce selection bias. ^{34,35} However, while our sensitivity analyses were consistent with the possibility of mild selection bias, we found little evidence of bias sufficiently large to invalidate our overall findings.

Use of beta blockers after myocardial infarction resulted in substantial reductions in mortality among older nursing home residents. At the same time, use of these agents resulted in worse functional outcomes among nursing home residents with substantial cognitive or functional deficits. In this highly vulnerable group, understanding the importance that individual patients place on avoiding death and on avoiding functional decline will be critical to guiding decision-making about use of these medications.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

Acknowledgments

The authors thank Irena Stijacic-Cenzer, MS (University of California, San Francisco) for her insights on analytic approaches and Kenneth Covinsky, MD MPH (University of California, San Francisco and the San Francisco VA Medical Center) for his guidance on study design and interpretation. Ms Stijacic-Cenzer and Dr. Covinsky both received support from the NIH R01 award that funded this work.

Support: Support was provided by the National Institute of Health (R01HL111032 and K24AG049057). Dr. Zullo is supported by an Agency for Healthcare Research and Quality award (5K12HS022998). The funders had no influence on the design and conduct of the study; collection, management, analysis, and interpretation of the data; preparation, review, or approval of the manuscript; and decision to submit the manuscript for publication.

Bibliography

 Amsterdam EA, Wenger NK, Brindis RG, et al. 2014 AHA/ACC guideline for the management of patients with non-ST-elevation acute coronary syndromes: executive summary: a report of the American College of Cardiology/American Heart Association Task Force on Practice Guidelines. Circulation. 2014; 130(25):2354–2394. [PubMed: 25249586]

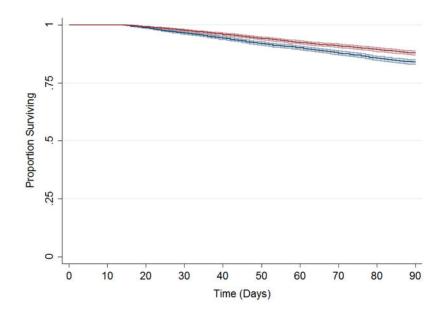
- 2. O'Gara PT, Kushner FG, Ascheim DD, et al. 2013 ACCF/AHA guideline for the management of ST-elevation myocardial infarction: a report of the American College of Cardiology Foundation/American Heart Association Task Force on Practice Guidelines. Circulation. 2013; 127(4):e362–425. [PubMed: 23247304]
- 3. Gundersen T, Abrahamsen AM, Kjekshus J, Ronnevik PK. Timolol-related reduction in mortality and reinfarction in patients ages 65–75 years surviving acute myocardial infarction. Prepared for the Norwegian Multicentre Study Group. Circulation. 1982; 66(6):1179–1184. [PubMed: 6128084]
- The Beta-Blocker Pooling Project (BBPP): subgroup findings from randomized trials in post infarction patients. The Beta-Blocker Pooling Project Research Group. Eur Heart J. 1988; 9(1):8– 16.
- 5. A randomized trial of propranolol in patients with acute myocardial infarction. I. Mortality results. JAMA. 1982; 247(12):1707–1714. [PubMed: 7038157]
