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## Two-Year Changes in Child Weight Status, Diet, and Activity by Neighborhood Nutrition and Physical Activity Environment

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### Abstract

**Objective**—To examine 2-year changes in weight status and behaviors among children living in neighborhoods differing on nutrition and activity environments.

**Methods**—Prospective observational Neighborhood Impact on Kids study in King County, WA and San Diego County, CA. Children 6–12 years old and a parent/caregiver completed Time 1 (n=681) and Time 2 (n=618) assessments. Children lived in neighborhoods characterized as 'high/favorable' or 'low/unfavorable' in nutrition and activity environments, respectively (four neighborhood types). Child BMI z-score and overweight/obesity status were primary outcomes, with diet and activity behaviors as behavioral outcomes.

**Results**—Adjusting for sociodemographics and Time 1 values, children living in two of the three less environmentally-supportive neighborhoods had significantly less favorable BMI z-score changes (+0.11, CI 0.01 to 0.21; +0.12, CI 0.03 to 0.21) and all three less supportive neighborhoods had higher overweight/obesity (RRs 1.41 – 1.49, CIs 1.13 to 1.80) compared to children in the most environmentally-supportive neighborhoods. Changes in daily energy intake and sedentary behavior by neighborhood type were consistent with observed weight status changes, with unexpected findings for physical activity.

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**Conclusions**—More walkable and recreation-supportive environments with better nutrition access were associated with better child weight outcomes and related behavior changes.

### Keywords

environmental factors; pediatrics; dietary recalls; accelerometry

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## Introduction

Childhood obesity remains highly prevalent in the U.S.[1] Ecological models highlight factors across levels of influence (e.g., familial, neighborhood) that may impact children's weight status and related behaviors [2–4]. Evidence linking neighborhood environment to children's weight status and eating and physical activity behaviors is more limited than for adults, is mostly cross-sectional, and is weak or inconsistent [5,6]. A systematic review examining associations between nutrition environment and childhood obesity indicated that most studies found no significant associations between supermarket or fast food restaurant availability in the home neighborhood and obesity among children, although associations were more common within lower income neighborhoods or lower income families [7]. Another review found more consistent associations between neighborhood built environments that promote walkability and lower child weight status [8]. However, a recent meta-analysis indicated that walkable built environments are more often negatively related to younger children's physical activity and more supportive of adolescents' physical activity [9]. While most home neighborhood environment studies focus on either nutrition or activity environments [7], the baseline Time 1 cross-sectional component of our Neighborhood Impact on Kids (NIK) Study indicated that children living in neighborhoods with environments favorable for both physical activity and nutrition (i.e., the High Physical Activity Environment and Nutrition Environment or High PAE/High NE neighborhoods) had the lowest rates of overweight/obesity [10] compared to other neighborhood types.

More longitudinal studies are needed to examine whether children's home neighborhood environments have lasting effects on the trajectory of children's weight status and related behaviors over time. Among the few existing prospective studies examining such associations, Leung and colleagues found that having more convenience stores in home neighborhoods was related to higher weight status 3 years later among girls, whereas more produce outlets was related to lower overweight/obesity in the future [11]. Others found that girls with more supermarkets in their home neighborhoods had lower BMI and lower odds of obesity 3 years later, but having more proximal limited-service restaurants was related to higher child weight status in the future [12]. Other research examined associations between a child's home neighborhood physical activity environment and weight status, with individual studies finding that more playable (e.g., proximal park space), walkable, and safer (e.g., further distance from high traffic streets) environments were related to healthier children's weight status trajectories [13–15]. These prospective studies appear promising, but there remain few such studies, most have not examined combinations of nutrition and activity environments, and studies have not included robust measures of children's diet and activity change.

The present analysis examined 2-year changes in child weight status and weight-related behaviors among children in the Neighborhood Impact on Kids study living in neighborhoods selected to differ on both nutrition and activity environments. We hypothesized that children living in the most favorable neighborhood environments for both nutrition and physical activity would have healthier weight status and behavioral changes over time, even after adjusting for neighborhood, household, and individual sociodemographic factors.

