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Publication Date

2024-10-01

DOI

10.1016/j.nuclphysa.2024.122924

Peer reviewed

Controlling volume fluctuations for studies of critical phenomena in nuclear collisions

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We generalize and extend the recently proposed method [1] to account for contributions of system size (or volume/participant) fluctuations to the experimentally measured moments of particle multiplicity distributions. We find that in the general case there are additional biases which are not directly accessible to experiment. These biases are, however, parametrically suppressed if the multiplicity of the particles of interest is small compared to the total charged-particle multiplicity, e.g., in the case of proton number fluctuations at top RHIC and LHC energies. They are also small if the multiplicity distribution of charged particles per wounded nucleon is close to the Poissonian limit, which is the case at low energy nuclear collisions, e.g., at GSI/SIS18. We further find that mixed events are not necessarily needed to extract the correction for volume fluctuations. We provide the formulas to correct pure and mixed cumulants of particle multiplicity distributions up to any order together with their associated biases.

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I. INTRODUCTION

One of the main goals of studying relativistic heavy-ion collisions is to explore the structure of the QCD phase diagram. Fluctuations of observed particles carrying quantum numbers of conserved charges, baryon number (B), electric charge and strangeness, represent a powerful tool for this endeavor as the cumulants of their distributions measure the derivatives of the grand-canonical partition function, and thus the pressure (P), with respect to the associated chemical potentials. For example, for a thermal system of volume V and temperature T, the cumulants of the net baryon number distribution, within the Grand Canonical Ensemble (GCE), are given by [2]

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$$\kappa_n[B] = \frac{\partial^n (\ln Z)}{\partial (\mu_B/T)^n} = \frac{V}{T} \frac{\partial^n P}{\partial (\mu_B/T)^n},$$

where Z is a GCE partition function and μ_B is a baryon chemical potential. Any nontrivial structures in the equation of state such as a possible phase transition [3–6] will result in potentially large derivatives of the pressure and thus in large values of the cumulants of conserved charges. In addition, as cumulants are derivatives of the pressure, they are accessible (at vanishing or small values of chemical potential) to Lattice QCD calculations [7, 8], which in principle enables a direct comparison of results from ab initio QCD calculations with experiment. For example, as pointed out in Ref. [9], the measurement of higher-order cumulants close to vanishing chemical potential may test the remnants of chiral criticality.

Measurements of fluctuations have meanwhile been carried out by many experiments. The STAR collaboration has measured cumulants of the net-proton number up to sixth order over the entire energy range available at RHIC [10, 11]. The HADES experiment has measured cumulants of proton number up to forth order at the low energy of $\sqrt{s_{\rm NN}} = 2.4 \, {\rm GeV}$ [12] and ALICE has measured the second- and third-order net-proton number cumulants at $\sqrt{s_{\rm NN}} = 2.76$ and 5.02 TeV [13, 14].

When comparing cumulants measured in experiment with those obtained from lattice QCD or other field theoretical calculations [15] one needs to be aware of several key differences. While theoretical calculations are typically done in the grand canonical ensemble where charges can be exchanged with a heat bath and are only conserved on the average, in experiment charges are explicitly conserved on event by event basis and one has to account for global as well as local charge conservation [16-20]. Also, in experiments one usually is restricted to the measurement of net protons whereas theory can only calculate cumulants of the net baryon number. In the presence of many pions this difference can be corrected for [21]. Finally, and this will be the topic of the present paper, in experiment the size of the particle emitting system is not constant. Even under the tightest centrality selection criteria, this gives rise to so-called volume fluctuations [22] or, equivalently, fluctuations of the number of wounded nucleons [23]. Moreover, there is a strong correlation between the event activity, i.e., charged particle multiplicity and the size of the system, e.g., within the model of independent particle sources [24]. To ensure accurate measurement of the fluctuation of conserved charges, it is crucial to separate the particles used for determining these fluctuations from those used for centrality determination. This separation is necessary to avoid autocorrelations, which can introduce artificial modifications to the fluctuation signals being measured [23]. One effective method to achieve this separation is to perform centrality selection using detectors that cover different rapidity intervals than those used for fluctuation analysis [12, 13]. Alternatively, another approach involves excluding the particles of interest, such as protons in net-proton analysis, from the centrality determination process [25]. However, even with this exclusion, some autocorrelation effects may persist. This persistence is partly due to strong correlations between pions and protons, which arise from the decays of baryonic resonances [23]. Moreover, volume fluctuations can significantly impact the measurements, particularly at lower energies. At these energies, the multiplicity of charged particles is predominantly made up of primordial protons, which can limit the resolution of centrality selection that can be achieved. This limitation underscores the importance of carefully considering the effects of volume fluctuations in the analysis. In Ref. [1] a novel and promising method based on event mixing has been proposed to experimentally determine and subtract the contributions to the cumulants caused by volume fluctuations. In the present work we will further elaborate on this topic, generalize the results, and provide the formulas for corrections of any higher-order cumulants.

This paper is organized as follows. In the next section we define the notation. We then present an analytical formulation of event mixing as proposed in Ref. [1]. We find that the cumulants of the the mixed events have additional bias terms which were assumed to vanish in the original work of [1], and we discuss the magnitude of these corrections for various scenarios. Next we extend our study to cumulants of higher order before we discuss and summarize our results.

II. NOTATION

In this paper we will mostly work within the wounded-nucleon model [24] to discuss volume or participant fluctuations. However, as we shall show later, the formalism can be easily applied also to the situation where one has generic volume fluctuations, as for example discussed in Refs. [22, 23]. Let us start with the expression of the particle number cumulants $\kappa_i[N]$ in the presence of wounded-nucleon fluctuations (for details see Appendix A):

$$\kappa_1[N] = \langle N_w \rangle \, \kappa_1[n] = \langle N_w \rangle \, \langle n \rangle = \langle N \rangle \tag{1}$$

$$\kappa_2[N] = \langle N_w \rangle \, \kappa_2[n] + \langle n \rangle^2 \, \kappa_2[N_w] = \bar{\kappa}_2[N] + \langle N \rangle^2 \, \frac{\kappa_2[N_w]}{\langle N_w \rangle^2} \tag{2}$$

$$\kappa_3[N] = \langle N_w \rangle \, \kappa_3[n] + 3 \, \langle n \rangle \, \kappa_2[n] \kappa_2[N_w] + \langle n \rangle^3 \, \kappa_3[N_w] = \bar{\kappa}_3[N] + 3 \, \langle N \rangle \, \bar{\kappa}_2[N] \frac{\kappa_2[N_w]}{\langle N_w \rangle^2} + \langle N \rangle^3 \frac{\kappa_3[N_w]}{\langle N_w \rangle^3} \tag{3}$$

$$\kappa_{4}[N] = \langle N_{w} \rangle \kappa_{4}[n] + 4 \langle n \rangle \kappa_{3}[n] \kappa_{2}[N_{w}] + 3\kappa_{2}^{2}[n] \kappa_{2}[N_{w}] + 6 \langle n \rangle^{2} \kappa_{2}[n] \kappa_{3}[N_{w}] + \langle n \rangle^{4} \kappa_{4}[N_{w}]
= \bar{\kappa}_{4}[N] + 4 \langle N \rangle \bar{\kappa}_{3}[N] \frac{\kappa_{2}[N_{w}]}{\langle N_{w} \rangle^{2}} + 3\bar{\kappa}_{2}^{2}[N] \frac{\kappa_{2}[N_{w}]}{\langle N_{w} \rangle^{2}} + 6 \langle N \rangle^{2} \bar{\kappa}_{2}[N] \frac{\kappa_{3}[N_{w}]}{\langle N_{w} \rangle^{3}} + \langle N \rangle^{4} \frac{\kappa_{4}[N_{w}]}{\langle N_{w} \rangle^{4}}$$
(4)

Here N refers to the particles of interest, say protons, and n to the number of these particles arising from one wounded nucleon; thus $\langle n \rangle$ is the average number of particles per wounded nucleon. The cumulants of the wounded-nucleon distribution are denoted by $\kappa_j[N_w]$ while the cumulants for the distribution of particles stemming from one wounded nucleon are $\kappa_j[n]$. The corresponding relations for cumulants of any order can be obtained with the provided software package [26].

The cumulants of interest are those at a fixed number of wounded nucleons. They reflect the true density fluctuations in a system at constant volume. We denote these cumulants for a system with fixed, i.e. non-fluctuating, number of $\langle N_w \rangle$ wounded nucleons as

$$\bar{\kappa}_i[N] = \langle N_w \rangle \, \kappa_i[n],$$

Below we will also deal with factorial cumulants, which we shall denote by C_j . Factorial cumulants, which measure the deviation from Poisson statistics, tell us about the true correlations in the system. As discussed in the Appendix B, they are linear combinations of the regular cumulants. For the first four orders we have

$$C_{1}[N] = \kappa_{1}[N] = \langle N \rangle,$$

$$C_{2}[N] = -\kappa_{1}[N] + \kappa_{2}[N],$$

$$C_{3}[N] = 2\kappa_{1}[N] - 3\kappa_{2}[N] + \kappa_{3}[N],$$

$$C_{4}[N] = -6\kappa_{1}[N] + 11\kappa_{2}[N] - 6\kappa_{3}[N] + \kappa_{4}[N].$$

The expressions for the particle number factorial cumulants are similar to Eqs. 1-4

$$C_1[N] = \langle N_w \rangle C_1[n] = \langle N_w \rangle \langle n \rangle = \langle N \rangle, \tag{5}$$

$$C_2[N] = \bar{C}_2[N] + \langle N \rangle^2 \frac{\kappa_2[N_w]}{\langle N_w \rangle^2},\tag{6}$$

$$C_3[N] = \bar{C}_3[N] + 3\langle N \rangle \bar{C}_2[N] \frac{\kappa_2[N_w]}{\langle N_w \rangle^2} + \langle N \rangle^3 \frac{\kappa_3[N_w]}{\langle N_w \rangle^3}, \tag{7}$$

$$C_{4}[N] = \bar{C}_{4}[N] + 4 \langle N \rangle \, \bar{C}_{3}[N] \frac{\kappa_{2}[N_{w}]}{\langle N_{w} \rangle^{2}} + 3 \bar{C}_{2}^{2}[N] \frac{\kappa_{2}[N_{w}]}{\langle N_{w} \rangle^{2}} + 6 \langle N \rangle^{2} \, \bar{C}_{2}[N] \frac{\kappa_{3}[N_{w}]}{\langle N_{w} \rangle^{3}} + \langle N \rangle^{4} \, \frac{\kappa_{4}[N_{w}]}{\langle N_{w} \rangle^{4}}. \tag{8}$$

79 Similar to the cumulants, we denote by

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$$\bar{C}_k[N] = \langle N_w \rangle \, C_k[n]$$

the factorial cumulants for a system at constant volume or number of wounded nucleons, $\langle N_w \rangle$.

III. MIXED EVENTS

In Ref. [1] a mixed event is constructed such that it has the same total multiplicity as a given real event but each particle (track) is drawn from a different event, so that, by construction, the mixed events follow the same total

multiplicity distribution as the original events. This is done in order to preserve volume flucutations as in real events. Since each particle (track) is chosen randomly from a random event, the distribution of particle species will follow a multinomial distribution with the Bernoulli probabilities $p_i = \langle N_i \rangle / \langle M \rangle$ for particles of type i. Here $\langle N_i \rangle$ denotes the mean number of particles of type i and $\langle M \rangle$ the mean total multiplicity. Hence, the probability to find A particles (successes) of type A and B particles of type B is given by the trinomial probability $B_3(A, B, M; p_A, p_B)$ and so on. Here M denotes the multiplicity of the event under consideration. Thus the distribution, $P_{mix}(A, B)$, of particles of species A and B in the mixed events is obtained by folding the multiplicity distribution $P_M(M)$ with a trinomial (in general multinomial) distribution:

$$P_{mix}(A, B) = \sum_{M} B_3(A, B, M; p_A, p_B) P_M(M)$$

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$$p_{\rm A} = \frac{\langle A \rangle}{\langle M \rangle}, \quad p_{\rm B} = \frac{\langle B \rangle}{\langle M \rangle}.$$

and

$$B_3(A, B, M; p_A, p_B) = \frac{M!}{A!B!(M - A - B)!} p_A^A p_B^B (1 - p_A - p_B)^{M - A - B}$$
(9)

The factorial cumulant-generating function for this distribution is

$$g_{F,mix}(z_{A}, z_{B}) = \ln \left[\sum_{A,B} P_{mix}(A, B) (z_{A})^{A} (z_{B})^{B} \right]$$

$$= \ln \left[\sum_{M} \left[h_{3}(z_{A}, z_{B}) \right]^{M} P_{M}(M) \right]$$

$$= G_{F,M}(h_{3}(z_{A}, z_{B}))$$
(10)

where z_A and z_B are auxiliary variables and $h_3\left(z_{\mathrm{A}},z_{\mathrm{B}}\right)$ defined as

$$h_3(z_A, z_B) = \sum_{A,B} B_3(A, B; M = 1; p_A, p_B)(z_A)^A(z_B)^B = (1 - p_A - p_B + p_A z_A + p_B z_B)$$

is the factorial moment-generating function for the trinomial distribution with one trial (M=1).

