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# Personalized Risk Assessment in Never, Light, and Heavy Smokers in a prospective cohort in Taiwan 


#### Abstract

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The objective of this study was to develop markedly improved risk prediction models for lung cancer using a prospective cohort of 395,875 participants in Taiwan. Discriminatory accuracy was measured by generation of receiver operator curves and estimation of area under the curve (AUC). In multivariate Cox regression analysis, age, gender, smoking pack-years, family history of lung cancer, personal cancer history, BMI, lung function test, and serum biomarkers such as carcinoembryonic antigen (CEA), bilirubin, alpha fetoprotein (AFP), and c-reactive protein (CRP) were identified and included in an integrative risk prediction model. The AUC in overall population was 0.851 ( $95 \% \mathrm{Cl}=0.840-0.862$ ), with never smokers 0.806 ( $95 \% \mathrm{CI}=0.790-0.819$ ), light smokers $0.847(95 \% \mathrm{CI}=0.824-0.871)$, and heavy smokers 0.732 ( $95 \% \mathrm{Cl}=0.708-0.752$ ). By integrating risk factors such as family history of lung cancer, CEA and AFP for light smokers, and lung function test (Maximum Mid-Expiratory Flow, MMEF $_{25-75 \%}$ ), AFP and CEA for never smokers, light and never smokers with cancer risks as high as those within heavy smokers could be identified. The risk model for heavy smokers can allow us to stratify heavy smokers into subgroups with distinct risks, which, if applied to low-dose computed tomography (LDCT) screening, may greatly reduce false positives.


Lung cancer is the leading contributor to cancer incidence and mortality worldwide ${ }^{1-3}$. The landmark National Lung Screening Trial (NLST) evaluated the benefits of low-dose computed tomography (LDCT) for screening of old-aged and heavy smokers ( $\geq 30$ pack-years) and found that annual screening by LDCT yielded a relative reduction of lung cancer mortality of $20 \%$ among those screened when compared to chest radiography, with a caveat of potential harms from high false positives, over-diagnoses, economic burden and repeated radiation ${ }^{4}$. The current recommendation for lung cancer screening by LDCT focused mainly on those heavy smokers with at least 30 pack years. However, not all lung cancer comes from heavy smokers. In fact, it was estimated that only about a quarter of currently diagnosed lung cancer patients in the U.S. meet the strict NLST eligibility criteria (age 55-74, $\geq 30$ pack-years smoking history) ${ }^{5}$. While focusing on heavy smokers is indeed a priority, a substantial portion of lung cancer cases continue to occur in light smokers and never smokers.

By design, light smokers and never smokers are not eligible for LDCT screening because they were assumed to have too low a risk for lung cancer. However, it is very likely that some of these individuals could have a risk of lung cancer similar to heavy smokers. To identify such individuals, additional risk factors working in tandem, other than smoking history, will be needed to create accurate risk models within light and never smokers. On the other hand, although LDCT could reduce mortality by $20 \%$, the high false positive rate ( $96.4 \%$ ) observed in the NLST calls for more accurate risk stratification among heavy smokers ${ }^{4}$. The UK Lung Screening (UKLS) Trial became the first trial to set up a threshold for pre-selecting screening population with an estimated risk of at least $5 \%$ of developing lung cancer in the next 5 years using the Liverpool Lung Project (LLP) risk model ${ }^{6}$.

[^0]The American Association for Thoracic Surgery (AATS) guidelines call for annual lung cancer screening with LDCT for those starting at age 50 years with a 20 pack-year history if there is an additional cumulative risk of developing lung cancer of $5 \%$ or greater in the next 5 years $^{7}$. Over the past decade, a concerted effort has been made to develop personalized risk prediction models for lung cancer ${ }^{8}$. Early reports yielded only modest discriminatory power with an area under the curve (AUC) of 0.72 or lower ${ }^{9-11}$. More recent models drawing on data collected by the Prostate, Lung, Colorectal, and Ovarian Cancer Screening Trial (PLCO) and the multi-center European Prospective Investigation into Cancer and Nutrition (EPIC) cohort which focused on smokers, have yielded improved discriminatory power with an AUC of $0.80-0.86$ in the modeling population ${ }^{12-14}$. These existing models have primarily incorporated only limited demographic factors (e.g., age, gender, and smoking history) and recognized clinical risk variables (e.g., chronic obstructive pulmonary disease (COPD) and pneumonia).

In this study, based on analyzing clinical, biomarker and other (e.g., lung function tests) data from a large prospective cohort in Taiwan, we developed integrative lung cancer prediction models for heavy smokers, light smokers and never smokers for 5-year and 10-year probability.

## Results

Characteristics of Cohort Participants. Among the 395,875 participants, there were a total of 1,117 incident lung cancer diagnoses. The mean ages were 40.4 for the whole cohort and 60.2 for the lung cancer cases. Categorization of the cohort by age group showed that the percentage of lung cancer cases increased from 0.07\% in those of age $<50$ years to $1.95 \%$ in those of age $\geq 70$ years. Over half ( $52 \%$ ) of the cohort was female and $38 \%$ of the lung cancer cases occurred in females; translating to sex-specific incidences of $0.21 \%$ for females and $0.32 \%$ for males. Owing to the high percentage ( $71 \%$ ) of never smokers in this cohort, $47 \%$ of the lung cancer cases occurred in never-smokers. Besides age, gender, and smoking, other variables associated with lung cancer included BMI, physical activity, and history of cancer (Table 1 and Supplemental Table 1).

