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# An Agent-Based Model of School Closing in Under-Vaccinated Communities During Measles Outbreaks 

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#### Abstract

The winter 2014-15 measles outbreak in the US represents a significant crisis in the emergence of a functionally extirpated pathogen. Conclusively linking this outbreak to decreases in the measles/mumps/rubella (MMR) vaccination rate (driven by anti-vaccine sentiment) is critical to motivating MMR vaccination. We used the NOVA modeling platform to build a stochastic, spatially-structured, individual-based SEIR model of outbreaks, under the assumption that $R_{0} \approx 7$ for measles. We show this implies that herd immunity requires vaccination coverage of greater than approximately $85 \%$. We used a network structured version of our NOVA model that involved two communities, one at the relatively low coverage of $85 \%$ coverage and one at the higher coverage of $95 \%$, both of which had 400 -student schools embedded, as well as students occasionally visiting superspreading sites (e.g. high-density theme parks, cinemas, etc.). These two vaccination coverage levels are within the range of values occurring across California counties. Transmission rates at schools and superspreading sites were arbitrarily set to respectively 5 and 15 times background community rates. Simulations of our model demonstrate that a 'send unvaccinated students home' policy in low coverage counties is extremely effective at shutting down outbreaks of measles.


## Author Keywords

Superspreaders, Epidemiological models, NOVA software platform, Herd immunity, R-zero.

## ACM Classification Keywords

I.6.1 SIMULATION AND MODELING

## 1. INTRODUCTION

A recent reemergence of measles in the United States provides us with an opportunity to think more deeply about vaccinations in modern societies and their role in providing protection against viral diseases. Measles is caused by a Morbillivirus (Family: Paramyxoviridea) that is expelled from an infected individual during coughing or sneezing. The virus can remain infectious for several hours in the form of aerosolized droplets or fomite residues [11]. As a consequence of near life-long immunity in individuals that have recovered from the disease, prior to large-scale vaccination programs in Europe and the US, measles outbreaks were seasonally linked to new cohorts of young children entering school for the first time [13]. The endemicity of this process in developed-world cities with relatively large populations (at least $250,000-500,000$ individuals) has been explained with the aid of mathematically sophisticated models [6]. Such seasonal outbreak patterns were eliminated in the US after 1981, through the implementation of the highly effective MMR (measles, mumps and rubella) blanket vaccination program [21].
The principle of vaccinating a population beyond its 'herd immunity' threshold lies at the heart of the continued success of the US MMR vaccination program [36, 4]. The threshold ratio of vaccinated to unvaccinated individuals for a herdimmune population-that is, a population where this ratio is sufficiently high to cause infections to rapidly fade-out rather than break-out-depends upon the basic reproductive number, $R_{0}$ (the expected number of cases caused by the index case) for the disease. The value of $R_{0}$ itself is influenced by the infectiousness of the pathogen, the lengths of time that infected individuals are infectious, and contact rates among individuals within the community. Thus, in essence, the efficacy of a vaccination program depends on both complex immunological and sociological processes.

The immunological process component involves optimal ages and dosages at which to administer vaccines of various kinds (e.g. attenuated live vaccine, as in MMR, or inactivated virus, as in Polio iPV), or segments of virus (influenza injections),
one dose, or primary dose plus boosters, and so on. MMR vaccinations may fail if the primary dose is given to infants protected by maternal antibodies transferred during breastfeeding [2]. Avidity testing can be used to assess the efficacy of vaccinations, but such tests are not widely utilized in measles epidemiology [32]. Thus, it can be misleading to take statistical data at face value regarding assumed levels of protection for given levels of a vaccination coverage [29].
The sociological process component enables misinformation to shape outbreak dynamics. For example, influenced by a now retracted study in The Lancet in 1998 that linked the MMR vaccine to autism [34], a significant minority of parents, clustered in particular geographic regions around the US, refuse MMR vaccinations for their children [23]. Even after the retraction of the study, these groups persist, and despite major efforts to remediate the damage, recent results confirm that these self-proclaimed "anti-vaxxers" respond negatively to educational campaigns, becoming more staunchly opposed to vaccination [30]. As a consequence of anti-vaxxer groups, including those opposed for religious purposes (e.g., the Amish in Ohio), measles outbreaks are now more likely in the US: the US Center for Disease Control (CDC) reported 911 cases for the decade 2001-2011 [1] (which is fewer than 8 cases per month), while close to 650 cases were reported in 2014, dominated by an outbreak in the Amish community in Ohio, and more than 100 for the month of January in 2015, dominated by cases in California linked to the so-called Disneyland outbreak [28, 8]. Thus, the sociological component includes a growing 'small world' phenomenon, with individuals making contact at high density entertainment venues that increasingly draw in patrons from considerable distances, and then serve as superspreading centers for highly contagious respiratory diseases, such as measles. Superspreading, defined as a process whereby a few individuals cause a disproportionately large number of secondary infections, has been considered in terms of individual and environmental heterogeneity [31]. Individuals with high pathogen shedding rates [26] or longer periods of infectiousness [25] may lead to greater numbers of secondary infections. Similarly, superspreading centers can be viewed as discrete environmental patches that provide opportunities for significantly higher contact rates than predicted by expected movement patterns [20].