## Method

### Neighborhood selection

Details about neighborhood selection are provided elsewhere [10,16]. Neighborhoods were defined as census block groups in King County (Seattle area), WA and San Diego County, CA. Activity and nutrition environment data were collected (e.g., park quality, presence of supermarkets) or captured from existing spatial data (e.g., street network). Neighborhoods considered high (i.e., more favorable or supportive) if physical activity environments (PAE) were higher than the corresponding county median summed z-score values of residential density, retail floor area ratio, land use mix, street connectivity (i.e., walkability [17]) and had at least one high quality park. Quality of parks was evaluated through the Environmental Assessment of Public Recreation Spaces (EAPRS) direct observation tool [18], with high quality parks defined as those with EAPRS scores > 200 (for the EAPRS tool and scoring see <http://www.seattlechildrens.org/research/child-health-behavior-and-development/saelens-lab/measures-and-protocols/>). Low PAE neighborhoods were below county median score on walkability and had no parks within the block group or within the ¼ mile buffer around the block group. Block groups were buffered to capture facilities such as parks that were just beyond their boundaries and easily accessible by residents near the boundary. High nutrition environment (NE) neighborhoods had a supermarket in the block group or nearby (<½ mile buffer) and fewer fast food restaurants ( 16 for King County or 31 for San Diego County). Low NE neighborhoods had either no supermarket nearby or had a proximal supermarket but many fast food restaurants (>16 for King County; >31 for San Diego County). The criteria for fast food restaurant differed by county because San Diego County had more than twice the number of fast food restaurants than King County, and the specific cut-offs were based pragmatically on providing a sufficient number of families to recruit. The neighborhood selection process resulted in four neighborhood types: High PAE/ NE, High PAE and Low NE, Low PAE/High NE, and Low PAE/Low NE. Maps of these counties and eligible block groups are provided elsewhere [16]. Environmental data were collected or derived from existing data sources in 2006–2007.

### Recruitment

Details about participant recruitment are provided elsewhere [10,16]. Households with children in eligible neighborhoods were identified through commercial marketing databases, mailed information about the study, and contacted by phone to determine eligibility and interest. Children 6–12 years old and one parent/caregiver were recruited, and initial (Time 1; T1) data collection occurred September 2007–January 2009. Time 2 data collection was scheduled for 2 years after T1 for each participant (September 2009–February 2011), with

average length of follow-up being 23.7 months ( $\pm 1.9$ ). Eligibility included a) living in an eligible neighborhood; b) able to do at least moderate intensity physical activity; c) not having any medical conditions or treatment associated with obesity or growth; d) child being >10th percentile BMI for age and sex based on parent report; and e) child not having significant eating disorder psychopathology, being on a medically prescribed dietary regimen, or having a mental health problem that would interfere with participation. Only one child per household participated, so if multiple children in a household were eligible and interested, the child with the nearest birthday to the recruitment date was enrolled.

## Participants

At Time 2 (T2), 642 of the 681 children available for cross-sectional T1 anthropometric analysis [10,16] completed the follow-up anthropometric measurement at a clinic or home visit, although 26 of these children had moved home residence since T1 and were excluded from the present longitudinal analyses. Children with any missing individual-level sociodemographic data (n=28) were also not included in analyses. The children with both T1 and T2 data came from 399 different census block groups.

Participant retention from T1 to T2 did not differ by neighborhood type (83.8% – 85.2%). Children lost to follow-up versus retained had participating parents with lower T1 age (39.7 versus 41.7,  $p < .003$ ), were more likely to be Hispanic (24.7% versus 15.9%,  $p < .036$ ), and had lower household income (<\$50K; 23.9% versus 12.8%,  $p < .022$ ). Children lost to follow-up had significantly higher T1 zBMI than those retained (0.63 versus 0.41,  $p < .03$ ) but the retained versus lost sample did not differ in T1 overweight/obesity rates or on any T1 eating or activity behaviors. T1 block groups not represented in T2 had significantly lower percentages of White populations (68% versus 76%;  $p < .008$ ) and lower resident median age (34.7 versus 37.0,  $p < .04$ ) relative to retained block groups.

## Measures

T1 and T2 anthropometric and behavioral measures were collected using the same procedures.

**Anthropometrics**—Child and parent height and weight were collected by a trained research assistant at the research clinic or at their home. Measures were obtained in triplicate to the nearest 0.1kg and 0.1cm, respectively, using a digital scale (Detecto 750; Detecto DR400C) and stadiometer (235 Heightronic digital stadiometer, portable SECA 214), with additional measurement until 3 of 4 consecutive measures were within 0.1kg and 0.5cm, respectively, and these most proximal values were averaged. Child BMI was calculated and standardized relative to CDC 2000 norms to determine Child BMI z-score [19]. Child overweight/obesity was defined as 85<sup>th</sup> BMI percentile for age and sex.

**Physical activity and sedentary behavior**—Child activity was measured by the GT1M Actigraph<sup>®</sup> accelerometer using 30 second epochs. Children were instructed to wear the accelerometer for seven consecutive days following the anthropometric visit. Children were asked to re-wear the accelerometer if <6 days of data (minimum of 10 waking hours per day) were recorded. Average minutes of moderate-to-vigorous physical activity (MVPA)

per valid day were determined using age-based 3-MET cut-points,[20] with a valid day defined as 8 or more valid hours, and a valid hour defined as having no more than 20 minutes of consecutive zero activity counts. Average sedentary minutes per valid day was the sum of epochs with activity counts 0–50 (per 30 second epoch) among valid hours within valid days. Only children with 3+ valid days were included in analyses.

