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Do employment centers matter? Consequences for commuting distance in the Los Angeles region, 2002-2019

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27 Abstract

The presence of employment centers provides the potential for reducing commuting 28 29 distance. However, employment centers have distinctive attributes, which may lead to varied impacts on commuting outcomes. We attempt to examine the effect of distance to the nearest 30 31 employment center on commuting distance while addressing the heterogeneity of employment centers and workers. We consider multiple attributes of employment centers 32 related to location, persistency, job density, industry diversity, and size. For workers, we 33 mainly focus on low- and high-income groups as they differ in several aspects, such as 34 availability of commute modes, options for housing, and preference for job opportunities. We 35 applied a fixed-effects model using data from 2002 to 2019 to capture within-tract variation. 36 37 Our analysis of the Los Angeles region shows that increasing proximity to the nearest employment center decreases commuting distance even after controlling for the job attributes 38 located in the neighborhood of workers. The results show that employment centers are not 39 equal in terms of their impact on commute distance and that their impact is different for 40 commuters from different income groups. Our results contribute to the literature by 41 42 deciphering the location and attributes of employment centers that may exert a greater impact on commuting distances. 43

44 **1. Introduction**

45 The urban spatial structure of global cities has changed substantially during recent decades. Starting from a monocentric form, the decentralization of employment has occurred with the 46 47 growth of metropolitan areas. A significant factor for understanding the decentralization process is whether employers co-locate and create agglomeration economies. If employers 48 co-locate, the urban spatial structure will evolve to a polycentric form with multiple 49 employment centers. Otherwise, employment would decentralize, showing a general pattern 50 of dispersion, where the polycentric form is rather viewed as a transitional stage (Glaeser et 51 52 al., 2001a). Recent studies on the spatial structure of US metropolitan areas are mixed regarding the decentralization process. Based on a comparison between 1990 and 2010, 53 54 Arribas-Bel & Sanz-Graciz (2014) find that the monocentric form dominates in US 55 metropolitan areas over time. Using the case of the Los Angeles region, Giuliano et al. (2019, 2022) suggest that the polycentric structure is persistent with relatively small changes over 56 57 time in the boundary and the number of jobs in employment centers. Relatedly, Li (2020) also 58 found that Chinese cities have evolved towards concentration while also showing some trend of decentralization. 59

60 The presence of multiple employment centers creates a potential for reducing commuting distance (Wang, 2000). In theory, employees benefit from agglomeration economies by 61 62 sharing job opportunities, and the decentralization of employment increases the possibility of 63 jobs-housing proximity and balance (e.g., Horner, 2002; Loo & Chow, 2011; Zhao et al., 2011). A few empirical findings also suggest that shorter distances to employment centers 64 (e.g., central business district (CBD), subcenters) are associated with shorter commutes (e.g., 65 66 Ding et al., 2017; Islam & Saphores, 2022; Zhu et al., 2022). However, the dispersion of employment can sometimes complicate commuting patterns, often leading to long-distance 67 commuting. Evidence of either cross- or inverse-commuting implies that workers do not 68 necessarily commute to the nearest employment center, further suggesting that workers might 69 be influenced by all existing centers as well as the considerable number of jobs not located 70 71 within employment centers (Koster & Rouwendal, 2013).

Employment centers have distinctive attributes, which may lead to different impacts on
commuting outcomes. Employment centers differ in size, density, persistence, the
composition of industries, and their geographical location. For instance, Giuliano et al. (2019)

75 showed that employment centers are specialized in certain industry types, with the 76 composition of industries also related to the peak density of centers. In addition, the heterogeneity in attributes of employees also impacts whether the commuting outcome of a 77 78 worker is sensitive to the presence of employment centers. Income level, for example, is associated with transportation mobility, residential location constraints, and the industry type 79 of jobs, which leads to a different relationship between the location of employment centers 80 81 and commuting distance (Hu & Schneider, 2017). To elaborate, the commuting distance of high income workers may be more sensitive to the distance to employment centers as there 82 are more professional jobs clustered in these areas. In contrast, low income workers may 83 easily find substitute employment outside of employment centers, which could weaken the 84 relationship between commuting distance and the distance to employment centers. 85

In this study, we attempt to examine the effect of distance to the nearest employment center 86 87 (i.e., either CBD or subcenter) on commuting distance while addressing the heterogeneity of employment centers and workers. We consider multiple attributes of employment centers 88 89 related to location, persistency, job density, level of industry diversity, and size. For workers, we mainly focus on low- and high-income groups as they differ in several aspects, 90 91 particularly their ability to have certain housing, job, and commuting options. We apply a 92 fixed effects model, which arguably provides stronger causal evidence given that it estimates how change within tracts over time impacts change in our outcome measure, compared to 93 94 typical regression models that compare across geographic units. Our analysis of the Los 95 Angeles region shows that increasing proximity to the nearest employment center reduces commuting distance even after controlling for the job attributes located in the neighborhood 96 97 of workers. The results show that employment centers are not equal in terms of their impact on commute distance and that their impact differs across commuters from high or low income 98 99 groups.

While testing the relationship between employment centers and commuting distance is not new, we revisit the relationship for three reasons. First, existing works have less considered whether the distance to the nearest employment center influences commuting outcomes. We explicitly explore whether the location and attributes of the nearest employment center and the distance to it influences commuting distance. Second, while existing literature reports mixed results on the effect of employment centers on commuting, their analysis typically relies on cross-sectional data. We revisit the hypothesis by using longitudinal data and fixed-

107 effects models that provide a more robust analytical approach. Lastly, the heterogeneity across employment centers is not a focus in the literature. While numerous studies have 108 suggested shorter commuting distances for workers living close to employment centers (e.g., 109 Ding et al., 2017; Islam & Saphores, 2022; Zhu et al., 2022), there is limited understanding of 110 the characteristics of employment centers that are influential and the type of workers that are 111 affected by the presence of centers. Overall, we fill the gap in the literature by testing the 112 113 relationship between employment centers and commuting with a focus on the heterogeneity among employment centers and workers using longitudinal data from 2002 to 2019. 114

115 **2. Literature Review**

116 2.1 Agglomeration economies and employment centers

Employment centers (e.g., CBD, subcenters) exist as a result of concentration, in which 117 they support the idea of agglomeration economies. In particular, the employment center is 118 119 one critical component for understanding the mechanisms of learning, matching, and sharing 120 among firms, households, as well as retailers (Duranton & Puga, 2004). While these centers 121 are generally located near transportation infrastructures, the employment centers further provide insights related to transportation and planning strategies. For instance, a polycentric 122 123 urban form with multiple employment centers suggest that transport infrastructures should be designed to connect the centers and planning regulations should permit mixed land uses in 124 125 those centers (Angel & Blei, 2016). If these employment centers no longer exist or include a relatively small portion of jobs within the region, transportation strategies become less 126 127 relevant with the location of those centers. As such, understanding the employment center(s) is one important way to identify the spatial distribution of activities in cities, examine the 128 129 connections between workplaces and workers, and devise transportation policies.

