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## Low-Fat Dietary Pattern and Breast Cancer Mortality in the Women's Health Initiative Randomized Controlled Trial

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#### ABSTRACT

#### **Purpose**

Earlier Women's Health Initiative Dietary Modification trial findings suggested that a low-fat eating pattern may reduce breast cancers with greater mortality. Therefore, as a primary outcome-related analysis from a randomized prevention trial, we examined the long-term influence of this intervention on deaths as a result of and after breast cancer during 8.5 years (median) of dietary intervention and cumulatively for all breast cancers diagnosed during 16.1 years (median) of follow-up.

#### **Patients and Methods**

The trial randomly assigned 48,835 postmenopausal women with normal mammograms and without prior breast cancer from 1993 to 1998 at 40 US clinical centers to a dietary intervention with goals of a reduction of fat intake to 20% of energy and an increased intake of fruits, vegetables, and grains (40%; n = 19,541) or to a usual diet comparison (60%; n = 29,294).

#### Results

In the dietary group, fat intake and body weight decreased (all P < .001). During the 8.5-year dietary intervention, with 1,764 incident breast cancers, fewer deaths occurred as a result of breast cancer in the dietary group, which was not statistically significant (27 deaths [0.016% per year] v61 deaths [0.024% per year]; hazard ratio [HR], 0.67; 95% CI, 0.43 to 1.06; P = .08). During the same period, deaths after breast cancer (n = 134) were significantly reduced (40 deaths [0.025% per year] v 94 deaths [0.038% per year]; HR, 0.65; 95% CI, 0.45 to 0.94; P = .02) by the dietary intervention. During the 16.1-year follow-up, with 3,030 incident breast cancers, deaths after breast cancer also were significantly reduced (234 deaths [0.085% per year] v 443 deaths [0.11% per year]; HR, 0.82; 95% CI, 0.70 to 0.96; P = .01) in the dietary group.

#### **Conclusion**

Compared with a usual diet comparison group, a low-fat dietary pattern led to a lower incidence of deaths after breast cancer.

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## INTRODUCTION

In the Women's Health Initiative Dietary Modification (WHI DM) trial, 48,835 postmenopausal women were randomly assigned to a dietary modification group or usual diet comparison group to assess low-fat dietary pattern effects on breast cancer incidence. As previously reported, the dietary modification program reduced fat intake; increased fruit, vegetable, and grain intake; and was associated with modest weight loss (3% loss after 1 year; -2.2 kg; P < .001). Dietary intake and body

weight differences between randomly assigned groups were maintained throughout the dietary intervention period.<sup>1,2</sup>

At the protocol-specified end of the dietary intervention, after 8.5 years (median) of follow-up, although the breast cancer incidence rate was lower in the dietary group, the difference was not statistically significant (hazard ratio [HR], 0.92; 95% CI, 0.84 to 1.01; P = .09). However, a significant reduction in estrogen receptor–positive (ER+), progesterone receptor–negative (PR–) cancers was observed (HR, 0.64; 95% CI, 0.49 to 0.84; P = .04). No suggestion of reduced breast cancer incidence existed during the limited

#### **ASSOCIATED CONTENT**



Data Supplement
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DOI: https://doi.org/10.1200/JCO.2016 72.0326 postintervention follow-up, but the significant reduction in ER+, PR- cancers continued (HR, 0.70; 95% CI, 0.56 to 0.88).<sup>3</sup> Although not statistically significant, during the dietary intervention, fewer deaths as a result of breast cancer were observed in the dietary group (HR, 0.77; 95% CI, 0.48 to 1.22).<sup>1</sup>

Taken together, these findings suggest that dietary modification reduces breast cancers associated with a greater risk of death, and this possibility has now been examined through 16.1 years (median) of cumulative follow-up. We present results of dietary modification influence on deaths as a result of breast cancer (breast cancer followed by death attributed to the cancer) and on the more general outcome of deaths after breast cancer (breast cancer followed by death as a result of any cause). These end points are examined first through the end of the 8.5-year (median) dietary intervention period and then cumulatively for all breast cancers diagnosed at any time throughout the 16.1-year (median) follow-up through 2013.

#### **PATIENTS AND METHODS**

#### **Participants**

Design details of the WHI DM trial, which was conducted at 40 US clinical centers with enrollment from 1993 through 1998, have been described previously. <sup>4,5</sup> Eligible participants were postmenopausal women between 50 and 79 years of age with no previous breast cancer, a dietary fat intake > 32% of total energy by food frequency questionnaire (FFQ), <sup>6</sup> and a mammogram not suspicious for cancer. The trial was approved by institutional review boards of the participating clinical centers, and participants provided written informed consent.