- Soumerai SB, McLaughlin TJ, Spiegelman D, Hertzmark E, Thibault G, Goldman L. Adverse outcomes of underuse of beta-blockers in elderly survivors of acute myocardial infarction. JAMA. 1997; 277(2):115–121. [PubMed: 8990335]
- Vitagliano G, Curtis JP, Concato J, Feinstein AR, Radford MJ, Krumholz HM. Association between functional status and use and effectiveness of beta-blocker prophylaxis in elderly survivors of acute myocardial infarction. J Am Geriatr Soc. 2004; 52(4):495–501. [PubMed: 15066062]
- 8. Krumholz HM, Radford MJ, Wang Y, Chen J, Heiat A, Marciniak TA. National use and effectiveness of beta-blockers for the treatment of elderly patients after acute myocardial infarction: National Cooperative Cardiovascular Project. JAMA. 1998; 280(7):623–629. [PubMed: 9718054]
- Gottlieb SS, McCarter RJ, Vogel RA. Effect of beta-blockade on mortality among high-risk and low-risk patients after myocardial infarction. N Engl J Med. 1998; 339(8):489–497. [PubMed: 9709041]
- Levy CR, Radcliff TA, Williams ET, Hutt E. Acute myocardial infarction in nursing home residents: adherence to treatment guidelines reduces mortality, but why is adherence so low? J Am Med Dir Assoc. 2009; 10(1):56–61. [PubMed: 19111854]
- 11. Rochon PA, Anderson GM, Tu JV, et al. Use of beta-blocker therapy in older patients after acute myocardial infarction in Ontario. CMAJ. 1999; 161(11):1403–1408. [PubMed: 10906894]
- Ko DT, Hebert PR, Coffey CS, Sedrakyan A, Curtis JP, Krumholz HM. Beta-blocker therapy and symptoms of depression, fatigue, and sexual dysfunction. JAMA. 2002; 288(3):351–357.
 [PubMed: 12117400]
- 13. Woolcott JC, Richardson KJ, Wiens MO, et al. Meta-analysis of the impact of 9 medication classes on falls in elderly persons. Arch Intern Med. 2009; 169(21):1952–1960. [PubMed: 19933955]
- 14. Sztramko R, Chau V, Wong R. Adverse drug events and associated factors in heart failure therapy among the very elderly. Can Geriatr J. 2011; 14(4):79–92. [PubMed: 23251319]
- 15. Fuat A, Hungin AP, Murphy JJ. Barriers to accurate diagnosis and effective management of heart failure in primary care: qualitative study. BMJ. 2003; 326(7382):196. [PubMed: 12543836]
- Francke AL, Smit MC, de Veer AJ, Mistiaen P. Factors influencing the implementation of clinical guidelines for health care professionals: a systematic meta-review. BMC Med Inform Decis Mak. 2008; 8:38. [PubMed: 18789150]
- Steinman MA, Dimaano L, Peterson CA, et al. Reasons for not prescribing guidelinerecommended medications to adults with heart failure. Med Care. 2013; 51(10):901–907.
 [PubMed: 23969589]
- 18. Rich MW, Chyun DA, Skolnick AH, et al. Knowledge Gaps in Cardiovascular Care of the Older Adult Population: A Scientific Statement From the American Heart Association, American College of Cardiology, and American Geriatrics Society. Circulation. 2016; 133(21):2103–2122. [PubMed: 27067230]