Employment centers may not exist if firms gain fewer benefits from clustering in certain 130 131 locations. Some researchers have argued that the rapid growth in ICT technology and reduced transport costs have made clustering unnecessary, implying a pivot toward decentralized 132 employment (Glaeser et al., 2001b; Mitchell, 1996). Relatedly, some studies suggest that 133 decentralization is a common pattern in US cities (e.g., Gordon & Richardson, 1996; Lee, 134 2007) as there are increasing numbers of jobs in non-center areas. However, there is also 135 136 research providing support for the persistence of the polycentric urban form in cities worldwide (Li & Derudder, 2022, Giuliano et al., 2019; 2022; Phelps & Ohashi, 2020). 137

Giuliano et al. (2022), for instance, find that the polycentric form is relatively persistent based on a comparison between 1990 and 2009 in Los Angeles. Through an investigation of the Los Angeles region, Kane et al. (2018) report that the number of centers increased from 46 to 53 between 1997 and 2014. Their study also shows that the percent of jobs within centers has slightly increased from 17.4% to 19.6%. Another study by Cortright (2015) find that the average employment growth rate in city centers was slightly larger than the peripheral areas based on an analysis of large metropolitan areas in the U.S.

One challenging task involved in understanding urban spatial structure when measuring 145 employment centers is the methodological approach. At least three methods have been 146 suggested for identifying areas with employment concentrations large enough to influence 147 148 rent, distribution of employment or population, or employment density: (1) using cutoffs for size and employment density (Giuliano & Small, 1991), (2) using density gradients 149 150 (McDonald, 1987), and (3) using nonparametric regression to identify centers based on the density surface (McMillen, 2001; Redfearn, 2007). Each methodological approach has its 151 152 advantages and disadvantages, in that different methods can result in over- or underestimation of employment centers and are often sensitive to the extent of the study area. While some 153 154 studies have suggested ways to quantify concentration and decentralization without 155 identifying employment centers (Hipp et al., 2022), we do not address them here.

Furthermore, our understanding of employment centers and urban spatial structure could 156 157 be different depending on the approach applied by researchers. The size and number of 158 centers may differ depending on how we define and identify employment centers. By applying the Giuliano & Small (1991) method to the Los Angeles region, Giuliano et al. 159 (2019) showed that the employment centers and the polycentric form have been quite 160 161 persistent between 1980 and 2010. In contrast, Kane et al. (2018) found that employment centers in the Los Angeles region have exhibited a great variation in shape, size, location, and 162 163 industrial composition over time. Given that there is no gold standard for identifying employment centers, most studies have selected an approach based on their research question. 164 165 See Yu et al. (2021) for a detailed description of how employment centers can be identified 166 differently using existing methodological approaches.

167 2.2 Employment centers and commuting distance

168

Employment centers attract trips and have a structuring influence on regions, further

having an impact on commuting outcomes. One early study by Cervero & Wu (1997) found
that workers in suburban employment centers experience shorter commutes in terms of trip
times and are more likely to commute by driving. While there are numerous studies on the

172 relationship between urban form and commuting, we only focus on commuting distance in

173 this literature review. Commuting time is affected by factors such as mode choice,

174 congestion, and transportation infrastructure (Wang, 2000), and thus provides limited insight

175 regarding whether employees actually co-locate.

In theory, the spatial dispersion of employment opportunities can either increase or decrease commuting distance; compared to the monocentric form, the dispersion of jobs can create an environment where commuters can live closer to work, while it simultaneously allows random commuting (i.e., cross-commuting or inverse-commuting leading to longer commute distances) (Bertaud, 2002; Ma & Banister, 2007; Ha et al., 2021). Relatedly, empirical findings in the literature are also mixed.

Studies on how the dispersion of jobs impacts commuting could be categorized into those 182 focusing on 1) the location of employment centers or 2) the polycentricity of urban form. A 183 184 few studies have shown that the distance to employment centers (e.g., CBD, subcenters) is positively associated with commuting distance (Ding et al., 2017; Kim et al., 2012; 185 Grunfelder & Nielsen, 2012). Similar results are also found in studies that focus on the 186 relationship between polycentricity and commuting distance (Veneri, 2010; Zhao et al., 187 188 2011). The results are slightly mixed when employment density is controlled in the model; Ding et al. (2017) showed that the distance to CBD is positively associated with commuting 189 190 distance, while employment density is not significantly related to commuting distance. Relatedly, Islam & Saphores (2022) showed that employment density and distance to CBD 191 192 both have impacts on commuting distance, while the effect of distance to subcenters is 193 insignificant.

The relationship between commuting distance and employment centers depends on several factors such as availability of faster transportation modes, increasing number of dual-worker households, availability of hybrid access to jobs, and preferences for housing locations as well as limited housing affordability (e.g., Islam & Saphores, 2022; Wolday et al., 2019; Schuetz, 2020). In addition, employment centers may have less effect on commuting in areas with an increasing number of jobs in non-center areas (Angel & Blei, 2016). For instance, for commuters in households with multiple workers, preferences for a certain neighborhood, and

unaffordable housing prices near the workplace, may lead to longer commutes even if they
live close to employment centers. Using the case of Paris, Aguilera (2005) shows that most
people residing in a subcenter work outside of the employment cluster, while the majority of
employees of a subcenter commute from distant locations. Other studies also suggest that the
co-location hypothesis is insignificant, in which polycentric cities rather increase the length
of commute trips (Guth, 2010; Grunfelder, 2015).

207 2.3 Heterogeneity in commuters and employment centers

Commuters are heterogeneous, having different socio-demographic characteristics (e.g., 208 gender, income, job industry, occupation). Numerous studies have shown the differences in 209 210 commuting among population groups with a focus on locations of residences and workplaces (Hanson & Pratt, 1988; Hu & Schneider, 2017; Sun et al., 2017; Kim et al., 2012; Maoh & 211 212 Tang, 2012). Here, we particularly focus on existing research that addresses heterogeneity in 213 workers by income given that income is associated with other factors such as age, job type, and preference for housing locations. In a study of the Chicago region, Wang et al. (2021) 214 215 show that there is heterogeneity in residential location preferences across income groups; for example, low income households are less likely to decentralize due to limited financial 216 217 capacity. The authors also suggest that high income households may value the urban amenities and job opportunities located in the regional center. Relatedly, Cervero & Wu 218 219 (1997) show that high housing prices in and near employment centers may lead to longer 220 commute distances. The authors also report that professional workers in suburban employment centers tend to live in nearby housing. Some low-income workers also have 221 limited travel modes available, and studies report that they tend to have longer commute 222 times even though their commuting distances are relatively short (Renne & Bennett, 2014). In 223 224 contrast, another study shows that the average commuting time in the United States was shorter for workers below 200% poverty level (25 min.) compared to workers above 350% 225 poverty level (28 min.) in 2020 (National Equity Atlas, 2023). Relatedly, Blumenberg & Ong 226 (2001) explain that low-income workers experience difficulties in finding job opportunities 227 far from their homes due to limited mobility. 228

Another important aspect of workers is the industry type of jobs. Employment centers can vary in their industry compositions. For instance, access to customers is more important in population-serving jobs, which are widely distributed across space, whereas professional

services are more likely to benefit from spatial clustering (Giuliano et al., 2019). On one 232 hand, Giuliano et al. (2019) show that industries such as information, professional and 233 business services, health care and social assistance are more likely to be located in 234 employment centers; on the other hand, the manufacturing and retail trade sectors are more 235 likely to be located outside centers. The income level of workers can differ according to 236 industry types, which implies that some workers benefit more from agglomeration 237 238 economies. Similarly, studies have shown that the agglomeration effect and co-location patterns are heterogeneous across creative employment groups and occupation (e.g., Cruz & 239 Teixeira, 2015; Kim et al., 2012). 240

Employment centers and commuters are heterogeneous, which may help explain the 241 242 relationship between polycentricity and commuting outcomes. Some employment centers are more specialized, having a large share of one or two industry sectors, while others tend to 243 244 have a mix of industries (Giuliano et al., 2019; Wang et al., 2021). High-income workers are more likely to have the skillsets and interests that benefit more from agglomeration 245 246 economies, while low-income workers providing general services may find more job opportunities outside the centers (Hu & Schneider, 2017; Lee & Clarke, 2019). In these 247 248 respects, the co-location theory may hold true for certain employment centers and 249 commuters; for instance, employment centers mainly consisting of high-skilled and 250 professional job opportunities may allow reduced commuting distance for only the qualified workers. In other words, there could be a set of employment centers and commuters that both 251 252 benefit from agglomeration economies by spatially clustering and experiencing shorter 253 commuting distance.

