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Neighborhoods and Racial/Ethnic Differences in Ideal Cardiovascular Health (the Multi-Ethnic Study of Atherosclerosis)

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Abstract

Using data from the Multi-Ethnic Study of Atherosclerosis baseline sample from 2000–2002 (N=5263; mean age=62) we examined cross-sectional racial/ethnic differences in ideal CVH, defined by the American Heart Association 2020 Impact Goals as a summary measure of ideal levels of blood pressure, fasting glucose, cholesterol, body mass index, diet, physical activity, and smoking. Using three different analytical approaches, we examined differences before and after adjustment for neighborhood socioeconomic, physical, and social environments. Significant racial/ethnic differences were present for all indicators of ideal CVH (excluding physical activity). Additional adjustments for neighborhood factors produced modest reductions in racial/ethnic differences. Future research is necessary to better understand the impact of neighborhood context on health disparities using longitudinal study designs.

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Keywords

Racial/ethnic health disparities; neighborhood environments; cardiovascular health

INTRODUCTION

Persistent and pervasive racial/ethnic differences in health are a major public health concern (Institute of Medicine (US) Committee on the Review and Assessment of the NIH's Strategic Research Plan and Budget to Reduce and Ultimately Eliminate Health Disparities, 2006; Kelley et al., 2005; Smedley et al., 2003). Racial/ethnic differences in the prevalence of cardiovascular disease (CVD) risk factors and incidence of and mortality due to cardiovascular diseases have been well documented, with blacks disproportionately burdened by adverse outcomes (Go et al., 2014; Mensah et al., 2005). Underlying causes of these differences remain poorly understood, but are most likely generated by multifactorial and multilevel causes that occur over the life-course. Much of the extant literature has focused on individual-level risk factors as explanations for racial/ethnic disparities (genetic, biological, socioeconomic, and psychosocial) (Kramer et al., 2004; Mensah and Dunbar, 2006; Safford et al., 2012). However, more attention has recently been placed on contextual factors such as neighborhood environments.

The increased interest in the intersection of race/ethnicity, neighborhood, and health has emerged in part from the increasing interest in the effects of neighborhoods on health generally. Studies have shown that living in socioeconomically disadvantaged neighborhoods is associated with increased morbidity and mortality, independent of individual-level factors (Borrell et al., 2004; Pickett and Pearl, 2001; Truong and Ma, 2006), and some of the most consistent evidence is from the area of cardiovascular disease (CVD) outcomes. (Borrell et al., 2004; Chaix, 2009; Diez Roux, 2003; Morenoff et al., 2007; Mujahid et al., 2011; Pickett and Pearl, 2001). In the area of CVD, studies have also shown that living in neighborhoods with poor physical environments (e.g. more unhealthy foods on average and fewer opportunities for physical activity) and social environments (e.g. less safety and social cohesion) are associated with increased risk of CVD risk factors such as obesity, diabetes, and hypertension, and fatal and non-fatal CVD (Christine et al., 2015; Diez Roux et al., 2016; Kaiser et al., 2016; Kershaw et al., 2015). Studies have also documented that blacks and other racial/ethnic minorities disproportionately reside in neighborhoods that are socioeconomically disadvantaged, and have poor access to health promoting resources. Combined, the aforementioned supports the hypothesis that neighborhood environments may be one underlying cause of racial/ethnic disparities in health. However, there is little empiric investigation explicitly testing this notion (Do et al., 2007; Dubowitz et al., 2008b; Morenoff et al., 2007; Mujahid et al., 2011; Robert and Reither, 2004).

Studies that have empirically investigated the contribution of neighborhood environments to racial/ethnic differences in health have most consistently examined estimates of differences before and after adjustment for neighborhood environments in a multivariable regression model without an explicit discussion of the major challenges and limitations in this approach (Do et al., 2007; Dubowitz et al., 2008b; Mujahid et al., 2011; Robert and Reither, 2004).