**Dietary quality and energy intake**—Child dietary intake was assessed by three random, 24-hour dietary recalls conducted by trained staff over the phone using the standard multiple-pass approach, after the anthropometric visit. A consensus recall approach was used with parents and children 8 years old, with older children reporting individually with parent assistance. Parent-child dyads used two-dimensional food models (Nutrition Consulting Enterprise) to assist with portion estimation during recalls. Recall data were analyzed using Nutrition Data System for Research (NDS-R) software (version 2.92). Only children with at least 2+ daily recalls were included in analyses. Child diet quality was assessed via the DASH score [21], based on the sum score (range 0–80) of eight food groups (grains, vegetables, fruits, dairy, meat, nuts/seed/legumes, fats/oils, and sweets), with a maximum component score of 10 for each food group and higher values indicating higher diet quality [21,22]. Diet quality and energy intake (total kcal/day) estimates were averages across available days of dietary recall.

### Sociodemographics

Demographic characteristics of participants' block groups were gathered from 2000 U.S. Census, including median resident age, average family size (among family households), number of residents reporting being White, and median household income. Parents reported their age, highest level of education, number of children in the household, and annual household income (in \$10K increments up to \$100K+), as well as child sex, age, race, and ethnicity at T1.

This study was approved by Institutional Review Boards at Seattle Children's Hospital, San Diego State University, and Emory University. Parents provided written consent, and children provided assent to participate.

### Analyses

Assessments were available at both T1 and T2 for 618 children, though sample sizes varied by outcome and were attenuated by missing values for covariates. There were 586 children in the continuous zBMI analysis, as 2 children with extreme zBMI changes (>3 SDs) from T1 to T2 were excluded, but 588 children were retained in the categorical weight status change analysis. Missing or insufficient data for diet and activity assessments resulted in sample sizes of 573 and 558 respectively for the analyses of eating and activity behaviors.

For bivariate analyses, analysis of variance (ANOVA) examined change in continuous outcomes across neighborhood type. Chi-square tests for trend were used to examine child weight status at T2 across neighborhood type overall and then separately by T1 weight status. Primary analyses were based on multivariable regression models. For continuous outcomes (zBMI, DASH scores, energy intake, MVPA and sedentary time) linear regression

was used. For child weight status, relative risk regression based on a modified Poisson regression approach was applied [23]. In all the regression models, T2 outcomes were treated as dependent variables, with corresponding T1 measures and baseline sociodemographics entered as predictors along with neighborhood type. Region (San Diego, Seattle) was entered, but was not significant in any model. Clustering by block group was not done because most participants were the only participants in their block group. Significance was defined as  $p < .05$ .

## Results

Parent age differed significantly by neighborhood type, but no other sociodemographic differences were found by neighborhood type (see Table 1). Child zBMI and behaviors at T1 and T2 are provided in Table 1. Across all models, values at T1 were significantly related to corresponding T2 values.

### Change in child zBMI and weight status

In unadjusted analysis, child zBMI did not differ by neighborhood type at T1 or T2. However, after controlling for T1 child zBMI and sociodemographics, children in low PAE / high NE neighborhood and high PAE / low NE neighborhoods had significantly higher average zBMI at T2 than children in the high PAE / high NE neighborhoods. In adjusted models, BMI z-score for children in the low PAE / low NE neighborhoods did not differ significantly from high PAE / high NE neighborhood children. Older children had lower T2 zBMI values (see Table 2).

Children in the high PAE / high NE neighborhoods were significantly less likely to have overweight/obesity at T2 compared to children in all other neighborhood types ( $p < .003$ ; Figure 1), with these differences sustained after controlling for T1 child weight status and sociodemographics. In the adjusted model, children in the three less-supportive neighborhood types were 41–49% more likely to have overweight/obesity at T2 than children in high PAE / high NE neighborhoods. Older children and girls were less likely to have overweight/obesity at T2, whereas Hispanic children were more likely to have overweight/obesity (Table 3).

When analyzed in models unadjusted for sociodemographics, but separately by child T1 weight status, findings were in the expected direction, but were not significantly different by neighborhood type. Among children having overweight/obesity at T1, the proportion of children transitioning to not having overweight/obesity at T2 was not statistically significant different by neighborhood type ( $p = .057$ ). The difference by neighborhood type in the proportion of children having overweight/obesity at Time 2 among those not having overweight/obesity at T1 was also not significant ( $p = .076$ ). Both these results however were in the expected direction, with the high PAE / high NE neighborhoods having the best outcomes (Figure 2). Among children having overweight/obesity at T1 ( $n = 153$ ), differences in child T2 weight status across neighborhood types did not reach statistical significance in the adjusted model, although estimates for neighborhood type were moderate effect sizes and in the expected direction (Table 3). Effects of child age and Hispanic ethnicity were significant in this model, as in the all children model. The relative risk for a child having

overweight/obesity at T2 for parent T1 BMI was significantly  $<1.00$  in the all child adjusted model and  $>1.00$  in the only child overweight/obesity at T1 adjusted model, although the magnitude of both of these estimates was small (Table 3).

### Child diet quality and energy intake

In bivariate analysis, there were no significant differences in DASH scores at T1 or T2 across neighborhood type (Table 2) or significant changes in DASH scores from T1 to T2 overall or differentially by neighborhood type.