Spatial heterogeneity in the effective vaccination rates alters the likelihood with which outbreaks occur. While herd immunity levels may exist at some of the originating sites of individuals visiting a superspreading center, unvaccinated individuals from those sites may carry the disease back to keep feeding a continuous low-level stream of cases at sites with vaccination rate above herd-immunity levels, as well as starting self-sustaining outbreaks at sites with vaccination rates below levels of herd immunity. Here we evaluate the efficacy of "stay-at-home regulations" for children who are not vaccinated in schools were outbreaks occur.

## 2. MODEL

The prevailing paradigm for modeling epidemics is to use systems of deterministic or stochastic differential or difference equations that divide the population into disease (e.g. susceptible, infected, recovered, immune) classes [12], as well as other demographic (e.g. age, sex) and behavioral classes (e.g. sexually active) and to fit transmission, recovery, and other relevant parameters using least-squares estimation (LSE) [3] or maximum likelihood estimation (MLE) methods [18], based on comparisons of model output and empirical data. Going back to the work of Frost and Reed, a second approach to modeling epidemics has been to follow transmission chains (incidence and offspring distributions, transmission trees and branches), generally considered within the framework of semi-Markov branching process [5]. While the first paradigm is most useful for large-scale epidemics involving infection of a significant fraction of the susceptible population (i.e., $>1 \%$ ), which is the case inter alia for influenzaHIV, tuberculosis, and measles in unvaccinated communities; the second is most useful for emerging diseases when the proportion infected is often very small (i.e., $<0.1 \%$ )-which includes inter alia, SARS [24], Ebola [15], and hantavirus, as well as measles in communities where vaccination rates are close to herd-immunity threshold levels [7, 8].
Given our interest in disease outbreak dynamics, the analysis we report here is based on a Markov chain approach, as discussed in Chowell et al. [12], but modified to allow for the implications and efficacy of key interventions that need to be evaluated. Our model of a measles outbreak in the USA explicitly includes barring unvaccinated school-aged individuals from attending schools when one or more individuals in the school have come down with the measles because of its viability as an affective outbreak mitigation strategy [9]. It also allows for a reduction in transmission rates over the course of the epidemics due to behavioral changes that reduce contact rates in the community between sick and uninfected individuals, as the community becomes more informed on how to interact with infectious individuals [22].

### 2.2 Transmission chain formulation

Most epidemiological models begin by dividing the population into 'susceptable' (S) and 'infected+infectious' (I) individuals, while elaborations discriminate between 'exposed but not yet infectious' (E) and 'infectious' (I), as well as 'removed' (R) [19], where the latter can be broken down into 'dead' (D) and 'recovered with some level of acquired immunity' (V) [16]. We emphasized here that we need not specify the size of the S-class involved, but rather assume some level of 'risk-of-infection' that is proportional to the number of infected individuals in the subpopulation of interest, where subpopulations form an interconnected meta-population, and 'risks-of-infection are subpopulation dependent. Thus we model the disease incidence rate in time period $[t, t+1]$ ( $t$ in our case will represent days) in subpopulation $j$ containing $I_{j}(t)$ infectious individuals, using Monte Carlo methods, from a Poisson distributions with mean