#### Random Assignment

Participants were randomly assigned to a low-fat dietary pattern intervention group or a usual diet comparison group in a two-to-three ratio for cost efficiency by using a permuted block algorithm stratified by clinical center and age decade, with allocations developed by the WHI Clinical Coordinating Center.

#### **Procedure**

Baseline characteristics were gathered by interview for medication use and by questionnaires for other variables. Body weight and height were measured at baseline and annually, with body mass index calculated as kilograms per meter squared. Physical activity was assessed by questionnaire at baseline, year 3, and year 6, with metabolic equivalent tasks (METs) calculated in MET-hours per week. Mammography screening was performed every 2 years or annually for the 16% who also participated in a WHI hormone therapy trial. Physicians outside the WHI directed breast cancer therapy.

The low-fat dietary program was designed to reduce fat intake to 20% of total energy and to increase vegetable, fruit, and grain intake.<sup>6,7</sup> Calorie restriction or weight loss were not intervention targets. Dietary group participants received 18 group sessions led by centrally trained certified nutritionists in year 1 and quarterly maintenance sessions throughout the dietary intervention period. Comparison group participants received written diet-related education materials. Participants provided a 4-day food record at baseline. In addition to a screening FFQ, FFQs were obtained after 1 year and thereafter from approximately one third of participants yearly in a rotating sample. Postintervention dietary intake was also based on 24-hour dietary recall obtained in a subset of 1,311 reconsenting participants. Information on adjuvant chemotherapy and radiation therapy was based on Medicare findings, whereas adjuvant endocrine therapy use was based on self-report.

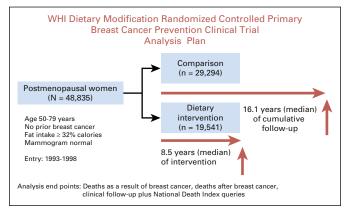
Participants with incident breast cancer continued to participate in subsequent dietary group meetings and activities. As such, dietary group participants diagnosed with breast cancer in the first year after random assignment would have approximately 1 year of nutritionist contact before and 7 years of contact after diagnosis. In contrast, participants diagnosed with breast cancer in year 8 after entry would have had more contact before and less after diagnosis. The protocol-specified trial completion date of March 31, 2005, was the end of the 8.5-year dietary intervention period when contact with study nutritionists ended. Subsequent outcome assessment required reconsent obtained from 84% versus 81% of surviving comparison and dietary group participants, respectively, for follow-up through 2010 and 86% of surviving participants from both groups for additional follow-up.

Outcome ascertainment was at 6-month intervals throughout the intervention period, with subsequent updates annually. Initial breast cancer self-reports were confirmed after medical record review by clinical center physician-adjudicators. Final adjudication was performed centrally at the Clinical Coordinating Center. Human epidermal growth factor receptor 2, ER, and PR status was determined at local laboratories. Cause of death was determined by centralized medical record or death certificate review and, in some cases, by report from participant relatives. All adjudicators were blind to random assignment. National Death Index queries completed through 2013 provided additional survival information regardless of reconsent status.

#### Statistical Analysis

The current study primary analysis end points, annualized rates of deaths as a result of and after breast cancer, during the dietary intervention period and cumulatively throughout follow-up were assessed by randomly assigned group by dividing the number of events by the corresponding person-time in each period. Incidence curves were generated, and HRs, 95% CIs, and *P* values were calculated from Cox proportional hazards regression models stratified by age at random assignment, random assignment status in the WHI hormone trials, and study period (time dependent). Analyses of deaths as a result of and after breast cancer included all 48,835 study participants measured from random assignment. The study design and the two analysis periods are outlined in Figure 1.

Because dietary effects in nine subgroups were investigated, fewer than one statistically significant interaction was expected by chance alone. The Kaplan-Meier method was used for graphical presentation of death rates. The amount of the dietary effect on death after breast cancer that



**Fig 1.** Participant flow for analyses of deaths as a result of and after breast cancer during the 8.3- and 8.5-year (mean and median) dietary intervention period and during the 16.5- and 16.1-year (mean and median) cumulative follow-up period in analyses that incorporated all 48,835 randomly assigned participants. Vertical arrows indicate the span of conducted National Death Index searches. WHI, Women's Health Initiative.

could be explained by its effect on PR status was assessed by the proportional reduction of the HR, relative to the null, after controlling for PR status. All statistical tests were two-sided. Analyses were performed with SAS 9.4 software (SAS Institute, Cary, NC). The current report updates, with long-term follow-up, dietary modification influence on deaths as a result of breast cancer and provides, for the first time in our knowledge, findings on deaths after breast cancer. These survival analyses were not protocol prespecified but provide the final and most accurate assessment of study findings.