Bivariate analysis indicated no significant differences in daily energy intake at T1 by neighborhood type, but there were differences in daily energy intake by neighborhood type at T2 (Table 1). Average energy intake increased significantly over time overall, with girls having lower energy intakes than boys. There significant differences in daily energy intake by neighborhood type at T2 after controlling for sociodemographics and T1 energy intake. Children in the Low PAE / Low NE neighborhoods had significantly higher increases over time in daily energy intake relative to children in High PAE / High NE neighborhoods (Table 2). Overall, girls had smaller increases in energy intake over time.

### Child physical activity and sedentary time

In bivariate analysis, average daily minutes of 3+METs physical activity did not differ by neighborhood type at T1 or T2 (Table 1), but decreased significantly overall from T1 to T2. Unexpectedly, physical activity decreased significantly less among children in the Low PAE / High NE compared to children in the High PAE / High NE neighborhoods in the adjusted model (Table 2). Older children and girls had significantly larger decreases in physical activity over time, whereas children from higher income households and from higher income neighborhoods had significantly smaller decreases in physical activity.

Average daily sedentary time did not significantly differ by neighborhood type at T1, but did significantly differ at T2 (Table 1) and increased significantly overall from T1 to T2 based on bivariate analysis. In the adjusted model, there was significant differential change by neighborhood type in sedentary time, with higher increases among children in Low PAE / Low NE and Low PAE / High NE neighborhoods compared to children in the High PAE / High NE neighborhoods. No sociodemographic factors were related to sedentary time changes (Table 2).

## Discussion

Present findings highlight the importance of having a favorable combination of physical activity and nutrition environments in neighborhoods where children live to support their healthy weight status trajectory and related behaviors. After adjusting for sociodemographics across multiple levels, children living in neighborhoods more conducive to walking, having at least one high quality park, and a favorable nutrition environment marked by the availability of at least one large grocery store without a high concentration of fast food restaurants, had significantly more favorable weight status changes than children in two of the three other less environmentally-supportive neighborhoods. Children in these most favorable neighborhood environments were nearly 50% less likely to have overweight/



obesity by T2 compared to children in all other neighborhood types when accounting for initial child weight status and sociodemographics. Although not statistically significant, partially because the study was not designed for subgroup analyses, the transitions between having and not having overweight/obesity also were in a similar direction by neighborhood type (i.e., most favorable child weight status outcomes in the High PAE / High NE neighborhoods), even though children in the most favorable neighborhoods initially had the lowest level of having overweight/obesity among neighborhood types [10]. The decrease over time in average child BMI z-score seen in the most supportive neighborhoods was more than half the average change seen among children having overweight/obesity provided intense weight management intervention [24]. By T2, the overweight/obesity rate in these most favorable neighborhoods was nearly half that of two of the three types of less favorable neighborhood types and approximately half the national prevalence for childhood overweight/obesity [25]. Present findings are consistent with some other longitudinal findings regarding neighborhood environment and child weight status trajectories, e.g., [26].

Differential changes by neighborhood type in daily energy intake and sedentary behavior were generally consistent with the observed neighborhood environment-child weight status change differences. Across time, children living in the least favorable neighborhood environments increased their relative adjusted daily average energy intake by >100 kcals more and their daily sedentary time by >50 minutes more than children living in the most favorable neighborhood environments. Children residing in neighborhoods with poorer activity environments but better nutrition environments also had a relative increase in sedentary behavior. Unexpectedly, one of the neighborhood types (low PAE / high NE neighborhoods) with less-supportive environments for physical activity had children with the smallest declines in physical activity over time. However, recent reviews suggest that macro-environmental factors that contribute to walkability and to used define a more supportive physical activity environment in the present study (e.g., greater land use mix) are not consistently or highly related to children's overall physical activity and have in some instances been related to lower physical activity among younger children [27]. Other aspects of the built environment or parental travel and activity patterns may influence children's physical activity changes over time. For example, analysis of other data from the Neighborhood Impact on Kids study found that micro-environment features immediately around a child's home, such as sidewalk and street-crossing qualities, were significantly associated with children's active travel, leisure activity, and accelerometer-assessed total physical activity [28]. Caregivers in low PAE neighborhoods with adequate time and financial resources may be transporting their children to public or private physical activity facilities and programs. Low PAE neighborhoods had higher median incomes, and families from these neighborhoods reported higher average household income compared high PAE neighborhood families. Higher household and neighborhood income were also related to lesser declines in physical activity from T1 to T2. More disadvantaged children living in low PAE may show different physical activity trajectories. The precipitous overall decline in children's average physical activity observed with age has been previously documented [29], and further inquiry is clearly needed to identify and implement strategies to prevent this decline, particularly among disadvantaged youth.