$$
\begin{equation*}
m_{j}(t)=\lambda_{j} I(t) \tag{1}
\end{equation*}
$$

to generate $E_{i}(t+1)$. (Note: Roman font $\mathrm{S}, \mathrm{I}$, and E name the class, while italic font $I_{j}(t)$ and $E_{j}(t)$ refer to the number of individuals at time $t$ in so-named class of subpopulation $j$ ).
Ignoring the population designation index $j$ for the moment, the more usual approach to characterizing transmission is to assume that, for some 'transmission intensity constant' $\beta>0$, the incidence rate is determined by the expression $m(t)=\beta S(t) I(t)$ in the case of density dependent transmission and by $m(t)=\beta S(t) I(t) / N(t)$ in the case of frequency-dependent transmission [17, 16, 27]. This approach has been generalized to assume that, for some population scaling constant $K>0$, transmission is more generally characterized by the following function, which includes both density-dependent $(K \rightarrow \infty)$ and frequency-dependent ( $K \rightarrow 0$ with appropriate rescaling of $\beta$ ) transmission as limiting cases:

$$
\begin{equation*}
m(t)=\beta S(t) \frac{I(t)}{1+N(t) / K} \tag{2}
\end{equation*}
$$

This latter characterization requires that both $S$ and $N$ are known when, by analogy (i.e., comparison of Eqns. 3 and 2 , with index $j$ ignored), the values of $S$ and $N$ are used to determine the value

$$
\begin{equation*}
\lambda(S, N)=\frac{\beta S}{1+N / K} \tag{3}
\end{equation*}
$$

An alternative approach is use the fact that when outbreaks occur $I \ll S$, or equivalently, $S \approx N$, so that $\lambda \approx$ $\frac{\beta}{1 / N+1 / K}$, or when $N \gg K, \lambda \approx K \beta$. Thus, in relatively large populations, the incidence rate becomes densityindependent and can be modeled by a stochastic Poisson process, following the approach we take below.

The model focuses on individuals in the population that become exposed to an individual in state I at time $t$, and follows their progress over time as they make transitions from states E to I to R (most of which are now in state V-i.e., immune, but a small percent transition to $\mathrm{D}-\mathrm{i} . \mathrm{e}$, death). In a totally naïve population, each individual in state I at time $t$ generates on average $\lambda$ individuals in state E at time $t+1$, using Poisson statistics. In a population in which a proportion $p_{v}$ are vaccinated, this expected number is modified by the value $\left(1-p_{v}\right)$ : specifically, if the Poisson drawing yields $r$ individuals to be assigned the state E, then each of these individuals is moved to state V with probability $p_{v}$. Thus, in short, the incidence rate at time $t+1$ will follow a Poisson distribution with expected value $\left(1-p_{v}\right) \lambda I(t)$. If we use the time variable $s$ to denote individuals at time $t$ that became infected with the disease at time $(t-s)$, then letting $U(t, s)$ represents the number of exposed individuals at time $t$ who were exposed $s$ days ago, it follows that the dependent variable $U(t, 0)$ is the incidence at time $t$ : i.e.,

$$
\begin{equation*}
U(t+1,0)=\left(1-p_{v}\right) \lambda I(t) \tag{4}
\end{equation*}
$$

For the sake of simplicity, assume the latent period $\left[0, s_{1}\right]$ is the same for all individuals. Similarly, assume all individuals have the same infectious period $\left[s_{1}, s_{1}+s_{2}-1\right]$. Thus, for all individuals exposed to the pathogen at any time $t=1,2, \ldots$, $s_{1}>0$ and $s_{2}>0$ are constants rather than random variables
across individuals. It then follows that these individuals are infected (exposed) but not yet infectious for the first $s_{1}-1$ periods of time, becoming infectious $s_{1}$ units after first exposure, and that they remain infections for $s_{2}-1$ periods of time. Thus all individuals make the transition to R status $s_{1}+s_{2}$ periods of time after first exposure. Under these assumptions of constant latent and infectious periods, and assuming that individuals enter the R state only after $s_{2}$ units of time, it follows that

$$
\begin{equation*}
I(t)=\sum_{k=1}^{s_{2}} U\left(t, s_{1}+k-1\right) \tag{5}
\end{equation*}
$$