The variables z_A and z_B facilitate the computation of factorial cumulants by taking the appropriate derivatives of the factorial cumulant-generating function for the multiplicity distribution $G_{F,M}$ $(h_3(z_A, z_B))$ (see Eq.(B2))

$$\left.C_{i,j}^{mix}[mix] = \left.\frac{\partial^{(i+j)}}{\partial(z_{\mathrm{A}})^{i}\partial(z_{\mathrm{B}})^{j}}G_{F,M}\left(h_{3}\left(z_{\mathrm{A}},z_{\mathrm{B}}\right)\right)\right|_{z_{\mathrm{A}}=z_{\mathrm{B}}=1} = p_{\mathrm{A}}^{i}p_{\mathrm{B}}^{j}C_{i+j}\left[M\right],$$

with $C_k[M]$ being the k^{th} -order factorial cumulant. Using the expression for the factorial cumulants of the multiplicity distribution derived in Appendix C, Eq. C5, we get within the wounded-nucleon model

$$C_{1}^{mix}[A] = \kappa_{1}^{mix}[A] = p_{A} \langle N_{w} \rangle \langle m \rangle$$

$$C_{2}^{mix}[A] = p_{A}^{2} C_{2}[M] = p_{A}^{2} \left[\kappa_{2} \left[N_{w} \right] \langle m \rangle^{2} + \langle N_{w} \rangle C_{2}[m] \right]$$

$$C_{1,1}^{mix}[A, B] = p_{A} p_{B} C_{2}[M] = p_{A} p_{B} \left[\kappa_{2} \left[N_{w} \right] \langle m \rangle^{2} + \langle N_{w} \rangle C_{2}[m] \right]. \tag{11}$$

For the corresponding cumulants up to second order we get accordingly

$$\kappa_{1}^{mix}[A] = C_{1}^{mix}[A] = p_{A} \langle N_{w} \rangle \langle m \rangle$$

$$\kappa_{2}^{mix}[A] = C_{2}^{mix}[A] + C_{1}^{mix}[A] = p_{A}^{2} \left[\kappa_{2} \left[N_{w} \right] \langle m \rangle^{2} + \langle N_{w} \rangle \left(\kappa_{2}[m] - \kappa_{1}[m] \right) \right] + p_{A} \langle N_{w} \rangle \langle m \rangle$$

$$= p_{A}^{2} \left[\kappa_{2} \left[N_{w} \right] \langle m \rangle^{2} + \langle N_{w} \rangle \left(\kappa_{2}[m] - \langle m \rangle \right) \right] + p_{A} \langle N_{w} \rangle \langle m \rangle$$

$$cov^{mix}[A, B] = C_{1,1}^{mix}[A, B] = p_{A}p_{B} \left[\kappa_{2} \left[N_{w} \right] \langle m \rangle^{2} + \langle N_{w} \rangle \left(\kappa_{2}[m] - \langle m \rangle \right) \right].$$
(12)

With $\langle m \rangle$ denoting the mean number of total particles emitted by a wounded nucleon, we get $\langle a \rangle = p_A \langle m \rangle$ and $\langle b \rangle = p_B \langle m \rangle$ for the mean number of particles per wounded nucleon of type A and B, respectively, and recover the results of Ref. [1]. For that we have to assume that the multiplicity distribution per wounded nucleon is Poissonian, i.e. that $C_2[m] = \kappa_2[m] - \langle m \rangle = 0$. This has been an implicit assumption in Ref. [1], which however is not valid in general as we shall discuss below.

The main benefit of the event mixing is to be able to relate the factorial cumulants of the various multiplicity distributions, as can be seen from Eq.11. All that enters is the second-order factorial cumulant, $C_2[M]$. The binomial probabilities, p_A and p_B , are in the sense trivial as they can be determined without any mixed events. Thus we may express the fluctuations of the wounded nucleons in terms of the factorial cumulant of the track multiplicity distribution

$$\langle N \rangle^2 \frac{\kappa_2[N_w]}{\langle N_w \rangle^2} = \frac{\langle N \rangle^2}{\langle M \rangle^2} \left(C_2[M] - \langle N_w \rangle C_2[m] \right) = \frac{\langle N \rangle^2}{\langle M \rangle^2} \left(C_2[M] - \bar{C}_2[M] \right), \tag{13}$$

where $\bar{C}_2[M] = \langle N_w \rangle C_2[m]$ is the second-order factorial cumulant for a system of $\langle N_w \rangle$ wounded nucleons without wounded nucleon fluctuations and N stands now for the multiplicity of the particles of interest, i.e., either A or B. While the factorial cumulant of the multiplicity distribution, $C_2[M]$, is accessible to experiment, that of a non-fluctuating system, $\bar{C}_2[M]$, is not. Let us, therefore define a bias term, Δ_2 , as

$$\Delta_2 \equiv \frac{\langle N \rangle^2}{\langle M \rangle^2} \bar{C}_2[M]. \tag{14}$$

In case of a Poissonian multiplicity distribution for one wounded nucleon the bias term vanishes, i.e., $\Delta_2 = 0$, since $\bar{C}_2[M] = \langle N_w \rangle C_2[m] = 0$ in this case, and we recover the results of Ref. [1]. Let us furthermore define the corrected cumulant, $\kappa_2^{corr}[N]$, which is based on measurable quantities only

$$\kappa_2^{corr}[N] = \kappa_2[N] - \frac{\langle N \rangle^2}{\langle M \rangle^2} C_2[M]. \tag{15}$$

Following Eq. 2 and using Eq. 6, the cumulant of the system without wounded nucleon fluctuations, $\bar{\kappa}_2[N]$, is given by

$$\bar{\kappa}_2[N] = \kappa_2[N] - \langle N \rangle^2 \frac{\kappa_2[N_W]}{\langle N_W \rangle^2} = \kappa_2^{corr}[N] + \Delta_2. \tag{16}$$

The bias, Δ_2 , while not directly measurable, may be constrained by a fit to the track multiplicity distribution within the wounded-nucleon model [24], as it is commonly done [11, 27, 28]. In addition, we note that for protons at very high collision energies we have $\langle N_p \rangle \ll \langle M \rangle$ so that Δ_2 is suppressed parametrically. This behavior can indeed be illustrated with simulations as presented in Sec. V. Since cumulants scale with the system size, or in our case with the number of wounded nucleons, $\langle N_w \rangle$, it is instructive to scale the (factorial) cumulants with the mean number of particles

$$\frac{\bar{\kappa}_2[N]}{\langle N \rangle} = \frac{\kappa_2[N]}{\langle N \rangle} - \frac{\langle N \rangle}{\langle M \rangle} \left(\frac{C_2[M]}{\langle N \rangle} - \frac{\bar{C}_2[M]}{\langle N \rangle} \right) = \frac{\kappa_2^{corr}[N]}{\langle N \rangle} + \frac{\Delta_2}{\langle N \rangle}$$
(17)

The scaled bias is then given by

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$$\frac{\Delta_2}{\langle N \rangle} = \frac{\langle N \rangle}{\langle M \rangle} \bar{c}_2[M],\tag{18}$$

where $\bar{c}_2[M] = \bar{C}_2[M]/\langle M \rangle$ is a scaled factorial cumulant. Typically, the scaled cumulants are of order unity, $\kappa_j[N]/\langle N \rangle \sim \mathcal{O}(1)$. In addition, the scaled factorial cumulants, $c_k[N] = C_k[N]/\langle N \rangle$, are expected to depend only weakly on the multiplicity. Therefore, the scaled bias should be much smaller than one, $\Delta_2/\langle N \rangle \ll 1$, for the volume correction to be reliable.

Finally, one may express the bias term Δ_2 also in terms of cumulants by using the relation between cumulants and factorial cumulants (see Appendix B), $C_2[M] = \kappa_2[M] - \langle M \rangle$ and so forth. This gives,

$$\Delta_2 = \frac{\langle N \rangle^2}{\langle M \rangle^2} \left(\bar{\kappa}_2[M] - \langle M \rangle \right) \tag{19}$$

A note of caution may be useful in this context. One might be inclined to express the fluctuations of the wounded nucleon directly using the cumulants of the multiplicity distribution, in which case one would get

$$\langle N \rangle^2 \frac{\kappa_2[N_w]}{\langle N_w \rangle^2} = \frac{\langle N \rangle^2}{\langle M \rangle^2} (\kappa_2[M] - \langle N_w \rangle \kappa_2[m]).$$

And since $\langle N_W \rangle \kappa_2[m] = \bar{\kappa}_2[M]$ is not directly accessible to experiment, one may further assign the bias to be $\Delta_2 = \frac{\langle N \rangle^2}{\langle M \rangle^2} \bar{\kappa}_2[M]$. This, however, would considerably overestimate its true value, Eq. 19, as cumulants always contain a "trivial" component proportional to the number of particles, which in principle is measurable.

IV. HIGHER-ORDER RESULTS

Let us now discuss the corrections for volume fluctuations up to fourth order. Given the discussion in the previous section the strategy is straightforward. First we express the fluctuations of the wounded nucleons in terms of factorial cumulants of the multiplicity distribution. Then we identify the parts which are experimentally accessible and those which are not. The latter will be the bias while the former will be subtracted from the expression for the cumulants in order to remove most of the effect of volume fluctuations. The terms involving cumulants of the wounded-nucleon distribution as they appear in the expressions for the cumulants as $\kappa_j[N_w]/\langle N_w \rangle^j$, see Eqs. (6-8) are:

$$\frac{\kappa_2[N_w]}{\langle N_w \rangle^2} = \frac{C_2[M] - \bar{C}_2[M]}{\langle M \rangle^2} \tag{20}$$

$$\frac{\kappa_3[N_w]}{\langle N_w \rangle^3} = -3 \frac{\bar{C}_2[M]}{\langle M \rangle^2} \frac{\kappa_2[N_w]}{\langle N_w \rangle^2} + \frac{C_3[M] - \bar{C}_3[M]}{\langle M \rangle^3} \tag{21}$$

$$\frac{\kappa_4[N_w]}{\langle N_w \rangle^4} = -6 \frac{\bar{C}_2[M]}{\langle M \rangle^2} \frac{\kappa_3[N_w]}{\langle N_w \rangle^3} - \frac{4\bar{C}_3[M]\langle M \rangle + 3\bar{C}_2[M]^2}{\langle M \rangle^4} \frac{\kappa_2[N_w]}{\langle N_w \rangle^2} + \frac{C_4[M] - \bar{C}_4[M]}{\langle M \rangle^4}$$
(22)

We note that binomial efficiency corrections do not affect the results as both, numerators and denominators of the right hand side of the above expressions, scale with the same power of the efficiency.