Cox Multivariate Analyses of Laboratory Test and Medical Examination Variables. We selected only those variables that were significant based on Cox multivariate regression analysis to develop risk prediction models for lung cancer (Supplemental Table 1). Significant risk factors among the laboratory test and medical examination variables included low MMEF ( $0-48 \mathrm{ml} / \mathrm{sec}$ versus $>103 \mathrm{ml} / \mathrm{sec}, \mathrm{HR}=2.26,95 \% \mathrm{CI}=1.84-2.78$ ), and increasing levels of bilirubin (highest quartile versus lowest quartile, $\mathrm{HR}=1.42,95 \% \mathrm{CI}=1.19-1.70$ ), AFP ( $\geq 1.8 \mathrm{ng} / \mathrm{ml}$ versus $<1.8 \mathrm{ng} / \mathrm{ml}, \mathrm{HR}=1.38,95 \% \mathrm{CI}=1.12-1.70$ ), $\mathrm{CEA}(>7.0 \mathrm{ng} / \mathrm{ml}$ versus $<1.5 \mathrm{ng} / \mathrm{ml}, \mathrm{HR}=7.68$, $95 \% \mathrm{CI}=6.07-9.73$ ), and CRP (highest versus lowest quartile, $\mathrm{HR}=1.55,95 \% \mathrm{CI}=1.20-1.99$ ). Serum bilirubin displayed a pattern of decreasing risk with increasing quartile levels, reaching a borderline significance when comparing the lowest to the highest quartile ( $\mathrm{HR}=1.16,95 \% \mathrm{CI}=0.96$ to 1.40 ). Additional analyses were performed in never smokers, light smokers (pack-years $<30$ ) and heavy smokers (pack-years $\geqq 30$ ). The significant variables for never smokers include age, gender, BMI, family history of lung cancer, AFP, MMEF, and CEA, the significant variables for light smokers include age, sex, smoking status, smoking pack-year, family history of lung cancer, AFP, and CEA, and the significant variables for heavy smokers include age, sex, smoking status, smoking intensity, BMI, MMEF, and CEA (Table 2).

Risk Modeling. We then developed an integrative risk prediction model for the overall cohort, named the MD Anderson - MJ Group Integrative Risk Assessment (MMIRA) model based on variables in Table 2. We generated a time-dependent ROC curve, which yielded an AUC of 0.851 ( $95 \% \mathrm{CI}=0.840$ to 0.862 ) (Fig. 1). We calculated the C-index using internal validation by splitting the overall dataset into equally sized training and validation sets (Supplemental Table 2). Concordance was excellent, with only minor attenuation when moving from the training set $(0.854)$ to the validation set $(0.848)$ for the overall. We were able to demonstrate good calibration agreement between the observed and predicted probability of no events within the 10 -year time frame (Supplemental Fig. S1). We also generated separate models in never smokers, light smokers, and heavy smokers (Fig. 1). The AUC of $0.806(95 \% \mathrm{CI}=0.790$ to 0.819$)$ and $0.847(95 \% \mathrm{CI}=0.824$ to 0.871$)$ showed excellent predictive power in never-smokers and light smokers. Excellent concordance was also observed with minor attenuation from training to validation set for never smokers ( 0.795 to 0.822 ), light smokers ( 0.830 to 0.868 ), and heavy smokers ( 0.733 to 0.744 ) (Supplemental Table 2). In addition, we generated separate models in former smokers and current smokers with AUCs of $0.873(95 \% \mathrm{CI}=0.829$ to 0.879$)$ and $0.875(95 \% \mathrm{CI}=0.864$ to 0.887$)$, respectively (Supplemental Fig. S2). Further analysis showed that the positive predictive value for overall, never smokers, light smokers, heavy smokers, former smokers, and current smokers were $0.67 \%, 0.43 \%, 0.48 \%, 2.88 \%$, $1.55 \%$, and $1.73 \%$, respectively (Supplementary Table 3).

Application of risk prediction model. We applied the MMIRA models developed in never smokers, light smokers, and heavy smokers to predict probability of developing lung cancer in 5 years and 10 years to hypothetical individuals of age 65 with a range of risk profiles (Fig. 2). For a 65 -year old never smoker with relatively low risk profile ( $\mathrm{BMI} \geq 30$, negative family history of lung cancer), the predicted risk of developing lung cancer was $0.11 \%$ in 5 -years and $0.26 \%$ in 10 -years. However, the predicted probability of developing lung cancer increased to the range of 0.22 to $11.58 \%$ in 5 -years and the range of $0.51 \%$ to $24.86 \%$ in 10 -years for the addition of one to five risk factors (BMI, positive family history of lung cancer, AFP, CEA, and MMEF). Similarly, for a 65 -year old person who is light smoker, the probability of developing cancer increases from $0.06 \%$ to $5.03 \%$ in 5 years, and from $0.15 \%$ to $11.27 \%$ in 10 years. For a 65 -year old person who is a heavy smoker, the probability of developing cancer increases from $0.16 \%$ to $3.53 \%$ in 5 years, and from $0.42 \%$ to $8.82 \%$ in 10 years.