The most efficient implementation of a homogeneous version of the model is to aggregate individuals by states

$$
Q_{i}(t) \equiv U(t, i-1), i=1, \ldots, s_{1}+s_{2}
$$

(i.e., time units since the exposure/transmission), and model how numbers in each state change over time using the following equations (where the notation $x \sim \operatorname{POISSON}[m]$ implies a random drawing from a Poisson distribution with parameter $m$ ):

$$
\begin{align*}
Q_{1}(t+1) & \sim \operatorname{POISSON}\left[\left(1-p_{v}\right) \lambda \sum_{k=1}^{s_{2}} Q_{k+s_{1}-1}(t)\right] \\
Q_{i}(t+1) & =Q_{i-1}(t) \quad i=2, \ldots, s_{1}+s_{2} \tag{6}
\end{align*}
$$

### 2.3 Individual-based NOVA model

A less efficient, but more comprehensive implementation that allows us to keep track of individuals as they may move through space or exhibit variation in susceptibility, length of latent period (i.e., variation in $s_{1}$ among individuals), length of infection (variation in $s_{2}$ ), and risk of mortality while ill (not included in the above model), all possibly as functions of genetic or individual level environmental factors, is to follow the progress of each individual recruited to the population using Eqn. 4 to link individuals to parents, rather than Eqn. 5 (which only counts incidence at an aggregated group level). Here we took this individual-based approach, because we wanted to keep track of the 'next-generation distribution' to obtain estimates of $R_{0}$ from this distribution, with the details of how to do this described elsewhere [25, 15]. We also wanted to follow individuals as they are influenced by spatial factors: in our case children in the local community environments with different vaccination coverage rates and including time spent at schools where transmission rates are higher. The constant values we used for the latent and infectious period designators were $s_{1}=7$ and $s_{2}=3$ days, based on data listed elsewhere (cf. Table 3 in [10]), with the assumption of constancy due to the fact that the natural variability in these numbers across individuals is not well characterized in specific communities. There are many reasons why $s_{2}$, in particular, is not well known, including the fact that in different communities detection of disease and implementation of treatment vary greatly. Thus the true infectious period is considerably longer than $s_{2}=3$ days, but the value we use represent an estimate of the 'effective' number of days during which individuals are available to transmit to other members of their local community prior to isolation and treatment.

Aside from this, estimates of the incidence rate parameter $\lambda$ will, to a large extent, vary collinearly with values of $s_{1}$ and $s_{2}$.

### 2.4 Spatial structure and incidence rates

Many different factors affect transmission and hence incidence rates: changes in the virulence of strains over time [33], the movement behavior of individuals within populations, the size of those populations, disease detection and treatment protocols that influence how early symptoms are recognized, and how strict patient isolation practices are. It is therefore unsurprising that estimates of $R_{0}$ for highly contagious diseases such as measles can vary by several hundred percent. We based our selection for the parameter $\lambda$ introduced in the Eqn. 1 on our expectation that in an unvaccinated population the value of $R_{0}$ should be somewhere in the range $[6.2,7.7]$ [29]. Following a heuristic procedure of trying out different values of $\lambda$, with $s_{1}=7$ and $s_{2}=3$, we found that $\lambda=2.2$ yielded a value of $R_{0}=6.6$ for the case $p_{v}=0$. Further, we ran the model ten times for each of the values of $p_{v}=0.05,0.1, \ldots, 0.75,0.80$, to obtain estimates of $R_{0}\left(p_{v}\right)$ from the next generation distribution that our simulations produced after running the model for 40 days after the introduction of a single index case (i.e., patient zero) into the population.
The results of these estimates are plotted in Fig. 1. In this figure we see that the regression line through the simulation data intersects the point $\left(p_{v}, R_{0}\right)=(1,0)$ as expected (i.e., the absence of susceptible individuals necessarily implies $R_{0}=0$ ). This regression line also indicates that $p_{v}=0.85$ is the herd-immunity threshold vaccination level, because $R_{0}(0.85)=1$. This estimate could have been made directly by drawing a line from the point $(1,0)$ to $(0,6.6)$ and calculating the point on this line that would yield $R_{0}=1$ or, equivalently, using the following well-known formula for critical coverage in a homogenous, well-mixed population: $1-1 / R_{0}=1-1 / 6.6=0.85$. The data generated from our simulations, however, shows that ten runs are sufficient to provide an excellent estimate of the value for each of the points plotted in Fig. 1.