Inserting these expressions into Eqs. (2-4) for the cumulants $\kappa_j[N]$, we can solve for the cumulants of the system with fixed number of wounded nucleons, namely the $\bar{\kappa}_j[N]$. The results are given in the following general form

$$\bar{\kappa}_j[N] = \kappa_j^{corr}[N] + \Delta_j[N] \tag{23}$$

with $\bar{\kappa}_j[N]$ the cumulant of order j for a system with fixed N_w nucleons, $\kappa_j^{corr}[N]$ the cumulant including the measurable corrections for volume fluctuations, and Δ_j the corresponding bias due to quantities that are not measurable. The second-order result we already derived in Sec. A, Eqs. (15) and (14), namely

$$\kappa_2^{corr}[N] = \kappa_2[N] - \frac{\langle N \rangle^2}{\langle M \rangle^2} C_2[M]$$

$$\Delta_2 = \frac{\langle N \rangle^2}{\langle M \rangle^2} \bar{C}_2[M].$$
(24)

For the third order we have

$$\kappa_3^{corr}[N] = \kappa_3[N] - \frac{3C_2[M]\kappa_2[N]\langle N \rangle}{\langle M \rangle^2} + \frac{3C_2[M]^2 \langle N \rangle^3}{\langle M \rangle^4} - \frac{C_3[M]\langle N \rangle^3}{\langle M \rangle^3}
\Delta_3 = \bar{C}_2[M] \left(\frac{3\kappa_2[N]\langle N \rangle}{\langle M \rangle^2} - \frac{3C_2[M]\langle N \rangle^3}{\langle M \rangle^4} \right) + \frac{\bar{C}_3[M]\langle N \rangle^3}{\langle M \rangle^3}.$$
(25)

And the fourth order result reads

$$\kappa_{4}^{corr}[N] = \kappa_{4}[N] - \left(\frac{6\kappa_{2}[N]\langle N \rangle^{2} \left(C_{3}[M]\langle M \rangle - 3C_{2}[M]^{2}\right)}{\langle M \rangle^{4}} + \frac{4C_{2}[M]\kappa_{3}[N]\langle N \rangle}{\langle M \rangle^{2}} + \frac{3C_{2}[M]\kappa_{2}(N)^{2}}{\langle M \rangle^{2}} + \frac{\langle N \rangle^{4} \left(-10C_{3}[M]C_{2}[M]\langle M \rangle + 15C_{2}[M]^{3}\right)}{\langle M \rangle^{6}} + \frac{C_{4}[M]\langle N \rangle^{4}}{\langle M \rangle^{4}}\right) \\
\Delta_{4} = \bar{C}_{2}[M] \left(-\frac{18C_{2}[M]\kappa_{2}[N]\langle N \rangle^{2}}{\langle M \rangle^{4}} + \frac{15C_{2}[M]^{2}\langle N \rangle^{4}}{\langle M \rangle^{6}} - \frac{4C_{3}[M]\langle N \rangle^{4}}{\langle M \rangle^{5}} + \frac{4\kappa_{3}[N]\langle N \rangle}{\langle M \rangle^{2}} + \frac{3\kappa_{2}[N]^{2}}{\langle M \rangle^{2}}\right) \\
+ \bar{C}_{3}[M] \left(\frac{6\kappa_{2}[N]\langle N \rangle^{2}}{\langle M \rangle^{3}} - \frac{6C_{2}[M]\langle N \rangle^{4}}{\langle M \rangle^{5}}\right) + \frac{\bar{C}_{4}[M]\langle N \rangle^{4}}{\langle M \rangle^{4}}.$$
(26)

The corresponding relations for correction and bias terms of any order can be obtained with the provided software package [26]. A Python package is provided to derive analytical formulas for both mixed and pure cumulants of multiplicity distributions, including participant/volume fluctuations. The correction formulas and their bias terms can be derived as well. The package, including a dedicated graphical user interface, can be downloaded via Ref. [26]. Equivalent expressions for the factorial cumulants, $C_n[N]$, and their related biases, $\Delta_{n,F}$, may then be obtained by

Equivalent expressions for the factorial cumulants, $C_n[N]$, and their related biases, $\Delta_{n,F}$, musing the relation between factorial cumulants and regular cumulants

$$C_n^{corr}[N] = \sum_{j=1}^n s(n,j) \,\kappa_j^{corr}[N],$$
$$\Delta_{n,F} = \sum_{j=1}^n s(n,j) \,\Delta_j,$$

with s(n,j) denoting Stirling numbers of the first kind (cf. Eq. B6)

The results for the corrected factorial cumulants are given in Appendix E.

V. SIMULATIONS

Experimental data are usually analyzed in centrality percentiles by introducing selection criteria on, e.g., the energy deposited in a forward detector system covering typically the projectile (target) spectator region or the multiplicity of charged particles emitted from the mid-rapidity region, with an acceptance reaching close to the projectile (target) rapidity regions in case of low beam energies [27, 28]. For the latter, care must be taken to ensure that the evaluated particles are not simultaneously used to determine the critical fluctuations [25]. The respective distributions, like e.g., the forward energy deposit or the charged particle multiplicity, are commonly modelled using the Glauber Monte Carlo Model [29]. The model provides event by event and for a given impact parameter the number of projectile/target nucleons which are "wounded" and responsible for the event activity (multiplicity), and those, which proceed nearly undisturbed into the phase space region covered by the forward detectors. To determine centrality using charged particle multiplicity the respective distribution is generally modelled assuming that particles are "produced" independently from distinct sources following a negative binomial distribution (NBD). Its probability mass function is defined as

$$P(n;\mu,k) = \frac{\Gamma(n+k)}{\Gamma(n+1)\Gamma(k)} \left(\frac{\mu}{k}\right)^n \left(\frac{\mu}{k} + 1\right)^{-(n+k)},\tag{27}$$

where μ denotes the mean of the NBD, while the combination of μ and k determines its higher-order cumulants

$$\kappa_n^{NBD} = \left. \frac{\partial^n \ln M(t)}{\partial t^n} \right|_{t=0},\tag{28}$$

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$$M(t) = \sum_{n=0}^{\infty} e^{tn} P(n; \mu, k) = \left(\frac{k}{k + \mu(1 - e^t)}\right)^k$$
 (29)

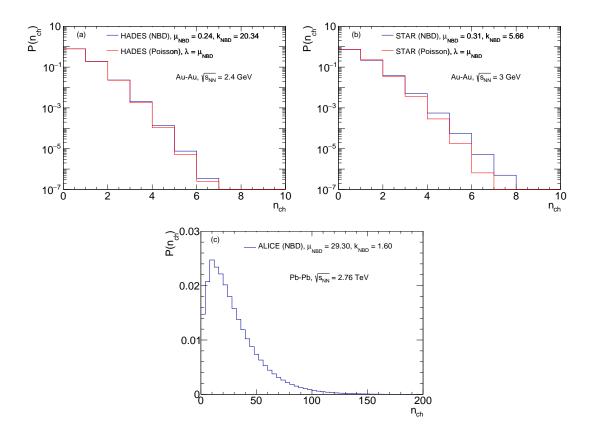


Figure 1. The NBD of charged particles per source (n_{ch}) adjusted to HADES (a) [27], STAR (b) [11], and ALICE (c) [28] is presented with blue lines. For the HADES and STAR data, corresponding Poisson distributions are also presented (red lines). The parameters used are listed in Table I.

being the moment-generating function of the NBD. The first four cumulants read

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$$\kappa_1^{\text{NBD}} = \mu, \tag{30}$$

$$\kappa_2^{\text{NBD}} = \frac{\mu(k+\mu)}{k},\tag{31}$$

$$\kappa_3^{\text{NBD}} = \frac{\mu(k+\mu)(k+2\mu)}{k^2},$$
(32)

$$\kappa_4^{\text{NBD}} = \frac{\mu(k+\mu)(k^2 + 6k\mu + 6\mu^2)}{k^3},$$
(33)

The parameters of the NBD are fixed in each experiment by the fitting procedure. In a first step the number of particle-emitting sources n_s is determined according to [11, 28]

$$n_{\rm s} = fN_w + (1 - f)N_{coll},$$
 (34)

where N_W and N_{coll} are the numbers of wounded nucleons and binary collisions, respectively. Sampling impact parameters according to $d\sigma = b db$ a list of number of sourses, n_s^i , is generated, with $i \in [1, \dots, N_{\text{Event}}]$. Then, for each event i, NBD is sampled n_s^i times and the parameters of the NBD, μ , k and f, are adjusted such that the obtained multiplicity distribution agrees with the corresponding experimental one. The mixing parameter f is introduced to improve the description by accounting also particles produced in hard (prompt) processes.

Panel (a) of Fig. 1 represents the NBD distribution as observed by the HADES experiment for Au+Au collisions at $\sqrt{s_{\mathrm{NN}}} = 2.4 \,\mathrm{GeV}$, with parameters $\mu = 0.24$, k = 20.34, and f = 1 taken from [27]. For comparison, we also present a Poisson distribution with the same mean, μ . Similar distributions from the STAR [11] and ALICE [28] experiments are presented in panels (b) and (c). Figure 1 shows that at the HADES energy the fitted NBDs are very close to Poisson distributions. Quantitatively this can be seen by evaluating the cumulants of the HADES NBD ($\mu = 0.24$,

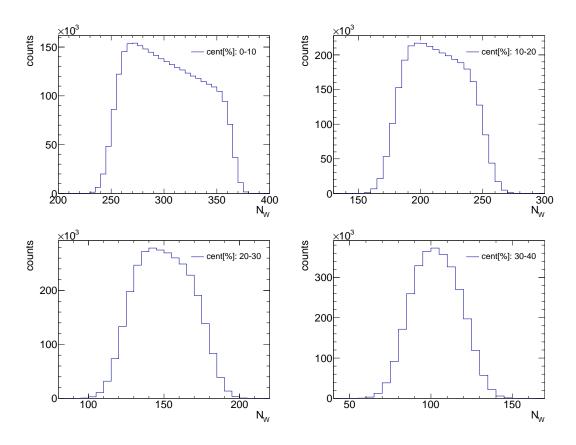


Figure 2. Distribution of wounded nucleons in Au+Au collisions at $\sqrt{s_{\rm NN}}=2.4$ GeV for four selected centrality classes, as obtained from the Glauber Monte Carlo simulations.

$$k = 20.34$$

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$$\kappa_1^{\text{NBD}}(\text{HADES}) = 0.24$$
(35)
$$\kappa_2^{\text{NBD}}(\text{HADES}) = 0.2428,$$
(36)
$$\kappa_3^{\text{NBD}}(\text{HADES}) = 0.2486,$$
(37)

$$\kappa_2^{\text{NBD}}(\text{HADES}) = 0.2428,\tag{36}$$

$$\kappa_3^{\text{NBD}}(\text{HADES}) = 0.2486,\tag{37}$$

$$\kappa_A^{\text{NBD}}(\text{HADES}) = 0.2602. \tag{38}$$

For a Poisson distribution all cumulants are equal to its mean and the HADES data are indeed close to fulfilling this condition. The statement, to a lesser extent, is also valid for the STAR Au+Au data at 3 GeV (see Fig. 1). The corresponding ALICE distribution, however, is much wider compared to the Poissonian baseline, but the ALICE NBD is obtained for very different acceptance than that used for the cumulant analysis. In Table I we also provide the NBD parameters as obtained by the STAR and ALICE collaborations for Au+Au and Pb+Pb collisions at $\sqrt{s_{\mathrm{NN}}}=3\,\mathrm{GeV}$ and 2.76 TeV, respectively.

experiments	μ	k	f
HADES	0.24	20.34	1
STAR	0.31	5.66	0.94
ALICE	29.3	1.6	0.8

Table I. NBD parameters as extracted from Glauber fits to particle distributions observed in different experiments. For simulations, the distributions should be folded within the experimental acceptance in which the cumulants are measured.

