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Quantitative Precipitation Estimation of Extremes in CONUS with Radar Data

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Key Points:

- The 5-year return values of extreme rainfall vary only minimally on spatial scales smaller than 100 kilometers.
- The gauge-only Mountain Mapper algorithm, used to spatially interpolate rain-gauge data in the operational Next-Generation Radar (NEXRAD) data products where radar data is unavailable, underestimates 5-year return values far from the locations of rain gauges.
- The Risser et al 2019 algorithm for spatially interpolating extreme value statistics between rain gauges represents these statistics more accurately than Mountain Mapper over the majority of CONUS.

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Abstract

Constructing an accurate, continental, in-situ-based, kilometer-scale, long-term record of the precipitation field and its spatiotemporal changes remains a significant challenge. Here we determine the extreme-value behavior of the NEXRAD Stage IV radar-based quantitative precipitation estimate (QPE). We find that the climatology of 5-year daily return values in CONUS East of the Rocky Mountains shows only slight variability on spatial scales smaller than ~ 100 km. In light of this finding, we test whether rain-gauge-only daily precipitation datasets can produce accurate extreme-value behavior at spatial scales finer than the spacing between gauges. We find that the 5-year daily return values are accurate at locations far from rain gauges only if the interpolation between gauges is carried out appropriately for extremes. Precipitation statistics derived from in-situ rain gauge data are therefore of sufficient spatial resolution to faithfully capture daily extremes over much of the eastern United States.

Plain Language Summary

Accurate measurement of the amount of precipitation that falls within a given region and time period is crucial for environmental modeling, climate change research, and resource and risk management. For all of those applications, it is desirable to understand not only how much precipitation falls on average, but also how much precipitation falls during an extreme event, such as a severe storm. Using data from weather radar, we show that certain statistical properties of extreme rainfall are highly correlated on spatial scales up to 100 kilometers over the eastern United States. This means that rain gauge networks, which have typical inter-gauge spacings of roughly 30 kilometers over the eastern United States, are dense enough to accurately measure these statistical properties. However, it's imperative to interpolate between the rain gauge measurements in a way that explicitly captures extremes if the application of interest requires capturing extremes accurately. Our research represents a step toward constructing an accurate, continental-scale, long-term, high-resolution precipitation dataset.

1 Introduction

Accurate measurement of the amount of precipitation that falls within a given region and time period is crucial for environmental modeling (e.g. Jones et al., 2001; Parra

47 et al., 2004; Abatzoglou, 2013), climate change research (e.g. Groisman et al., 1999; Alexan-
48 der et al., 2006; Schär et al., 2016), and resource and risk management (e.g. Rosenzweig
49 et al., 2002; Schumann, 2011; Vogel et al., 2019). All of these applications require un-
50 derstanding not only the seasonal-mean or annual-mean precipitation but also the ex-
51 treme tail of the daily or sub-daily precipitation distribution. Precise measurements of
52 the total rainfall over a specific area at scales of tens of kilometers or less, such as a city
53 or watershed, are often also needed.

54 Estimating the true spatiotemporal distribution of precipitation from observational
55 data is known as quantitative precipitation estimation (QPE), and is currently obtained
56 from three main data sources: satellite, rain gauges, and ground-based radar. Each source
57 provides unique advantages subject to specific limitations. Satellite observations provide
58 spatially continuous measurements, but are subject to severe uncertainty because pre-
59 cipitation must be inferred from cloud top height or temperature as derived from microwave
60 and/or infrared spectra (Iguchi et al., 2009; Tapiador et al., 2012). This uncertainty gen-
61 erally leads to overestimation of extreme precipitation events relative to gauge- or radar-
62 based estimates (AghaKouchak et al., 2011; Mehran et al., 2014). Over the contiguous
63 United States (CONUS), where the density of both rain gauges and radar stations is high,
64 satellite-based QPE products tend to perform compared unfavorably to other estimates.
65 For example, Timmermans et al. (2019) found significant biases in the representation of
66 daily precipitation extremes from satellite-based gridded QPEs compared with rain gauge
67 estimates. Satellite products are therefore not considered further here.

68 Rain gauges provide the most accurate and temporally continuous point measure-
69 ments of precipitation despite errors from undercatch, variance in management quality,
70 and changes in location or equipment (see Tapiador et al., 2012, for a recent review). They
71 also provide the longest time record of any precipitation measurement by far. However,
72 gauges yield point measurements only, and one must interpolate spatially between them
73 to estimate precipitation over an area. Ground-based radar observations provide very
74 high native spatial and temporal resolution. Each Weather Surveillance Radar 88 Doppler
75 Radar (WSR-88D) radar stations in CONUS (NOAA, 2006) completes a full scan of the
76 sky every ~ 10 minutes, and the station's preprocessing algorithm bins the scan into 1
77 km range by 1° azimuth sections, amounting to sub-hourly precipitation estimates on
78 a $\lesssim 4$ km grid (Fulton et al., 1998). However, the relationship between radar reflectiv-
79 ity and precipitation rate is degenerate and differs for different types of storms, and there-

80 fore it must thus be determined empirically via comparison with rain gauge data (Fulton
81 et al., 1998; Young et al., 2000). Multisensor estimates combine the high spatial and tem-
82 poral coverage of the radar data with the high fidelity of the gauge data and hence rep-
83 resent the state-of-the-art in operational QPE. However, experiments to validate these
84 QPEs often take place over small regions and on timescales shorter than a few months
85 (e.g., Willie et al., 2017; Spies et al., 2018; B.-C. Seo et al., 2018).