We also assigned risk scores to each risk factor based on the strength of the association (Supplementary Table 4). The higher HR a risk factor conferred, the higher the risk score was assigned to the risk factor. For example, in the age category, age 50-59 was the reference group and the assigned score was 0 , age $<50$ was protective with an

| Variables | Level | Overall |  |  |  | Never smoking |  |  |  | Light smoking |  |  |  | Heavy smokers |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\mathrm{N}=395,875$ |  | LC= 1,117 |  | N=281,111 |  | LC= 525 |  | $\mathrm{N}=90,486$ |  | LC= 156 |  | N=2,4278 |  | LC = 436 |  |
|  |  | N | \% | n | \% | N | \% | n | \% | N | \% | n | \% | N | \% | n | \% |
| Age, mean(SD) |  | 40(13.8) |  |  |  | 40(13.8) |  |  |  | 37(11.3) |  |  |  | 57(10.1) |  |  |  |
| Age | <50 | 293,040 | 74\% | 201 | 18\% | 207,061 | 74\% | 123 | 23\% | 79,619 | 88\% | 57 | 37\% | 6,360 | 26\% | 21 | 5\% |
|  | 50-59 | 55,772 | 14\% | 260 | 23\% | 41,874 | 15\% | 146 | 28\% | 5,539 | 6\% | 26 | 17\% | 8,359 | 34\% | 88 | 20\% |
|  | 60-69 | 34,531 | 9\% | 412 | 37\% | 23,988 | 9\% | 167 | 32\% | 3,673 | 4\% | 47 | 30\% | 6,870 | 28\% | 198 | 45\% |
|  | $\geq 70$ | 12,532 | 3\% | 244 | 22\% | 8,188 | 3\% | 89 | 17\% | 1,655 | 2\% | 26 | 17\% | 2,689 | 11\% | 129 | 30\% |
| Sex | Male | 191,120 | 48\% | 691 | 62\% | 92,763 | 33\% | 133 | 25\% | 75,272 | 83\% | 142 | 91\% | 23,085 | 95\% | 416 | 95\% |
|  | Female | 204,755 | 52\% | 426 | 38\% | 188,348 | 67\% | 392 | 75\% | 15,214 | 17\% | 14 | 9\% | 1,193 | 5\% | 20 | 5\% |
| Smoking status | Never | 281,111 | 71\% | 525 | 47\% |  |  |  |  |  |  |  |  |  |  |  |  |
|  | Former | 23,750 | 6\% | 127 | 11\% |  |  |  |  | 18,167 | 20\% | 33 | 21\% | 5,583 | 23\% | 94 | 22\% |
|  | Current | 89,867 | 23\% | 461 | 41\% |  |  |  |  | 71,406 | 80\% | 121 | 79\% | 18,461 | 77\% | 340 | 78\% |
| Smoking intensity* | <1 | 86,089 | 75\% | 372 | 63\% |  |  |  |  | 76,304 | 84\% | 143 | 92\% | 9,785 | 40\% | 229 | 53\% |
| (pack/per day) | $\geq 1$ | 28,675 | 25\% | 220 | 37\% |  |  |  |  | 14,182 | 16\% | 13 | 8\% | 14,493 | 60\% | 207 | 47\% |
| Pack-year* | <15 | 66,516 | 58\% | 68 | 11\% |  |  |  |  | 66,516 | 74\% | 68 | 44\% |  |  |  |  |
|  | 15-29.9 | 23,970 | 21\% | 88 | 15\% |  |  |  |  | 23,970 | 26\% | 88 | 56\% |  |  |  |  |
|  | $\geq 30$ | 24,278 | 21\% | 436 | 74\% |  |  |  |  |  |  |  |  |  |  |  |  |
| BMI (kg/m2) | $<25$ | 291,339 | 74\% | 781 | 70\% | 213,289 | 76\% | 362 | 69\% | 63,087 | 70\% | 101 | 65\% | 14,963 | 62\% | 318 | 73\% |
|  | 25-29.9 | 89,563 | 23\% | 312 | 28\% | 57,869 | 21\% | 147 | 28\% | 23,471 | 26\% | 53 | 34\% | 8,223 | 34\% | 112 | 26\% |
|  | $\geq 30$ | 14,801 | 4\% | 24 | 2\% | 9,842 | 4\% | 16 | 3\% | 3,884 | 4\% | 2 | 1\% | 1,075 | 4\% | 6 | 1\% |
| Physical activity | Inactive | 206,056 | 52\% | 546 | 50\% | 145,219 | 52\% | 240 | 47\% | 47,992 | 53\% | 78 | 51\% | 12,845 | 53\% | 228 | 53\% |
|  | Low active | 89,154 | 23\% | 177 | 16\% | 65,183 | 23\% | 89 | 17\% | 20,224 | 22\% | 24 | 16\% | 3,747 | 15\% | 64 | 15\% |
|  | Fully active | 100,458 | 25\% | 374 | 34\% | 70,587 | 25\% | 186 | 36\% | 22,237 | 25\% | 50 | 33\% | 7,634 | 32\% | 138 | 32\% |
| Family history of lung cancer | No | 379,272 | 96\% | 1,059 | 95\% | 269,537 | 96\% | 497 | 95\% | 86,590 | 96\% | 146 | 94\% | 23,145 | 95\% | 416 | 95\% |
|  | Yes | 16,603 | 4\% | 58 | 5\% | 11,574 | 4\% | 28 | 5\% | 3,896 | 4\% | 10 | 6\% | 1,133 | 5\% | 20 | 5\% |
| Personal cancer history | No | 390,895 | 99\% | 1,091 | 98\% | 277,098 | 99\% | 512 | 98\% | 89,911 | 99\% | 155 | 99\% | 23,886 | 98\% | 424 | 97\% |
|  | Yes | 4,980 | 1\% | 26 | 2\% | 4,013 | 1\% | 13 | 2\% | 575 | 1\% | 1 | 1\% | 392 | 2\% | 12 | 3\% |