In the following we test the proposed method using two different simulations referred to as Model A and Model B. While the sampling of wounded nucleons is the same for both models, in Model A we sample different particle species independently while in Model B we introduce correlations between pions and protons via resonance production and decay. For both models A and B, we use the Glauber model to extract the distributions of wounded nucleons

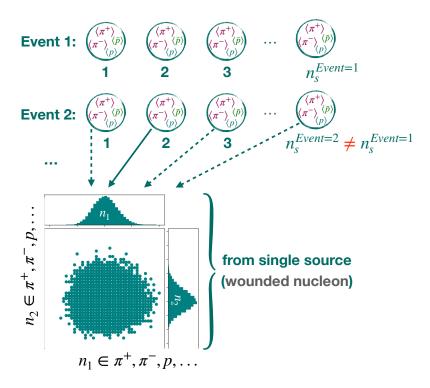


Figure 3. Simulation of particle production within the model of independent sources. The circles indicate the individual sources sampled according to Eq. 34. n_1 and n_2 show particle species used to sample emission from a single source. The distributions per single source can be chosen arbitrarily. See models A and B discussed below.

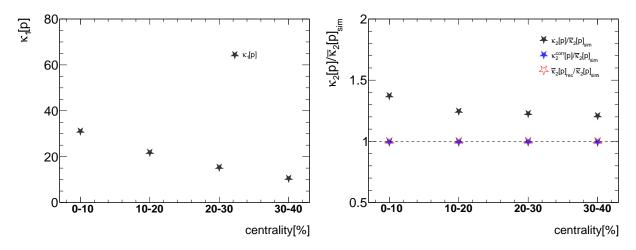


Figure 4. Left panel: Mean number of simulated protons used in model A as a function of centrality, adjusted to multiplicities measured by HADES in Au-Au collisions at $\sqrt{s_{\rm NN}} = 2.4\,{\rm GeV}$ [12]. Right panel: Second-order cumulants of protons in model A including volume fluctuations (black stars), corrected with Eq. 24 (blue stars) and reconstructed with Eq. 23 (open red stars). The results are normalized to $\bar{\kappa}_2[p]_{sim}$, corresponding to the second-order cumulants of protons in the absence of volume fluctuations.

corresponding to four different centrality classes in Au+Au collisions at $\sqrt{s_{\rm NN}}=2.4\,{\rm GeV}$. They are presented in Figure 2. Particles are produced by independent sources. (cf. Eq. 34). We note that other sources of correlation, such as those stemming from the reconstruction of closely spaced tracks and/or non-binomial efficiencies, were not considered in these simulations. The simulation process is schematically illustrated in Fig. 3.

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We will concentrate on the HADES data but also briefly discuss the high-energy limits, which concern the ALICE and STAR experiments. For the HADES data the extracted number of binary collision is zero (f = 1, see Table I), the number of sources per event are tehrefore sampled exclusively from the wounded nucleon distributions presented

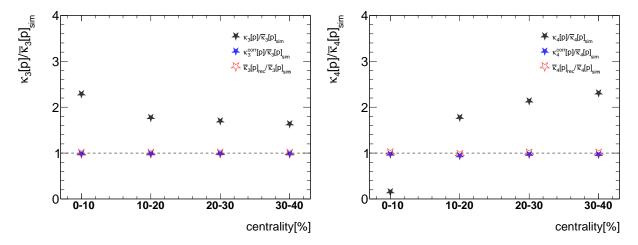


Figure 5. Left panel: Third order cumulants of protons in model A including volume fluctuations (black stars), corrected with Eq. 25 (blue stars) and reconstructed with Eq. 23 (open red stars). The results are normalized to $\bar{\kappa}_3[p]_{\rm sim}$, corresponding to the second-order cumulants of protons in the absence of volume fluctuations. Right panel: Fourth order cumulants of protons in model A including volume fluctuations (black stars), corrected with Eq. 26 (blue stars) and reconstructed with Eq. 23 (open red stars). The results are normalized to $\bar{\kappa}_4[p]_{\rm sim}$, corresponding to the fourth order cumulants of protons in the absence of volume fluctuations.

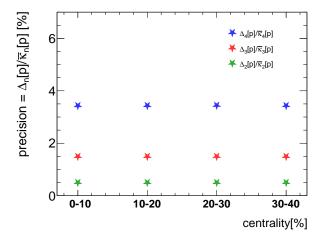


Figure 6. The normalised bias terms in model A.

s4 in Fig. 2.

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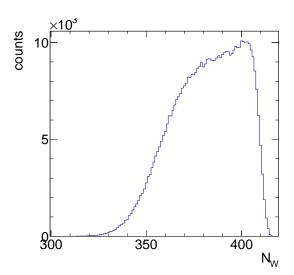
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VI. MODEL A

In model A we first generate the charged-particle multiplicity for individual events using the NDB distribution as extracted by experimental measurements. In doing so we sample the NBD distribution n_s times. Different particle species are then taken as fractions of the total number of charged particles. For example, from a randomly sampled NBD distribution a respective fraction is assigned to protons. From the remaining charged particles another fraction is assigned to positively charged pions and the rest is taken as negatively charged pions. These fractions are chosen such that the overall probability of having protons, positively and negatively charged pions correspond to 75%, 9% and 16% of all charged particles, respectively, based on the actual HADES measurement in Au+Au collisions (see [30] and references therein). In addition, we account for acceptance effects, because the NDB parameters are obtained in different acceptance than that used for the fluctuation analyses. We therefore fold the entire NDB distribution with a binomial distribution such that the experimentally measured mean multiplicities of particles in the acceptance used for fluctuation analysis are reproduced. Volume fluctuations are naturally accounted for as for each event the number of sources n_s are randomly sampled from the corresponding distributions.



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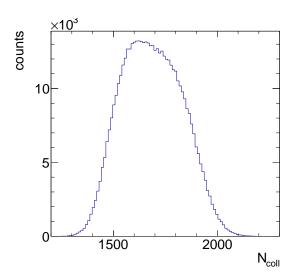


Figure 7. Number of wounded nucleons (left panel) and binary collisions (right panel) as generated with a Glauber Monte Carlo simulation using input from the ALICE experiment [28]. The selection corresponds to the 5% most central Pb-Pb collisions at $\sqrt{s_{\rm NN}} = 2.76$ TeV.

The simulated mean numbers of protons are shown in the left panel of Fig. 4 for the four centrality classes. In the right panel of Fig. 4 the reconstructed second-order cumulants of protons are presented, normalized to the expected true cumulant, $\bar{\kappa}_2[p]_{sim}$. The black stars represent the results which include volume fluctuations. The values κ_2^{corr} as calculated using Eq. 24 are shown with blue stars, while the red stars are evaluated using Eq. 23. The results for the third and fourth order cumulants are shown in Fig. 5. The corresponding normalized biases $\Delta_n/\bar{\kappa}_n$ are presented in Fig. 6. We find the normalized biases to be very small, of the order of a few percent, so that the corrected cumulants, κ_n^{corr} are very close to their expected true values, $\bar{\kappa}_n$. As already discussed, this is to be expected since the multiplicity distribution per wounded nucleon in the present model is close to Poisson.

To investigate the high-energy limit, we also apply the method to ALICE data. In doing so we first run Glauber Monte Carlo simulations for Pb-Pb collisions at $\sqrt{s_{\rm NN}} = 2.76$ TeV. The input parameters are taken from Ref. [28], which selects different centrality classes by introducing sharp cuts on the charged-particle distributions. The distributions of wounded nucleons and binary collisions corresponding to the 5% most central collisions are presented in Fig. 7 [23]. The reconstructed mean number of wounded nucleons and binary collisions, corresponding to the 5% most central collisions are $\langle N_W \rangle \approx 382$ and $\langle N_{coll} \rangle \approx 1685$, respectively, consistent with the numbers given in [28]. With these numbers one can estimate a mean number of particle emitting sources, yielding $\langle n_s \rangle \approx 642$ (see Eq. 34). The corresponding mean number of charged particles can be estimated as $\langle N_{ch} \rangle = \langle n_s \rangle \times \mu_{NBD} \approx 18811$. On the other hand, the total number of charged particles measured inside the ALICE acceptance is about 1601 [31]. We thus folded the ALICE NBD distribution with a binomial with the acceptance factor of $\epsilon = 1601/18811 \approx 8.5\%$ to obtain the distribution within the experimental acceptance. The so obtained NDB distribution from ALICE is presented in Fig. 8. We further note that only the acceptance in rapidity is considered. Fluctuation analyses are performed within a finite momentum range. Inclusion of the latter will further reduce the discrepancy between NBD and the corresponding Poisson distribution shown with the red histogram in Fig. 8. Finally using the NBD distribution presented in Fig. 8, and measured proton number, $\langle N_p \rangle \approx 35$ [31], we estimated $\Delta_2[p] \approx 1.2$. This corresponds to a bias of $\Delta_2[p]/\bar{\kappa}_2[p] \approx 3.3\%$ (cf. Eq. 24).

VII. MODEL B

In model B we introduce correlations between charged particles, specifically pions and protons by generating resonances. Especially for HADES energies most of the observed pions are believed to originate from decays of Delta resonances. Therefore, the effect of such decay correlations, while no treated quantitatively here, needs to be taken

¹ We note that binomial folding of the NDB distribution, with the acceptance factor ϵ , changes only the parameter μ of the original NDB distribution ($\mu \to \epsilon \mu$), while the parameter k stays unchanged.

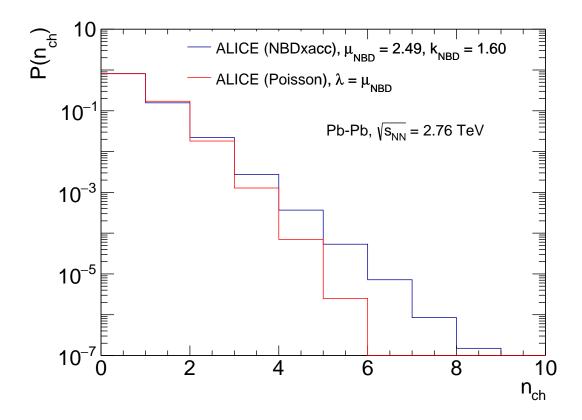


Figure 8. NDB distributions from ALICE as measured by fitting the signal amplitude in VZEROs (left panel) and after folding within the experimental acceptance 8.5% (see the text for details). The red histogram corresponds to a Poisson distribution with the same mean, μ as the folded NBD distribution. The original NBD distribution from ALICE is presented in the panel (c) of Fig. 1.

into account. Specifically, this is done by generating resonances from each source and letting them decay into two different particle species. Moreover, the resonances are generated from a Poisson distribution. In addition we produce independent particles from each source as well, sampled also from a Poisson distribution. Schematic illustration of the model for a single source is given in Fig. 9. Fluctuations of sources are introduced like in the model A. The input parameters for model B are given in Table II.

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particles	mean numbers per source	
resonances	0.03	
independent protons	0.23	
other particles	0.21	

Table II. Parameters for model B are mean numbers of different particles species per source (see [30] and references therein). In addition, each resonance decays into one proton and one pion. Numbers of resonances, independent protons and other particles are sampled from independent Poisson distributions.

By construction the simulated protons, pions and resonances follow a Poisson distribution, however, the distribution of the total number of particles does not, due to the correlation between pions and protons introduced via the resonance decay. Indeed, particle production through resonance decays enhances their fluctuations (see Appendix F).

In experiments measurements are performed inside the finite acceptance by imposing selection criteria in momentum space, e.g., on rapidity and/or transverse momentum of particles. Moreover, such conditions typically lead to different acceptances for different particle species. In order to study the impact of the finite acceptance on the presented formalism, we introduce arbitrary rapidity distributions for protons, pions and other particles as illustrated in the right panel of Fig. 9. To this end we generate rapidity values for protons, pions and other particles from the corresponding distributions presented in Fig. 9.