There were some important study limitations. Neighborhood was defined by census block groups, without more fine-grained consideration of environmental features more immediately around children's homes or perceptions of environment. This study was limited to two metropolitan areas in the western U.S. with relatively higher costs of living and the sample was more racially white and higher income than the neighborhoods in which they lived, with most parents having high education levels. Lack of sufficient racial/ethnic and income diversity within neighborhood types prevented exploration of interactions between sociodemographics and environment. Prior studies have found higher childhood obesity rates in more disadvantaged neighborhoods or households [30,31], with some evidence that built environment differences may explain these disparities [10,32]. Environmental factors and their quality are not equally distributed [33] and emerging evidence suggests that activity and nutrition environments may interact with socioeconomic factors to impact healthy behaviors and weight status [34]. Socioeconomic factors may be important in explaining disparities between different racial/ethnic groups in child weight status, with activity and nutrition environments perhaps contributing differently or to a lesser extent to weight-related behaviors in disadvantaged populations [35]. More research is needed to understand and then intervene on environmental inequities [36]. Present analyses also did not examine how neighborhood environment may be impacting children's within-home environments or how much exposure children or their parent(s)/caregiver(s) had to their neighborhood environment, although prior analyses from this study did find relationships between parenting practices/rules around eating and children's diet quality [37]. Unlike Chen and colleagues [12], this study did not examine changes to environments over time. More natural experiments are needed to evaluate changes in policies, systems, and environments designed to, or that may inadvertently, affect children's weight outcomes [38]. However, such policy, systems, and environmental changes at the neighborhood level are not often implemented by researchers or institutions (e.g., schools), but rather by decision-makers within local governments. The Institute of Medicine has specific recommendations for local governments to improve neighborhood environments and policies to support healthy child weight [39]. The present study did not examine mechanisms by which environment has influence. Others speculate mechanisms could include cueing/prompting by environment (e.g., presence/signage of fast food restaurants prompts visitation), the convenience of facilities in environments in which to engage in desired behaviors (e.g., having a high quality park available for active recreation), and environment establishing social norms (e.g., seeing others being active in the park may prompt other children to be active there) [3,40] It is not known whether actual objective environment or whether perceptions of environment are predominant, with evidence that these may have independent effects on health behaviors including physical activity [41].

Results add credibility and urgency to recommendations to improve both nutrition and physical activity environments as components of solutions to the global childhood obesity epidemic [42]. The present definition of a salutogenic environment that includes walkability, high quality parks, proximal supermarkets, and relatively few fast food restaurants was shown to predict more favorable trajectories in child weight status, energy intake, and sedentary behavior. Creating more neighborhoods with these characteristics requires actions by municipal governments (zoning), transportation departments (street design), parks

departments, multiple components of the food industry, and multiple government agencies that regulate food retail. Though making such changes might appear to be too complex to be feasible, in fact communities and states across the US are making changes like these, and they have been associated with reductions in childhood obesity (see <https://stateofobesity.org/progress/>). However, more evidence regarding the impact of environmental and policy interventions for childhood obesity prevention, particularly outside of school settings is needed [38,43], including evaluating the interactive effect of programmatic and environmental interventions.