#### **RESULTS**

Although no ongoing contact with study nutritionists occurred during postintervention, the percent energy from fat remained modestly lower in the dietary group for several years.<sup>3</sup> The participant flow for those diagnosed with breast cancer during the dietary intervention period and throughout the cumulative follow-up period is outlined in Appendix Figure A1 (online only).

For a subset of participants (n = 1,764) with breast cancer diagnosed during the dietary intervention period, selected intervention-related characteristics are outlined at baseline and serially in Appendix Table A1 (online only). Dietary intakes by randomly assigned group at baseline and 1 year in participants with breast cancer were similar to those in the overall study population, and significant differences remained between groups after breast cancer diagnosis for changes in percent energy from fat ( $-11.5\% \ v - 1.5\%; P < .001$ ) and intake of fruits, vegetables, and grains (P < .001) on the basis of the first FFQ obtained 1.6 years (mean) after diagnosis. Recreational physical activity after breast cancer, available in a subset of 330 participants, was 12.8 versus 10.8 MET-hours/week (mean) for the dietary and comparison groups, respectively (P = .41). Physical activity was not assessed postintervention.

As reported previously, during the 8.5-year dietary intervention period, 8% fewer breast cancers were observed in the dietary group.  $^{1,3}$  In the current report, of all 3,030 breast cancers diagnosed throughout the 16.1 years of cumulative follow-up, the HR for breast cancer incidence was 0.97 (95% CI, 0.90 to 1.04; Fig 1). Most breast cancer characteristics were balanced between randomly assigned groups. Similar to findings during the dietary intervention period, a larger fraction of cancers in the dietary group were ER+ and PR- (P = .05; Table 1).

During the 8.5 years of dietary intervention, with 1,764 incident breast cancers, fewer deaths occurred as a result of breast cancer in the dietary group, which was not statistically significant (27 deaths [0.016% per year] v 61 deaths [0.024% per year]; HR, 0.67; 95% CI, 0.43 to 1.06; P = .08; Figs 2 and 3A). During the same period, deaths after breast cancer (n = 134) were significantly reduced (40 deaths [0.025% per year] v 94 deaths [0.038% per year]; HR, 0.65; 95% CI, 0.45 to 0.94; P = .02) by the dietary intervention (Fig 3B). During the 16.1 years of cumulative followup with 3,030 incident breast cancers, somewhat fewer deaths occurred as a result of breast cancer in the dietary group (111 deaths [0.035% per year] v 185 deaths [0.039% per year]; HR, 0.91; 95% CI, 0.72 to 1.15; Fig 4A), and deaths after breast cancer were significantly reduced (234 deaths [0.085% per year] v 443 deaths

**Table 1.** Characteristics of Invasive Breast Cancer by Randomly Assigned Group in the Women's Health Initiative Dietary Modification Trial During the 16.1-Year (Median) Cumulative Follow-up Period

	Intervention (n = 1,176)		Comparison (n = 1,854)		
Characteristic	No.	%	No.	%	P*
Histology					.89
Ductal	752	64.4	1,202	65.3	
Lobular	116	9.9	177	9.6	
Ductal and lobular	155	13.3	248	13.5	
Other	145	12.4	213	11.6	
ER status					.59
Positive	945	85.0	1,499	85.7	
Negative	167	15.0	250	14.3	
PR status					.08
Positive	812	73.9	1,226	70.9	
Negative	287	26.1	504	29.1	
ER/PR status					.05
ER+, PR+	793	73.6	1,200	70.5	
ER+, PR-	141	13.1	281	16.5	
ER-, PR-	144	13.4	221	13.0	
HER2 overexpression	146	15.3	202	13.7	.28
Triple-negative tumor	76	8.1	133	9.1	.37
Stage					.90
Local	870	75.5	1,371	75.2	
Regional or distant	283	24.5	451	24.8	
Grading					.40
Well differentiated	309	29.0	486	28.7	
Moderately differentiated	454	42.6	772	45.7	
Poorly differentiated	302	28.4	433	25.6	
Tumor size, cm					.98
< 1	307	28.6	504	29.6	
1 to < 2	482	44.8	727	42.7	
≥ 2	286	26.6	472	27.7	
Positive lymph nodes					.98
None	789	75.5	1,252	75.8	
1 to 3	195	18.7	296	17.9	
≥ 4	61	5.8	103	6.2	

Abbreviations: ER, estrogen receptor; HER2, human epidermal growth factor receptor 2; PR, progesterone receptor.