In the left panel of Fig. 10 we present mean multiplicities of protons produced via resonances (red circles) and

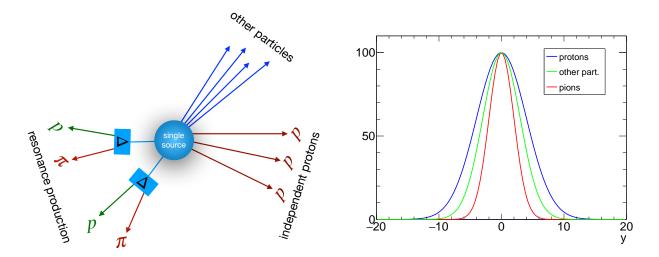


Figure 9. Left panel: Schematic illustration of model B. From a single source two resonances and three protons are produced. The resonances further decay to protons and pions. In simulations both, resonances and independent protons are sampled from Poisson distributions. The corresponding mean values of resonances and independent protons are provided in Table II. Right panel: Schematic rapidity distributions for different particles in order to study acceptance effects.

independently (blue circles), while the black circles correspond to the total number of protons. In addition, we produce pions from resonances, and, by construction, their mean values are equal to those of protons from resonances. The right panel of Fig. 10 shows the second-order cumulants of protons divided by the expected value $\bar{\kappa}_2[p]_{sim}$. The black stars represent those including participant (volume) fluctuations, $\kappa_2[p]/\bar{\kappa}_2[p]_{sim}$. The corrected cumulants $\kappa_2^{corr}[p]/\bar{\kappa}_2[p]_{sim}$ (see Eq. 24) are shown with blue symbols, while the open red stars represent the true reconstructed values of fluctuations of protons $\bar{\kappa}_2[p]/\bar{\kappa}_2[p]_{sim}$ as calculated using Eq. 23. Similar results for the third and fourth order cumulants of protons are presented in Fig. 11 (see Eqs. 25, 26, 23). In Fig. 12 the normalized cumulants as a function of cumulant order are presented for the 10% most central collisions. The right panel of Fig. 12 corresponds to the full acceptance, while in the right panel the results in the finite acceptance, delimited as |y| < 1, are presented. One clearly observes that in the finite acceptance the precision of the method is significantly better. In Fig. 13 we show the magnitude of the corresponding normalized biases, $\Delta_n/\bar{\kappa}_n$, for the full acceptance (left panel) and for |y| < 1 (right panel). While the biases for the full acceptance may at first sight appear rather large ($\sim 40\%$) one should realize that for the most central events the uncorrected fourth order cumulant is more than a factor of 50 larger in magnitude than the true cumulants. In other words the proposed corrections, while not perfect are a huge improvement of the measurement. The situation gets better for the limited acceptance.

VIII. DISCUSSION AND SUMMARY

- We have shown that using mixed events to determine the contributions of wounded nucleon or volume fluctuations is equivalent to extracting the latter from the track multiplicity distribution. However event mixing may offer an advantage since it allows to generate an almost arbitrarily large event ensemble with the same multiplicity distribution, and thus eliminate possible constraints due to limited event statistics. In either case, not all contributions can be accessed by direct measurement. The remaining terms lead to biases, Δ_k , which depend on the multiplicity distribution per wounded nucleon. These biases are, however, parametrically suppressed by powers of $\langle N \rangle / \langle M \rangle$ depending on the order of the cumulants. The biases are also small if the multiplicity distribution per wounded nucleon is close to Poisson. We suggest constraining these biases in experiments with fits to the observed multiplicity distribution within the wounded-nucleon model. Here, we refer to the Glauber Monte Carlo fits performed by each experiment. One of the outcomes of such a fitting procedure is the distribution of charged particles per source, as presented in Fig. 1 for the HADES case, which can be used to estimate the bias terms.
- We have worked here within the wounded nucleon model to formulate volume fluctuations. Alternatively, one may introduce generic volume fluctuations as done e.g. in [22]. It is easy to show (see Appendix D) that the resulting expressions for the corrected cumulants, κ_j^{corr} , and the biases, Δ_j , are identical to those derived here, i.e. Eqs. (24-26).

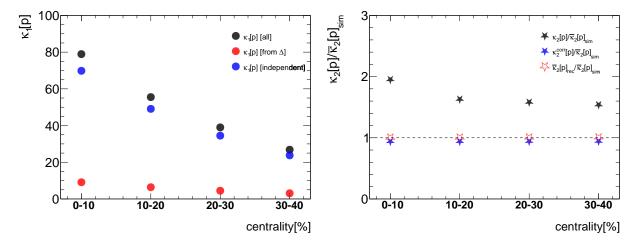


Figure 10. Left panel: Mean number of protons produced via resonances and independently are presented with the red and blue circles respectively. The black circles represent total multiplicity of protons. Pions are produced only via resonances. Right panel: Reconstructed second-order cumulants of protons including participant fluctuations (black stars). Corrected values for cumulants $\kappa_2^{corr}[p]$, i.e., without the bias term $\Delta_2[p]$ are presented with blue stars, while red stars represent fully corrected, against volume fluctuation. The results are normalized to the true second order cumulant, $\bar{\kappa}_2[p]_{sim}$.

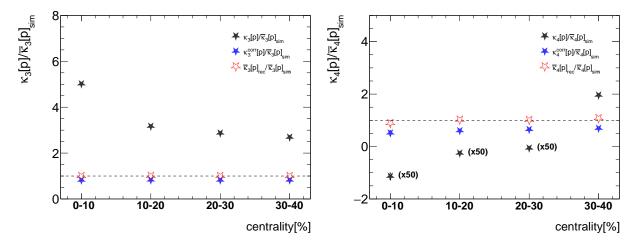


Figure 11. Left panel: Reconstructed third-order cumulants of protons including participant fluctuations (black stars) for model B. Corrected values for cumulants $\kappa_3^{corr}[p]$, i.e. without the bias term $\Delta_3[p]$ are presented with blue stars, while red stars represent fully corrected, against volume fluctuations, values $\bar{\kappa}_3[p] = \kappa_3^{corr}[p] + \Delta_3[p]$. Right panel: Similar to the left panel but for the fourth-order cumulants. Note that the values for the fourth-order cumulants with volume fluctuations (black stars) need to be multiplied by 50 for the first three centrality classes. The results are normalized to the true third or fourth order cumulant, $\bar{\kappa}_3[p]_{sim}$, $\bar{\kappa}_4[p]_{sim}$, respectively

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- We note that one gets similar expressions for the fluctuations from the wounded nucleons, Eqs. (20-22), in terms of the factorial cumulants of, for example, pions instead of the total track multiplicity. This has the advantage that the corrections do not involve the particles of interest, protons, in our case.
- We have checked that the proposed method also works if the multiplicity distribution is determined for a different acceptance than the particle distribution of interest. In this case all quantities in the expression for the corrected cumulants, Eqs.(24-26) involving the multiplicity should be evaluated in the multiplicity acceptance while all quantities involving the particles of interest, such as $\langle N \rangle$ or $\kappa_j[N]$ should be determined in the particle acceptance.
- We have verified that the proposed method is not affected by potentially different rapidity distributions for different particle types.
- We did not consider separate contributions from produced and stopped protons in the finally measured fluctu-

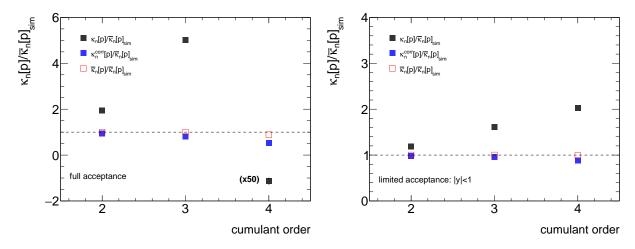


Figure 12. Left panel: Cumulants of protons in the full acceptance, presented for the 10% most central collisions (cf. Figs. 10 and 11). Right panel: Similar plot for the 10% most central collisions, but inside the finite acceptance delimited with the |y| < 1 criterion. The results are normalized to the true cumulants, $\bar{\kappa}_n[p]_{sim}$.

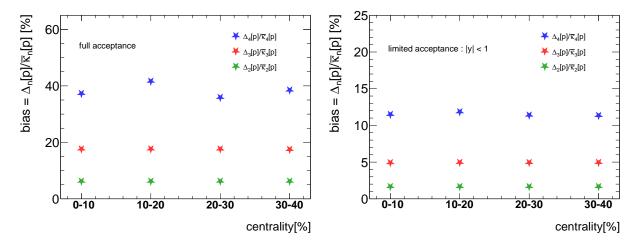


Figure 13. The bias terms for model B, in the full (left panel) and finite (right panel) acceptances. The acceptance, |y| < 1 is introduced using rapidity distributions of pions and protons as shown in Fig. 9.

ation signals. This is an important topic which goes beyond our studies presented in this work.

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- We note that the correction term and bias for the second-order cumulant depends on the properties of the multiplicity distribution only while those for higher-order cumulants also involve the (uncorrected) cumulants of interest (at a lower order), $\kappa_n[N]$ (see Eq. (25,26).
- The proposed method is also applicable for mixed cumulants, such as the covariance between two particle species. The relevant formulas for mixed cumulants between species A and B up to $\kappa_{2,2}[A,B]$ are given in Appendix G.
 - The corresponding relations for correction and bias terms for mixed cumulants of any order can also be obtained with the provided software package [26].
- Here we have not explicitly discussed corrections for net proton cumulants. Since the expressions for the cumulants of the net-particle distribution, in the presence of wounded nucleon fluctuations, are the same as Eqs. (1-(4)) with $N \to N \bar{N}$ [22, 23] our formulas above can be readily applied. For example for the second

order cumulant of the net-particle number distribution, we obtain using Eq.(24)

$$\kappa_2^{corr}[N - \bar{N}] = \kappa_2[N - \bar{N}] - \frac{\left(\langle N \rangle - \langle \bar{N} \rangle\right)^2}{\langle M \rangle^2} C_2[M]$$

$$\Delta_{2,N-\bar{N}} = \frac{\left(\langle N \rangle - \langle \bar{N} \rangle\right)^2}{\langle M \rangle^2} \bar{C}_2[M].$$
(39)

Alternatively and as a cross check, one can use the expression for the mixed cumulants provided in Appendix G:

$$\kappa_2[N - \bar{N}] = \kappa_2[N] + \kappa_2[\bar{N}] - 2\kappa_{1,1}[N, \bar{N}] \tag{40}$$

so that (see Eq. (G8))

$$\kappa_2^{corr}[N - \bar{N}] = \kappa_2[N - \bar{N}] - \frac{\left(\langle N \rangle - \langle \bar{N} \rangle\right)^2}{\langle M \rangle^2} C_2[M] \tag{41}$$

with the bias

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$$\Delta_{2,N-\bar{N}} = \frac{\left(\langle N \rangle - \langle \bar{N} \rangle\right)^2}{\langle M \rangle^2} \bar{C}_2[M] \tag{42}$$

For systems at vanishing baryon chemical potential, such as those created at very high collision energies around mid-rapidity, $\langle N \rangle = \langle \bar{N} \rangle$, the corrected second-order cumulant is identical to the measured one (as discussed in [22, 23]) and the bias vanishes.

In a similar way the correction formulas for mixed cumulants and the corresponding bias terms can be expressed for net particles. In this case the replacements $A \to A - \bar{A}$ and $B \to B - \bar{B}$ should be performed (see Eqs. G8-G14).

In summary, we have presented a method to correct experimentally measured particle number cumulants for the effect of volume fluctuations. The essential idea is to extract the contribution from the volume fluctuations from the distribution of charged particles which, after appropriate re-scaling, may be subtracted from the measured cumulants of interest. Our proposed method is not exact as there remains a bias or remnant which can not be accessed directly from experiment. However, we have shown by model calculations that this bias is very small compared to the contribution from participant fluctuations and we hence consider our method an important step towards measuring the true dynamical fluctuations of the system.

ACKNOWLEDGMENTS

V.K would like to thank GSI and the Institute for Nuclear Theory at the University of Washington for their kind hospitality and stimulating research environment. V.K. has been supported by the U.S. Department of Energy, Office of Science, Office of Nuclear Physics, under contract number DE-AC02-05CH11231, by the INT's U.S. Department of Energy grant No. DE-FG02-00ER41132, and by the ExtreMe Matter Institute EMMI at the GSI Helmholtzzentrum für Schwerionenforschung, Darmstadt, Germany.