These challenges include how best to: deal with the non-independence of observations within neighborhoods (Hubbard et al., 2010; Subramanian and O'Malley, 2010), account for the neighborhood context when all relevant factors may not be measured in a given study, and account for insufficient within-neighborhood sample size and variation in exposures and outcomes. A recent review by Schempf et al. (Schempf and Kaufman, 2012) provides an in-depth discussion of these issues, a comparison of analytic approaches that may address these issues, and compares the strengths and limitations of each analytic approach. Studies providing similar comparisons across analytic approaches remain scarce in the literature. (Morenoff et al., 2007; Schempf and Kaufman, 2012).

Given the increased attention to and importance of primordial prevention, as underscored by the recent American Heart Association 2020 Strategic Impact Goals to reduce cardiovascular mortality by 20% and increase ideal cardiovascular health by 20% (Lloyd-Jones et al., 2010), exploring the contribution of neighborhood environments to racial/ethnic differences in CVD risk in a multi-ethnic study is timely (Havranek et al., 2015). Thus, the overall goal of this study was to determine if cross-sectional racial/ethnic differences in ideal cardiovascular health (CVH) indicators were reduced after adjustment for neighborhood environments, and if results were robust to statistical methods of neighborhood adjustment.

METHODS

Study population

MESA is a prospective study of 6,814 adults of diverse racial/ethnic backgrounds (self-identified race/ethnicity as non-Hispanic white, non-Hispanic black, Hispanic, and non-Hispanic Chinese) from six study sites (Los Angeles County, California; Chicago, Illinois; Baltimore City and County, Maryland; St. Paul, Minnesota; Forsyth County, North Carolina; New York City, New York) (Bild et al., 2002). Study participants were 45–84 years of age and free from clinical cardiovascular disease at baseline (August 2000 to July 2002). A detailed description of the study recruitment procedures and methods has been previously described (Bild et al., 2002).

Our cross-sectional analyses are restricted to the MESA participants who provided consent to participate in the ancillary MESA Neighborhood Study (N=6191). Institutional review board approval was received at each of the MESA study sites.

Study variables

Study outcomes—We examined seven indicators of ideal CVH as defined by the American Heart Association 2020 Impact Goals (Lloyd-Jones et al., 2010) using either study questionnaire or clinical examination at baseline (2000–2002). Each indicator was categorized as poor, intermediate, and ideal based on established criteria and are defined in Table 1 (Lloyd-Jones et al., 2010). Cholesterol and fasting blood glucose were measured from a 75mL fasting blood sample obtained at the baseline clinical examination. Blood pressure (BP) was measured as the average of the second and third readings after five minutes resting in a seated position. BMI was assessed using measurements of height and weight obtained during clinical examination ($BMI = \text{height, m} / \text{weight, kg}^2$). Smoking was

based on self-reported data from study questionnaires and combined questions on whether a participant smoked in the past 30 days and the lifetime number of cigarettes smoked to create categories of never, current, and former smoking. We estimated minutes of moderate and vigorous exercise from walking, conditioning, and leisure-time activities based on participant's assessment of the time and frequency spent on activities during a typical week in the past month (Ainsworth et al., 1999; Bertoni et al., 2009). We estimated five components of healthy diet using a 120-item food frequency questionnaire (Block et al., 1990), including high intake of healthy foods (fruits and vegetables; fish; whole grains) and low intake of unhealthy foods/nutrients (sugar sweetened beverages; sodium) similar to prior studies (Lloyd-Jones et al., 2010; Rasmussen-Torvik et al., 2013).

We also created three summary measures of ideal CVH as the sum of each individual component score for health behaviors, health factors, and overall cardiovascular health. For the overall CVH measure, which ranged in value between 0 and 14, we created three categories based on the distribution of the data: poor (score of 0–8); intermediate (score of 9–10); and ideal (score of 11–14). In analyses we consider the 7 indicators of ideal CVH and 3 summary measures as dichotomous variables (1=ideal; 0=intermediate/poor).

Key covariates of interest

Neighborhood-level covariates—Neighborhoods were defined as census tracts based on previous work indicating good agreement across individuals residing within the same tract in relation to our neighborhood characteristics of interest (Mujahid et al., 2007). Three neighborhood-level socioeconomic indicators were derived from the 2000 U.S. Census based on methods previously described (Mujahid et al., 2008). The indicators combine 16 variables representing the dimensions of family structure, area crowding, residential stability, education, employment, occupation, and income/wealth) based on a principle components analysis with varimax orthogonal rotation to reduce the potential for multicollinearity in analytic models. The three factors account for 70 % of the total variance.