[0.11% per year]; HR, 0.82; 95% CI, 0.70 to 0.96; P = .01) in the dietary group (Fig 4B).

The HR for deaths after breast cancer was unchanged by the addition of baseline weight and time-dependent weight change to the analysis. Analyses that added a time-dependent variable for PR status resulted in an HR of 0.87 (95% CI, 0.74 to 1.02) during the cumulative follow-up period, which corresponded to a 29% explanation of the benefit observed in the unadjusted analyses (Figs 3B and 4B).

Breast cancer therapy information from Medicare was available in 1,403 of 3,030 participants; no significant differences between the dietary and comparison groups were seen for radiation therapy (64%  $\nu$  65%, respectively) or adjuvant chemotherapy (25%  $\nu$  27%, respectively). Endocrine adjuvant therapy, which was based on self-report (n = 673), also was similar in the randomly assigned groups (71%  $\nu$  68% users, respectively).

Two of nine subgroups examined for dietary modification influence on deaths after breast cancer were significant at the .05 level

<sup>\*</sup>P values are based on  $\chi^2$  tests of association for most tumor characteristics. Trend tests were used for ordinal tumor characteristics: grading, tumor size, and positive lymph nodes.

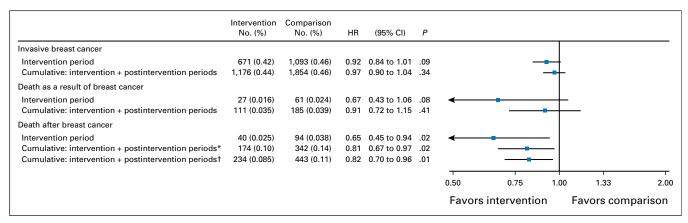


Fig 2. Breast cancer results by study period. Forest plot and summary statistics of the dietary modification influence on invasive breast cancer incidence, deaths as a result of breast cancer, and deaths after breast cancer by study period (dietary intervention period, 8.5 years; cumulative follow-up [intervention + postintervention] period, 16.1 years). (\*) Includes deaths as a result of any cause that occurred after breast cancer diagnosis during the intervention period only. (†) Includes deaths as a result of any cause that occurred after breast cancer diagnosis during the intervention period.

(waist circumference and percent energy from fat [on the basis of case-only analysis in participants with a 4-day food record baseline assessment]; Fig 5). In participants with breast cancer, information

on cause of death was available in 668 of 677. Breast cancer was the most common cause (38%; n = 257) followed by other cancers (20%; n = 132) and cardiovascular disease (18%; n = 120).

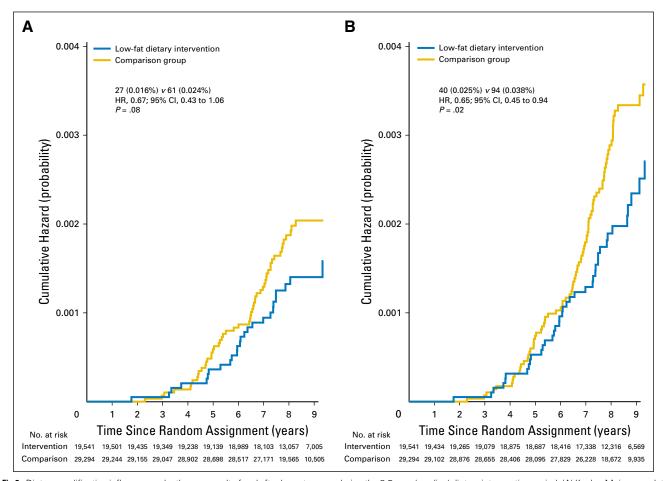


Fig 3. Dietary modification influence on deaths as a result of and after breast cancer during the 8.5-year (median) dietary intervention period. (A) Kaplan-Meier cumulative hazard estimates for death as a result of breast cancer (breast cancer followed by death attributed to the cancer) during the 8.5-year (median) dietary intervention period and (B) Kaplan-Meier cumulative hazard estimates for death after breast cancer (breast cancer followed by death as a result of any cause) during the dietary intervention period among all 48,835 trial participants, with 1,764 breast cancers measured since random assignment. Summary statistics are from a Cox proportional hazards regression model stratified by age-group and random assignment in the hormone therapy trials. The P value corresponds to a two-sided score (log-rank test). HR, hazard ratio.