Appendix A: Wounded-nucleon model

Here we briefly discuss the wounded-nucleon model following the Appendix of Ref. [32]. The wounded-nucleon model assumes that particles are produced by independent sources, called wounded nucleons or participants. Therefore, the probability to find A particles of type A and B particles of type B can be written as

$$P(A,B) = \sum_{w} W(w) \sum_{a_1, \dots a_w} \sum_{b_1, \dots b_w} p(a_1, b_1) \dots p(a_w, b_w) \, \delta_{A, \sum_{k=1}^w a_k} \delta_{B, \sum_{k=1}^w b_k}.$$
(A1)

Here W(w) denotes the probability to have w wounded nucleons, and p(a,b) is the probability to have a particles of type A and b particles of type B from one wounded nucleon. The moment-generating function, $h(t_A, t_B)$ is then

$$H(t_{A}, t_{B}) = \sum_{A,B} e^{t_{A}A} e^{t_{b}B} P(A, B)$$

$$= \sum_{w} W(w) \sum_{a_{1}, \dots a_{w}} \sum_{b_{1}, \dots b_{w}} p(a_{1}, b_{1}) \cdots p(a_{w}, b_{w}) e^{t_{A} \sum_{k=1}^{w} a_{k}} e^{t_{B} \sum_{k=1}^{w} b_{k}}$$

$$= \sum_{w} W(w) \sum_{a_{1}, b_{1}} p(a_{1}, b_{1}) e^{t_{A}a_{1} + t_{b}b_{1}} \cdots \sum_{a_{w}, b_{w}} p(a_{w}, b_{w}) e^{t_{A}a_{w} + t_{B}b_{w}}$$

$$= \sum_{w} W(w) \left[\sum_{a, b} p(a, b) e^{t_{A}a + t_{B}b} \right]^{w}$$

$$= \sum_{w} W(w) [h_{w}(t_{A}, t_{B})]^{w}$$

$$= \sum_{w} W(w) e^{w g_{w}(t_{A}, t_{B})}$$
(A2)

where $h_w(t_A, t_B) = \sum_{a,b} p(a,b) e^{t_A a + t_B b}$ is the moment-generating function and $g_w(t_A, t_B) = \ln [h_w(t_A, t_B)]$ the cumulant-generating function for one wounded nucleon, respectively. The cumulant-generating function, $G(t_A, t_B) = \ln [H(t_A, t_B)]$, is then given by

$$G(t_A, t_B) = \ln \left[H(t_A, t_B) \right] = \ln \left[\sum_{w} W(w) e^{w g_w(t_A, t_B)} \right] = G_W(g_w(t_A, t_B))$$
(A3)

where $G_W(t) = \ln \left[\sum_w W\left(w\right) e^{w\,t} \right]$ is the cumulant-generating function for the wounded-nucleon distribution, W(w).

We note, that $g_w(0,0) = G_W(0) = 0$ by construction. The cumulants are then obtained as

$$\kappa_{j,k}[A,B] = \left. \frac{\partial^{(j+k)}}{\partial t_A^j \partial t_B^k} G(t_A, t_B) \right|_{t_A = t_B = 0}.$$
(A4)

For example, denoting $\kappa_{i,j}[a,b]$ as the cumulants of the distribution from one wounded nucleon, we have

$$\kappa_1[A] = \frac{\partial}{\partial t_A} \left. G(t_A, 0) \right|_{t_A = 0} = \frac{dG_w}{dg_w} \left. \frac{dg_w(t_A, 0)}{dt_A} \right|_{t_A = 0} = \left. \frac{dG_w}{dg_w} \right|_{g_w = 0} \left. \frac{dg_w(t_A, 0)}{dt_A} \right|_{t_A = 0} \\
= \kappa_1[N_w] \kappa_1[a] = \langle N_w \rangle \langle a \rangle \tag{A5}$$

where $\kappa_1[n] = \langle a \rangle$ denotes the mean number of particles of type A per wounded nucleon and $\kappa_1[w] = \langle N_w \rangle$ the mean number of wounded nucleons. The second-order cumulant is

$$\begin{split} \kappa_{2}[A] &= \frac{\partial^{2}}{\partial t_{A}^{2}} \left. G(t_{A}, 0) \right|_{t_{A} = 0} \\ &= \frac{d^{2}G_{w}}{dg_{w}^{2}} \left. \left(\frac{dg_{w}(t_{A}, 0)}{dt_{A}} \right)^{2} \right|_{t_{A} = 0} + \frac{dG_{w}}{dg_{w}} \left. \frac{d^{2}g_{w}(t_{A}, 0)}{dt_{A}^{2}} \right|_{t_{A} = 0} \\ &= \kappa_{2}[N_{w}] \kappa_{1}[a]^{2} + \kappa_{1}[N_{w}] \kappa_{2}[a] = \kappa_{2}[N_{w}] \left\langle a \right\rangle^{2} + \left\langle N_{w} \right\rangle \kappa_{2}[a] \end{split} \tag{A6}$$

The covariance is

$$cov[A, B] = \frac{\partial^{2}}{\partial t_{A} \partial t_{B}} G(t_{A}, t_{B}) \Big|_{t_{A}, t_{B} = 0} = \frac{\partial}{\partial t_{B}} \left(\frac{dG_{w}}{dg_{w}} \frac{\partial g_{w}(t_{A}, t_{B})}{\partial t_{A}} \right) \Big|_{t_{A}, t_{B} = 0}$$

$$= \frac{d^{2}G_{w}}{dg_{w}^{2}} \frac{\partial g_{w}(t_{A}, t_{B})}{\partial t_{A}} \frac{\partial g_{w}(t_{A}, t_{B})}{\partial t_{B}} \Big|_{t_{A}, t_{B} = 0} + \left(\frac{dG_{w}}{dg_{w}} \frac{\partial^{2}g_{w}(t_{A}, t_{B})}{\partial t_{A} \partial t_{B}} \right) \Big|_{t_{A}, t_{B} = 0}$$

$$= \kappa_{2}[N_{W}] \kappa_{1}[a] \kappa_{1}[a] + \kappa_{1}[N_{w}] cov[a, b]$$

$$= \kappa_{2}[N_{w}] \langle a \rangle \langle b \rangle + \langle N_{w} \rangle cov[a, b]$$
(A7)

Using the relation between the cumulant and factorial cumulant-generating function, Eq. B3, the factorial cumulant generating function is given by

$$G_F(z_A, z_B) = G(\ln(z_A), \ln(z_B)) = G_W(g_w(\ln(z_A), \ln(z_B))) = G_W(g_{F,w}(z_A, z_B)),$$
 (A8)

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$$g_{F,w}\left(z_{\mathrm{A}},z_{\mathrm{B}}\right) = ln\left[\sum_{a,b} p\left(a,b\right) z_{\mathrm{A}}^{a} z_{\mathrm{B}}^{b}\right]$$

the factorial cumulant-generating function for the distribution of one wounded nucleon, p(a,b). The structure is the same as for the cumulant-generating function, except that now the argument of the wounded-nucleon cumulant-generating function is the factorial cumulant-generating function, $g_{F,w}$. Thus the factorial cumulants are easily obtained by simply replacing all the cumulants of the particle distribution for a given wounded nucleon, $\kappa_{i,j}[a,b]$ with the corresponding factorial cumulants, with $C_1[X] = \kappa_1[X] = \langle X \rangle$

$$C_{1}[A] = \frac{\partial}{\partial z_{A}} G_{F}(z_{A}, 1)|_{z_{A}=1} = \kappa_{1}[N_{w}]C_{1}[a] = \langle N_{w} \rangle \langle a \rangle$$

$$C_{2}[A] = \kappa_{2}[N_{w}] \langle a \rangle^{2} + \langle N_{w} \rangle C_{2}[a]$$

$$C_{1,1}[A, B] = cov[A, B]$$
(A9)

Appendix B: Cumulant and factorial cumulant-generating functions

Given a multiplicity distribution for particles of type A and B, P(A,B) the generating functions for cumulants, $g(t_A,t_B)$, and factorial cumulants, $g_F(z_A,z_B)$ are given by

$$g(t_A, t_B) = \ln \left[\sum_{A,B} P(A, B) e^{t_A A} e^{t_B B} \right]$$
(B1)

$$g_F(z_A, z_B) = \ln \left[\sum_{A,B} P(A, B) (z_A)^A (z_B)^B \right].$$
 (B2)

By construction, $g(t_A = 0, t_B = 0) = 0$ and $g_F(z_A = 1, z_B = 1) = 0$. Cumulants of order (i, j), $\kappa_{j,k}[A, B]$, are then obtained through

$$\kappa_{j,k}[A,B] = \left. \frac{\partial^{(j+k)}}{\partial t_A^j \partial t_B^k} g(t_A, t_B) \right|_{t_A = t_B = 0},$$

while the factorial cumulants, $C_{j,k}[A,B]$, are given by

$$C_{j,k}[A,B] = \left. \frac{\partial^{(j+k)}}{\partial z_{\rm A}^j \partial z_{\rm B}^k} g_F\left(z_{\rm A}, z_B\right) \right|_{z_{\rm A}=z_{\rm D}=1}.$$

The generating functions are related through

$$g_F(z_A, z_B) = g[\ln(z_A), \ln(z_B)]$$
 (B3)

or vice versa

$$g(t_A, t_B) = g_F\left(e^{t_B}, e^{t_B}\right) \tag{B4}$$

These relations may also be used to convert cumulants into factorial cumulants and vice versa [33]. For example, for the diagonal cumulants, $\kappa_n[A]$ we have

$$\kappa_n[A] = \sum_{j=1}^n S(n,j) C_j[A]$$
(B5)

where S(n,j) denotes the Stirling numbers of the second kind. The inverse relation is

$$C_n[A] = \sum_{j=1}^n s(n,j) \kappa_j[A], \tag{B6}$$

with s(n,j) denoting the Stirling numbers of the first kind. For the first four orders this evaluates to

$$\kappa_1 = C_1
\kappa_2 = C_1 + C_2
\kappa_3 = C_1 + 3C_2 + C_3
\kappa_4 = C_1 + 7C_2 + 6C_3 + C_4$$
(B7)

and

$$C_{2} = \kappa_{2} - \kappa_{1}$$

$$C_{3} = 2\kappa_{1} - 3\kappa_{2} + \kappa_{3}$$

$$C_{4} = -6\kappa_{1} + 11\kappa_{2} - 6\kappa_{3} + \kappa_{4}$$
(B8)

Appendix C: Multiplicity Distribution

The multiplicity distribution, P(M), is given by summing over all (charged) particles,

$$P_{M}(M) = \sum_{A,B,X} P(A,B,X) \,\delta_{M,A+B+X},\tag{C1}$$

340 where

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$$P(A, B, X) = \sum_{w} W(w) \sum_{a_{1}, \dots a_{w}} \sum_{b_{1}, \dots b_{w}} \sum_{x_{1}, \dots x_{w}} p(a_{1}, b_{1}, x_{1}) \cdots p(a_{w}, b_{w}, x_{w}) \ \delta_{A, \sum_{k=1}^{w} a_{k}} \delta_{B, \sum_{k=1}^{w} a_{k}} \delta_{X, \sum_{k=1}^{w} x_{k}}$$

is the distribution of particles of type A, B and all others, denoted by X. The distribution for particles A and B defined in Appendix A are then given by $P(A,B) = \sum_{X=0}^{\infty} P(A,B,X)$ while that for the particles per wounded nucleons are given by $P(a,b) = \sum_{x=0}^{\infty} P(a,b,x)$