We calculated two physical environment indicators to represent physical activity and healthy food environment. Using data from the National Establishment Time Series Data (NETS) (Walls & Associates, 2012), we calculated kernel densities of all indoor and outdoor recreational facilities (excluding parks) around MESA participants' home addresses at baseline year of enrollment. Facilities were defined using established Standardized Industrial Codes (SIC) based on prior work (Gordon-Larsen et al., 2006; Powell et al., 2007). We supplemented kernel densities with resident reports of the physical activity environment. MESA participants and an informant sample of non-MESA area residents were asked four items on whether or not their neighborhood (defined as the area within about a 20 minute walk or 1 mile from the home) was conducive to physical activity (Echeverria et al., 2004; Mujahid et al., 2007). We averaged across these items to create an overall score (Cronbach's alpha=0.65). The kernel density and survey measure of physical activity environment were each standardized and aggregated to create an overall physical activity indicator for which higher scores represent a better physical activity environment. We used a similar process to create an overall indicator of neighborhood healthy foods (Auchincloss et al., 2012). Neighborhood social environment was characterized using survey items assessing

neighborhood aesthetic quality (n=3 items), safety (n=2 items), and social cohesion (n=4 items). These items were shown to have good measurement properties based on prior work (Cronbach's alpha=0.86) (Mujahid et al., 2011; Mujahid et al., 2008). We created an overall measure by aggregating responses across these items and empirical Bayes estimation techniques as previously described (Mujahid et al., 2008).

Individual-level covariates—All additional study covariates were collected during the baseline questionnaire (2000–2002). Self-reported race/ethnicity was categorized as white, non-Hispanic, black, non-Hispanic, Hispanic. Chinese were excluded due to limited sample size and limited geographic coverage across study sites. Additional covariates obtained from the study questionnaire included study site, age (continuous), sex (male, female), education (less than a high school diploma; high school diploma; some college; college graduate and beyond), and family income (less than \$24,999; \$25,000–\$49,999; \$50,000–\$74,999; \$75,000 or more).

Analysis plan and methods

All analyses were performed using Stata version 12. In descriptive analyses, we examined bivariate associations between each of the three summary measures of ideal CVH and study covariates. To examine the contribution of neighborhood context to racial/ethnic differences in CVH, we contrasted race coefficients in a neighborhood unadjusted logistic model (Model 1 below) to race coefficients from a series of models (models 2–4 below) that adjusted for neighborhood context in various ways (Morenoff et al., 2007; Schempff and Kaufman, 2012).

Model 1: Neighborhood unadjusted

$$\text{logit}(Y_i) = \beta_0 + \beta_1 \text{BLACK} + \beta_2 \text{HISPANIC} + \sum_{k=3}^t \beta_k X_{ki} + \varepsilon_i$$

Where $\text{logit}(Y_i)$ is the log odds of person i having ideal levels of the cardiovascular health, β_1 , and β_2 are coefficients on binary indicator variables for black and Hispanic race/ethnicity respectively, and β_k are the coefficients for the X_{ki} s, a vector of individual-level covariates. In this model, we estimated racial/ethnic differences in ideal cardiovascular health, independent of only individual-level confounders (study site, gender, race/ethnicity, age, education, income). There is no adjustment for neighborhood-level factors and thus, we refer to this model as a neighborhood unadjusted model. Because neighborhood-level factors were unaccounted for in these models, it is considered a naïve model representing initial estimates of black-white differences in ideal cardiovascular health when neighborhood context is not taken into account.