| COPD | No | 259,217 | 77\% | 532 | 51\% | 182,532 | 77\% | 284 | 58\% | 63,322 | 83\% | 84 | 60\% | 13,363 | 59\% | 164 | 39\% |
|  | Yes | 20,748 | 6\% | 172 | 16\% | 13,354 | 6\% | 52 | 11\% | 4,592 | 6\% | 21 | 15\% | 2,802 | 12\% | 99 | 24\% |
|  | Restricted lung | 55,009 | 16\% | 346 | 33\% | 39,993 | 17\% | 156 | 32\% | 8,662 | 11\% | 36 | 26\% | 6,354 | 28\% | 154 | 37\% |
| Lung function |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| (1) FEV1 (\%) | 0-61 | 17,821 | 5\% | 143 | 13\% | 12,156 | 5\% | 42 | 8\% | 3,480 | 4\% | 15 | 10\% | 2,185 | 9\% | 86 | 20\% |
|  | 62-73 | 26,408 | 7\% | 115 | 11\% | 19,037 | 8\% | 55 | 11\% | 5,041 | 6\% | 17 | 11\% | 2,330 | 10\% | 43 | 10\% |
|  | 74-90 | 130,574 | 37\% | 328 | 30\% | 94,443 | 37\% | 158 | 31\% | 28,851 | 35\% | 44 | 29\% | 7,280 | 31\% | 126 | 30\% |
|  | 91-105 | 126,604 | 35\% | 301 | 28\% | 87,987 | 35\% | 157 | 31\% | 31,343 | 38\% | 47 | 31\% | 7,274 | 31\% | 97 | 23\% |
|  | >106 | 55,741 | 16\% | 201 | 18\% | 38,656 | 15\% | 99 | 19\% | 12,951 | 16\% | 28 | 19\% | 4,134 | 18\% | 74 | 17\% |
| (2) MMEF ( $\mathrm{ml} / \mathrm{sec}$ ) | 0-48 | 15,598 | 4\% | 174 | 16\% | 10,113 | 4\% | 63 | 12\% | 2,927 | 4\% | 19 | 13\% | 2,558 | 11\% | 92 | 22\% |
|  | 49-54 | 11,722 | 3\% | 61 | 6\% | 8,014 | 3\% | 23 | 4\% | 2,543 | 3\% | 8 | 5\% | 1,165 | 5\% | 30 | 7\% |
|  | 55-77 | 103,091 | 29\% | 317 | 29\% | 72,785 | 29\% | 141 | 27\% | 23,397 | 29\% | 34 | 22\% | 6,909 | 30\% | 142 | 33\% |
|  | 78-103 | 147,469 | 41\% | 349 | 32\% | 105,481 | 42\% | 179 | 35\% | 34,297 | 42\% | 64 | 42\% | 7,691 | 33\% | 106 | 25\% |
|  | >103 | 79,651 | 22\% | 193 | 18\% | 56,163 | 22\% | 109 | 21\% | 18,568 | 23\% | 27 | 18\% | 4,920 | 21\% | 57 | 13\% |
| Bilirubin (mg/dl) | $\mathrm{M}: \leq 0.68 ; \mathrm{F} \leq 0.56$ | 100,354 | 26\% | 310 | 28\% | 66,330 | 24\% | 117 | 23\% | 26,279 | 30\% | 47 | 31\% | 7,745 | 33\% | 146 | 34\% |
|  | $\begin{gathered} \text { M: } 0.69-0.87 \text {; } \\ \text { F: } 0.57-0.70 \end{gathered}$ | 95,562 | 25\% | 301 | 27\% | 67,021 | 24\% | 158 | 30\% | 22,087 | 25\% | 43 | 28\% | 6,454 | 27\% | 100 | 23\% |
|  | $\begin{aligned} & \text { M: 0.88-1.11; } \\ & \text { F: } 0.71-0.90 \end{aligned}$ | 97,384 | 25\% | 286 | 26\% | 70,491 | 26\% | 123 | 24\% | 21,471 | 24\% | 41 | 27\% | 5,422 | 23\% | 122 | 28\% |
|  | $\begin{gathered} \mathrm{M}: \geq 1.12 ; \\ \mathrm{F}: \geq 0.91 \end{gathered}$ | 93,105 | 24\% | 212 | 19\% | 70,402 | 26\% | 122 | 23\% | 18,536 | 21\% | 23 | 15\% | 4,167 | 18\% | 67 | 15\% |
| AFP (ng/ml) | <1.8 | 82,806 | 22\% | 100 | 9\% | 64,516 | 24\% | 64 | 12\% | 16,862 | 19\% | 8 | 5\% | 1,428 | 6\% | 28 | 6\% |
|  | $\geq 1.8$ | 299,179 | 78\% | 1,000 | 91\% | 206,728 | 76\% | 450 | 88\% | 70,576 | 81\% | 147 | 95\% | 21,875 | 94\% | 403 | 94\% |
| CEA (ng/ml) | $<1.5$ | 253,091 | 65\% | 315 | 28\% | 200,170 | 72\% | 226 | 43\% | 47,212 | 53\% | 33 | 21\% | 5,709 | 24\% | 56 | 13\% |
|  | 1.5-2.5 | 86,595 | 22\% | 298 | 27\% | 53,648 | 19\% | 138 | 26\% | 25,533 | 29\% | 49 | 32\% | 7,414 | 31\% | 111 | 25\% |
|  | 2.6-4.2 | 37,278 | 10\% | 266 | 24\% | 17,968 | 6\% | 93 | 18\% | 12,514 | 14\% | 44 | 28\% | 6,796 | 28\% | 129 | 30\% |
|  | 4.3-7.0 | 10,588 | 3\% | 139 | 12\% | 4,060 | 1\% | 25 | 5\% | 3,317 | 4\% | 19 | 12\% | 3,211 | 13\% | 95 | 22\% |
|  | $>7.0$ | 2,757 | 1\% | 98 | 9\% | 993 | 0\% | 43 | 8\% | 687 | 1\% | 10 | 6\% | 1,077 | 4\% | 45 | 10\% |
| CRP (mg/L) | 0-1 | 259,145 | 71\% | 551 | 56\% | 187,169 | 72\% | 278 | 59\% | 59,607 | 72\% | 85 | 63\% | 12,369 | 56\% | 188 | 50\% |
| Continued |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |


| Variables | Level | Overall |  |  |  | Never smoking |  |  |  | Light smoking |  |  |  | Heavy smokers |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\mathrm{N}=395,875$ |  | LC= 1,117 |  | $\mathrm{N}=281,111$ |  | LC= 525 |  | $\mathrm{N}=90,486$ |  | LC= 156 |  | $\mathrm{N}=2,4278$ |  | LC= 436 |  |
|  |  | N | \% | n | \% | N | \% | n | \% | N | \% | n | \% | N | \% | n | \% |
|  | 1.1-3 | 66,919 | 18\% | 227 | 23\% | 46,419 | 18\% | 113 | 24\% | 15,002 | 18\% | 28 | 21\% | 5,498 | 25\% | 86 | 23\% |
|  | 3,1-10 | 30,639 | 8\% | 132 | 13\% | 20,681 | 8\% | 53 | 11\% | 6,693 | 8\% | 13 | 10\% | 3,265 | 15\% | 66 | 18\% |
|  | >10 | 8,499 | 2\% | 69 | 7\% | 5,537 | 2\% | 25 | 5\% | 1,900 | 2\% | 8 | 6\% | 1,062 | 5\% | 36 | 10\% |

Table 1. Cohort characteristics. $\mathrm{LC}=$ lung cancer. $\mathrm{BMI}=$ Body mass index. COPD $=$ chronic obstructive pulmonary disease. FEV1 = forced expiratory volume in 1 second. MMEF $=$ maximum midexpiratory flow. $\mathrm{AFP}=$ alpha-fetoprotein. $\mathrm{CEA}=$ carcinoembryonic antigen. $\mathrm{CRP}=\mathrm{C}$-reactive protein. *Among smokers.
assigned score of -4 , whereas the assigned scores for age $60-69$ and age $\geq 70$ were 2 and 3 , respectively. The risk scores for all cohort participants ranged from -4 to 19 for overall cohort: for never smokers, -5 to 17 ; for light smokers, -5 to 14 ; and for heavy smokers, -3 to 12 . Figure 3 depicts the probability of developing lung cancer in 5 and 10 years as a function of increasing risk scores. For example, for never smokers with a score of 15, the corresponding risk would be $8.42 \%$ and $18.48 \%$ in 5 and 10 years, respectively (Fig. 3B). Similarly, we could use risk scores to stratify light smokers into 20 categories with the 5 -year lung cancer probability ranging from $0.00 \%$ to $7.39 \%$ and stratify heavy smokers into 16 categories with the 5 -year lung cancer probability ranging from $0.02 \%$ to $7.48 \%$ (Fig. 3C,D).

## Discussion

It was estimated that $26.7 \%$ of lung cancer cases occurred among heavy smokers who meet NLST eligibility criteria in the U.S. ${ }^{5}$. The growing desire to extend LDCT screening beyond heavy smokers is understandable, particularly among the overwhelming majority of Asian women who were inflicted with lung cancer but never smoked ( $70 \%$ to $90 \%$ ). In this cohort, more than $70 \%$ of lung cancer occurred in female never smokers. These high lung cancers in Asian women came mainly from second hand smoke from their fathers, brothers, and spouses, living in a small enclosed space, thus the second hand smoking rate could reach $75 \%$ in their earlier years throughout their life. The LDCT screening results in heavy smokers has elevated the public expectation for targeted screening for high-risk groups other than heavy smokers. In this paper, we have developed robust risk prediction models for never and light smokers in Asia, in addition to more accurately identify higher risk subjects in heavy smokers. As new findings, other than the smoking information and family history, four clinically common biomarkers, CEA, AFP, CRP and bilirubin, as well as a specific lung function test, were found to be uniquely useful in identifying high risk individuals. These biomarkers and the lung function test divided never smokers, light smokers, and heavy smokers into distinct groups with a range of 5 -year lung cancer probability. Never-smokers with risk scores of 14 and above (Fig. 3C) and light smokers with risk scores of 13 and above would have an absolute cancer risk above $5 \%$ in five years, a risk threshold level suggested by UKLS and AATS to start the LDCT screening ${ }^{6,7}$.