The moment-generating function is then given by (proceeding analogously to Eq. A2:

$$\begin{split} H_{M}(t) &= \sum_{M} P_{M}(M) \, e^{tM} = \sum_{M,A,B,X} P(A,B,X) \delta_{M,A+B+X} \, e^{tM} = \sum_{A,B,X} P(A,B,X) \, e^{t(A+B+X)} \\ &= \sum_{w} W\left(w\right) \sum_{a_{1},\cdots a_{w}} \sum_{b_{1},\cdots b_{w}} \sum_{x_{1},\cdots x_{w}} p\left(a_{1},b_{1},x_{1}\right) \cdots p\left(a_{w},b_{w},x_{w}\right) \, e^{t\sum_{k=1}^{w} a_{k}} e^{t\sum_{k=1}^{w} b_{k}} e^{t\sum_{k=1}^{w} x_{k}} \\ &= \sum_{w} W\left(w\right) \sum_{a_{1}} \sum_{b_{1}} \sum_{x_{1}} p\left(a_{1},b_{1},x_{1}\right) e^{t(a_{1}+b_{1}+c_{1})} \cdots \sum_{a_{w}} \sum_{b_{w}} \sum_{x_{w}} p\left(a_{w},b_{w},x_{w}\right) e^{t(a_{w}+b_{w}+c_{w})} \\ &= \sum_{w} W\left(w\right) \left[\sum_{a} \sum_{b} \sum_{x} p\left(a,b,x\right) e^{t(a+b+c)}\right]^{w} \\ &= \sum_{w} W\left(w\right) \left[h_{m,w}(t)\right]^{w} \\ &= \sum_{w} W\left(w\right) e^{wg_{m,w}(t)} \end{split}$$

where, $h_{m,w}(t)$ the moment-generating function and $g_{m,w}(t) = \ln [h_{m,w}(t)]$ the cumulant-generating function of the multiplicity distribution for *one* wounded nucleon, $p(m) = \sum_{a,b,x} p(a,b,x) \delta_{m,a+b+x}$. The cumulant-generating function, $G_M(t)$, for the multiplicity distribution, P(M), is then given by

$$G_M(t) = \ln\left[H_M(t)\right] = \ln\left[\sum_{w} W(w) e^{wg_{m,w}(t)}\right] = G_W(g_{m,w}(t))$$
 (C2)

The cumulants of the multiplicity distribution are given by (following the analogous Eqs. A5 and A6)

$$\kappa_1[M] = \langle N_w \rangle \langle m \rangle
\kappa_2[M] = \kappa_2[N_w] \langle m \rangle^2 + \langle N_w \rangle \kappa_2[m]$$
(C3)

where $\kappa_i[m]$ denote the cumulants of the multiplicity distribution of *one* wounded nucleon and $\kappa_i[N_W]$ those of the wounded nucleon distribution. Analogous to Eq. A8 the factorial cumulant-generating function is readily obtained

$$G_{F,M}(z) = G_M(\ln(z)) = G_W(g_{m,w}(\ln(z))) = G_W(g_{F,m,w}(z)),$$
 (C4)

with

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$$g_{F,m,w}(z) = \ln \left[\sum_{m} p(m) z^{m} \right],$$

the factorial cumulant-generating function for the multiplicity distribution of one nucleon, p(m). Again, the factorial cumulants are obtained by replacing the cumulants of the distribution p(m), $\kappa_i[m]$ with the corresponding factorial cumulants, $C_i[m]$, in Eq. C3 by the factorial cumulants

$$C_1[M] = \langle N_w \rangle \langle m \rangle$$

$$C_2[M] = \kappa_2[N_w] \langle m \rangle^2 + \langle N_w \rangle C_2[m]$$
(C5)

Appendix D: Wounded Nucleon vs Volume Fluctuations

Here we will discuss the relation between wounded nucleon fluctuations [23] and so-called volume fluctuations as they are discussed e.g. in [22]. Following Ref. [22] the cumulant-generating function is given by

$$\Phi(t) = \ln \left[\int dV P(V) e^{V\xi(t)} \right] = \chi^{V}(\xi(t))$$
(D1)

with $\chi^{V}(t)$ the cumulant-generating function for the distribution of volumes, P(V), and

$$\xi(t) = \frac{1}{V} \ln \left[\sum_{N} p(N; V) e^{Nt} \right]$$
 (D2)

the generating function for scaled cumulants, κ/V , given for the distribution of particles at fixed volume, p(N;V).

Then, for a fixed volume V, the scaled cumulants are given by

$$\frac{\kappa_j}{V} = \frac{\partial^j}{\partial t^j} \, \xi(t)|_{t=0} \,. \tag{D3}$$

For the wounded nucleon model we have (see Appendix A)

$$G(t) = \ln \left[\sum_{w} W(w) e^{w g_w(t)} \right] = G_W(g_w(t))$$
 (D4)

with $G_W(t)$ the cumulant-generating function for the wounded nucleon distribution, W(w), and

$$g_w(t) = \ln\left[\sum_n p(n)e^{nt}\right] \tag{D5}$$

the generating function for the distribution of particles for one wounded nucleon. The cumulants per wounded nucleons for a fixed number of wounded nucleons, N_w , are given by

$$\frac{\kappa_j[N]}{N_w} = \kappa_j[n] = \frac{\partial^j}{\partial t^j} g_w(t)|_{t=0}.$$
 (D6)

Comparing the above expressions, one finds that the cumulants for volume fluctuations can be obtained from those for the wounded-nucleon number by the following replacements

$$\begin{split} \kappa_j[N_w] \to \kappa_j[V] \\ \bar{\kappa}_j[N] = \langle N_w \rangle \, \kappa_j[N] \to \langle V \rangle \, \frac{\kappa_j}{V} \end{split}$$

Indeed comparing the second-order cumulants for both scenarios we have

$$\kappa_{2}[N] = \langle N_{w} \rangle \, \kappa_{2}[n] + \langle n \rangle^{2} \, \kappa_{2}[N_{w}] = \bar{\kappa}_{2}[N] + \langle N \rangle^{2} \, \frac{\kappa_{2}[N_{w}]}{\langle N_{w} \rangle^{2}}$$

$$\kappa_{2}[N] = \langle V \rangle \, \frac{\kappa_{2}}{V} + \left(\frac{\kappa_{1}}{V}\right)^{2} \kappa_{2}[V] = \bar{\kappa}_{2}[N] + \langle N \rangle^{2} \, \frac{\kappa_{2}[V]}{\langle V \rangle^{2}}$$

where in the second line we used $\kappa_1 = \langle N \rangle$ and $\bar{\kappa}_2 = \langle V \rangle \frac{\kappa_2}{V}$. Obviously, analogous replacements also hold for the factorial cumulants

$$\bar{C}_j[N] = \langle N_w \rangle \, C_j[n] \to \langle V \rangle \, \frac{C_j}{V}$$

with C_i/V the volume scaled factorial cumulants.

Appendix E: Results for factorial cumulants

Here we provide the formulas for the corrected factorial cumulants, C_k^{corr} and c_k^{corr} , and the associated biases, $\Delta_{k,F}$ and $\delta_{k,F}$. Both the factorial cumulants and the biases are related to the corresponding cumulants via the linear relation Eqs. B7 and B8. The corrected factorial cumulants and the associated biases are

$$C_2^{corr} = C_2[N] - \frac{\langle N \rangle^2}{\langle M \rangle^2} C_2[M]$$
 (E1)

$$C_3^{corr} = C_3[N] - \frac{3C_2[M]C_2[N]\langle N \rangle}{\langle M \rangle^2} + \frac{3C_2[M]^2 \langle N \rangle^3}{\langle M \rangle^4} - \frac{C_3[M]\langle N \rangle^3}{\langle M \rangle^3}$$
 (E2)

$$C_4^{corr} = C_4[N] - \left(\frac{6C_2[N]\langle N \rangle^2 \left(C_3[M]\langle M \rangle - 3C_2[M]^2\right)}{\langle M \rangle^4} + \frac{4C_2[M]C_3[N]\langle N \rangle}{\langle M \rangle^2}\right)$$
(E3)

$$+\frac{3C_{2}[M]C_{2}[N]^{2}}{\langle M \rangle^{2}} + \frac{\langle N \rangle^{4} \left(-10C_{3}[M]C_{2}[M]\langle M \rangle + C_{4}[M]\langle M \rangle^{2} + 15C_{2}[M]^{3}\right)}{\langle M \rangle^{6}}$$
 (E4)

$$\Delta_{2,F} = \frac{\langle N \rangle^2}{\langle M \rangle^2} \bar{C}_2[M] \tag{E5}$$

$$\Delta_{3,F} = \bar{C}_2[M] \left(\frac{3C_2[N]\langle N \rangle}{\langle M \rangle^2} - \frac{3C_2[M]\langle N \rangle^3}{\langle M \rangle^4} \right) + \frac{\bar{C}_3[M]\langle N \rangle^3}{\langle M \rangle^3}$$
 (E6)

$$\Delta_{4,F} = \bar{C}_2[M] \left(\frac{4C_3[N]\langle N \rangle}{\langle M \rangle^2} + \frac{3C_2[N]^2}{\langle M \rangle^2} - \frac{18C_2[M]C_2[N]\langle N \rangle^2}{\langle M \rangle^4} - \frac{4C_3[M]\langle N \rangle^4}{\langle M \rangle^5} + \frac{15C_2[M]^2\langle N \rangle^4}{\langle M \rangle^6} \right)$$
(E7)

$$+ \bar{C}_3[M] \left(\frac{6C_2[N]\langle N \rangle^2}{\langle M \rangle^3} - \frac{6C_2[M]\langle N \rangle^4}{\langle M \rangle^5} \right) + \frac{\bar{C}_4[M]\langle N \rangle^4}{\langle M \rangle^4}$$
 (E8)

For the scaled factorial cumulants we have

$$c_2^{corr} = c_2[N] - \frac{\langle N \rangle}{\langle M \rangle} c_2[M] \tag{E9}$$

$$c_3^{corr} = c_3[N] - 3\frac{\langle N \rangle}{\langle M \rangle}c_2[M]c_2[N] + \left(\frac{\langle N \rangle}{\langle M \rangle}\right)^2 \left(3c_2[M]^2 - c_3[M]\right)$$
 (E10)

$$c_4^{corr} = c_4[N] - \frac{\langle N \rangle}{\langle M \rangle} \left(3c_2[M]c_2[N]^2 + 4c_2[M]c_3[N] \right)$$
(E11)

$$+\left(\frac{\langle N\rangle}{\langle M\rangle}\right)^2 \left(18c_2[M]^2c_2[N] - 6c_3[M]c_2[N]\right) \tag{E12}$$

$$+\left(\frac{\langle N\rangle}{\langle M\rangle}\right)^3 \left(-15c_2[M]^3 + 10c_2[M]c_3[M] - c_4[M]\right) \tag{E13}$$

$$\delta_{2,F} = \frac{\langle N \rangle}{\langle M \rangle} \bar{c}_2[M] \tag{E14}$$

$$\delta_{3,F} = 3 \frac{\langle N \rangle}{\langle M \rangle} c_2[N] \bar{c}_2[M] + \left(\frac{\langle N \rangle}{\langle M \rangle}\right)^2 (\bar{c}_3[M] - 3c_2[M] \bar{c}_2[M]) \tag{E15}$$

$$\delta_{4,F} = \frac{\langle N \rangle}{\langle M \rangle} \left(3c_2[N]^2 \bar{c}_2[M] + 4c_3[N] \bar{c}_2[M] \right)$$
(E16)

$$+\left(\frac{\langle N\rangle}{\langle M\rangle}\right)^2 \left(6c_2[N]\bar{c}_3[M] - 18c_2[M]c_2[N]\bar{c}_2[M]\right) \tag{E17}$$

$$+ \left(\frac{\langle N \rangle}{\langle M \rangle}\right)^{3} \left(15c_{2}[M]^{2} \bar{c}_{2}[M] - 6c_{2}[M]\bar{c}_{3}[M] - 4c_{3}[M]\bar{c}_{2}[M] + \bar{c}_{4}[M]\right)$$
(E18)

Appendix F: Particle production through cluster decays

Let us assume that particles are produced via clusters and that each cluster further decays into two particles.

Moreover, clusters are generated from a Poisson distribution. As each cluster decays into two particles the probability of measuring k particles is equivalent to measuring k/2 clusters and can be presented as:

$$p(k;\langle N_{cl}\rangle) = e^{-\langle N_{cl}\rangle} \frac{\langle N_{cl}\rangle^{k/2}}{(k/2)!}$$
 (F1)

The corresponding moment-generating function reads:

$$M(t) = \sum_{k/2=0}^{\infty} e^{tk} e^{-\langle N_{cl} \rangle} \frac{\langle N_{cl} \rangle^{k/2}}{(k/2)!} = e^{\langle N_{cl} \rangle \left(e^{2t} - 1 \right)}, \tag{F2}$$

where $\langle N_{cl} \rangle$ denotes mean number of clusters produced.