3. Data and methods

255 **3.1 Los Angeles region**

The Los Angeles region is well known for its high level of polycentricity, with the presence of multiple employment centers. Here, we refer to the Los Angeles-Long Beach, CA Combined Statistical Area (CSA) as the Los Angeles region, which includes five counties: Los Angeles, Orange, Riverside, San Bernardino, and Ventura. The CSAs are identified for adjacent metro- and micropolitan areas with significant commuting flows indicating their interdependence. The five-county region in Los Angeles has been examined by multiple studies, such as Giuliano et al. (2019) and Kane et al. (2018). The region accommodated a

263 population of approximately 18 million and 7 million jobs, which is the second largest US

264 CSA unit following New York-Newark, NY-NJ-CT-PA CSA. The Los Angeles region has

been widely studied in previous research to examine hypotheses regarding agglomeration

economies and co-location (e.g., Kane et al., 2018; Giuliano et al., 2019). Empirical findings

267 from the literature based on the Los Angeles region show consensus and some conflicts

268 regarding the changes in urban spatial structure and employment centers. For instance,

269 Gordon & Richardson (1996) showed that the percentage of jobs in centers dropped over

time, further suggesting that agglomeration economies are declining.

271 Previous studies suggest that there are more than 30 employment centers in the Los 272 Angeles region. Giuliano & Small (1991) first suggested that there were 35 centers based on 273 the two-cutoffs approach (i.e., more than 10 jobs per acre and more than ten-thousand jobs total) using data from 1980. Another finding from Forestall & Greene (1997) found 120 274 275 centers based on a more generous approach. Most recently, Giuliano et al. (2019) identified 48 centers (95%/10K cutoffs) and 13 centers (99%/20K cutoffs) in 2009 by applying the 276 277 Giuliano & Small method, which shows that the results are sensitive to cutoff values. Kane et al. (2018) applied a non-parametric identification approach and found 53 centers in 2014. The 278 279 differences in the results can mainly be attributed to the data source for employment, the 280 spatial unit of analysis, and the identification approach. By using different approaches, interpretation of the persistence of urban form may differ: Giuliano et al. (2019) found a 281 persistent polycentric structure, while Kane et al. (2018) suggested that the boundaries and 282 283 industrial compositions of centers vary greatly across time.

284 Our study area includes the five counties that are within the Los Angeles region. There are 3924 census tracts located within our study area, which is our spatial unit of analysis in the 285 statistical models. We selected our study area to contribute to the long history of research on 286 urban spatial structure and commuting in this area. Additionally, the Los Angeles region is 287 288 unique since it exhibits one of the most polycentric structures, and it is a place that has experienced dynamic changes in terms of urban spatial structure and employment 289 290 decentralization. According to the Longitudinal Employer-Household Dynamics (LEHD), the 291 number of workers in the region increased from 13.7 million to 17.3 million between 2002 292 and 2019. Figures A1 to A4 illustrate the distribution of residential and workplace locations 293 for low and high income workers. For instance, we see low-income workers' residential 294 locations more concentrated in areas proximate to the downtown areas.

296 **3.2 Data**

295

We used the annual Origin Destination (OD) Employment Statistics data (2002 to 2019) from the LEHD. This data provides the aggregated number of workers based on their residence and workplace at the census block level. Based on the OD data from this source, we estimated the average commuting distance for each census tract by calculating the network distance between the centroids of census tracts. In doing so, we excluded data that have their origin or destination located outside of our study area. By using the Workplace Area Characteristics (WAC), we identified the location and characteristics of employment centers.

304 3.3 Identification of employment centers

305 While there is no perfect approach that ensures objectivity (Yu et al., 2021), we apply the 306 Giuliano & Small (1991) approach which allows us to identify employment centers in a more 307 consistent way over time (e.g., see Kane et al. (2018) which identifies employment centers in 308 the Los Angeles region using Redfearn's (2007) approach). In this approach, we identify 309 areas that are adjacent based on the cutoff settings for employment density and employment size. While we have several options for identifying employment centers, we use the 95th 310 311 percentile value for the density cutoff and 10,000 for the size cutoff. Using stricter cutoff values generally results in a smaller number of centers. Since we are interested in a larger 312 study area, we use a more generous approach. The spatial unit of analysis is also a factor that 313 may lead to different results. A recent study by Giuliano et al. (2019) suggests that 314 administrative units are inappropriate as they vary in shape and size across time and space. To 315 address this issue, we follow the method suggested by Giuliano et al. (2019) that uses a one-316 square-mile regular hexagon as the spatial unit. 317

We first created one-square-mile regular hexagons across our study area. We next merged census blocks to the hexagon that contained the block centroid and estimated the number of jobs for each hexagon. The LEHD data provides the number of jobs at a fine-grained scale, which provides more precision when aggregating to hexagons as our spatial unit. We used hexagons that exceeded the employment density criterion, and then used two cutoffs to identify the employment centers. We then used the inverse distance function to identify the employment centers, as suggested by McDonald & Prather (1994). Here, we apply stepwise

regression models to test each of the identified centers from the cutoffs, excluding centers that did not show a significant effect on the density gradient. We iterated this process for 18 different time periods (2002-2019) using the LEHD data.

328 **3.4 Variables**

Our dependent variable is the average commuting distance based on the LEHD OD data. Based on the OD data at the census tract level, we estimated the network distance between the origin and destination. We then estimated the average commuting distance for each census tract by using the number of workers as weights, and then log transforming this measure. We calculated this value annually for 18 years from 2002 to 2019.

Our main explanatory variable is the logged distance to the nearest employment center. 334 After identifying employment centers as described above, we estimated the network distance 335 to the edge of the nearest center from the centroid of each census tract. Furthermore, we have 336 337 a set of variables that address the characteristics of the nearest employment center. We identified the location of the nearest employment center and created dummy variables for 338 339 whether each center is located within Los Angeles City, Los Angeles County, and Orange County. These variables allow us to test if employment centers located in more dense and 340 centralized areas might affect commuting outcomes differently given their unique position 341 342 and development history (Giuliano et al., 2007; Giuliano et al., 2019). Since the employment centers are identified by using a hexagonal spatial unit, there are some cases where a center is 343 included in more than one administrative boundary. In this case, we only created a dummy 344 variable for the county that includes the largest share. For instance, if 40% of the center is 345 located within Los Angeles County and the remainder is located across Orange County, we 346 347 designated this center as being in Orange County.