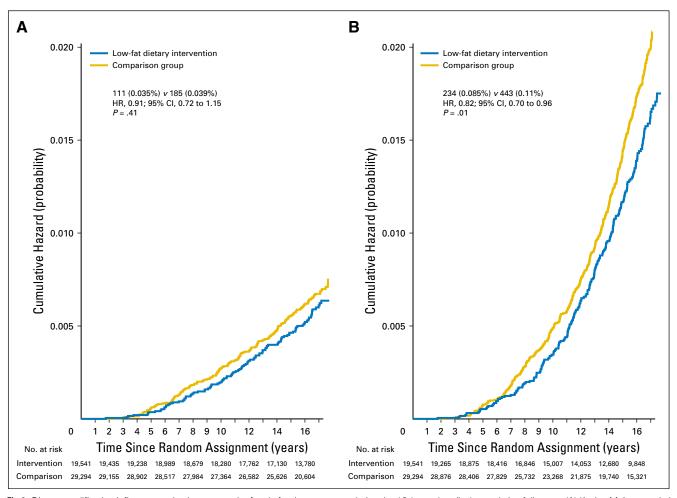


Fig 4. Dietary modification influence on deaths as a result of and after breast cancer during the 16.1-year (median) cumulative follow-up. (A) Kaplan-Meier cumulative hazard estimates for death as a result of breast cancer during the 16.1-year (median) cumulative follow-up and (B) Kaplan-Meier cumulative hazard estimates for death after breast cancer during the 16.1-year (median) cumulative follow-up among all 48,835 trial participants, with 3,030 breast cancers measured since random assignment. Summary statistics are from a Cox proportional hazards regression model stratified by age-group, random assignment in the hormone therapy trials, and study period (time dependent). The *P* value corresponds to a two-sided score (log-rank test). HR, hazard ratio.

#### DISCUSSION

With long-term follow-up of the WHI DM trial, deaths after breast cancer were significantly reduced in the dietary group both during the dietary intervention period and throughout the 16.1-year cumulative follow-up period. The lower risk of poor prognosis, ER+, PR- breast cancers <sup>10</sup> in the dietary group contributed to the favorable dietary effect on death after breast cancer. The findings on deaths as a result of and after breast cancer presented here arose from standard intention-to-treat comparisons between randomly assigned groups in analyses that included all 48,835 participants. WHI DM is a primary prevention trial; however, in adjuvant breast cancer trials, death as a result of all causes measured from diagnosis (breast cancer overall survival) is considered an important clinical end point. <sup>11,12</sup> An additional factor that potentially influenced deaths after breast cancer could be a favorable dietary influence on mortality as a result of other causes, including cardiovascular disease.

Although fewer deaths occurred as a result of breast cancer in the dietary group, the difference was not statistically significant and perhaps was limited by event number. However, the HR for dietary influence on deaths as a result of breast cancer during the dietary intervention period was comparable to that seen for deaths after breast cancer (HR, 0.67 and 0.65, respectively). In addition, accurate determination of cause of death after breast cancer can be difficult given the potential interaction between common comorbidities and cancer therapies, such as cardiac disease after anthracycline, trastuzumab, or radiation therapy. Thus, the actual breast cancer mortality rate likely lies between the attributed rate, on the basis of medical record review, and the rate that considered all deaths after breast cancer.

The subgroup analyses identified postmenopausal women as more likely to benefit from the dietary intervention. An intervention effect on deaths after breast cancer was more likely if a woman had either a lifestyle ( $\geq$  36.8% energy from fat) or a consequence of lifestyle ( $\geq$  88 waist circumference) that has been associated with adverse breast cancer outcome before dietary modification. Because subgroup analysis identified no benefit on deaths after breast cancer for dietary group participants, with < 27.9% calories from fat at entry, a notional threshold for effect may be > 28% calories from fat, which is not far from the current average US intake of postmenopausal women of 33%. A modest reduction in fat intake with minimal weight loss represents an easily achievable goal by many.