86 A variety of studies make use of gauge-based QPE and solve the aforementioned
87 spatial interpolation problem in various ways, sometimes making use of elevation cor-
88 rections or models of climatology (e.g. Daly et al., 1994, 2015; Schaake et al., 2004; Sheri-
89 dan et al., 2010; Livneh et al., 2013). We refer to this class of gauge-based analysis of
90 extremes as “grid-then-fit” techniques because they interpolate at the native temporal
91 scale (e.g., daily) and then calculate statistical properties of the interpolated data. These
92 approaches tend to underestimate extreme precipitation, especially at small (0.25°) scales
93 (Sun & Barros, 2010; Gervais et al., 2014; Behnke et al., 2016). To rectify this issue, Risser
94 et al. (2019) have developed a statistical “fit-then-grid” technique in which Generalized
95 Extreme-Value (GEV) statistics (see Coles et al., 2001) are calculated at individual rain
96 gauges, the GEV parameters are spatially interpolated, and then the gridded GEV dis-
97 tributions are reconstructed from these interpolated parameters. This method implic-
98 itly assumes that the parameters of the GEV distribution vary smoothly in space such
99 that high-quality inference about extremes can be made in between stations. The op-
100 timal gauge interpolation technique depends on both the grid resolution and the appli-
101 cation of interest (Chen & Knutson, 2008; Gervais et al., 2014), and best practices for
102 interpolating to smaller scales than the inter-gauge spacing have not been established.

103 This study seeks to determine whether the extreme statistics of daily precipitation
104 vary smoothly between rain gauges over the CONUS, testing the assumption of Risser
105 et al. (2019) over that domain, and to evaluate the accuracy of this novel fit-then-grid
106 technique as compared with standard grid-then-fit algorithms. To these ends, we con-
107 sider the GEV statistics of a dataset at very high (4 km) spatial resolution, namely the
108 NEXRAD Stage IV daily dataset (Fulton et al., 1998; D.-J. Seo & Breidenbach, 2002;
109 Lin & Mitchell, 2005; Lin, 2011), a radar-based multisensor QPE, from 2002-2019. Stage IV
110 is available at hourly, six-hourly, and daily frequencies. We focus on daily maxima here
111 to facilitate comparison with the GHCN-D network, the most extensive network of rain
112 gauges in the CONUS, and to test the results of Risser et al. (2019) directly. Stage IV

113 has been evaluated extensively in the literature, including via the use of percentile-based
114 metrics to capture extreme value behavior (Prat & Nelson, 2015; Nelson et al., 2016).
115 However, percentile-based metrics have been shown to produce different results depend-
116 ing on the specific metric used (Schär et al., 2016). McGraw et al. (2019) previously em-
117 ployed GEV statistics in a comparison between rain gauges and Stage IV data at hourly,
118 3-hourly, 6-hourly, and daily frequencies. However, they only reported those statistics
119 at the locations of ~ 500 rain gauges and did not consider spatial variability in their GEV
120 fits. Our paper is the first (to our knowledge) to publish GEV statistics at every grid
121 cell in Stage IV.

122 We describe our data processing and GEV fitting in Section 2. In Section 3, we use
123 this new data product to explore whether the high spatial resolution of this QPE pro-
124 vides new information on the climatology of extremes at finer spatial scales than acces-
125 sible using gauge-only estimates. In Section 4, our product is compared with the Risser
126 et al. (2019, hereafter R19) gauge-only interpolation technique, as well as with the Moun-
127 tain Mapper algorithm (Schaake et al., 2004), a more conventional gridded QPE that
128 incorporates the Parameter-elevation Relationships on Independent Slopes Model (PRISM)
129 climatological model (Daly et al., 2015) and is widely used for operational weather anal-
130 yses. Specifically, Mountain Mapper is the official rainfall product distributed by the California-
131 Nevada, Colorado Basin, and Northwest River Forecast Centers of the National Oceanog-
132 raphic and Atmospheric Administration (NOAA). We contextualize our findings within
133 existing literature on the spatial scales of extremes in Section 5, then summarize our work
134 in Section 6.

135 **2 Data Processing**

136 We computed and made use of three distinct extreme-value datasets in this paper;
137 these are summarized below.

138 *Stage IV GEV:* We downloaded the 4-km-resolution daily NEXRAD Stage IV grid-
139 ded multisensor QPE for every day between 1 January 2002 and 31 December 2019, to-
140 taling 6573 days (18 years) on a 881×1121 grid. We compared the Stage IV daily mea-
141 surements to Global Historical Climatology Network Daily (GHCN-D) rain gauge data
142 (Menne et al., 2012) in grid cells that contained a GHCN station. These validation steps
143 are outlined in Supporting Information S1. Stage IV was found to agree very well with

144 GHCN-D in both means and extremes in grid cells that contained a gauge, meaning that
145 the normalization of the radar data to nearby rain gauges in the Stage IV processing pipeline
146 preserves information about extremes.