There have been attempts to use additional data to improve the discriminative performance of risk stratification and participant selection for LDCT screening, most notably the LLP model, which added 4 history questions (history of pneumonia, personal history of cancer, asbestos exposure, and family history of lung cancer) and has been applied to the UKLS trial. Our study added more risk factors including laboratory biomarkers and a lung function test and was able to identify those with cancer risk exceeding the $5 \%$ threshold in 5 -year probability set up by the UKLS and AATS ${ }^{6,7}$. Our model incorporated several unique predictors of lung cancer risk, including MMEF as an index of airway obstruction, and the serum markers CEA, bilirubin, AFP, and CRP. These covariates have not been integrated into current lung cancer risk prediction models (Supplemental Table 5) because such data are often unavailable for population-based cohort studies.

Spirometry has been used to demonstrate airflow obstruction, and can also suggest restrictive ventilator impairment. In lung cancer risk screening, spirometry is particularly valuable among never smokers. Different parameters of lung function test have been suggested to evaluate their relationships with lung cancer risks, such as COPD, $\mathrm{FEV}_{1} \%$ or Forced Expiratory Flow in the middle half of FVC ( $\mathrm{MMEF}_{25-75 \%}$ ). Incremental reduction of $\mathrm{FEV}_{1}$ values has been strongly associated with lung cancer risk, independent of smoking ${ }^{15}$. COPD, as defined by the GOLD [Global Initiative for Chronic Obstructive Lung Disease] criteria with FEV 1/FVC $<0.7$, was also known to be a strong risk factor for lung cancer, in both smokers and never smokers. Consistent with the literature, we found, in our cohort, either FEV1 or COPD a strong risk factor for lung cancer. However, in our final comparison among parameters, $\mathrm{MMEF}_{25-75 \%}$ turned out to be the most sensitive indicator for lung cancer risk after a multivariate analysis. Therefore, the risk score sheet relied on the values of $\mathrm{MMEF}_{25-75 \%}$ in our modeling. The most likely explanation for reduced lung function as a lung cancer risk is that it reflects airway inflammation, a prodromal phase for lung cancer risk. Airway inflammation could present itself as either obstructive or restrictive impairment, with the more impairment the higher the risk. Another possibility is that the reduced lung function may impair the ability to clear inhaled carcinogens from their airways, which could lead to increased contact time between carcinogens and airway epithelial cells. These mechanisms probably facilitated $\mathrm{MMEF}_{25-75 \%}$ as a lung cancer risk not only for smokers but also for never smokers, a feature important in our search for high risk individuals among never smokers. In our cohort, the lowest $8 \% \mathrm{MMEF}_{25-75 \%}$ of overall subjects had doubled their cancer risks and contributed $22 \%$ of all lung cancer cases, resulting in a multivariate adjusted HR at 2.06 .

The CEA glycoprotein is an established tumor marker for colorectal cancer, and has been evaluated as a prognostic or predictive marker for lung cancer ${ }^{16,17}$. In our cohort, high CEA, e.g. $>7 \mathrm{ng} / \mathrm{ml}$, showed marked increase in lung cancer risk, with adjusted HR 12.82 for never smokers and 4.21 for light smokers. This level of CEA,