Next, we created dummy variables indicating whether the center 1) is persistent across the 348 18 years of our analysis, 2) has high density, 3) has a high level of industrial diversity, 4) has 349 350 a high employment to population ratio, and 5) has a large size. Since the size and shape 351 change slightly across time, we designated centers as persistent if more than half of the center 352 area was consistently identified as an employment center. For the other four variables, we created dummy variables by focusing on the top quartile. For the level of industrial diversity, 353 354 we estimated the entropy index based on seven industry types according to the NAICS code¹. We here note that employment center(s) located in downtown area(s) may have different 355

effects on commuting distance. However, we do not treat them distinctively in our models
since there are limited number of employment center(s) in the downtown areas. Furthermore,
the main features of CBD's are arguably captured in several of our measures, including large
size, persistence, and high density.

For socio-demographics, we used the ratio of younger adults and the ratio of low-income 360 workers living in each census tract using the LEHD dataset. These two variables control for 361 age and income which are known to have effects on commuting outcomes (Ding & Bagchi-362 Sen, 2019; Ha et al., 2020). Middle-aged and workers from high income households show 363 364 longer commute distances (Axisa et al., 2012; Mercado & Páez, 2009). However, we do not 365 directly test this relationship but rather the association between the socio-demographic 366 composition and commute distances of the census tracts due to the aggregated nature of our data. In addition, we included residential neighborhood factors, mainly addressing the job 367 368 density within 3 km and the level of industry diversity for each census tract measured by using the entropy index. One of the main reasons for this approach was to control for the 369 370 effect of neighborhood attributes related to employment before testing the relationship between the distance to the nearest employment center and commuting distance. 371

372 **3.5 Methods**

373 We used a series of tract fixed effects (FE) models to estimate the relationship between commuting distance and distance to nearest employment center. Our outcome variable is the 374 log-transformed commuting distance measured at the census tract level; we applied log-375 transformation to adjust the right-skewed distribution and improve the linear relationship 376 with our measures. We have three types of explanatory variables: (1) socio-economic factors 377 (SF), (2) residential neighborhood factors (NF), and (3) the distance to nearest employment 378 center (DNC). Additionally, we have two types of variables to assess the interaction effects: 379 (1) location (LC) and (2) attributes (AC) of employment centers. We have three population 380 groups (all workers, low-income workers (i.e., jobs with earnings \$1250 per month or less), 381 and high-income workers (i.e., jobs with earnings greater than \$3333 per month)). For each 382 383 population group, we tested four models: (1) without interaction variables, (2) with interaction effects of center locations, (3) with interaction effects of center attributes, and (4) 384 385 with interaction effects of both center locations and attributes. The full model can be written 386 as:

389

 $ln(commuting distance)_t$

$$= \beta_1 SF_t + \beta_2 NF_t + \beta_3 DNC_t + LC_t + \beta_4 DNC_t \times LC_t + AC_t$$

$$+ \beta_5 DNC_t \times AC_t + \beta_6 N + \varepsilon_t$$

where t indicates the time of the data and N is a vector of indicator variables for all tracts inthe study area (the fixed effects).

The fixed effects model allows us to estimate within effects when units - in our case, 392 393 census tracts – are measured repeatedly (Firebaugh et al., 2013). The tract fixed effects demean commuting distances for each census tract, and therefore the only variation we are 394 395 estimating is whether the dependent variable is either below or above the mean value of each tract. This allows us to examine the relationship between the changes in independent 396 397 variables and changes in the dependent variable within each tract, rather than the assessing the relationship between independent and dependent variables across units. We used 398 399 frequency weights for the model based on the number of workers in each census tract; for the models of low- or high-income workers, we created weights based on the number of either 400 low- or high-income workers living within each census tract. The number of workers living 401 402 in each census tract varies greatly, which makes it appealing for weighting samples. Robust 403 standard errors were used with the jackknife function, and the analyses were performed using STATA 17. It should be noted that we did not consider using the random effects model, as it is 404 limited in its ability to provide reliable estimates of causal interest (Gunasekara et al., 2014). 405 406 We also report the results using a pooled linear regression model to provide information on how the results differ from the fixed-effects models. 407

408

409 **4. Results**

410 **4.1 Descriptive statistics**

411 **4.1.1 Average commuting distance**

The average commuting distance in the Los Angeles region was relatively stable, although it showed some increase between 2002 and 2008; the average commuting distance of all workers increased from 28.9 km to 31.5 km over the study period. As shown in Figure 1, high-income workers tended to commute longer distances compared to the low-income workers. Between 2002 and 2019, the average commuting distance slightly increased for

- 417 high-income workers from 30.7 km to 32.1 km, whereas it increased from 28.1 km to 30.9
- 418 km for low-income workers. In 2019, for example, the difference in average commuting



419 distance between the two commuter groups was 1.2 km.

420

421 Figure 1. Trends of average commuting distance (km) in Los Angeles region, 2002-2019.

422 **4.1.2 Employment centers in Los Angeles region**

Table 1 shows the descriptive characteristics of employment centers in the Los Angeles 423 region from 2002 to 2019 and Figure 2 compares the spatial distribution of employment 424 centers between 2002 and 2019. The number of employment centers was quite consistent, 425 ranging from 73 to 78. Comparing 2002 and 2019, the percentage of employment within 426 centers increased from 36.7% to 37.4%, which suggests that agglomeration economies are 427 persistent. Job density increased by 23.6% in areas defined as employment centers, while it 428 429 increased by 24.5% in areas not defined as employment centers. We note that the employment centers in the region contain a significant number of jobs; in 2019, for example, 430 the total area of employment centers within the urbanized area in the region accounts for only 431 4.7%, while they contain 37.4% of the employment. The median density of employment 432 centers also increased from 6901.6 to 7562.0 jobs per square-km between 2002 and 2019, 433 which also suggests that jobs tend to cluster more over time. The median entropy index of 434 industry composition and the employment to population ratio did not show significant 435 changes over time. The distribution of the attributes of employment centers are also shown in 436 437 Figure 3.

438 One thing to note is that the number of employment centers identified in this paper is

slightly larger than those from other recent studies. For instance, Giuliano et al. (2021) found 439 48 employment centers within the Los Angeles region in 2009. Another study by Kane et al. 440 (2018) identified 53 centers using the non-parametric estimation approach. The difference in 441 the results may come from several factors. First, different data sources were applied to 442 identify the spatial distribution of jobs. Specifically, Giuliano et al. (2021) use the National 443 Establishment Time-Series (NETS) data, whereas Kane et al. (2018) use the point-level 444 establishment data provided from Reference USA. The coverage of the data may affect the 445 results; for instance, using the 95th percentile cutoff based on employment density is sensitive 446 to the number of spatial units for analysis. Second, the employment center identification 447 approach also matters. As we have seen in the literature, the location, size, and shape of 448 employment centers vary across studies, which may be the reason for the differences in our 449 results. Finally, all these studies use arbitrary spatial units (e.g., hexagon or grid). Unlike 450 administrative spatial units, the arbitrary units may differ slightly according to how the 451 researchers have created them, which in turn may influence the outcomes. Nonetheless, the 452 453 number of employment centers does not change radically across the different strategies.





Figure 2. Los Angeles region employment centers, 2002 and 2019.