 Gerety MB, Chiodo LK, Kanten DN, Tuley MR, Cornell JE. Medical treatment preferences of nursing home residents: relationship to function and concordance with surrogate decision-makers. J Am Geriatr Soc. 1993; 41(9):953–960. [PubMed: 8204138]

- Harris-Kojetin L, Sengupta M, Park-Lee E. Long-term care providers and services users in the United States: Data from the National Study of Long-Term Care Providers, 2013–2014. National Center for Health Statistics. Vital Health Stat. 2016; 3(38)
- 21. Zullo AR, Lee Y, Daiello LA, et al. Beta-blocker use among United States nursing home residents after myocardial infarction: a national study. J Amer Geriatr Soc. 2016 (In Press).
- 22. Hirdes JP, Frijters DH, Teare GF. The MDS-CHESS scale: a new measure to predict mortality in institutionalized older people. J Am Geriatr Soc. 2003; 51(1):96–100. [PubMed: 12534853]
- 23. Morris JN, Fries BE, Morris SA. Scaling ADLs within the MDS. J Gerontol A Biol Sci Med Sci. 1999; 54(11):M546–553. [PubMed: 10619316]
- 24. Briesacher BA, Soumerai SB, Field TS, Fouayzi H, Gurwitz JH. Nursing home residents and enrollment in Medicare Part D. J Am Geriatr Soc. 2009; 57(10):1902–1907. [PubMed: 19702612]
- 25. Leal JR, Laupland KB. Validity of ascertainment of co-morbid illness using administrative databases: a systematic review. Clin Microbiol Infect. 2010; 16(6):715–721. [PubMed: 19614717]
- Lix LM, Yan L, Blackburn D, Hu N, Schneider-Lindner V, Teare GF. Validity of the RAI-MDS for ascertaining diabetes and comorbid conditions in long-term care facility residents. BMC Health Serv Res. 2014; 14:17. [PubMed: 24423071]
- 27. Mor V, Intrator O, Unruh MA, Cai S. Temporal and Geographic variation in the validity and internal consistency of the Nursing Home Resident Assessment Minimum Data Set 2.0. BMC Health Serv Res. 2011; 11:78. [PubMed: 21496257]
- Saczynski JS, Andrade SE, Harrold LR, et al. A systematic review of validated methods for identifying heart failure using administrative data. Pharmacoepidemiol Drug Saf. 2012; 21(Suppl 1):129–140. [PubMed: 22262599]
- 29. Gruber-Baldini AL, Zimmerman SI, Mortimore E, Magaziner J. The validity of the minimum data set in measuring the cognitive impairment of persons admitted to nursing homes. J Am Geriatr Soc. 2000; 48(12):1601–1606. [PubMed: 11129749]
- 30. Austin PC. A comparison of 12 algorithms for matching on the propensity score. Stat Med. 2014; 33(6):1057–1069. [PubMed: 24123228]
- 31. Stuart EA. Matching methods for causal inference: A review and a look forward. Stat Sci. 2010; 25(1):1–21. [PubMed: 20871802]
- 32. Austin PC. Comparing paired vs non-paired statistical methods of analyses when making inferences about absolute risk reductions in propensity-score matched samples. Stat Med. 2011; 30(11):1292–1301. [PubMed: 21337595]
- 33. Fine JP, Gray RJ. A proportional hazards model for the subdistribution of a competing risk. J Am Stat Assoc. 1999; 94(446):496–509.
- 34. Hernan MA, Hernandez-Diaz S, Robins JM. A structural approach to selection bias. Epidemiology. 2004; 15(5):615–625. [PubMed: 15308962]
- 35. Weisskopf MG, Sparrow D, Hu H, Power MC. Biased Exposure-Health Effect Estimates from Selection in Cohort Studies: Are Environmental Studies at Particular Risk? Environ Health Perspect. 2015; 123(11):1113–1122. [PubMed: 25956004]
- 36. Chrischilles EA, Schneider KM, Schroeder MC, et al. Association Between Preadmission Functional Status and Use and Effectiveness of Secondary Prevention Medications in Elderly Survivors of Acute Myocardial Infarction. J Am Geriatr Soc. 2016; 64(3):526–535. [PubMed: 26928940]
- 37. Gnjidic D, Bennett A, Le Couteur DG, et al. Ischemic heart disease, prescription of optimal medical therapy and geriatric syndromes in community-dwelling older men: A population-based study. Int J Cardiol. 2015; 192:49–55. [PubMed: 25988541]
- 38. Ko DT, Hebert PR, Coffey CS, et al. Adverse effects of beta-blocker therapy for patients with heart failure: a quantitative overview of randomized trials. Arch Intern Med. 2004; 164(13):1389–1394. [PubMed: 15249347]
- 39. Barron AJ, Zaman N, Cole GD, Wensel R, Okonko DO, Francis DP. Systematic review of genuine versus spurious side-effects of beta-blockers in heart failure using placebo control:

- recommendations for patient information. Int J Cardiol. 2013; 168(4):3572–3579. [PubMed: 23796325]
- 40. Bangalore S, Messerli FH, Cohen JD, et al. Verapamil-sustained release-based treatment strategy is equivalent to atenolol-based treatment strategy at reducing cardiovascular events in patients with prior myocardial infarction: an INternational VErapamil SR-Trandolapril (INVEST) substudy. Am Heart J. 2008; 156(2):241–247. [PubMed: 18657652]
- 41. Kamiyama T, Muratani H, Kimura Y, Fukiyama K, Omae T. Divergent effects of beta-blocker therapy on self-estimated and objectively scored activities of daily living. Hypertens Res. 1999; 22(2):85–93. [PubMed: 10487324]
- 42. Park H, Satoh H, Miki A, Urushihara H, Sawada Y. Medications associated with falls in older people: systematic review of publications from a recent 5-year period. Eur J Clin Pharmacol. 2015; 71(12):1429–1440. [PubMed: 26407688]
- 43. Foebel AD, Liperoti R, Gambassi G, et al. Prevalence and correlates of cardiovascular medication use among nursing home residents with ischemic heart disease: results from the SHELTER study. J Am Med Dir Assoc. 2014; 15(6):410–415. [PubMed: 24559641]
- 44. Fried TR, McGraw S, Agostini JV, Tinetti ME. Views of older persons with multiple morbidities on competing outcomes and clinical decision-making. J Am Geriatr Soc. 2008; 56(10):1839–1844. [PubMed: 18771453]
- 45. Fried TR, Tinetti M, Agostini J, Iannone L, Towle V. Health outcome prioritization to elicit preferences of older persons with multiple health conditions. Patient Educ Couns. 2011; 83(2): 278–282. [PubMed: 20570078]
- 46. Centers for Medicare & Medicaid Services. Nursing Home Data Compendium 2015 Edition. 2015. https://www.cms.gov/Medicare/Provider-Enrollment-and-Certification/CertificationandComplianc/Downloads/nursinghomedatacompendium_508-2015.pdf. Accessed 14 July, 2016
- 47. Rathore SS, Mehta RH, Wang Y, Radford MJ, Krumholz HM. Effects of age on the quality of care provided to older patients with acute myocardial infarction. Am J Med. 2003; 114(4):307–315. [PubMed: 12681459]

A. Time to death



B. Time to re-hospitalization

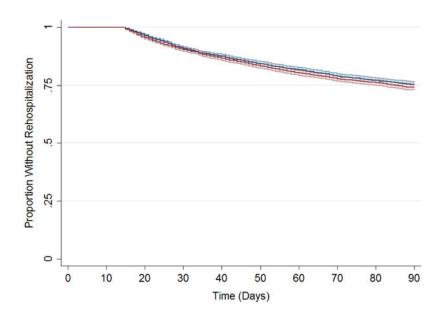


Figure 1. Association between beta-blocker use and death and re-hospitalization

Panel A shows time to death among beta blocker users and non-users. Panel B shows time to re-hospitalization in the 2 groups. There are no events in the first 14 days after hospital discharge because subjects who left the nursing home for any reason in the first 14 days after hospital discharge were excluded from analysis.

Red lines are beta blocker-users; blue lines are non-users. Shaded areas are the 95% confidence intervals around each survival curve.

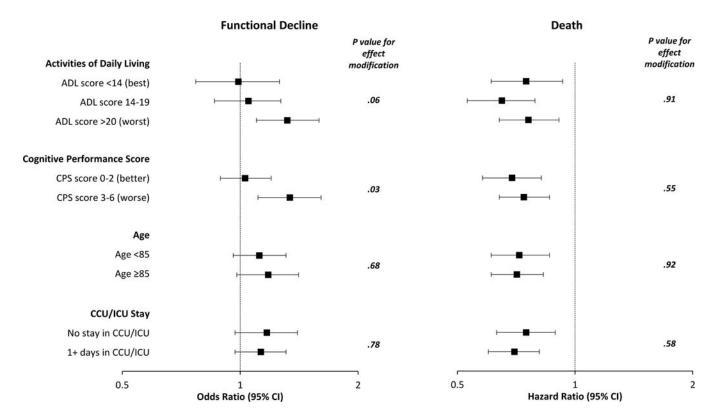


Figure 2. Impact of beta blockers on functional decline and death: subgroup analyses

P values show the significance of effect modification on the multiplicative scale. Values for additive effect modification are as follows, and are expressed as relative excess risk due to interaction (RERI). For the outcome of functional decline, RERI (95% CI) for moderate ADL dependence is 0.11 (-0.36 to 0.58), P=0.65 and for high ADL dependence is 0.66 (0.20 to 1.13), p<0.01, indicating positive additive interaction for high ADL dependence; RERI (95% CI) for worse cognitive performance score is 0.08 (-0.12 to 0.29), P=.42; RERI (95% CI) for higher age is -0.14 (-0.38 to 0.11), P=.27; RERI (95% CI) for ICU/CCU stay is -0.03 (-0.29 to 0.24), P=0.85. For the outcome of death, RERI (95% CI) for moderate ADL dependence is -0.35 (-0.70 to 0.01), P=0.05, indicating potential negative additive interaction, and for higher ADL dependence is -0.19 (-0.54 to 0.17), p=0.31; RERI (95% CI) for worse cognitive performance score is -0.15 (-0.42 to 0.12), P=.29; RERI (95% CI) for higher age is 0.00 (-0.21 to 0.22), P=.97; RERI (95% CI) for ICU/CCU stay is -0.05 (-0.26 to 0.14), P=0.60.