| AUC |  | Overall | Never Smokers | Light Smokers | Heavy Smokers |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 0.851 | 0.806 | 0.847 | 0.732 |
| Variables |  | HR(95\% CI) | HR(95\% CI) | HR(95\% CI) | HR(95\% CI) |
| Age |  | 1.07(1.07-1.08) | 1.07(1.06-1.08) | 1.08(1.07-1.10) | 1.08(1.07-1.09) |
| Sex | Male | 0.88(0.74-1.04) | 0.71(0.58-0.87) | 1.58(0.91-2.77) | 1.11(0.71-1.74) |
|  | Female | 1.00 | 1.00 | 1.00 | 1.00 |
| Smoking | Never | 1.00 |  |  |  |
|  | $<30$ pack-year | 1.14(0.92-1.42) |  |  |  |
|  | $\geq 30$ pack-year | 2.60(2.15-3.14) |  |  |  |
| Smoking status | Former |  |  | 1.00 | 1.00 |
|  | Current |  |  | 1.64(1.08-2.48) | 1.48(1.16-1.89) |
| Smoking pack year | <15 |  |  | 1.00 |  |
|  | 15~30 |  |  | 1.44(1.03-2.02) |  |
| Smoking intensity | $<1$ |  |  |  | 1.00 |
| (pack/per day) | $\geq 1$ |  |  |  | 1.46(1.20-1.79) |
| BMI | $<25$ | 2.11(1.35-3.30) | 1.91(1.13-3.20) |  | 2.48(1.11-5.58) |
| (kg/m2) | 25-29.9 | 1.79(1.13-2.83) | 1.55(0.91-2.64) |  | 2.02(0.89-4.59) |
|  | $\geq 30$ | 1.00 | 1.00 |  | 1.00 |
| Family history of lung cancer | No | 1.00 | 1.00 | 1.00 |  |
|  | Yes | 1.44(1.08-1.94) | 1.66(1.13-2.45) | 2.00(1.05-3.81) |  |
| Personal cancer history | No | 1.00 |  |  |  |
|  | Yes | 1.55(1.02-2.34) |  |  |  |
| Bilirubin (mg/dl) | $\begin{gathered} \text { Male: } \leq 0.68 ; \\ \text { Female: } \leq 0.56 \end{gathered}$ | 1.16(0.96-1.40) |  |  |  |
|  | $\begin{aligned} & \text { Male:0.69-0.87; } \\ & \text { Female:0.57-0.70 } \end{aligned}$ | 1.06(0.88-1.29) |  |  |  |
|  | $\begin{aligned} & \text { Male:0.88-1.11; } \\ & \text { Female:0.71-0.90 } \end{aligned}$ | 1.08(0.89-1.31) |  |  |  |
|  | $\text { Male }: \geq 1.12 ; \text { Female }: \geq$ | 1.00 |  |  |  |
| AFP ( $\mathrm{ng} / \mathrm{ml}$ ) | <1.8 | 1.00 | 1.00 | 1.00 |  |
|  | $\geq 1.8$ | 1.35(1.08-1.69) | 1.37(1.05-1.80) | 2.12(1.03-4.33) |  |
| MMEF (ml/sec) | 0-48 | 2.00(1.58-2.52) | 1.87(1.36-2.57) |  | 2.06(1.47-2.87) |
|  | 49-54 | 1.64(1.20-2.26) | 1.46(0.93-2.29) |  | 1.71(1.10-2.67) |
|  | 55-77 | 1.49(1.22-1.82) | 1.24(0.96-1.60) |  | 1.63(1.19-2.22) |
|  | 78-103 | 1.41(1.16-1.71) | 1.19(0.93-1.51) |  | 1.33(0.96-1.83) |
|  | >103 | 1.00 | 1.00 |  | 1.00 |
| CEA (ng/ml) | $<1.5$ | 1.00 | 1.00 | 1.00 | 1.00 |
|  | 1.5-2.5 | 1.30(1.09-1.56) | 1.24(0.99-1.55) | 1.45(0.92-2.28) | 1.15(0.83-1.59) |
|  | 2.6-4.2 | 1.64(1.35-1.99) | 1.82(1.40-2.36) | 1.88(1.18-3.01) | 1.25(0.91-1.72) |
|  | 4.3-7.0 | 2.26(1.79-2.86) | 1.76(1.14-2.71) | 2.16(1.20-3.89) | 1.79(1.27-2.52) |
|  | $>7.0$ | 5.70(4.39-7.41) | 12.82(9.10-18.07) | 5.29(2.55-10.95) | 2.55(1.70-3.83) |
| CRP (mg/L) | 0-1 | 1.00 |  |  |  |
|  | 1.1-3 | 1.02(0.87-1.20) |  |  |  |
|  | 3.1-10 | 0.96(0.78-1.17) |  |  |  |
|  | $>10$ | 1.38(1.07-1.78) |  |  |  |

Table 2. Cox regression lung cancer prediction models from MD Anderson and MJ group Integrative Risk Assessment (MMIRA) in overall, never smokers, light smokers, and heavy smokers. BMI = Body mass index. $\mathrm{AFP}=$ alpha-fetoprotein. $\mathrm{MMEF}=$ maximum midexpiratory flow. $\mathrm{CEA}=$ carcinoembryonic antigen. CRP $=$ C-reactive protein.
constituting $1 \%$ to $4 \%$ of the cohort subjects, served as an excellent screening biomarker in our prediction model. The AFP tumor marker is most commonly used to aid screening and diagnosis of liver cancer and monitor response to treatment ${ }^{18}$, but an increased level of AFP has also been associated with other malignancies. In this cohort, mild elevation of AFP, $\geq 1.8 \mathrm{ng} / \mathrm{ml}$, was associated with $37 \%$ increase in lung cancer among never smokers.

Elevated level of serum CRP, a systemic marker of chronic inflammation, has been consistently associated increased risk of lung cancer ${ }^{19,20}$. With CRP greater than $10 \mathrm{mg} / \mathrm{L}$, lung cancer risk increased by $54 \%$ in this cohort. For CRP at that level, there were $2 \%$ of the cohort and $7 \%$ of lung cancer cases.

We had reported the elevated lung cancer risk from low serum bilirubin, which has anti-oxidant properties. Others also reported that relatively low serum bilirubin was associated with higher risks of lung cancer and


Figure 1. Discriminatory accuracy of lung cancer prediction models from MD Anderson - MJ Group Integrative Risk Assessment (MMIRA) in overall, heavy, light and never smokers for lung cancer risk. Discriminatory accuracy for predicting lung cancer risk within 10 years was assessed by receiver operator characteristics (ROC) analysis calculating area under the curve (AUC).


Figure 2. Application of the integrative risk prediction models in never smokers, light smokers, and heavy smokers (smoking pack-year $\geq \mathbf{3 0}$ ) to predict absolute risk of developing lung cancer in 5 year and 10 years for hypothetical individuals, with different risk profiles by the addition of different risk factors.

COPD in a cohort study ${ }^{21}$. This risk remained in our integrative prediction model for the overall group, but not subgroups we examined with multivariate analysis.

It is difficult to compare performance metrics between published risk prediction models for lung cancer as each have been developed in different populations with varying lengths of follow-up time. In an independent case-control study used to compare the early Bach, Spitz and LLP models, differences in model sensitivity and specificity were highlighted and only moderate discriminatory power (AUC $=0.66-0.69$ for all models) was found ${ }^{22}$. The NLST trial defined high-risk criteria based solely on age and smoking history. It has been estimated that if the PLCO risk prediction model had been used to select individuals for LDCT screening in the NLST trial, 12 additional deaths attributable to lung cancer could have been prevented ${ }^{14}$. Using similar calculations, we estimate that an additional one death could have been avoided if updated $\mathrm{PLCO}_{\mathrm{M} 2012}$ model has been used, and an additional eight deaths due to lung cancer could have been avoided if our MMIRA model had been applied.