Year	Number of emp.	umberTotal% ofMean jobMean jobDensitf emp.emp.density indensity insquare		Density (square-kr	Employmer n)	nt per	Average entropy index of industry composition			Average employment to population ratio				
	centers	within centers	within centers	center areas	non- center areas	25%	50%	75%	25%	50%	75%	25%	50%	75%
2002	74	2,428,504	36.7	8,497.6	693.1	5,362.6	6,901.6	9,921.6	0.69	0.76	0.83	4.65	7.76	14.92
2003	76	2,373,249	35.7	8,458.0	706.7	5,343.3	6,915.1	9,407.4	0.71	0.77	0.83	4.59	7.53	14.48
2004	77	2,448,917	36.1	8,727.7	716.7	5,160.2	7,146.3	10,190.1	0.70	0.76	0.84	4.75	7.72	13.43
2005	78	2,491,027	35.8	8,960.7	738.3	5,380.3	7,252.5	10,247.8	0.70	0.77	0.85	4.36	7.19	13.59
2006	77	2,537,495	35.9	9,043.4	749.1	5,370.1	7,350.1	9,999.9	0.69	0.78	0.84	4.80	8.39	14.38
2007	78	2,589,488	36.3	9,144.0	751.6	5,540.6	7,732.3	10,093.8	0.66	0.77	0.83	4.67	7.34	14.36
2008	75	2,564,495	35.8	9,139.6	760.4	5,594.2	7,960.6	10,134.8	0.68	0.78	0.85	4.19	7.12	14.55
2009	75	2,529,593	36.6	9,015.2	724.5	5,420.2	7,268.8	9,746.7	0.66	0.77	0.85	4.69	7.49	15.77
2010	78	2,595,133	36.9	8,762.0	735.6	4,906.5	7,074.0	9,444.2	0.67	0.76	0.84	4.76	7.45	14.67
2011	77	2,657,576	37.4	8,818.1	738.0	5,216.4	6,950.1	9,600.7	0.69	0.76	0.85	4.72	7.43	14.25
2012	77	2,659,009	37.4	8,822.9	738.4	5,288.5	6,652.2	9,911.8	0.69	0.77	0.85	4.61	8.07	13.71
2013	77	2,663,214	36.5	9,152.4	767.4	5,472.5	6,839.9	9,944.8	0.67	0.76	0.85	4.66	7.98	14.05
2014	76	2,793,134	37.3	9,188.7	779.3	5,626.9	6,812.1	9,698.5	0.71	0.77	0.85	4.56	7.28	13.15
2015	74	2,842,030	37.1	9,512.2	799.1	6,091.2	7,349.4	10,151.2	0.68	0.77	0.84	4.56	7.26	13.77
2016	73	2,953,457	37.3	9,885.1	823.3	6,339.8	7,435.4	10,425.3	0.69	0.77	0.85	4.72	7.34	14.74
2017	77	2,987,113	37.2	9,911.6	836.6	6,320.7	7,522.4	9,814.8	0.69	0.75	0.84	4.65	7.06	12.62
2018	76	3,059,272	37.4	10,064.2	849.9	6,380.1	7,616.7	10,301.9	0.70	0.76	0.84	4.29	7.11	13.20
2019	77	3,111,324	37.4	10,504.8	863.2	6,391.2	7,562.0	10,550.1	0.70	0.77	0.83	4.37	7.54	12.68

457 Table 1. Descriptive characteristics of job density and employment centers in Los Angeles region, 2002-2019.



Figure 3. Distribution of employment center attributes, 2002-2019.

4.1.3 Correlation analysis 461

462 Table A1 and A2 show the correlation among our main variables (see Appendix). Table A1 shows that the job density and the industry mix of jobs measured at the neighborhood level 463 464 are negatively associated with the distance to the nearest employment center. This is an 465 expected result not only because generally there are a smaller number of jobs as the distance increases from employment centers but because those jobs tend to be less diverse often 466 dominated by retail or basic service providers in non-center locations. The correlation 467 coefficient for job density and industry mix of jobs within 3 km was 0.206, statistically 468 significant at the confidence level of 95%. Table A2 shows the correlation among the 469 attributes of employment centers. Persistency is positively correlated with job density and the 470 employment to population ratio. Employment centers with high industry diversity showed 471 negative correlations with job density and the employment to population ratio. Lastly, the size 472 of employment centers did not show significant correlation with other attributes. The results 473 474 generally show that the attributes of employment centers have low levels of correlation.

Average commuting distance models 4.2 475

4.2.1 All workers 476

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477 We next examine the estimates of the fixed effects models. Table 2 shows the model results; the outcome variable is logged commuting distance, and all workers are considered. 478 479 Starting with Model 1-1, census tracts with an increasing percentage of younger adult workers (i.e., aged 20-34) experience larger increases in commuting distance. And tracts with 480 481 an increasing percentage of low-income workers experience a larger decrease in commuting 482 distance. These results were all statistically significant and consistent across the four models presented in Table 2. Considering residential neighborhood factors, census tracts with 483 increasing job density within 3 km and decreasing industry diversity experienced decreasing 484 commuting distances. Nonetheless, even after controlling for socio-economic and residential 485 neighborhood factors, the distance to nearest employment center showed a significant and 486 positive sign. This result indicates that workers that experienced decreasing distance to an 487 employment center (either because the center is new, or the center expanded closer to them) 488 tend to experience decreasing commuting distances, which supports the co-location theory. 489 Moving to the second and third models in Table 2, model 1-2 includes the interaction terms

related to the location of employment centers. Here, the coefficient of the distance to nearest 491 employment center (b=0.001) can be interpreted as the base value for employment centers 492 located within counties other than Los Angeles or Orange County. Employment centers 493 located in Los Angeles County have greater influence on commuting distance. For example, 494 the coefficient of distance to nearest employment center for census tracts that have their 495 nearest center located in Los Angeles County is 0.004 (= 0.001 + 0.003). In contrast, the 496 497 interaction term for Orange County centers showed a negative sign which diminishes the relationship between distance to the nearest employment center and commuting distances. 498 This result suggests that the proximity to the nearest center does not contribute to shorter 499 commuting distances in Orange County. Model 1-3 estimates the effect of interaction terms 500 501 relevant to center attributes. The results show that employment centers that are persistent, have high job density and industry mix, have higher employment to population ratio, and 502 have larger size tend to have a greater impact on reducing commuting distance. These 503 504 findings imply that employment centers are heterogeneous in terms of their impact on 505 commuting distances of workers living close to them.

Model 1-4 estimates the results when simultaneously including the interaction terms 506 507 related to both center location and attributes. The coefficient of the distance to nearest employment center reduced to smaller than 0.001 and statistically insignificant. Since we 508 509 include multiple interaction terms, the coefficients should be interpreted with caution. For 510 instance, if the nearest employment center is located in Los Angeles County with high job 511 intensity, the overall effect of the distance to nearest employment center is not near 0. Also, we control for job density within 3 km (logged) based on the workers' residential area so the 512 results here can be viewed as evidence of the net contribution of employment centers that 513 have diverse effects on nearby residents. The results from Model 1-4 are similar to those 514 estimated from Models 1-2 and 1-3. Finally, model 1-5 is based on the pooled OLS model 515 which shows similar results to the fixed effects models along with a much higher R-square 516 517 value, which is common in models comparing across units, rather than within units.

518

519 **4.2.2 Low-income workers**

520 We next explore the model results for low-income workers (see Table 3). In general, the 521 estimates are mostly consistent with the results for all workers, while there are a few

differences to highlight. For low-income workers, the coefficient of distance to nearest
employment centers also showed a positive sign even after controlling for socio-economic
and residential neighborhood factors (see Model 2-1). This result clearly suggests that lowincome workers commute shorter distances if they live near an employment center.