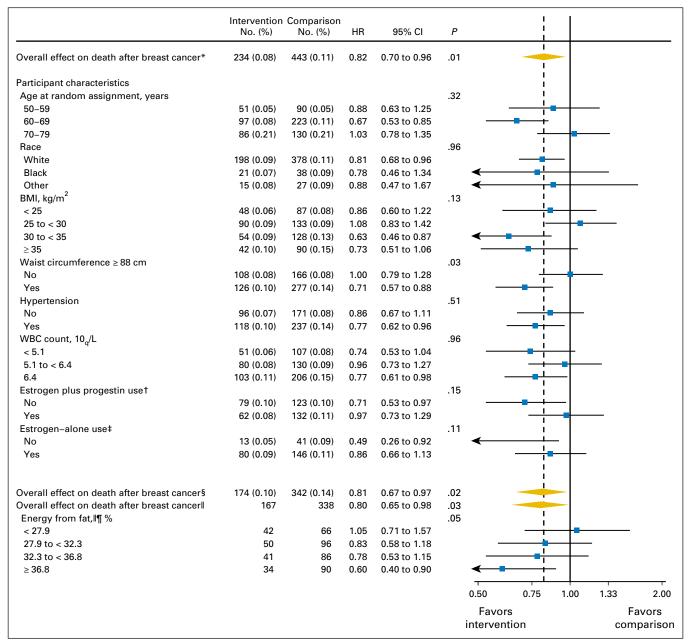


Fig 5. Subgroup analysis by a forest plot of hazard ratios (HRs), dietary intervention versus comparison, for death after breast cancer by select subgroups of participants with an invasive breast cancer diagnosis that occurred anytime throughout the cumulative follow-up period. Annualized percentages are shown for all subgroups except percent energy from fat, where the numbers of deaths with 4-day food records were reported. Because of expense, baseline food records were scored for only breast cancers that occurred during the intervention period. For this analysis HRs were estimated in a case-only analysis by using logistic regression of random assignment stratified by the four ordinal groups of dietary fat. For rare outcomes (eg, < 5% incidence during study follow-up), case-only analyses provided HRs and corresponding 95% CIs that were essentially equivalent to those that would arise if the subgroups were available for the entire randomly assigned cohort. To illustrate, the corresponding case-only estimate for the overall effect on death after breast cancer (HR, 0.80) and the usual Cox proportional hazards regression estimate that includes deaths as a result of any cause that occurred after breast cancer diagnosis during the intervention period only (HR, 0.81) are provided. P values correspond to a test of the interaction between randomly assigned group and subgroup. (\*) Includes deaths as a result of any cause that occurred after breast cancer diagnosis during the intervention or postintervention periods. (†) Among participants who did not report a prior hysterectomy. Ever used estrogen plus progestin includes random assignment into the Women's Health Initiative conjugated equine estrogens plus medroxyprogesterone acetate trial. (‡) Among participants who reported a prior hysterectomy. Ever used estrogen alone includes random assignment into Women's Health Initiative conjugated equine estrogens alone trial. §Includes deaths as a result of any cause that occurred after breast cancer diagnosis during the interv

A substantial number of observational studies have addressed the issue of dietary intake and breast cancer incidence and outcome, but the findings were mixed, 15-17 with both negative 18 and positive 19 findings that perhaps reflected observational study limitations. These limitations include dietary intake measurement error, limited variation in dietary intake within study populations, and common reliance

on a single dietary intake made before diagnosis. The uncertainty raised by the observational study findings provided a major impetus to the conduct of the current randomized clinical trial.

The WHI low-fat dietary program reduced fat and increased fruit and vegetable intake, changes that potentially influence breast cancer, <sup>20-22</sup> and their effects on study findings cannot be separated. Although adjustment for weight change during the intervention did not attenuate the dietary effect on death after breast cancer, the WHI DM trial evaluated a low-fat eating pattern, not the influence of any individual factor on study outcomes. With respect to other lifestyle factors, when measured in a subgroup, physical activity did not significantly differ between randomly assigned groups.

At entry, women in the WHI DM trial had no prior breast cancer and a normal screening mammogram. Nonetheless, results from a modeling exercise suggest that the WHI dietary intervention may have primarily affected undiagnosed preclinical breast cancers. The model<sup>23</sup> incorporates cancer doubling time, mammography detection threshold, and subclinical breast cancer prevalence. When applied to the WHI estrogen plus progestin trial, which entered a similarly aged population, after a mean 5.6 years of follow-up, 94% of the breast cancers were estimated to be preclinical cancers, with only 6% de novo cases.<sup>23</sup>

In the WHI DM trial, after 16.1 years of follow-up, no overall reduction in breast cancer incidence was seen, but a smaller fraction of cancers in the dietary group were ER+, PR-, which explains 29% of the difference in deaths after breast cancer. Overall, 9% fewer deaths occurred as a result of breast cancer (n = 296; P = .41), and 18% fewer deaths occurred after breast cancer (n = 677; P = .01) in the dietary group. One interpretation could be that dietary intervention–reduced breast cancers are more likely to be fatal and favorably influence other causes of death that may or may not have been specific to the cancer therapy experience of patients with breast cancer. Because the WHI low-fat dietary intervention had at least short-term effects of lowering systolic blood pressure, reducing metabolic syndrome risk, and lowering cholesterol-targeted and hypertension medicine,  $^{24,25}$  it could have had an influence on other chronic diseases.