Figure 1. A regression line through estimates of $R_{0}$ obtained via simulation for populations with different levels of vaccination coverage $\left(p_{v}\right)$. Note that $p_{v}=0.85$ corresponds to $R_{0}=1$, implying that a vaccination rate of $85 \%$ coverage is needed to achieve the herd-immunity threshold.

Following the above procedure for estimating a suitable value for $\lambda$, once $s_{1}$ and $s_{2}$ had been selected from reports in the
literature, and a target value of $R_{0}$ identified, we designed a spatial configuration that could be used to test the efficacy of a 'send unvaccinated students home' policy for schools to control outbreaks in situations where the vaccination coverage in the school is relatively low. In particular, we set up two communities, one vaccinated around the herd-immunity threshold level (low coverage: 85\%) and one vaccinated well above the herd-immunity threshold level (high coverage: 95\%). For


Figure 2. The spatial structure of our stochastic model depicting probabilities of daily movements of individuals between home community and school as well as visits to superspreading sites, such as entertainment centers. The black and green transition rates correspond to the stochastic matrix $T_{0}$ provided in the text. When the purple values are substituted for the green values in the under-vaccinated community (i.e., $85 \%$ vaccination coverage) we obtain a 'send home policy' effect, modeled by stochastic matrix $T_{h}$. The implications of these transition rates for time spent at different sites are discussed in the text. The solid blue input vector represents a small probability that during the course of an outbreak infected individuals may be imported into the system from an origin other than the two explicitly modeled communities. The incidence rate parameter values, indicated in red, are $\lambda=2.2$ within the community, but 5 times this rate within schools and 15 times this rate at the superspreading site.
purposes of comparison, we refer to the data provided in Fig. 3. These data reflect a $92.3 \%$ for the recommended 2 doses of MMR. True protection is likely somewhat higher because of individual who have received a 1-dose vaccination. The values specified in Fig. 2, when infected individuals in

| All | Public | Private |  |
| :--- | :---: | :---: | :---: |
| Number of Schools | 7,684 | 5,852 | 1,832 |
| Number of Students | 533,680 | 491,905 | 41,775 |
| All Required Immunizations | $90.2 \%$ | $90.6 \%$ | $85.4 \%$ |
| Conditional Entrants | $6.5 \%$ | $6.3 \%$ | $8.5 \%$ |
| Permanent Medical Exemptions | $0.19 \%$ | $0.18 \%$ | $0.29 \%$ |
| Personal Belief Exemptions | $3.15 \%$ | $2.92 \%$ | $5.88 \%$ |
| 4+ DTP | $92.2 \%$ | $92.5 \%$ | $88.6 \%$ |
| 3+ Polio | $92.6 \%$ | $93.0 \%$ | $88.5 \%$ |
| 2+ MMR | $92.3 \%$ | $92.7 \%$ | $87.6 \%$ |
| 3+ Hep B | $94.8 \%$ | $95.0 \%$ | $91.8 \%$ |
| 1+Vari (or physician-documented disease) | $95.3 \%$ | $95.5 \%$ | $92.1 \%$ |

Figure 3. A graphical image of Table 1 from the 2013-2014 Kindergarten Immunization Assessment Results, California Department of Public Health, Immunization Branch, available from CDPH (download pdf)
under-vaccinated communities are not sent home, correspond to the transition matrix (order of variables, following Fig. 2, is 'superspreading sites', ' $95 \%$ community', 'school in $95 \%$ community', ' $85 \%$ community', 'school in $85 \%$ community