One notable finding is that the distance to nearest employment center did not show a 526 significant association with commuting distance in Model 2-4. However, as described earlier, 527 we should not interpret this coefficient solely, but with consideration of the interaction terms. 528 529 For instance, if the nearest center is in Los Angeles County and if it has a high employment to 530 population ratio, the coefficient of distance to nearest center should be interpreted as 0.009 (= -0.000 + 0.006 + 0.003) and significant. Similar to the models for all workers, low-income 531 workers tend to commute shorter distances if they live close to an employment center located 532 in Los Angeles County, exhibiting higher job density, industry mix, and employment to 533 534 population ratio. The coefficient for persistency did not show statistical significance.

535 **4.2.3 High-income workers**

536 Lastly, Table 4 shows the model estimates for high-income workers. While most of the results are again consistent with the model estimates for all workers, high-income workers 537 living in census tracts with more low-income workers tend to commute longer distances. For 538 539 our variable of interest, high-income workers living in census tracts with an increasing distance to the nearest employment center experience increases in commuting distance. The 540 results were slightly different when it comes to the location dummy variables; the dummy 541 variables all showed a negative sign which suggests that the effect of the distance to nearest 542 employment center is the greatest for those located in areas other than Los Angeles and 543 Orange County. The interaction terms with the attributes of nearest employment center all 544 showed a positive and significant coefficient; employment centers that are persistent, have 545 high job density, diversity, and employment to population ratio with larger size exert more 546 547 influence on commuting distances for high-income workers.

548 Table 2. Fixed-effect models (1-1 to 1-4) and pooled OLS model (1-5) (DV: logged average commuting distance, All workers)

Variables	Model	1-1	Model	-2	Model 1-3		Model 1-4		Model 1-5	
	Coef.	(sig.)	Coef.	(sig.)	Coef.	(sig.)	Coef.	(sig.)	Coef.	(sig.)
Socio-economic factors										
Ratio of age 20-34	1.141	***	1.138	***	1.142	***	1.140	***	0.777	***
Ratio of low income (\$1,250/month or less)	-0.123	***	-0.128	***	-0.117	***	-0.121	***	-0.612	***
Neighborhood factors										
ln (job density within 3 km)	-0.036) ***	-0.036	***	-0.034	***	-0.034	***	-0.082	***
Industry mix level of jobs within 3 km	0.068) ***	0.067	***	0.066	***	0.066	***	-0.186	***
Distance to nearest employment center (DNC)	0.002	***	0.001	***	< 0.001		< 0.001		0.066	***
Location of nearest employment center										
Los Angeles City			-0.048	***			-0.028	***	-0.271	***
Los Angeles County			-0.055	***			-0.051	***	-0.368	***
Orange County			0.012	***			0.019	***	-0.165	***
DNC * Los Angeles City			0.003	***			0.002	***	-0.037	***
DNC * Los Angeles County			0.003	***			0.002	***	0.021	***
DNC * Orange County			-0.001	***			-0.001	***	-0.035	***
Attributes of nearest employment center										
Persistent from 2002 to 2019					0.005	***	0.001	***	-0.003	***
High job density					-0.030	***	-0.028	***	0.123	***
High industry diversity					-0.013	***	-0.016	***	0.032	***
High emp. to pop. ratio					-0.033	***	-0.035	***	-0.067	***
Large size					-0.033	***	-0.029	***	-0.461	***
DNC * Persistent from 2002 to 2019					< 0.001	***	< 0.001	***	-0.003	***
DNC * High job density					0.001	***	0.001	***	-0.022	***
DNC * High industry diversity					0.001	***	0.001	***	0.001	***
DNC * High emp. to pop. ratio					0.002	***	0.002	***	0.085	***
DNC * Large size					0.002	***	0.001	***	0.047	***
Constant	10.097	7 ***	10.138	***	10.109	***	10.138	***	1.053	***
Model statistics										
R-squared									0.702	
Within	0.062	2	0.063		0.064		0.065			
Between	0.466	5	0.469		0.472		0.470			
Overall	0.409)	0.413		0.412		0.412			

* p < 0.1; ** p < 0.05; *** p < 0.01

550 Table 3. Fixed-effect models (2-1 to 2-4) and pooled OLS model (2-5) (DV: logged average commuting distance, Low income workers)

Variables	Model	2-1	Model	2-2	Model 2-3		Model 2-4		Model 2-5	
	Coef.	(sig.)	Coef.	(sig.)	Coef.	(sig.)	Coef.	(sig.)	Coef.	(sig.)
Socio-economic factors										
Ratio of age 20-34	1.514	1 ***	1.509	***	1.516	***	1.510	***	1.258	***
Ratio of low income (\$1,250/month or less)	-0.864	1 ***	-0.870	***	-0.848	***	-0.851	***	-1856	***
Neighborhood factors										
ln (job density within 3 km)	-0.030	5 ***	-0.036	***	-0.034	***	-0.034	***	-0.100	***
Industry mix level of jobs within 3 km	0.107	7 ***	0.105	***	0.105	***	0.103	***	-0.223	***
Distance to nearest employment center (DNC)	0.003	3 ***	-0.001<		< 0.001		-0.001<		0.030	***
Location of nearest employment center										
Los Angeles City			-0.086	***			-0.057	***	-0.204	***
Los Angeles County			-0.134	***			-0.126	***	-0.352	***
Orange County			0.027	***			0.039	***	-0.125	***
DNC * Los Angeles City			0.005	***			0.003	***	-0.014	***
DNC * Los Angeles County			0.006	***			0.006	***	0.001	***
DNC * Orange County			-0.002	***			-0.003	***	-0.019	***
Attributes of nearest employment center										
Persistent from 2002 to 2019					0.019	***	0.014	***	0.030	***
High job density					-0.030	***	-0.027	***	0.054	***
High industry diversity					-0.039	***	-0.041	***	0.024	***
High emp. to pop. ratio					-0.055	***	-0.058	***	-0.167	***
Large size					-0.045	***	-0.044	***	-0.115	***
DNC * Persistent from 2002 to 2019					-0.001	***	-0.001<		-0.006	***
DNC * High job density					0.001	***	0.001	***	-0.015	***
DNC * High industry diversity					0.003	***	0.003	***	0.001	***
DNC * High emp. to pop. ratio					0.003	***	0.003	***	0.020	***
DNC * Large size					0.003	***	0.002	***	0.011	***
Constant	10.172	2 ***	10.259	***	10.200	***	10.268	***	1.131	***
Model statistics										
R-squared									0.701	
Within	0.080	5	0.088		0.089		0.091			
Between	0.453	3	0.458		0.459		0.480			
Overall	0.382	2	0.391		0.386		0.409			

* p < 0.1; ** p < 0.05; *** p < 0.01

551 Table 4. Fixed-effect models (3-1 to 3-4) and pooled OLS model (3-5) (DV: logged average commuting distance, High income workers)