The WHI dietary intervention findings can be compared with three randomized primary breast cancer prevention trials that evaluated tamoxifen. <sup>26-28</sup> In the P-1 trial in 13,388 participants, after a median 7 years of follow-up, tamoxifen reduced breast cancer incidence by 43%, but 10% more deaths occurred after breast cancer in the tamoxifen group (n = 240; not significant).<sup>26</sup> In the Royal Marsden Prevention Trial in 2,471 participants, after a median 13 years of follow-up, breast cancer incidence was somewhat lower in the tamoxifen group (22% lower, P = .1), with 27 deaths after breast cancer in each randomly assigned group.<sup>27</sup> In the International Breast Cancer Intervention Trial 1 in 7,154 participants, after a median 16 years follow-up, tamoxifen reduced breast cancer incidence by 29% (P < .001), but 19% more deaths occurred as a result of breast cancer (n = 57; P = .8) and 10% more deaths after breast cancer (n = 389; P = 0.4) in the tamoxifen group.<sup>28</sup> Favorable effects on mortality may emerge from these trials with longer follow-up, and differences in study populations and conduct exist compared with the WHI DM trial. However, one interpretation of the current tamoxifen findings could be that

tamoxifen reduced breast cancers that were less likely to be fatal and did not favorably influence other causes of death.<sup>29</sup>

To our knowledge, only two other full-scale randomized trials examined dietary change and breast cancer incidence. A low-fat intervention was evaluated in 4,690 women with high breast density but found no influence on breast cancer incidence (HR, 1.19; 95% CI, 0.91 to 1.55). Compared with the WHI DM, the study population was younger (mean age, 47 years), commonly premenopausal, and had lower a lower body mass index (mean, 23.3 kg/m²). In the Prevención con Dieta Mediterránea trial, 4,152 postmenopausal women at high cardiovascular risk were randomly assigned to two Mediterranean diets or a control condition (advised to reduce dietary fat). As a secondary end point, breast cancer incidence (n = 35) was significantly lower in one Mediterranean diet compared with the control condition. However, control group counseling was limited, and participants did not materially reduce fat intake.

Strengths of the current study include the randomized design, a large population of 48,835 postmenopausal women, 3,030 incident breast cancers confirmed by medical record review and 677 deaths after breast cancer, serial screening mammography, and long-term follow-up. National Death Index supplemental follow-up provided additional survival information independent of reconsent status. Although no established biomarkers existed for dietary fat intake, the sustained lower body weight in the dietary group supports dietary intake differences between randomly assigned groups. Study limitations are associated with post hoc analyses and lack of comprehensive breast cancer therapy information.

In summary, women randomly assigned to a low-fat dietary pattern had a nonsignificantly lower risk of death as a result of breast cancer and a significantly reduced risk of death after breast cancer. Future studies of other lifestyle interventions on breast cancer incidence and outcome could incorporate some form of a low-fat dietary pattern as a base.

# AUTHORS' DISCLOSURES OF POTENTIAL CONFLICTS OF INTEREST

Disclosures provided by the authors are available with this article at jco.org.

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Financial support: Rowan T. Chlebowski

**Administrative support:** Rowan T. Chlebowski, Garnet L. Anderson, Barbara V. Howard, Ross L. Prentice

**Provision of study materials or patients:** Rowan T. Chlebowski, Garnet L. Anderson, Cynthia A. Thomson, JoAnn E. Manson, Dorothy Lane, Ross L. Prentice

**Collection and assembly of data:** Rowan T. Chlebowski, Garnet L. Anderson, JoAnn E. Manson, Barbara V. Howard, Ross L. Prentice