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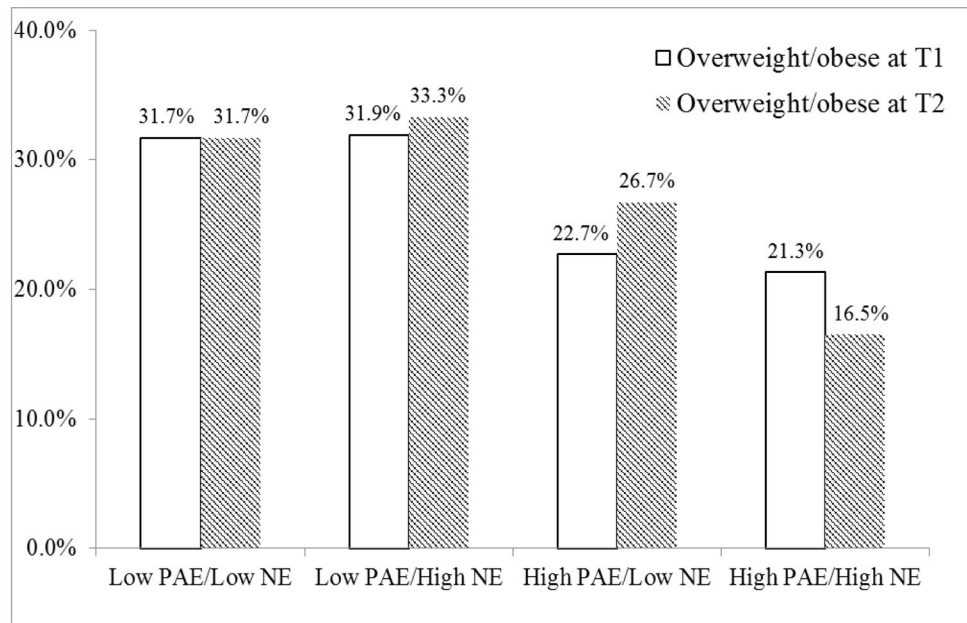
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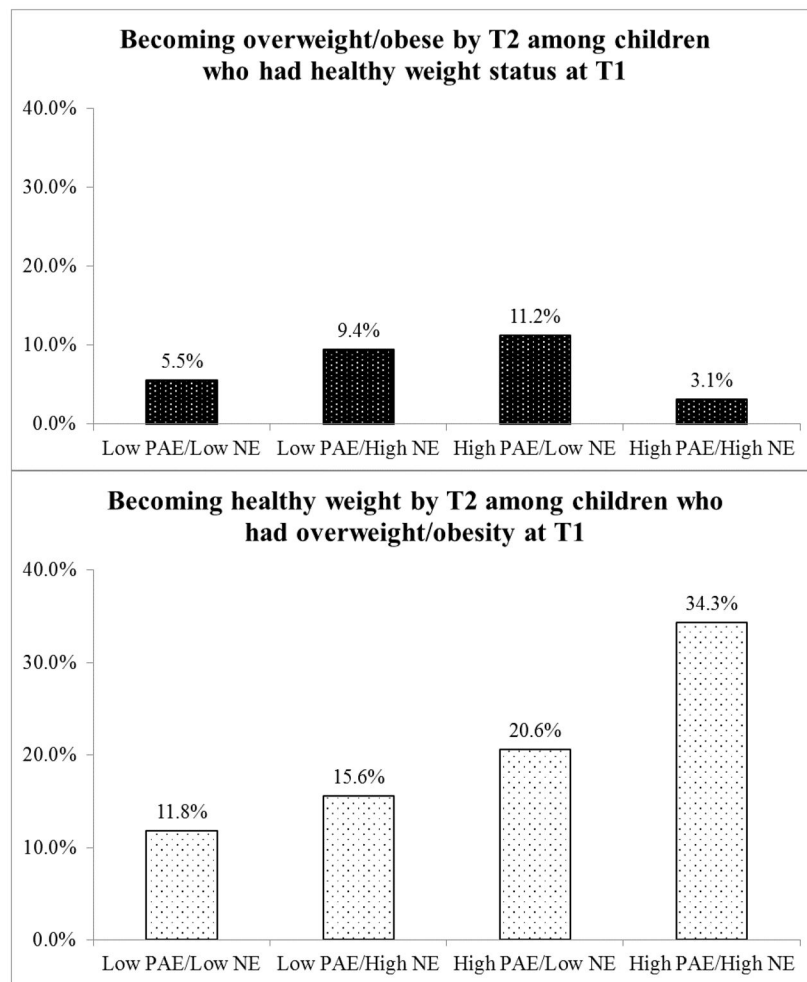
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### Study Importance

- Evidence is limited regarding the impact of neighborhood built and nutrition environment on the change in children's weight status over time.
- Evidence is further limited examining associations of neighborhood environments with robust measures of changes in children's weight-related behaviors.
- The present study finds that children living in neighborhoods that are supportive in both physical activity and nutrition environments have more healthy changes in child weight status, energy intake, and sedentary behavior over time.
- Changing neighborhood activity and nutrition environments has the potential to improve children's weight status and related behaviors.



**Figure 1.** Proportion of children having overweight/obesity at T1 and T2 by neighborhood type; significant differences by neighborhood type at T2 ( $p < .003$ ); analysis of T1 differences were reported previously [10].



**Figure 2.** The percentage of children transitioning between T1 and T2 from having healthy weight status ( $n=435$ ) to having overweight/obesity ( $p=.057$  for differences by neighborhood type) and from having overweight/obesity ( $n=153$ ) to having healthy weight status ( $p=.076$  for differences by neighborhood type)



**Table 1**Longitudinal sample characteristics and child anthropometrics and behaviors by neighborhood type (n=618<sup>\*</sup>)

	Low PAE/ Low NE neighborhoods	Low PAE/ High NE neighborhoods	High PAE/ Low NE neighborhoods	High PAE/ High NE neighborhoods
<b>Neighborhood Characteristics</b>				
# of block groups	n=108	n=85	n=91	n=115
Average percent White	78.3%	76.9%	73.1%	75.7%
Median resident age (years)	38.1 (6.8)	37.1 (6.2)	35.4 (5.7)	37.2 (7.0)
Average family size	3.2 (0.4)	3.2 (0.3)	3.1 (0.5)	3.1 (0.5)
Median household income (\$)	73449 (22783)	64690 (23035)	54190 (19676)	56329 (19142)
<b>Child, Parent, or Household Characteristics</b>				
# of children <sup>*</sup>	161	143	150	164
Baseline child age (years)	9.1 (1.5)	9.0 (1.6)	9.0 (1.6)	9.3 (1.5)
Child sex (% female)	50.9%	50.3%	49.3%	54.3%
Child ethnicity (% Hispanic)	12.2%	16.4%	17.7%	17.4%
Child race (%)				
- White	80.1%	85.3%	82.0%	84.1%
- Black or African American	5.0%	1.4%	1.3%	1.8%
- Asian or Pacific Islander	5.6%	6.3%	2.0%	1.8%
- Other or multiple races	9.3%	6.9%	14.7%	12.2%
Parent sex (% female)	85.7%	83.8%	84.0%	92.0%
Baseline parent age (years)	41.5 (5.8)	41.4 (5.6)	41.1 (5.8)	42.8 (5.6)
Baseline highest parent education in household				
- Up to some college	24.8%	27.3%	23.0%	18.8%
- Completed college	39.2%	40.3%	37.8%	41.3%
- More than college degree	35.9%	32.4%	39.1%	40.0%
T1 parent BMI	27.3 (6.3)	27.1 (5.2)	26.5 (5.8)	26.5 (5.8)
Baseline # of children <18 years old in household	2.4 (1.0)	2.4 (0.9)	2.4 (0.9)	2.3 (0.8)
Baseline household income				
- <50K	8.6%	12.5%	16.2%	14.0%
- 50–100K	36.2%	34.6%	41.2%	38.9%
- >100K	55.3%	52.9%	42.6%	47.1%