Variables	Model	3-1	Model 3	3-2	Model 3	-3	Model 3	-4	Model 3	-5
	Coef.	(sig.)	Coef.	(sig.)	Coef.	(sig.)	Coef.	(sig.)	Coef.	(sig.)
Socio-economic factors										
Ratio of age 20-34	0.903	3 ***	0.901	***	0.905	***	0.904	***	1.646	***
Ratio of low income (\$1,250/month or less)	0.457	7 ***	0.452	***	0.464	***	0.461	***	0.458	***
Neighborhood factors										
ln (job density within 3 km)	-0.023	3 ***	-0.023	***	-0.021	***	-0.021	***	-0.093	***
Industry mix level of jobs within 3 km	0.034	4 ***	0.033	***	0.032	***	0.032	***	-0.051	***
Distance to nearest employment center (DNC)	0.002	2 ***	0.002	***	0.001	***	0.002	***	0.021	***
Location of nearest employment center										
Los Angeles City			0.020	***			0.012	***	-0.380	***
Los Angeles County			0.011	***			0.008	***	-0.397	***
Orange County			0.017	***			0.021	***	-0.184	***
DNC * Los Angeles City			-0.002	***			-0.001	***	0.002	***
DNC * Los Angeles County			-0.002	***			-0.001	***	0.017	***
DNC * Orange County			-0.003	***			-0.002	***	-0.007	***
Attributes of nearest employment center										
Persistent from 2002 to 2019					0.007	***	0.005	***	0.027	***
High job density					-0.032	***	-0.030	***	0.010	***
High industry diversity					-0.010	***	-0.011	***	-0.016	***
High emp. to pop. ratio					-0.039	***	-0.040	***	-0.168	***
Large size					-0.035	***	-0.033	***	-0.126	***
DNC * Persistent from 2002 to 2019					< 0.001	***	< 0.001	***	0.001	***
DNC * High job density					0.001	***	0.001	***	-0.013	***
DNC * High industry diversity					0.001	***	0.001	***	0.002	***
DNC * High emp. to pop. ratio					0.002	***	0.003	***	0.022	***
DNC * Large size					0.002	***	0.002	***	0.010	***
Constant	9.980) ***	9.998	***	9.995	***	9.993	***	1.041	***
Model statistics										
R-squared									0.645	
Within	0.052	2	0.053		0.054		0.055			
Between	0.377	7	0.357		0.376		0.338			
Overall	0.323	3	0.305		0.319		0.286			

* p < 0.1; ** p < 0.05; *** p < 0.01

553 **4.2.4 Summary**

554 Our findings suggest that the distance to the nearest employment center is significantly associated with longer commuting distances. The results are meaningful since we attempt to 555 556 identify the relationship after controlling for socio-economic and residential neighborhood factors. Furthermore, our fixed effects models focus exclusively on change within tracts, and 557 do not compare across tracts which allows us to infer causal relationships. The within R-558 squared values of all models ranged from 0.05 to 0.09 in our models while it was greater for 559 the models based on low-income workers. Our independent variables explain around 5-9% of 560 the variation in the changes in commuting distances over time. While the within R-squared 561 for our fixed effects models were somewhat low, this is a common feature of fixed effects 562 models. The low explanatory power of our fixed effects models suggest that only a limited 563 portion of the variance in commuting distance is explained by the changes in our explanatory 564 variables over time; however, the model results still indicate the significant relationships. 565 566 Furthermore, the R-squared values for the OLS versions of our models comparing across 567 units were nearly .70, which shows a quite high explained variance, highlighting the limited amount of variability that there is to explain within units. 568

569 Figures 4 and 5 summarize the results of the fixed-effects models and present the estimated effects of the distance to nearest employment center on commuting distance; Figure 4 570 571 assumes the location of the employment center as Los Angeles County and Figure 5 assumes 572 it as other counties than Los Angeles and Orange County. The figures show that the effect of the nearest employment center on commuting distance differs by the characteristics of 573 employment centers as well as by different income groups. As we will discuss in more depth 574 575 shortly, these figures show that: 1) employment centers have stronger effects on commuting distance in Los Angeles County; 2) Los Angeles employment centers impact low income 576 workers more strongly than high income workers; and 3) in more distant counties, 577 employment centers tend to more strongly impact commute distances of high income workers 578 compared to low income workers. 579



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Figure 4. Estimated effects of distance to nearest employment center on commuting distance. Location of employment center assumed as Los Angeles County.





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Figure 5. Estimated effects of distance to nearest employment center on commuting distance. Location of employment center assumed as counties other than Los Angeles and Orange County.

We note that the relationship tested in this paper is subject to the self-selection issue. In 588 particular, residential self-selection implies that individuals may choose to live in certain 589 areas due to constraints, preferences, and modal availability (Cao, 2009). Low income 590 workers who could not afford high housing prices near employment centers may instead 591 reside in distant areas (Cao, 2009). It is also possible that some low income workers live 592 close to the employment centers because they either cannot afford cars for long-distance 593 commutes or have more job availabilities in employment centers (Bohte et al., 2009). For 594 high income workers, they may choose where they live according to their lifestyle 595 596 preferences, which may lead to a weak relationship between commute distance and the distance to the nearest employment center. This component may partly explain the relatively 597 lower explanatory power of the models for high-income workers. 598

599 5. Discussion and conclusions

600 This paper investigated how distance to the nearest employment center is related to commuting distance in the Los Angeles region. By using longitudinal data on commuting and 601 602 employment locations, we were able to determine that commuters experiencing an 603 increasingly near employment center tend to have decreasing commuting distances. The relationship was statistically significant even after controlling for socio-economic and 604 neighborhood factors. In particular, we applied a fixed effects model, which captures 605 variation within tracts as a more rigorous test of this relationship. We have further examined 606 607 the heterogeneity across employment centers and workers by applying interaction terms and testing multiple fixed effects models. Between 2002 and 2019, we find that the polycentric 608 urban form in the Los Angeles region has been consistent, where more than 35% of the jobs 609 within the region are located in employment centers. Among large US metropolitan areas, the 610 Los Angeles region has the largest share of jobs in employment centers (Angel & Blei, 2016), 611 612 which makes a unique case to explore the effect of employment centers on commuting.

613 One notable finding is that not all employment centers are equal. Based on the interaction terms related to location and attributes of employment centers, the results suggest that some 614 615 centers exert more influence on commuting distance. Compared to other places in the region, for instance, the proximity to employment centers was found to have a greater effect on 616 617 reducing commute lengths in Los Angeles County. Related to the attributes of centers, the 618 centers that were persistent throughout the period from 2002 to 2019 showed greater influence on commuting distance than non-persistent ones. In addition, centers with a higher 619 density, higher employment to population ratio, and larger size showed greater effects on 620 reducing commuting distance of nearby residents. The results imply that employment centers 621 are heterogeneous in terms of their location and attributes, with different impacts on 622 commuting distance of workers living near them. 623

Employment centers that have high job density and a high employment to population ratio imply that there are a greater number of jobs located within the center. These employment centers may have characteristics that lead to further employment growth, which in turn can exert more influence on commuting behavior. For instance, Giuliano et al. (2011) show that accessibility is a critical component for employment centers to grow; employment centers with greater labor force accessibility may attract more workers from locations in proximity.

In addition, employment centers with greater industry mix may provide potential to reduce commuting distances especially for households with multiple workers. Moreover, persistent employment centers are likely to have attracted more jobs and provide better accessibility to workers with accumulated infrastructure levels which may contribute to shorter commuting distances.