*ADL score <14 corresponds to independence or requiring limited assistance with ADLs; ADL score 14–19 corresponds to requiring extensive assistance; and ADL score 20 or above corresponds to extensive dependence on others to perform ADLs.

*CPS score 0–2 corresponds to normal to mildly impaired cognition including mild dementia. CPS score 3–6 corresponds to moderate or severe cognitive impairment (roughly equivalent to a Folstein Mini Mental State Exam score of 14/30 or lower).

Table 1

Characteristics of beta blocker users and non-users: before and after propensity score-based matching

	n (%)			
	Before matching (original cohort)		After matching (analytic cohort)	
Characteristic	Beta blocker users (N=8,953)	Beta blocker non-users (N= 6,767)	Beta blocker users (N=5,496)	Beta blocker non-users (N=5,496)
Age, mean (SD) years	83 (8)	84 (8)	84 (8)	84 (8)
Female sex	6,304 (70.4)	4,836 (71.5)	3,901 (71.0)	3,887 (70.7)
Race				
Caucasian	7,232 (80.8)	5,597 (82.7)	4,485 (81.6)	4,497 (81.8)
African-American	1,158 (12.9)	756 (11.2)	644 (11.7)	646 (11.8)
Other	563 (6.3)	414 (6.1)	367 (6.7)	353 (6.4)
Chronic conditions				
Diabetes	2,855 (31.9)	1,942 (28.7)	1,567 (28.5)	1,582 (28.8)
Heart failure	4,534 (50.6)	3,051 (45.1)	2,554 (46.7)	2,562 (46.6)
COPD	2,218 (24.8)	1,942 (28.7)	1,498 (27.3)	1,504 (27.4)
Depression	1,101 (12.3)	838 (12.4)	660 (12.0)	622 (11.3)
Elixhauser comorbidity score, median, (IQR)	3 (2–4)	3 (2–4)	3 (2–4)	3 (2–4)
ADL status prior to hospitalization *				
Independent to limited assistance required	3,054 (34.1)	2,347 (34.7)	1,834 (33.4)	1,866 (34.0)
Extensive assistance required	3,050 (34.1)	2,188 (32.3)	1,801 (32.8)	1,778 (32.4)
Extensive dependency	2,849 (31.8)	2,232 (33.0)	1,861 (33.9)	1,852 (33.7)
Cognitive status prior to hospitalization *				
Intact or borderline intact	2,790 (31.2)	1,961 (29.0)	1,580 (28.8)	1,585 (28.8)
Mild to moderate dementia	4,609 (51.5)	3,505 (51.8)	3,294 (59.9)	3,305 (60.1)
Moderately severe to very severe dementia	1,554 (17.4)	1,301 (19.2)	622 (11.3)	606 (11.0)
CHESS score prior to hospitalization, mean (SD)	0.6 (0.8)	0.6 (0.8)	0.6 (0.8)	0.6 (0.8)
Symptoms, geriatric prior to hospitalization	<u> </u>			-
Dizziness, vertigo, or syncope	103 (1.2)	82 (1.2)	54 (1.0)	55 (1.0)
Falls	1,843 (20.6)	1,515 (22.4)	1,193 (21.7)	1,187 (21.6)
Dyspnea	621 (6.9)	645 (9.5)	461 (8.4)	455 (8.3)
Number of medications prior to hospitalization	11 (8–15)	12 (9–15)	11 (8–15)	12 (8–15)
Medication use prior to hospitalization *				
Statins	2,584 (28.9)	1,944 (28.8)	1,559 (28.4)	1,580 (28.8)
Antiplatelets	1,453 (16.2)	1,165 (17.2)	914 (16.6)	916 (16.7)