Model 2: Fixed Effects

$$\text{logit}(Y_{ij}) = \beta_0 + \beta_1 \text{BLACK}_{ij} + \beta_2 \text{HISPANIC}_{ij} + \sum_{k=3}^t \beta_k X_{kij} + \sum_{l=1}^{n-1} \gamma_l \text{Neighborhood}_{lij} + \varepsilon_{ij}$$

Where, $\text{logit}(Y_{ij})$ is the log odds of the i^{th} person in the j^{th} neighborhood having ideal levels of cardiovascular health, $\beta_0, \beta_1, \beta_2, \beta_k$, and X_{kij} have the same specifications as Model 1, and γ_l is a vector of $n-1$ coefficients, each corresponding to an indicator variable for each of the $n-1$ neighborhoods ($\text{Neighborhood}_{ij} = 1$ if person i lives in the j^{th} neighborhood, and is 0 otherwise). In this model we estimated racial/ethnic differences in the odds of ideal cardiovascular health, adjusted for individual-level covariates. This model is often considered the gold standard for accounting for group-level factors as it provides a summary of all measured and unmeasured group-level covariates. By controlling for all known and unknown characteristics of neighborhood environments, estimates of racial/ethnic differences represent the “weighted” average of within-neighborhood race/ethnic differences. However, there are also several limitations that warrant mention. First, this approach requires sufficient within-neighborhood sample size to ensure variation in study outcomes. For example, any neighborhood that does not have at least one participant with ideal levels of CVH and at least one participant without ideal levels of CVH is dropped from this analysis. This limitation is of particular concern given the levels of racial residential segregation that exists within the U.S. (Williams and Collins, 2001). A second limitation is when using this approach, one cannot estimate the specific features of neighborhoods that may account for reductions in racial/ethnic differences: the model represents all of the neighborhood features, both measured and unmeasured, by an indicator variable. As a result of this, there is also no way to examine potential interactions between neighborhood-level factors and other individual-level factors. A final limitation is the potential loss of statistical efficiency due to loss in degrees of freedom when there are a large number of clusters or limited within-cluster variation (Schempf and Kaufman, 2012).

Model 3: Random Effects (intercept)

$$\text{logit}(Y_{ij}) = \beta_{0j} + \beta_1 \text{BLACK}_{ij} + \beta_2 \text{HISPANIC}_{ij} + \sum_{k=3}^t \beta_k X_{kij} + \sum_{l=1}^p \gamma_l W_{lj} + \varepsilon_{ij}$$

$$\beta_{0j} = \gamma_{00} + \mu_{0j} \quad \mu_{0j} \sim N(0, \tau_0^2)$$

Where $\text{logit}(Y_{ij})$ is the log odds of ideal cardiovascular health for the i^{th} person in the j^{th} neighborhood, β_{0j} is the intercept for each neighborhood j , β_1 and β_2 are coefficients on indicator variables for black and Hispanic race/ethnicity for the i^{th} person in the j^{th} neighborhood, β_k are coefficients on X_{kij} , a vector of individual-level covariates, and γ_l are coefficients on W_{lj} , a vector of neighborhood-level covariates (neighborhood physical, social, and socioeconomic indicators). The intercept β_{0j} is modeled as a random variable for each neighborhood j , where γ_{00} is the average log odds of ideal cardiovascular health averaged across all neighborhoods and μ_{0j} is the deviation from the average log odds of ideal cardiovascular health for each neighborhood j and is assumed to have a normal distribution. The random effects or multi-level model has a number of benefits. Primarily, the ability to adjust for neighborhood and individual level variables allows for an assessment of both individual and neighborhood level variation in the data. Also, the ability to evaluate whether racial disparities in ideal cardiovascular health vary across neighborhoods provides an opportunity to investigate cross-level interactions. However, a limitation of this approach is that it assumes that the exposure, in this case race/ethnicity, is not conditionally correlated

with the neighborhood random effects (Schempf and Kaufman, 2012). This is an assumption that is difficult to empirically evaluate, but theoretically easy to violate given the degree of racial residential segregation in the U.S.

Model 4: Hybrid Fixed Effects (cluster mean adjustment)

$$\begin{aligned} \text{logit}(Y_{ij}) = & \beta_{0j} + \beta_{w1} \text{BLACK}_{ij} + \beta_{w2} \text{HISPANIC}_{ij} + \sum_{k=3}^t \beta_k X_{kij} + \sum_{l=1}^p \gamma_l W_{lj} + \varepsilon_{ij} \\ \beta_{0j} = & \gamma_{00} + (\beta_{b1} - \beta_{w1}) \overline{\text{BLACK}}_j + (\beta_{b2} - \beta_{w2}) \overline{\text{HISPANIC}}_j + \mu_{0j} \quad \mu_{0j} \sim N(0, \tau_0^2) \end{aligned}$$