Our models for light smokers ( $\mathrm{AUC}=0.847$ ) and never-smokers $(\mathrm{AUC}=0.808)$ had excellent predictive power. By calculating a risk score based on risk factor profile, the 5 -year lung cancer probability of a light smoker ranged from $0.00 \%$ to over $7.39 \%$, and the probability of a never smoker ranged from $0.01 \%$ to over $15.82 \%$. Never-smokers with risk scores of 14 and above and light smokers with risk scores of 13 and above would have

B) Never smokers



Figure 3. Predicted probability of lung cancer risk based on the risk scores from MMIRA models (A) in overall population, $(\mathbf{B})$ in never smokers, (C) in light smokers, and (D) in heavy-smokers. Blue line - predicted probability in 5-year. Red line- predicted probability in 10 year.
an absolute cancer risk above $5 \%$ in five years, a threshold level suggested by UKLS and AATS to start the LDCT screening ${ }^{6,7}$. Thus, our prediction model is able to stratify light smokers and never smokers into subgroups with dramatically different probability of developing lung cancer, with a portion, albeit small, of them as high as those in heavy smokers. Clinicians and patients can consult our score sheet in making an informed decision for assessing the risks and benefits of screening with LDCT. As false positives and over-diagnosis had been a problem for using LDCT in heavy smokers, so will be the challenges for screening any group other than heavy smokers. However, individuals can make better-informed decisions based on his/her absolute risk of developing lung cancer.

Our study has a couple of limitations. First, although we have a relatively high level of discrimination, external validation of the models is required to determine predictive ability in an independent population. Nevertheless, internal calibration and bootstrap analysis of goodness of fit showed excellent agreement between predicted and observed events and between the two randomly selected subcohorts. Secondly, as the MJ model has not been validated in a non-Asian population, we do not know if it will function with same predictive power across other racial/ethnic groups.

In summary, using a very large prospective cohort of an Asian population, we have demonstrated the power of incorporating routine laboratory test data and medical evaluation variables into prediction algorithms for lung cancer. Our models should improve selection of high-risk individuals for targeted screening strategies. Additional studies are necessary to validate these results in independent cohorts and to extend the findings to other ethnic populations.

## Methods

Study Population and Data Collection. All subjects were recruited by the MJ Health Group, Taiwan, to participate in a national health-screening program. The current analysis was conducted after a median 7.3 years (range $=0 \sim 11.9$ years) of follow-up from 1996 to 2007. Details of the screening program have been reported previously ${ }^{23}$. In brief, each subject completed a comprehensive health history questionnaire to collect medical history and epidemiological data. Participants underwent hands-on physical examinations and submitted to a panel of 103 blood and medical tests including lung function tests. Informed consent was obtained from all participants. The study was approved by Institutional Review Boards at the National Health Research Institute in Taiwan and MD Anderson Cancer Center. All the methods of subject recruitment, data collection, and experiments were performed in accordance with relevant guidelines and regulations.

Ascertainment of lung cancer. The national ID of each cohort participant was matched to the National Cancer Registry and National Death File in order to assess outcomes and events. As of 2008, the cohort had registered 1,117 new cases of lung cancer and 799 lung cancer deaths.

Laboratory test and lung function test. Serum biomarkers CEA, bilirubin, AFP, and CRP were tested using the Abbott ARCHITECT ci8200. Airway obstruction was measured in a standard spirometry test and obstruction expressed as FEV1\% or MMEF (maximum mid-expiratory flow).

Statistical Analysis. Risk predictors were subjected to stepwise Cox proportional hazards regression analysis to identify significant predictors in multivariate models. Hazard ratios (HRs) and 95\% confidence intervals (CIs) were estimated for each variable. Continuous variables, including MMEF, CEA, bilirubin, AFP, and CRP, were assessed by quartile or other cut-points as guided by cohort distribution. To evaluate the discriminatory accuracy of the risk prediction models, we assessed the goodness of fit by calculating the concordance index (C-index) and the area under the curve (AUC) from the receiver operating characteristic (ROC) curve analysis. We examined the goodness of fit for 10 -year risk prediction in each of the training, validation, and full datasets. Bootstrap resampling was performed 100 times to generate the $95 \%$ CI for AUC. We created risk scores based on weighted sum of the identified risk factors in each model and the weights were based on the coefficients, $\beta \mathrm{i}$, from the multivariate Cox regression model following the procedure by Sullivan et al. 2004 ${ }^{24}$. For example, the increase in risk associated with a 5 -year increase in age was first estimated as the constant $B$, followed by the calculation of risk score rounded to the nearest integer using the $\beta \mathrm{ii} / \mathrm{B}$. All statistical tests were two-sided and only $P$-values less than 0.05 were considered statistically significant.

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## Author Contributions

X.W., C.P.W., J.G. and S.M.L. developed the concept and designed the study. M.T., X.P., M.H., S.W. and Y.Y. were involved in data collection and data analyses. X.W., C.P.W., Y.Y., C.W., J.A.R., W.C., C.H., S.C. and J.G. were involved in data interpretation. C.K.T. provided study materials and supervised data collection and assembly of data. X.W., C.P.W., S.C. and J.G. drafted the manuscript.

## Additional Information

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