Steinman et al.

n (%) Before matching (original cohort) After matching (analytic cohort) Beta blocker Beta blocker users Beta blocker users Beta blocker non-users Characteristic non-users (N=8,953)(N=5,496)(N=5,496)(N=6,767)Warfarin 992 (11.1) 938 (13.9) 707 (12.9) 723 (13.2) Psychotropics 5,400 (60.3) 4,367 (64.5) 3,547 (64.5) 3,482 (63.4) 6 (4–9) 6 (4–9) 6 (4–9) 6 (4-9) Length of hospital stay for AMI, median (IQR) Number of days in ICU / CCU None 3,385 (37.8) 3,277 (48.4) 2,374 (43.2) 2,361 (43.0) 1 to 2 2,425 (27.1) 1,589 (23.5) 1,376 (25.0) 1,396 (25.4) 1,739 (31.6) 3,143 (35.1) 1,901 (28.1) 1,746 (31.8) 3 or more Nursing home care pathway after hospitalization Skilled nursing facility (SNF) benefit 6,714 (75.0) 4,569 (67.5) 3,894 (70.9) 3,867 (70.4) 2,239 (25.0) 2,198 (32.5) 1,602 (29.2) 1,629 (29.6) Long-term care **Nursing Home Facility Characteristics** Ownership For profit 6,488 (72.5) 4,909 (72.5) 4,019 (73.1) 3,991 (72.6) 1,983 (22.2) 1,451 (21.4) 1,151 (20.9) 1,195 (21.7) Non-profit Government 482 (5.4) 407 (6.0) 326 (5.9) 310 (5.6) Size <100 beds 1,375 (15.4) 871 (12.9) 1,535 (27.9) 1,521 (27.7) 100-200 beds 5,258 (58.7) 3,951 (58.4) 3,206 (58.3) 3,220 (58.6) >200 beds 2,320 (25.9) 1,945 (28.7) 755 (13.7) 755 (13.7) Quality indicators 2.8 (0-6.5) % of residents restrained, median (IQR) 3.1 (0.4-6.9) 2.9 (0.4-6.6) 3.0 (0.3-6.7) 0.74(1.1) No. of quality-of-life deficiencies, mean 0.73(1.1)0.73 (1.0) 0.75 (1.1) % of residents with pressure sores, mean 7.2 (4.5) 7.0 (4.3) 7.1 (4.6) 7.0 (4.3) (SD) Staffing Direct care hours/resident/day, mean (SD) 3.4 (0.8) 3.4 (0.8) 3.4 (0.7) 3.4 (0.8)

Page 15

^{*}ADL status was measured by the Morris 28-point ADL score, and categorized as 0–14 (independent to limited assistance required), 15–19 (extensive assistance required), and 20 (extensive dependency). Cognitive status was measured by Cognitive Performance Scale (CPS) and trichotomized as 0–1 (intact to borderline intact), 2–3 (mild to moderate dementia), and 4–6 (moderately-severe to very severe dementia). Psychotropics include antidepressants, antipsychotics, antianxiety medications, and sedative/hypnotics.

Steinman et al. Page 16

Table 2

Impact of beta blockers on functional decline, death, and rehospitalization

Outcome	Odds Ratio / Hazard Ratio for beta blocker users vs. non-users* (95% CI)	Number needed to treat (NNT) / number needed to harm (NNH) (95% CI)
Functional decline	1.14 (1.02 – 1.28)	NNH 52 (32 – 141)
Death	0.74 (0.67 – 0.83)	NNT 26 (19 – 39)
Re-hospitalization	1.06 (0.98 – 1.14)	NNH 82 (NNH 250 to ∞ to NNT 36) [†]

 $NNH = number \ needed \ to \ harm; \ NNT = number \ needed \ to \ treat. \ NNH \ and \ NNT \ calculated \ as \ 1/(control \ event \ rate - intervention \ event \ rate).$

 $^{^{*}}$ Odds ratio for functional decline; hazard ratio for death and re-hospitalization