The cumulants of total particle number k can be computed as:

$$\kappa_n[k] = \left. \frac{\partial^n}{\partial t^n} ln(M(t)) \right|_{t=0}$$
 (F3)

For the first two cumulants one gets:

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$$\kappa_1[k] = 2\langle N_{cl} \rangle \tag{F4}$$

$$\kappa_2[k] = 4\langle N_{cl} \rangle$$
(F5)

One clearly sees from Eqs. F4 and F5 that $\kappa_1(k) \neq \kappa_2(k)$, i.e the total number of particles does not follow a Poisson distribution, although the clusters do. Moreover, one observes that particle production through clusters enhances fluctuations. In general, for clusters following a Poisson distribution and decaying into m particles, the cumulants of total particle number can be written as:

$$\kappa_n[k] = m^n \langle N_{cl} \rangle \tag{F6}$$

We note that Eq. F6 also applies to particle production through resonance decays.

Appendix G: Mixed Cumulants

Here we provide the relevant formulas for mixed cumulants. Given the generating function, Eq. A3, the mixed cumulants for particles of type A and B are given by (see Eq. A4)

$$\kappa_{j,k}[A,B] = \frac{\partial^{(j+k)}}{\partial t_A^j \partial t_B^k} g(t_A, t_B) \bigg|_{t_A = t_B = 0}$$
(G1)

The explicit formulas for the four lowest-order mixed cumulants are:

$$\kappa_{1,1}[A,B] = \bar{\kappa}_{1,1}[A,B] + \langle A \rangle \langle B \rangle \frac{\kappa_2 [N_W]}{\langle N_W \rangle^2}$$
(G2)

$$\kappa_{2,1}[A,B] = \bar{\kappa}_{2,1}[A,B] + (2\langle A \rangle \bar{\kappa}_{1,1}[A,B] + \langle B \rangle \bar{\kappa}_{2,0}[A,B]) \frac{\kappa_2[N_W]}{\langle N_W \rangle^2} + \langle A \rangle^2 \langle B \rangle \frac{\kappa_3[N_W]}{\langle N_W \rangle^3}$$
(G3)

$$\kappa_{1,2}[A,B] = \bar{\kappa}_{1,2}[A,B] + (\langle A \rangle \bar{\kappa}_{0,2}[A,B] + 2\langle B \rangle \bar{\kappa}_{1,1}[A,B]) \frac{\kappa_2 [N_W]}{\langle N_W \rangle^2} + \langle A \rangle \langle B \rangle^2 \frac{\kappa_3 [N_W]}{\langle N_W \rangle^3}$$
(G4)

$$\kappa_{2,2}[A,B] = \bar{\kappa}_{2,2}[A,B] + \left(\langle A \rangle^2 \bar{\kappa}_{0,2}[A,B] + 4\langle A \rangle \langle B \rangle \bar{\kappa}_{1,1}[A,B] + \langle B \rangle^2 \bar{\kappa}_{2,0}[A,B]\right) \frac{\kappa_3 \left[N_W\right]}{\langle N_W \rangle^3} \\
+ \left(2\langle A \rangle \bar{\kappa}_{1,2}[A,B] + 2\langle B \rangle \bar{\kappa}_{2,1}[A,B] + 2\bar{\kappa}_{1,1}[A,B]^2 + \bar{\kappa}_{0,2}[A,B]\bar{\kappa}_{2,0}[A,B]\right) \frac{\kappa_2 \left[N_W\right]}{\langle N_W \rangle^2} \\
+ \langle A \rangle^2 \langle B \rangle^2 \frac{\kappa_4 \left[N_W\right]}{\langle N_W \rangle^4} \tag{G5}$$

where, analogous to the notation for the regular cumulants, $\bar{\kappa}_{j,k}[A,B]$ denotes the mixed cumulant for constant number of wounded nucleons $\langle N_W \rangle$. Note, that $\kappa_{j,0}[A,B] = \kappa_j[A]$ and $\kappa_{0,j}[A,B] = \kappa_j[B]$ correspond to the regular cumulant for particles of type A and B respectively. The first order mixed cumulant, $\kappa_{1,1}[A,B] = cov[A,B]$ is also referred to as the covariance between the distributions of particles A and B. In order to obtain the corrected mixed cumulants we proceed in the same fashion as for the regular cumulant. We express the terms involving cumulants of the wounded nucleons, $\kappa_i[N_W]/\langle N_W \rangle^i$ in terms of the factorial cumulants of the multiplicity distribution (See Eqs.

20-22) and solve for the mixed cumulants with fixed number of wounded nucleons, $\bar{\kappa}_{j,k}[A,B]$. Again, the results are given in the form

$$\bar{\kappa}_{i,k}[A,B] = \kappa_{i,k}^{corr}[A,B] + \Delta_{i,k} \tag{G6}$$

where $\kappa_{j,k}^{corr}[A,B]$ are the cumulants including the measurable corrections and $\Delta_{j,k}$ are the corresponding biases due to quantities which are not directly measurable.

$$\kappa_{1,1}^{corr}[A,B] = \kappa_{1,1}[A,B] - \frac{\langle A \rangle \langle B \rangle}{\langle M \rangle^2} C_2[M] \tag{G7}$$

$$\Delta_{1,1} = \frac{\langle A \rangle \langle B \rangle}{\langle M \rangle^2} \bar{C}_2[M] \tag{G8}$$

$$\kappa_{2,1}^{corr}[A,B] = \kappa_{2,1}[A,B] - \frac{\langle B \rangle C_2[M]\bar{\kappa}_{2,0}[A,B]}{\langle M \rangle^2} - \frac{2\langle A \rangle C_2[M]\kappa_{1,1}[A,B]}{\langle M \rangle^2} + \frac{\langle A \rangle^2 \langle B \rangle \left(2C_2[M]^2 - \langle M \rangle C_3[M]\right)}{\langle M \rangle^4}$$
(G9)

$$\Delta_{2,1} = \frac{1}{\langle M \rangle^4} \left[2\langle A \rangle \langle M \rangle^2 \bar{C}_2[M] \kappa_{1,1}[A, B] + \langle B \rangle \langle M \rangle^2 \bar{C}_2[M] \bar{\kappa}_{2,0}[A, B] - \langle A \rangle^2 \langle B \rangle \left(\bar{C}_2[M] \left(\bar{C}_2[M] + C_2[M] \right) - \langle M \rangle \bar{C}_3[M] \right) \right]$$
(G10)

$$\kappa_{1,2}^{corr}[A,B] = \kappa_{1,2}[A,B] - \frac{\langle A \rangle C_2[M] \bar{\kappa}_{0,2}[A,B]}{\langle M \rangle^2} - \frac{2\langle B \rangle C_2[M] \kappa_{1,1}[A,B]}{\langle M \rangle^2} + \frac{\langle B \rangle^2 \langle A \rangle \left(2C_2[M]^2 - \langle M \rangle C_3[M]\right)}{\langle M \rangle^4}$$
(G11)

$$\Delta_{1,2} = \frac{1}{\langle M \rangle^4} \left[2 \langle B \rangle \langle M \rangle^2 \bar{C}_2[M] \kappa_{1,1}[A, B] + \langle A \rangle \langle M \rangle^2 \bar{C}_2[M] \bar{\kappa}_{0,2}[A, B] - \langle A \rangle \langle B \rangle^2 \left(\bar{C}_2[M] \left(\bar{C}_2[M] + C_2[M] \right) - \langle M \rangle \bar{C}_3[M] \right) \right]$$
(G12)

$$\kappa_{2,2}^{corr}[A,B] = \kappa_{2,2}[A,B] + \frac{2\langle A \rangle^2 C_2[M]^2 \bar{\kappa}_{0,2}[A,B]}{\langle M \rangle^4} + \frac{2\langle B \rangle^2 C_2[M]^2 \bar{\kappa}_{2,0}[A,B]}{\langle M \rangle^4} \\
- \frac{C_2[M] \bar{\kappa}_{0,2}[A,B] \bar{\kappa}_{2,0}[A,B]}{\langle M \rangle^2} - \frac{\langle A \rangle^2 C_3[M] \bar{\kappa}_{0,2}[A,B]}{\langle M \rangle^3} - \frac{\langle B \rangle^2 C_3[M] \bar{\kappa}_{2,0}[A,B]}{\langle M \rangle^3} \\
+ \frac{12\langle A \rangle \langle B \rangle C_2[M]^2 \kappa_{1,1}[A,B]}{\langle M \rangle^4} - \frac{2C_2[M] \kappa_{1,1}[A,B]^2}{\langle M \rangle^2} - \frac{2\langle A \rangle C_2[M] \kappa_{1,2}[A,B]}{\langle M \rangle^2} \\
- \frac{2\langle B \rangle C_2[M] \kappa_{2,1}[A,B]}{\langle M \rangle^2} - \frac{4\langle A \rangle \langle B \rangle C_3[M] \kappa_{1,1}[A,B]}{\langle M \rangle^3} - \frac{10\langle A \rangle^2 \langle B \rangle^2 C_2[M]^3}{\langle M \rangle^6} \\
+ \frac{8\langle A \rangle^2 \langle B \rangle^2 C_3[M] C_2[M]}{\langle M \rangle^5} - \frac{\langle A \rangle^2 \langle B \rangle^2 C_4[M]}{\langle M \rangle^4} \tag{G13}$$

$$\begin{split} \Delta_{2,2} &= \bar{C}_2[M]^2 \left(-\frac{\langle A \rangle^2 \bar{\kappa}_{0,2}[A,B]}{\langle M \rangle^4} - \frac{\langle B \rangle^2 \bar{\kappa}_{2,0}[A,B]}{\langle M \rangle^4} + \frac{3C_2[M] \langle A \rangle^2 \langle B \rangle^2}{\langle M \rangle^6} \right) \\ &+ \bar{C}_2[M] \left(-\frac{C_2[M] \langle A \rangle^2 \bar{\kappa}_{0,2}[A,B]}{\langle M \rangle^4} - \frac{C_2[M] \langle B \rangle^2 \bar{\kappa}_{2,0}[A,B]}{\langle M \rangle^4} + \frac{\bar{\kappa}_{0,2}[A,B] \bar{\kappa}_{2,0}[A,B]}{\langle M \rangle^2} \right. \\ &- \frac{2\bar{C}_3[M] \langle A \rangle^2 \langle B \rangle^2}{\langle M \rangle^5} - \frac{12C_2[M] \langle A \rangle \langle B \rangle \kappa_{1,1}[A,B]}{\langle M \rangle^4} + \frac{2\langle A \rangle \kappa_{1,2}[A,B]}{\langle M \rangle^2} + \frac{2\kappa_{1,1}[A,B]^2}{\langle M \rangle^2} \\ &+ \frac{2\langle B \rangle \kappa_{2,1}[A,B]}{\langle M \rangle^2} + \frac{6C_2[M]^2 \langle A \rangle^2 \langle B \rangle^2}{\langle M \rangle^6} - \frac{2C_3[M] \langle A \rangle^2 \langle B \rangle^2}{\langle M \rangle^5} \right) \\ &+ \bar{C}_3[M] \left(\frac{\langle A \rangle^2 \bar{\kappa}_{0,2}[A,B]}{\langle M \rangle^3} + \frac{\langle B \rangle^2 \bar{\kappa}_{2,0}[A,B]}{\langle M \rangle^3} + \frac{4\langle A \rangle \langle B \rangle \kappa_{1,1}[A,B]}{\langle M \rangle^3} - \frac{4C_2[M] \langle A \rangle^2 \langle B \rangle^2}{\langle M \rangle^5} \right) \\ &+ \frac{\bar{C}_2[M]^3 \langle A \rangle^2 \langle B \rangle^2}{\langle M \rangle^6} + \frac{\bar{C}_4[M] \langle A \rangle^2 \langle B \rangle^2}{\langle M \rangle^4} \end{split} \tag{G14}$$

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