147 As noted by other authors (Prat & Nelson, 2015; Nelson et al., 2016), Stage IV is
148 a fundamentally heterogeneous dataset. The product from the three western NOAA River
149 Forecast Centers (California-Nevada, Northwest, and Colorado Basin, hereafter “West-
150 ern RFCs”) differs substantially from that of the nine other RFCs comprising the CONUS
151 (hereafter “Eastern RFCs”). Specifically, the Western RFCs produce their QPE using
152 the gauge-only Mountain Mapper technique discussed below and do not incorporate radar
153 data at all, while the Eastern RFCs use the radar-inclusive procedures outlined in Fulton
154 et al. (1998) and Lin and Mitchell (2005). Data from the Western RFCs are therefore
155 not actually made using a multisensor technique, and so are for the most part not con-
156 sidered further in this paper.

157 In each grid cell of Stage IV, we extracted seasonal maximum precipitation amounts
158 for each season (DJF, MAM, JJA, and SON), and then fit the GEV distribution to the
159 18 seasonal maxima over our period of record. This is a fairly short period of record over
160 which to apply GEV statistics; however, in this work we draw our conclusions from the
161 5-year return values only, which are well sampled by 18 years of data. To assess whether
162 our GEV fits provided an adequate representation of the data, we performed a 2-sided
163 Kolmogorov-Smirnov (K-S) test to quantify the likelihood that the observed seasonal max-
164 ima were drawn from the GEV distribution. We found a p -value of < 0.05 in at least
165 94% of grid cells in the Eastern RFCs in all seasons, meaning that the data were plau-
166 sibly drawn from the GEV distribution. We used the GEV fit to generate 5-year return
167 value estimates in each grid cell for each season. Following R19, we assessed the errors
168 in our fit parameters using a bootstrap resampling technique: the seasonal maxima at
169 each grid cell were resampled with replacement and then re-fit 250 times, and the stan-
170 dard deviation of the fit parameters in those 250 fits were used to define the $1-\sigma$ error
171 on the parameters.

172 *R19 GEV:* We used the same extreme-value dataset as R19, which is based on GHCN-D
173 rain gauge measurements, but extracted 5-year return values instead of 20-year return
174 values as in that paper. To create a mean climatology from the R19 analysis, the exact
175 same procedures described in R19 were applied to GHCN measurements of seasonal av-

176 erage daily precipitation (instead of seasonal maximum daily precipitation), and an ordi-
 177 nary least-squares fit was applied at each station such that the spatially-interpolated
 178 parameters were the mean and variance instead of the GEV parameters. We refer to this
 179 mean climatology hereafter as the “R19 mean”.

180 *Mountain Mapper:* The Mountain Mapper dataset (Schaake et al., 2004) is a gauge-
 181 only gridded precipitation product at 4-km resolution that is widely used in operational
 182 weather analysis. It interpolates from rain gauges to the 4-km grid using an inverse-square
 183 weighting scheme, incorporating also the PRISM climatological model (Daly et al., 2015).
 184 However, the official Mountain Mapper product is not archived at NOAA. Instead, we
 185 have created several versions of the dataset using an identical procedure to Schaake et
 186 al. (2004). The creation of our own versions of the dataset is beneficial for direct com-
 187 parison to R19 and Stage IV for three reasons. First, we have used the same rain gauge
 188 network as Risser, namely stations from the GHCN-D network in CONUS with at least
 189 66.7% nonmissing values over our time period (8097 stations). Second, we can interpo-
 190 late the gauge network onto any grid we choose. Finally, we can force the long-term mean
 191 of Mountain Mapper to equal any chosen climatology; versions of Mountain Mapper con-
 192 strained to the R19 mean and to the mean of our 18-year slice of Stage IV are employed
 193 in this paper. We compute GEV statistics and their errors at each grid cell in the same
 194 way as for Stage IV. Our computations using the Mountain Mapper procedure are dis-
 195 cussed further in Supporting Information S2.

196 **3 Spatial Scales of Extremes**

197 If the hypothesis that the extreme statistics of daily precipitation vary smoothly
 198 between rain gauges over the CONUS is true, then a spatial power spectrum of the 5-year
 199 return value map should show little power at $\lesssim 50$ -km scales. We thus used a wavelet de-
 200 composition to compute a spatial power spectrum of the 5-year return values in the Stage IV
 201 dataset using a 2-D continuous wavelet transform. Following the procedure outlined in
 202 Torrence and Compo (1998), we used Morlet wavelets with non-dimensional frequency
 203 $\omega_0 = 6$, and 40 widths equally spaced on a log scale from 4 km to ~ 2000 km.¹ Maps
 204 of the power on various representative spatial scales are shown for JJA and DJF in Fig-

¹ Using a Morlet wavelet in this way is mathematically identical to a “short-time” Fourier transform using a Gaussian window function