We also observed differences between low- and high-income workers. The commuting 635 distances of low-income workers were not necessarily associated with the distance to the 636 nearest employment center; the effect was the greatest when the employment center was 637 638 located in Los Angeles County. For high-income workers, their commuting distance was positively associated with the distance to the nearest employment center while the effect size 639 640 increased when the employment center is located in counties other than Los Angeles and 641 Orange county. This result suggests that many high-income workers living within Los 642 Angeles and Orange county might not commute to an employment center in proximity. Overall, when considering the location of employment centers, those located nearer the 643 644 region center more strongly impact commuting distance of low-income workers, while the ones located in the peripheral areas have greater impact on high-income workers. 645

Employment centers are an outcome of both economies and diseconomies of 646 agglomeration. While there have been efforts to encourage employment center growth at the 647 local level, existing studies show mixed results of their success. For instance, Agarwal (2015) 648 649 shows that policy measures such as expenditure on development, growth control, and 650 business fees do not show a significant effect on employment center growth. Instead, the 651 author suggests that facilitating access to the labor force may provide potential to indirectly 652 encourage employment center growth. While our results show that some characteristics of employment centers are associated with greater effects on commuting distance, it would not 653 be appropriate to assume that an employment center can be easily reshaped in a certain way 654 655 by a single policy measure or initiative. That said, by highlighting the heterogeneous nature of the benefits of employment centers, this study encourages policymakers to refine their 656 657 understanding of the workings of their employment centers and carefully monitor how the 658 centers evolve over time. It is also important to pay attention to who gains and who loses 659 since not all workers will equally benefit from employment centers as shown in this study 660 through a comparison of high-income and low-income workers. In some circumstances, it 661 would be desirable to provide more affordable housing units near employment centers for

662 more sustainable and inclusive place making.

Empirical evidence highlights the benefits of short commutes. Short-distance commuters 663 exhibit better job performance and contribute to greater economic growth of employers (Ma 664 & Ye, 2019). Commuters with longer commute lengths experience reduced satisfaction and 665 subjective well-being (Manaugh & El-Geneidy, 2013; Nie & Sousa-Poza, 2018). Moreover, 666 shorter commuting length is a desirable goal for cities as it is conducive to reducing the 667 negative externalities of transport. Unfortunately, it is not an easy task to accomplish. Low-668 income workers may have limited ability to choose their home and workplace, while high-669 670 income workers have other factors to consider when selecting their residential area. Moreover, we have seen growth in commuting distance in the past, which have led to ideas 671 672 such as encouraging the overall connectivity within regions rather than encouraging 673 transportation strategies that focus on improving access to employment centers (e.g., Angel & 674 Blei, 2016). However, our results imply that employment centers have the potential to reduce commute lengths depending on their characteristics and the income level of workers. For 675 676 cities, we further suggest that improving access to employment centers that have a greater possibility to affect commuting outcomes could be beneficial. 677

We note limitations and methodological issues of this study. First, we acknowledge that our 678 results do not address the self-selection issue. The residential location of workers is not 679 randomly assigned but is a self-selected result which may affect the relationship between 680 681 commuting and distance to the nearest employment center. Second, earning categories defined in the LEHD data are not adjusted for inflation over time. This could result in 682 workers being classified in a different income bin; in other words, the percentage of low-683 684 wage category decreases, and the percentage of high-wage category increases year by year as 685 a result of inflation. While we acknowledge this data limitation, we use the LEHD data because of its primary advantages. The LEHD data is updated annually making it possible to 686 687 examine changes over time, particularly for understanding residential and workplace location and the commute outcomes. With this data issue, our results on the change in commute 688 689 distances for low-wage workers would represent a poorer population over time due to 690 inflation. Similarly, the results for high-wage workers would include more "mid-income" 691 workers over time because of inflation. Nonetheless, we suggest that our results contribute to 692 the literature since we find differences between the low and high-wage workers. Lastly, we 693 agree that there are several ways to identify employment centers within a region and the

results may differ when using other methodological approaches. In our case, we applied the approach suggested by Giuliano & Small (1991) not only because it is the most widely used in the literature, but also because it relies on a fixed threshold which allows us to explore the changes over time in a more robust way.

698 Nevertheless, this study contributes to the literature on urban spatial structure, polycentricity, and commuting, particularly by deciphering the location and attributes of 699 700 employment centers that may exert a greater impact on commuting distances. Furthermore, 701 the results have implications for understanding the agglomeration effect and co-location theory in a polycentric metropolitan area. Future studies may explore other attributes of 702 employment centers that might impact commuting outcomes. In addition, other dimensions 703 such as socio-economic characteristics of commuters and other commuting outcomes such as 704 mode choice and travel time could be further investigated. Since the Los Angeles region has 705 706 an exceptional urban spatial structure with a high degree of polycentricity, the results may not 707 be generalizable to other regions that have a smaller number of employment centers with 708 more concentration. Lastly, we also note that the COVID-19 pandemic may result in different 709 relationships between commuting and employment centers as the number of workers working 710 from home has dramatically increased. While the time scope of our results is limited to 2002 -2019, future work may further investigate how the pandemic has reshaped the employment 711 712 centers and their association with commuting.

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722 Appendix

723 Table A1. Correlation between distance to nearest employment center and job related

724 measures at the neighborhood level

Variables	(a)	(b)	(c)
(a) Dist. to nearest emp. center	-	-	-
(b) Job density within 3km	-0.412 *	-	-
(c) Industry mix of jobs within 3 km	-0.197 *	0.206 *	-

Table A2. Correlation among attributes of employment centers

(a)	(b)	(c)	(d)	(e)
-	-	-	-	-
0.271 *	-	-	-	-
0.017	-0.175 *	-	-	-
0.315 *	0.379 *	-0.124 *	-	-
-0.040	-0.048	0.035	-0.172	-
	(a) - 0.271 * 0.017 0.315 * -0.040	(a) (b) - - 0.271 * - 0.017 -0.175 * 0.315 * 0.379 * -0.040 -0.048	(a) (b) (c) - - - 0.271 * - - 0.017 -0.175 * - 0.315 * 0.379 * -0.124 * -0.040 -0.048 0.035	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$





Figure A1. Distribution of low income workers' residential density. Colors are generated by
quintiles, darker colors indicating higher density.



734

Figure A2. Distribution of high income workers' residential density. Colors are generated by
 quintiles, darker colors indicating higher density.





Figure A3. Distribution of low income workers' workplace density. Colors are generated by
quintiles, darker colors indicating higher density.



741

Figure A4. Distribution of high income workers' workplace density. Colors are generated by
quintiles, darker colors indicating higher density.

745 **References**

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ⁱ The seven types of industry are identified based on the NAICS codes identified in the LODES data:

(1) 11 (Agriculture, Forestry, Fishing and Hunting), 21 (Mining, Quarrying, and Oil and Gas Extraction), 22 (Utilities), 23 (Construction),

(2) 31-33 (Manufacturing),

(3) 42 (Wholesale Trade), 44-45 (Retail Trade), 48-49 (Transportation and Warehousing),

(4) 51 (Information), 52 (Finance and Insurance), 53 (Real Estate and Rental and Leasing), 54
(Professional, Scientific, and Technical Services), 55 (Management of Companies and Enterprises), 56 (Administrative and Support and Waste Management and Remediation Services),

(5) 61 (Educational Services), 62 (Health Care and Social Assistance),

(6) 71 (Arts, Entertainment, and Recreation), 72 (Accommodation and Food Services),

(7) 81 (Other Services [except Public Administration]), 92 (Public Administration).