Where $\text{logit}(Y_{ij})$ is the log odds of ideal cardiovascular health for the i^{th} person in the j^{th} neighborhood, β_{0j} is the intercept for each neighborhood j , β_1 and β_2 are coefficients on indicator variables for black and Hispanic race/ethnicity for the i^{th} person in the j^{th} neighborhood, β_k are coefficients on X_{kij} , a vector of individual-level covariates, and γ_l are coefficients on W_{lj} , a vector of neighborhood-level covariates (neighborhood physical, social, and socioeconomic indicators). The intercept β_{0j} is modeled as a random variable for each neighborhood j and is a function of the proportion of each racial/ethnic group in that specific neighborhood. Here, γ_{00} is the average log odds of ideal cardiovascular health averaged across all neighborhoods, $\overline{\text{BLACK}}_j$ and $\overline{\text{HISPANIC}}_j$ are the proportion of people in the j^{th} neighborhood who are black or Hispanic respectively, $\beta_{b1} - \beta_{w1}$ is the difference between the between-neighborhood association and within-neighborhood association between the proportion black in a given neighborhood j and the neighborhood-specific intercept, and μ_{0j} is an error term for each neighborhood j and is assumed to have a normal distribution with variance τ_0^2 .

Using a hybrid fixed effects model allows for a slightly different interpretation of coefficients in the logistic regression model because it accounts for between-neighborhood variation in race/ethnicity that contributes to neighborhood-level confounding. Here, the estimated coefficients on the race/ethnicity variables, β_{w1} and β_{w2} can be interpreted as the difference in ideal CVH between black and white (or Hispanic and white) people, holding the distribution of race/ethnicity (proportion black and proportion Hispanic) in the neighborhood constant (and holding the individual (X_k)- and neighborhood (W)- level confounders constant). Hybrid fixed effects models combine the strengths of the fixed effects and random effects models. The benefit from FE of being able to calculate marginal effects transfers to this model, but here we are able to make that estimate without ignoring all between-neighborhood variation, which is an advantage of using a RE model (Desai and Begg, 2008; Rabe-Hesketh and Skrondal, 2008; Schempf and Kaufman, 2012).

For all aforementioned models, we examined correlations among neighborhood features and assessed multicollinearity. There were modest to strong correlations between neighborhood indicators. The strongest correlations were for social environment and socioeconomic environment ($r=-.78$), physical activity and food environment ($r=0.73$), and physical activity and socioeconomic environment ($r=-0.65$). However, after assessing the variance inflation factors (all values less than 5) and tolerance (all values greater than .10) no violations of multicollinearity were found. For ease of interpretation and a useful way of comparing

changes in model estimates of race/ethnic differences in ideal CVH across our models, we calculated the largest percent reduction as the change in estimates between Model 1, the model unadjusted for neighborhoods, compared to whichever Model 2–4 that resulted in the largest decrease in estimate based on the formula below where β_u is the estimate from Model 1, and β^* is the estimate from one of Models 2–4.

$$LPC = \frac{|\beta_u - \beta^*|}{|1 - \beta_u|} \times 100$$

RESULTS

Among the 6191 MESA Neighborhood participants that provided consent to participate in the MESA Neighborhood Study, we excluded 728 Chinese participants for whom sample sizes were too small to examine neighborhood context. Additional exclusions included address errors (N=133) and missing information on study covariates (N=67) for a final analytic sample of N=5,263 representing 1033 neighborhoods with an average of 5 study participants per neighborhood (range 1 to 250). Among this analytic sample, the mean age was 62.0 (SD= 10.1), 52.7% were female, 34.3% had less than a high school education, and 26.3% had a family income of less than \$25,000 per year. Ideal cardiovascular health, experienced by 4.1% of our analytic sample, varied by age, education, income, study site, and all neighborhood measures (excluding one of three neighborhood SES factor scores). The prevalence of ideal CVH decreased with age and was highest among those with the highest levels of education and family income (at least a college bachelor's degree/earning $\geq$ \$75,000). As seen in previous studies (Unger et al., 2014), ideal CVH increases with increasing neighborhood physical, social, and socioeconomic environments (Table 2).