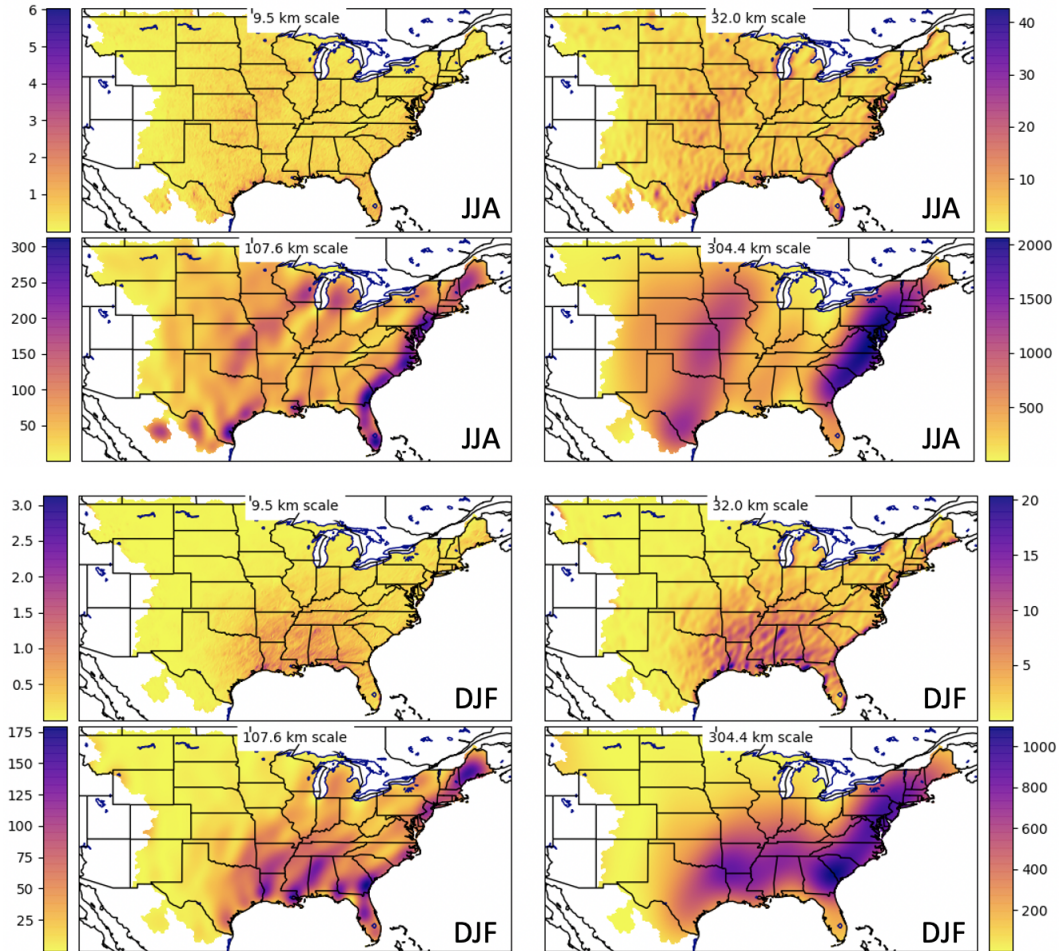


Figure 1. Wavelet decomposition of Stage IV 5-year return value map for **(Top:)** DJF and **(Bottom:)** JJA at four representative spatial scales. The colormap denotes power spectral density in arbitrary units.

205 re 1 to help visualize the wavelet decomposition. The power spectrum of our 5-year re-
 206 turn value map is presented in Figure 2. (The same maps and power spectra are shown
 207 for the MAM and SON seasons in the Supporting Information.) The spatial scales s plot-
 208 ted on the x-axis are nearly equal to the Fourier wavelength λ for this choice of wavelet
 209 (formally $\lambda = 1.02s$ following Torrence & Compo, 1998), and should be interpreted in
 210 the same manner as a Fourier wavelength, namely as the combined length scale of a pos-
 211 itive and negative fluctuation about the mean. Note that substantial edge effects obscure
 212 any useful information at scales larger than $s = 1000$ km, so these are not plotted.

To aid in understanding the implications of Figure 2, we have overplotted the power spectrum of a test dataset that contains pure white noise within the entire domain (the