Data analysis and interpretation: All authors

Manuscript writing: All authors

Final approval of manuscript: All authors

Accountable for all aspects of the work: All authors

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#### **AUTHORS' DISCLOSURES OF POTENTIAL CONFLICTS OF INTEREST**

#### Low-Fat Dietary Pattern and Breast Cancer Mortality in the Women's Health Initiative Randomized Controlled Trial

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#### **Appendix**

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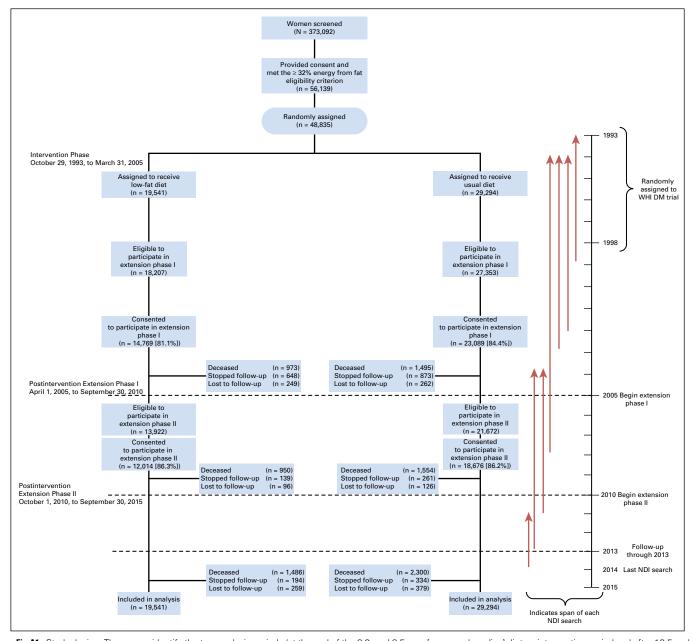


Fig A1. Study design. The arrows identify the two analysis periods (at the end of the 8.3- and 8.5-year [mean and median] dietary intervention period and after 16.5 and 16.1 [mean and median] years of cumulative follow-up). NDI, National Death Index; WHI DM, Women's Health Initiative Dietary Modification.

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 Table A1.
 Intervention-Related Characteristics of 1,764 Women Diagnosed With Breast Cancer During the Dietary Intervention Period in the Women's Health Initiative Dietary Modification Trial\*

Characteristic	Intervention (n = 671)		Comparison $(n = 1,093)$		
	Mean	SD	Mean	SD	Р
Baseline					
Age at screening, years	62.7	6.7	62.8	6.7	.71
Physical functioning (RAND 36-Item Health Survey)	82.1	19.0	80.9	18.9	.18
Total energy expenditure from physical activity, MET-h/week	10.2	11.6	9.3	11.2	.12
Gail model 5-year risk	1.9	1.1	1.9	1.2	.22
Intervention related at baseline					
Baseline weight, kg	77.5	16.5	77.6	15.8	.96
Dietary energy, kcal	1,781.8	694.4	1,769.5	675.0	.71
Energy from fat, %	37.6	4.9	37.8	5.0	.38
Fruit and vegetable consumption, median servings/day	3.8	1.8	3.6	1.8	.08
Grains, median servings/day	4.7	2.5	4.7	2.4	.69
Intervention related at year 1†					
Change in weight, kg	-3.2	9.2	-0.1	7.6	< .001
Change in energy intake, kcal/day	-282.9	611.1	-173.3	566.2	< .001
Change in energy from fat, %/day	-13.9	7.5	-2.5	6.1	< .001
Change in fruit and vegetable consumption, median servings/day	1.5	2.1	0.2	1.5	< .001
Change in grains consumption, median servings/day	0.4	2.8	-0.5	2.2	< .001
Intervention related at last measure before breast cancer diagnosis					
Time to diagnosis, years	4.5	2.5	4.5	2.5	.87
Age at breast cancer diagnosis, years	67.5	6.9	67.6	6.9	.66
Change in weight, kg	-1.2	10.0	0.5	8.5	< .001
Change in energy intake,‡ kcal/day	-297.0	596.2	-184.9	600.0	< .001
Change in energy from fat,‡ %/day	-11.5	8.2	-1.5	6.9	< .001
Change in fruit and vegetable consumption,‡ median servings/day	1.4	2.2	0.2	1.7	< .001
Change in grains consumption, median servings/day	-0.1	2.7	-0.7	2.3	< .001

Abbreviations: MET, metabolic equivalent task; SD, standard deviation.
\*Gail MH, Costantino JP, Pee D, et al: Projecting individualized absolute invasive breast cancer risk in African American women. J Natl Cancer Inst 99:1782-1792, 2007 †Change from baseline to year 1 (n = 1,591).

<sup>‡</sup>Change from baseline to last measure before breast cancer diagnosis (n = 1,530 for change in dietary characteristics; 87% of breast cancers).