	Low PAE/ Low NE neighborhoods	Low PAE/ High NE neighborhoods	High PAE/ Low NE neighborhoods	High PAE/ High NE neighborhoods
<b>Child Anthropometrics and Behaviors</b>				
T1 child BMI z-score	0.51 (0.99)	0.45 (1.05)	0.36 (0.94)	0.32 (0.85)
T2 child BMI z-score	0.46 (1.04)	0.46 (1.03)	0.39 (0.99)	0.23 (0.92)
T1 child diet quality (DASH score)	41 (8)	42 (7)	42 (7)	42 (7)
T2 child diet quality (DASH score)	41 (7)	41 (8)	41 (8)	41 (7)
T1 child energy intake (kcal/day)	1768 (348)	1685 (388)	1717 (414)	1793 (455)
T2 child energy intake (kcal/day)	2028 (483)	1878 (469)	1878 (460)	1902 (422)
T1 child 3+ METs moderate-to-vigorous physical activity (minutes/day)	146 (49)	144 (51)	150 (55)	143 (57)
T2 child 3+ METs moderate-to-vigorous physical activity (minutes/day)	119 (65)	121 (73)	106 (43)	103 (43)
T1 child daily sedentary time (minutes/day)	393 (72)	396 (66)	399 (68)	399 (79)
T2 child daily sedentary time (minutes/day)	498 (157)	479 (145)	441 (76)	445 (87)

Note. Values are mean (standard deviation) or percentages as indicated; PAE = physical activity environment; NE = nutrition environment; non-percentage values are mean (standard deviation) unless otherwise noted;

\* some sample sizes are lower than this because of missing sociodemographic data (e.g., parental refusal to report on household income)

**Table 2**

Linear regression model results for child Time 2 BMI z-score and related behaviors, adjusting for multi-level sociodemographics and corresponding T1 BMI z-score and related behaviors

	BMI z-score (n=586)		Energy Intake (kcal/day) (n=573)		3+ MET MVPA (mins/day) (n=558)		Sedentary activity (mins/day) (n=558)	
	Estimate	(95% CI)	Estimate	(95% CI)	Estimate	(95% CI)	Estimate	(95% CI)
<b>Neighborhood type<sup>d</sup></b>								
- Low PAE / Low NE	0.06	(-0.04, 0.16)	116.4*	(15.6, 217.2)	7.3	(-2.8, 17.4)	50.2*	(22.2, 78.2)
- Low PAE / High NE	0.11*	(0.01, 0.21)	-8.2	(-111.1, 94.7)	12.3*	(2.2, 22.5)	38.3*	(10.7, 66.5)
- High PAE / Low NE	0.12*	(0.03, 0.21)	-15.9	(-113.0, 81.1)	-4.7	(-14.3, 5.0)	-9.9	(-36.8, 17.0)
<b>Child age</b>	-0.03*	(-0.05, -0.005)	23.2	(-0.9, 47.2)	-8.9*	(-12.1, -5.7)	5.0	(-2.5, 12.4)
<b>Child sex (girls)<sup>b</sup></b>	-0.02	(-0.08, 0.05)	-236.3*	(-307.7, -165.0)	-13.7*	(-21.0, -6.5)	-6.3	(-25.5, 12.9)
<b>Child race (non-White)<sup>c</sup></b>	0.04	(-0.05, 0.13)	-43.3	(-134.6, 48.0)	0.5	(-8.4, 9.3)	-1.7	(-26.3, 23.0)
<b>Child ethnicity (Hispanic)<sup>d</sup></b>	0.09	(-0.01, 0.20)	-28.3	(-136.1, 79.4)	7.3	(-3.1, 17.8)	14.0	(-15.2, 43.2)
<b>Parent age</b>	0.001	(-0.006, 0.008)	3.7	(-3.5, 10.9)	-0.2	(-0.9, 0.5)	0.09	(-1.9, 2.1)
<b>Parent education</b>	-0.004	(-0.04, 0.04)	14.6	(-26.5, 55.8)	-2.3	(-6.3, 1.8)	4.6	(-6.7, 15.9)
<b>Parent BMI at T1</b>	0.00003	(-0.001, 0.001)	0.1	(-0.5, 0.7)	-0.02	(-0.10, 0.06)	-0.02	(-0.24, 0.19)
<b>Number of children in household</b>	-0.009	(-0.05, 0.03)	2.9	(-37.9, 43.8)	2.1	(-1.9, 6.2)	3.6	(-7.7, 14.8)
<b>Annual household income</b>	0.007	(-0.008, 0.023)	4.5	(-11.9, 21.0)	2.3*	(0.7, 3.9)	1.6	(-2.9, 6.1)
<b>Number of neighborhood residents reporting being White race</b>	-0.0004	(-0.0001, 0.0001)	0.32	(-0.02, 0.08)	0.002	(-0.002, 0.007)	0.01	(-0.004, 0.023)
<b>Average family size in neighborhood</b>	-0.08	(-0.195, 0.035)	18.1	(-100.3, 136.4)	5.7	(-6.0, 17.4)	22.5	(-10.1, 55.1)
<b>Median age of neighborhood residents</b>	-0.005	(-0.012, 0.003)	-1.2	(-8.8, 6.4)	0.2	(-0.5, 1.0)	-0.2	(-2.3, 1.9)
<b>Median household income</b>	-1.15e-6	(-3.06e-6, 7.65e-7)	0.001	(-0.001, 0.003)	0.0002*	(0.00001, 0.00004)	-0.0001	(-0.001, 0.0004)