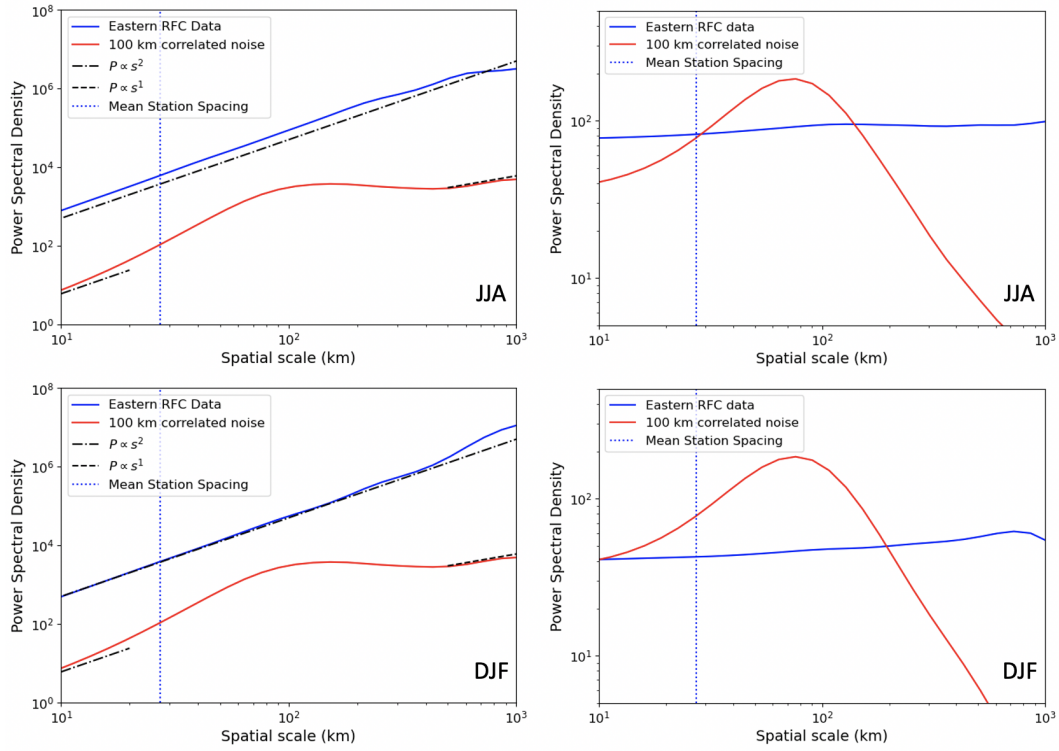


Figure 2. Seasonal wavelet power spectrum of the Stage IV 5-year return value map for **Top:** JJA and **Bottom:** DJF over the eastern RFCs only (blue lines). Vertical lines show the mean spacing of 27 km between GHCN stations in the eastern RFC domain. The power spectrum of white noise correlated at the 100 km scale is also shown (red line). The left panels show the raw power spectra; the right panels show the spectra after being divided by the $P = s^2$ line.

eastern RFCs), correlated at the 25-pixel (100-km) scale; that is, we made a map of pure Gaussian noise then oversampled it by a factor of 25. A log-linear power spectrum takes the form

$$S_\nu(f) = cf^{-\beta} \quad (1)$$

213 where f is the spatial frequency and β is the spectral scaling. The correlated noise test
 214 spectrum can be interpreted as transitioning between $\beta = 2$ at the smallest length scales,
 215 where the map is highly autocorrelated, and $\beta = 1$ at the largest length scales, where
 216 the map is completely uncorrelated and looks like pure white noise. Note that white noise
 217 is not spectrally flat on a log-linear scale, but instead follows a $\beta = 1$ scaling. In be-
 218 tween these two regimes is stored all the information content in the map, and as such,
 219 the power spectrum is strongly peaked at 100 km length scales. This can be seen most
 220 clearly after the spectra have been divided by the $P = s^2$ line in the right panel of Fig-
 221 ure 2. The 5-year return value maps in both JJA and DJF show similar behavior to the
 222 correlated noise map, with strong autocorrelation ($P \propto s^2$) at small spatial scales but
 223 with a broader, less prominent spike in power that begins near 200-km scales and con-
 224 tinues out to 800-km scales. This means that 5-year return values are strongly autocor-
 225 related at $s < 200$ km, confirming the hypothesis that extreme statistics of daily pre-
 226 cipitation vary smoothly between rain gauges over the CONUS. The power spectral den-
 227 sity is maximized at very large length scales of ~ 800 km. The strong autocorrelation at
 228 small scales is present in each of the four major Köppen-Geiger climate classes within
 229 the eastern CONUS, as shown in the Supplementary Material.

230 4 QPE Product Comparison

231 Section 3 validated the implicit assumption of the R19 fit-then-grid technique that
 232 extreme statistics of daily rainfall vary smoothly between rain gauges. It is *a priori* un-
 233 clear, though, whether this fit-then-grid algorithm is actually more accurate than a stan-
 234 dard grid-then-fit algorithm when applied to an identical set of rain gauges and given
 235 an identical mean climatology. We set up this test by comparing the extreme-value be-
 236 havior between the Mountain Mapper and R19 datasets. To ensure a direct comparison,
 237 we used a version of Mountain Mapper constructed such that its long-term seasonal mean
 238 (\bar{Y} in Supporting Information S2) was equal to the R19 mean. As Figure 3 shows, the
 239 return values are substantially different between the two datasets, with Mountain Map-
 240 per underestimating R19 by greater than 10% over much of the CONUS in both DJF