2')

$$
T_{0}=\left(\begin{array}{ccccc}
0.2 & 0.01 & 0 & 0.01 & 0 \\
0.4 & 0.74 & 0.75 & 0 & 0 \\
0 & 0.25 & 0.25 & 0 & 0 \\
0.4 & 0 & 0 & 0.74 & 0.75 \\
0 & 0 & 0 & 0.25 & 0.25
\end{array}\right)
$$

and when individuals are sent home to the matrix

$$
T_{h}=\left(\begin{array}{ccccc}
0.2 & 0.01 & 0 & 0.01 & 0 \\
0.4 & 0.74 & 0.75 & 0 & 0 \\
0 & 0.25 & 0.25 & 0 & 0 \\
0.4 & 0 & 0 & 0.94 & 0.95 \\
0 & 0 & 0 & 0.25 & 0.05
\end{array}\right)
$$

The dominant eigenvectors that characterize the stable probability distributions associated with stochastic matrices $T_{0}$ and $T_{h}$ respectively (i.e., the eigenvectors corresponding the eigenvalue that has value 1) are $(0.009,0.3715,0.124,0.3715,0.124)^{\prime} \quad$ and ( $0.010,0.415,0.138,0.415,0.022)^{\prime}$.

### 2.5 Accessing the model

As mentioned above, the model was built using the NOVA modeling platform, which is downloadable for free at the novamodeler website [35, 15, 14]. The NOVA file for the model is available at the Getz Lab Website (Nova Models download: Measles, Spring Simulation Conference, 2016). A web-based implementation of the model be accessed at the Nova OnLine Model Library (username and password are both 'numerus').

## 3. RESULTS

The first study we undertook was to identify a value for the transmission parameter $\lambda$ and then a characterize how the disease outbreak threshold, as represented by $R_{0}\left(p_{v}\right)$, varied with vaccination coverage parameter $p_{v} \in[0,1]$. We settled on the value $\lambda=2.2$, which we confirm corresponds to $R_{0}(0)=6.6$ (unvaccinated population) and $R_{0}(0.85)=1$ (herd immunity threshold for vaccination coverage), as illustrated in Fig. 1.

We then used our simulation model to compare outbreak sizes in a five compartment system that has the structure depicted in Fig. 2: viz., two communities, one with $85 \%$ vaccination coverage the other with $95 \%$ vaccination coverage, both containing schools where students experience a disease transmission hazard that is five times the background community rate and both sending individuals to superspreading sites (represented by a single spatial compartment, but may in fact be a collection of sites) for limited periods of time where the disease transmission hazard is 15 times that of the background community and, hence, three times that at schools. We note, however, that in the school environment we took account of the fact that the population size is relatively small compared to the community at large. So in the school environment we reduce the risk of transmission by a factor $\left(S_{\text {school }} / N_{\text {school }}(t)\right.$, where we set $N_{\text {school }}=400$ and $S_{\text {school }}(t)$ is the number of individuals at that school that remained susceptible in not yet being infected or sent home by time $t$ ). From the dominant eigenvector associated with the matrices $T_{0}$ we confirm that under a 'no action' policy infectious individuals in both communities spend approximately $74 \%$ of their time in the local
community/home environment (incidence rate is given by $\lambda$ ), $25 \%$ at school (incidence rate $5 \lambda$ ), and $1 \%$ at superspreading centers/sites (incidence rate $15 \lambda$ ). Similarly, from the dominant eigenvector associated with the matrix $T_{h}$ it follows that under a 'send unvaccinated students home' policy, during the course of an outbreak, infectious individuals in the first community spend their time, as specified for $T_{0}$ while, individuals in the second community now spend approximately $94 \%$ of their time at in the local community/home environment, only $5 \%$ at school, and again approximately $1 \%$ of their time at superspreading centers/sites.
We run our spatial model a hundred times for each of the two cases: 'no action' and a 'send unvaccinated students home' policy. In each case, each run terminated either with the fading of the outbreak (no infectious cases possible) or after 200 days, which ever came first. The average and standard deviation of the results obtained from each set of one hundred runs are reported in Table 1; density plots of resulting distributions of case number are illustrated in Fig. 4.