	BMI z-score (n=586)		Energy Intake (kcal/day) (n=573)		3+ MET MVPA (mins/day) (n=558)		Sedentary activity (mins/day) (n=558)	
	Estimate	(95% CI)	Estimate	(95% CI)	Estimate	(95% CI)	Estimate	(95% CI)
<b>Corresponding Time 1 value</b>	0.94*	(0.91, 0.98)	0.3*	(0.2, 0.4)	-0.5*	(-0.6, -0.04)	0.6*	(0.4, 0.7)
Sample size	586		573		558		558	

Note. Estimates are unstandardized beta coefficients; models also included site (Seattle versus San Diego area) and amount of time between measurement visits; sample sizes vary based on differences in data availability by outcome; PAE = Physical activity environment; NE = Nutrition environment;

<sup>a</sup>Referent group = High PE / High NE neighborhoods;

<sup>b</sup>Referent group = boys;

<sup>c</sup>Referent group = White;

<sup>d</sup>Referent group = Non-Hispanic;

\* p<.05

**Table 3**

Relative risks of child having overweight/obesity at Time 2 adjusting for multi-level sociodemographics and Time 1 weight status

	Among all children (n=588)		Among only children who had overweight/obesity at T1 (n=153)	
	RR estimate	95% CI	RR estimate	95% CI
<b>Neighborhood type<sup>d</sup></b>				
- Low PAE / Low NE	1.41*	(1.13, 1.70)	1.25	(1.00,1.49)
- Low PAE / High NE	1.49*	(1.19,1.78)	1.25	(0.99,1.51)
- High PAE / Low NE	1.49*	(1.18,1.80)	1.17	(0.90,1.44)
<b>Child age</b>	0.88*	(0.80,0.96)	0.91*	(0.85,0.97)
<b>Child sex (girls)<sup>b</sup></b>	0.77*	(0.56,0.99)	0.96	(0.81,1.11)
<b>Child race (non-White)<sup>c</sup></b>	1.21	(0.98,1.45)	1.15	(0.99,1.31)
<b>Child ethnicity (Hispanic)<sup>d</sup></b>	1.45*	(1.19,1.71)	1.18	(1.05,1.32)
<b>Parent age</b>	1.00	(0.98,1.02)	1.00	(0.98,1.01)
<b>Parent education</b>	0.95	(0.86,1.03)	0.98	(0.91,1.04)
<b>Parent BMI at T1</b>	0.998*	(0.997,<1.000)	1.01*	(1.00,1.03)
<b>Number of children in household</b>	0.97	(0.86,1.09)	0.98	(0.88,1.08)
<b>Annual household income</b>	1.01	(0.97,1.05)	1.00	(0.97,1.03)
<b>Number of neighborhood residents reporting being White race</b>	1.00	(1.00,1.00)	1.00	(1.00,1.00)
<b>Average family size in neighborhood</b>	0.86	(0.54,1.17)	0.84	(0.60,1.08)
<b>Median age of neighborhood residents</b>	1.00	(0.98,1.02)	1.01	(0.99,1.02)
<b>Median household income</b>	1.00	(1.00,1.00)	1.00	(1.00,1.00)
<b>Child who had overweight/obesity at T1</b>	11.74*	(11.38,12.1)	—	—

Note. Models also included site (Seattle versus San Diego area) and time between measurement visits; PAE = Physical activity environment; NE = Nutrition environment;

<sup>a</sup>Referent group = High PE / High NE neighborhoods;

<sup>b</sup>Referent group = boys;

<sup>c</sup>Referent group = White;

<sup>d</sup>Referent group = Non-Hispanic;

\* p<.05