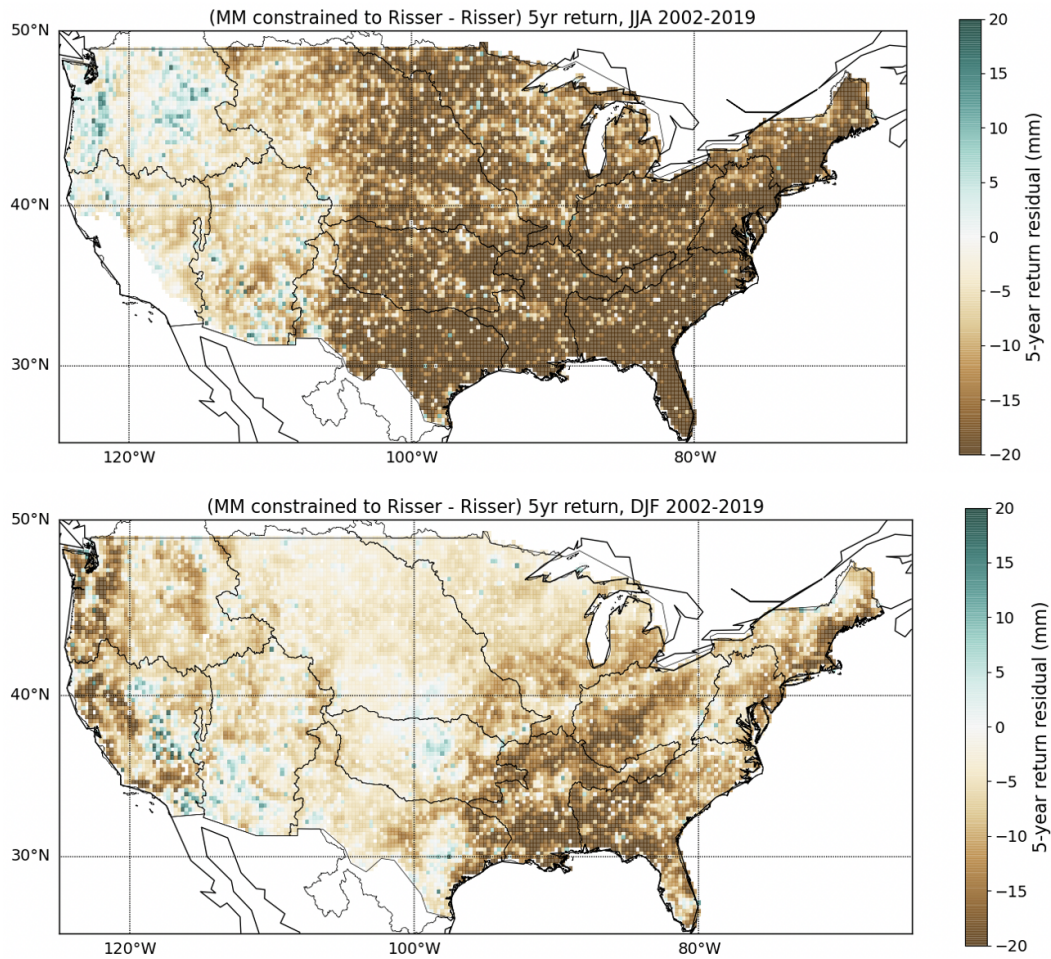


Figure 3. Difference between 5-year return value from R19 and our Mountain Mapper implementation constrained to the R19 mean climatology for **Top:** JJA and **Bottom:** DJF. The extremes in Mountain Mapper are lower in magnitude than in R19 over the majority of CONUS in both seasons.

241 and JJA. An assessment of the statistical significance of this difference is given in Sup-
 242 porting Information S3.

243 We next evaluated the R19 and Mountain Mapper 5-year return values against Stage IV.
 244 This comparison is somewhat difficult to probe directly because the long-term means of
 245 R19 and Stage IV are not strictly equal, so differences in extremes may be partially caused
 246 by differences in the long-term means of those datasets. To get around this, we computed
 247 return values from a version of Mountain Mapper that is forced to equal the long-term
 248 means of Stage IV, and compared both this Mountain Mapper version and Stage IV it-
 249 self with R19 (see Figure 4). In so doing, any differences are isolated to the treatment

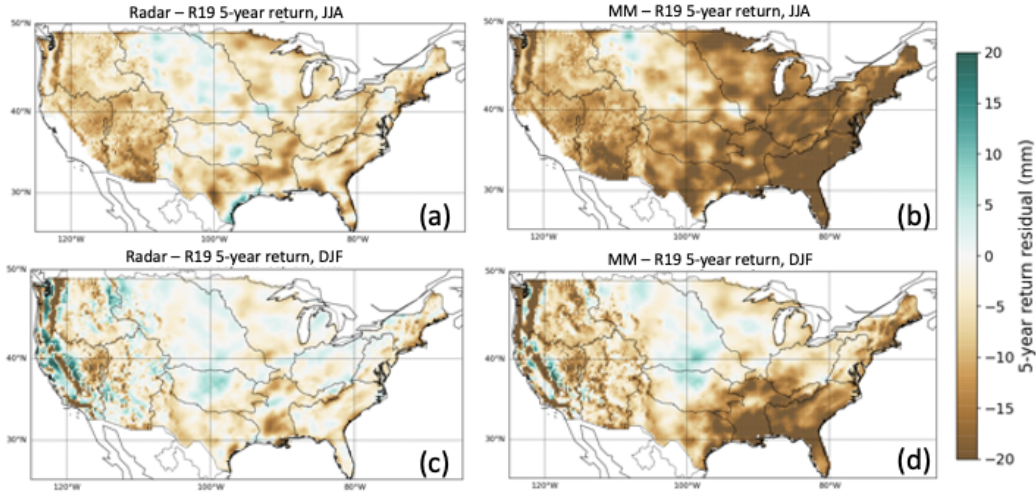


Figure 4. Difference between 5-year return value from (a) R19 and Stage IV in JJA, (b) R19 and Mountain Mapper in JJA, (c) R19 and Stage IV in DJF, and (d) R19 and Mountain Mapper in DJF. Here the Mountain Mapper datasets have been constrained to the Stage IV mean climatology. The R19 extremes agree more closely with Stage IV than Mountain Mapper in both JJA and DJF (i.e., the discrepancies are smaller in panels a and c than in panels b and d), validating the ability of the R19 technique to interpolate extremes to smaller spatial scales.