Table 1. Comparison of number of cases under 'no action' versus 'send unvaccinated students home' policy.

| Communities | No action | Send home |
| :--- | :---: | :---: |
| 85\% vaccination coverage | $348 \pm 403$ | $2.4 \pm 3.5$ |
| 95\% vaccination coverage | $42 \pm 50$ | $1.6 \pm 1.5$ |
| To superspreader site | $4.9 \pm 6.1$ | $0.2 \pm 0.6$ |

## 4. DISCUSSION

When modeling highly heterogeneous and stochastic processes, while endeavoring to maintain generalizability beyond the particular set of circumstances under considerion, a number of simplifying assumptions must be made. At each step in this process, uncertainty inevitably increases. Here, we present a model that allows us to assess the efficacy of 'send-home' policies as a way of extinguishing an outbreak of a highly transmissible disease, such as measles. This model makes underlying assumptions regarding the value of $R_{0}$, [29], vaccination rates obtained from the California Department of Health, transmission rates, various contact rates, and infectiousness period duration, amongst others. While the model was calibrated to values found in the literature, there remains a great deal of uncertainty, as each measles outbreak appears to follow its own unique set of parameter values. Nonetheless, the strength of the methods presented here lie in their ability to be easily adjusted based on information gathered from current outbreaks. Thus, while our model is undeniably a simplified realization of any outbreak, the model provides evidence for the considerable efficacy of a 'send unvaccinated students home' policy during outbreaks of measles in communities that are 'close to' versus 'well above' the herd-immunity vaccination coverage threshold.
Interestingly, our policy of sending home students during possible outbreaks had a significant impact on the number of cases resulting in communities with both relatively high ( $95 \%$ ) and relatively low ( $85 \%$ ) vaccination rates. The mean total number of cases (obtained by adding columns in Table 1) without implementing a 'send unvaccinated students home' policy was 395 , yet when 4 out of 5 unvaccinated children


Figure 4. Probability density plots of $\log$ number of cases from 100 runs of the model for each of the with and without implementation of the 'send unvaccinated students home' policy cases: A. low vaccination rate community ( $85 \%$ ); B. high vaccination rate community ( $95 \%$ ) (note: the abscissa scale is different from case A .
are sent home in the low vaccinated community alone (i.e., the policy was not applied to the high vaccinated community), the total mean number of cases dropped below 5. The mean duration of these outbreaks (data not shown in Table 1) was also significantly different between the two cases: for the no action case it was 251 days compared to 93 days when the 'send unvaccinated students home' policy was implemented in the low vaccination community alone.
The data, supported by our model, strongly suggest that the 2014-15 measles outbreak in California occurred as a result of the variable vaccination coverage across different communities and the mixing of these communities at superspreading centers. While herd immunity levels may be met at origin locations of visitors to major superspreading centers, such as Disneyland, the ephemeral populations that form each day may have a cumulative vaccination rates below the average for California as a whole. Over time, with high enough turnover rates and population sizes, rare disease recurrences are bound to be introduced when superspreading assemblages are below elevated herd immunity thresholds that are associated with high density tourist aggregation or entertainment centers, which then act as superspreading centers. Conventional wisdom in epidemiology pushes vaccination as the so-
lution to herd immunity, but emerging evidence shows that anti-MMR vaccine populations only become less likely to vaccinate after intervention campaigns [30]. Rather than attempting to reform the attitudes of the substantial minority of people who oppose vaccination for reasons varying from religious objection to distrust of the medical or political establishment, we offer an easily implemented, and politically neutral, mitigation technique. By sending students without proof of vaccination home from school at a success rate of around $80 \%$ (interventions are rarely $100 \%$ successful) when an outbreak is imminent or present, the number of resulting cases dramatically declines. This policy is effective in communities with relatively high vaccination rates ( $95 \%$ ), as well as communities near the herd immunity threshold (85\%).

While the 2014 measles epidemic has been effectively suppressed, its reappearance in the United States after apparent eradication suggests that future outbreaks should not be unexpected, nor should we be unprepared if one occurs. In addition, the methods applied here can be easily adapted to epidemics of other pathogens, particularly those that are highly transmissible, like influenza virus or coronavirus. In these cases, superspreading centers are likely to play a similarly important role in the spread of the outbreak. Gaining insight into the mechanisms underlying this process is vital, and the conceptual model outlined above offers an opportunity to do so, in addition to simulating the impact of various 'send unvaccinated students home' policies. Our empirical model of the regional processes also provides a relative risk surface for future disease outbreaks that may be especially useful in the case of another epidemic emerging from a superspreading center.

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