250 of extremes. This Mountain Mapper version is found to underestimate extremes rela-
 251 tive to Stage IV over large portions of the Eastern RFCs, whereas R19 agrees more closely.

252 The difference between Mountain Mapper and R19 is attributable to the grid-then-
 253 fit approach taken by Mountain Mapper: using an inverse-square weighting scheme to
 254 interpolate between grid points makes it unlikely for extremes to occur at grid points far
 255 from any one rain gauge. This hypothesis is confirmed by considering the difference be-
 256 tween Mountain Mapper and Stage IV as a function of distance from the nearest rain
 257 gauge over the eastern RFCs. The 5-year return values from Mountain Mapper agree
 258 well with Stage IV at distances $\lesssim 10$ km from the nearest gauge, but begin to underes-
 259 timate Stage IV at larger distances in both DJF and JJA (Figure 5). It is important to
 260 note that the spatial averaging inherent in the Mountain Mapper technique is not a de-
 261 ficiency *per se*, and is in fact the appropriate way to measure the spatial average of ex-
 262 tremes over a large grid box (Gervais et al., 2014) for comparison to climate models at
 263 ~ 100 km resolution. However, we have shown that the Risser technique provides a more

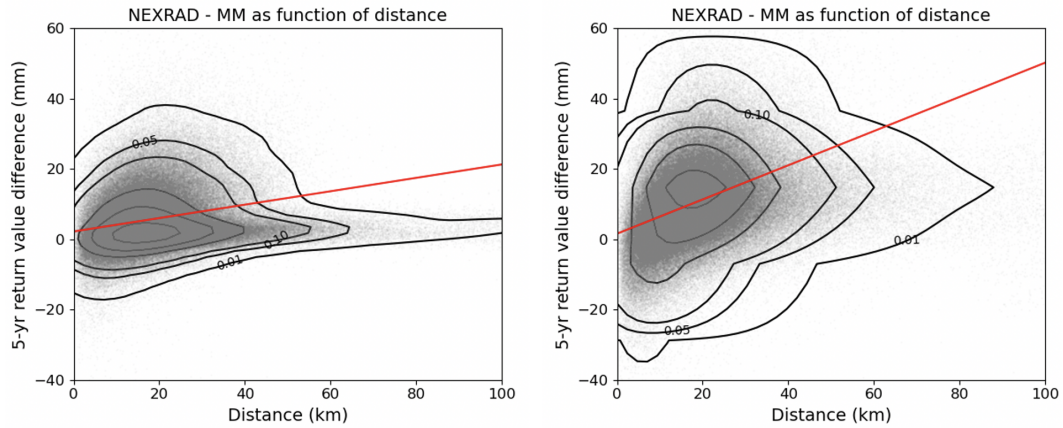


Figure 5. Five-year return value difference between Stage IV and Mountain Mapper as a function of distance from the nearest rain gauge used in Mountain Mapper for **Left:** DJF and **Right:** JJA over the eastern CONUS. The small grey points denote individual grid cells; the contours describe the cumulative density of points. The red line shows a least-squares linear fit to the grey points. The return values agree in grid cells near rain gauges, but Mountain Mapper begins to underestimate Stage IV as the distance from a gauge is increased.

264 accurate estimate of rainfall extremes at ~ 25 km scales, assuming the radar-aided Stage IV
 265 dataset to be a “ground truth”.

266 5 Discussion

267 The long correlation lengths of 5-year return values in the eastern CONUS derived
 268 in Section 3 are perhaps unsurprising in the context of the dynamical systems that pro-
 269 duce extreme precipitation in that region. In the central and eastern United States, ex-
 270 treme precipitation is most often associated with one of three categories of storm: mesoscale
 271 convective systems (MCSs), landfalling tropical cyclones (TCs), and synoptic forcing events
 272 (i.e., extratropical cyclones). MCSs are organized groups of thunderstorms that produce
 273 distinct circulations at scales longer than 100 km and persist over timescales of 3 hours
 274 to 1 day (Parker & Johnson, 2000; Houze, 2004; Feng et al., 2019). These systems ac-
 275 count for over half of extreme rainfall events at 24-h duration in the warm season in these
 276 regions (Schumacher & Johnson, 2006; Stevenson & Schumacher, 2014). Landfalling tropi-
 277 cal cyclones (TCs) also contribute substantially in the summer and fall in the eastern
 278 and southeastern United States, especially in coastal regions (Shepherd et al., 2007; Knight
 279 & Davis, 2009; Miniussi et al., 2020). In the cool season, extreme precipitation results

280 primarily from strong synoptic forcing events (Maddox et al., 1979; Schumacher & John-
281 son, 2006; Stevenson & Schumacher, 2014). Although synoptic forcing events occur with
282 nearly unchanging frequency throughout the year, MCSs and TCs are much more sel-
283 dom present in DJF (Stevenson & Schumacher, 2014), and extremes in DJF tend to be
284 of lower magnitude than in JJA (Maddox et al., 1979; Stevenson & Schumacher, 2014).
285 Individual storms of these types all tend to produce heavy precipitation over length scales
286 of 100 km or more. The long correlation lengths in the statistics of extreme precipita-
287 tion presented here can thus be partially attributed to the long correlation lengths of in-
288 dividual events. This interpretation is in good agreement with Touma et al. (2018), who
289 used indicator semivariograms to assess the correlation scales of 90th percentile rainfall
290 days over CONUS. Although that analysis was split into more climatological regions, their
291 North, Northeast, South, and Southeast regions all display DJF length scales within 1σ
292 of 300 km.

293 Previous studies (e.g. Kursinski & Mullen, 2008) have shown, perhaps in appar-
294 ent tension with the above, that individual extreme storms can be highly localized in both
295 space and time, with heavy precipitation falling over spatial scales of ~ 50 km or less. Im-
296 portantly, though, the spatial statistics of rainfall depend strongly on the time cadence
297 considered. In a case study of the Cévennes-Vivarais region of France, Lebel et al. (1987)
298 and Kirstetter et al. (2010) showed that the decorrelation distance of rainfall amounts
299 increases with lengthening temporal scale from hourly to daily cadence. We wish to stress
300 that our results are only valid at the daily cadence we considered; the assumption of smoothly-
301 varying GEV statistics between rain gauges, and therefore the R19 technique, may not
302 be justified at sub-daily cadences.

303 It is interesting to consider Figure 2 in terms of the fractal properties of rainfall
304 explored by Lovejoy (1982) and Lovejoy and Mandelbrot (1985). Those authors describe
305 the spatial structure cloud and rain areas according to $N \sim L^{-D}$, where N is the ex-
306 tent to which a fractal fills space as measured at scale L . Bies et al. (2016) explored the
307 relationship between the fractal and power spectrum interpretations of scaling fields, find-
308 ing that the fractal dimension D and β in Equation 1 are related according to $D = 1 + (4 - \beta)/2$
309 for a 2-dimensional field. (In the terminology of Bies et al. (2016), we measured here a
310 “surface β ” and the cited papers use a “coastal edge D ”.) We find $\beta \approx 2$ (which leads
311 to $D \approx 2$) for this process up to scales of a few hundred km; that is, the extreme pre-
312 cipitation field is 2-dimensional. This is another way of interpreting the high level of au-

313 tocorrelation on small scales. $D = 2$ is larger than the fractal dimension found by Lovejoy
314 and Mandelbrot (1985) for cloud and rain areas, meaning that the spatial scales over which
315 extreme statistics vary in Stage IV are larger than the spatial scales of individual pre-
316 cipitation events.

317 **6 Conclusions**

318 We have tested the assumption that the climatology of extremes varies only min-
319 imally at length scales smaller than the average inter-rain-gauge spacing of ~ 30 km in
320 the eastern CONUS. We find that this assumption is valid: 5-year daily return values
321 are strongly autocorrelated at scales up to at least 100 km in both DJF and JJA. We
322 also find that the fit-then-grid algorithm of R19 substantially improves the fidelity of daily
323 extreme statistics compared with the grid-then-fit Mountain Mapper technique. On both
324 4-km and 25-km scales, the grid-then-fit Mountain Mapper technique underestimates ex-
325 tremes relative to the more spatially complete multi-sensor Stage IV QPE in the east-
326 ern United States, whereas the Risser et al. (2019) technique measures extremes more
327 accurately than Mountain Mapper at 25 km scales. Taken together, these findings show
328 that rain gauge observations are sufficient to capture the large majority of the extreme-
329 value information in the climatology of the true rain field, but only if interpolated ap-
330 propriately for the application of interest. This paper improves confidence that appropriately-
331 constructed gauge-only gridded products provide an accurate historical record of daily
332 extreme statistics beyond the years in which radar data are available, an important step
333 toward creating an accurate, continental-scale, in-situ-based, long-term precipitation record
334 for use in hydrological modeling, resource management, and climate change studies. As
335 the resolution of global circulation models continues to increase into the future, QPEs
336 will be required at finer and finer scales, and standard gauge-interpolation techniques
337 will fail to accurately represent precipitation within these grid boxes. The human im-
338 pacts of extreme events are felt at human scales, e.g. homes (10m), farms (1 km), and
339 watersheds (10 km). Our work moves toward casting measurements of extremes into a
340 risk framework at those scales.

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365 **Data Availability Statement**

366 Datasets for this research are available at the following locations:

- 367 • The NEXRAD Stage IV data product is described in Lin (2011) and can be ac-
368 cessed online at <https://data.eol.ucar.edu/dataset/21.093>
- 369 • The GHCN-Daily data is described in Menne et al. (2012) and can be accessed
370 online at [https://www.ncei.noaa.gov/access/metadata/landing-page/bin/
371 iso?id=gov.noaa.ncdc:C00861](https://www.ncei.noaa.gov/access/metadata/landing-page/bin/iso?id=gov.noaa.ncdc:C00861).
- 372 • The CNRFC gridded QPE product can be accessed online at [https://www.cnrfc
373 .noaa.gov/arc_search.php](https://www.cnrfc.noaa.gov/arc_search.php)

374

- The PRISM NORM81 climatology is described in Daly et al. (2015) and can be

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accessed online at <http://www.prism.oregonstate.edu/recent/>

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