Figure 1 shows racial/ethnic differences in ideal cardiovascular health indicators. Significant differences were present for all components and summary measures (all p's $\leq$ 0.01). Blacks had the lowest prevalence of ideal blood pressure and smoking, but highest prevalence of ideal cholesterol. Hispanics had the lowest prevalence of ideal BMI, glucose, physical activity, and diet. Blacks had the lowest prevalence of all summary measures (ideal health factors, health behaviors, and overall cardiovascular health).

Table 3 shows the odds ratios in ideal cardiovascular health factors comparing blacks and whites and Hispanics and whites. In initial models, the one-level logistic regression models, blacks as compared to whites were less likely to have ideal blood pressure (OR=0.510, 95% C.I.: 0.438–0.594) and fasting blood glucose (OR=0.581, 95% C.I.: 0.502, 0.673), independent of age, study site, gender, education, and family income. Adjustment for neighborhood-level factors resulted in a 21.45% and 16.2% reduction in black-white differences respectively. Similar patterns were documented in Hispanic-white differences in blood pressure (OR=0.600, 95% C.I.: 0.498, 0.724) and fasting blood glucose (OR=0.542, 95% C.I.: 0.450, 0.652) with 10.5% and 26.0% reductions. For cholesterol, blacks (OR=1.236; 95% C.I. 1.066–1.419) had a higher odds of ideal cholesterol levels compared to whites, independent of age, study site, gender, education, and family income. There were no significant Hispanic-white differences in ideal cholesterol.

Racial/ethnic differences in ideal cardiovascular health behaviors are presented in Table 4. Blacks (OR=0.401, 95% C.I. 0.339–0.475) and Hispanics (OR=0.417, 95% C.I. 0.324–0.520) were less likely to have ideal levels of BMI compared to whites and ideal levels of diet (OR=0.446 for blacks; OR=0.349 for Hispanics), independent of gender, age, education, income, and study site. Adjustment for neighborhood context resulted in minimal reductions in estimates for BMI (largest percent reduction: 2.3% reduction for blacks; 13.2% reduction in whites). Blacks were also less likely to have ideal smoking levels as compared to whites, independent of individual-level covariates (OR=0.713, 95% C.I.: 0.584, 0.872) and differences were reduced after adjusting for neighborhood context (largest percent reduction 31.4%). Alternatively, Hispanics were more likely to have ideal levels of smoking compared to whites, independent of individual-level covariates. There were no significant racial/ethnic differences in physical activity, independent of individual-level variables.

In examining racial/ethnic differences in three cardiovascular summary measures, significant black-white differences and Hispanic-white differences were present for all measures (Table 5). Blacks and Hispanics were less likely to have ideal health factors (OR=0.613, 95% C.I.: 0.505, 0.744 for blacks; OR=0.680, 95% C.I. 0.536, 0.862) for Hispanics), health behaviors (OR=0.362, 95% C.I. 0.292, 0.448 for blacks; OR=0.441, 95% C.I.: 0.334, 0.581 for Hispanics) and overall cardiovascular health (OR=0.324, 95% C.I.: 0.216, 0.487 for blacks; OR=0.403, 95% C.I. 0.245, 0.664) as compared to whites, independent of individual-level covariates. Adjustment for neighborhood context slightly reduced these differences, although they remained statistically significant (largest percent reduction range: 5.6%–20.2% for blacks; 8.2%–28.1% for Hispanics).

In comparing the methods of adjusting for neighborhood context across models: fixed effect, random effects, and hybrid fixed effects models, the largest percent reduction was most consistently found with the hybrid fixed effect approach. The fixed effects model appears to be the least stable as some estimates of racial/ethnic differences were increased instead of decreased after adjustment for neighborhood-level factors. However, given the relatively modest differences in estimates across the three modeling approaches, our results were robust.

DISCUSSION

In a multi-ethnic cohort of middle-aged adults, we found significant black-white and Hispanic-white differences for all ideal health factors (blood pressure, cholesterol, and glucose), three of four ideal health behaviors (diet, body mass index, and smoking) and all ideal CVH summary measures (health factors, health behaviors, and overall ideal cardiovascular health), independent of sociodemographic characteristics. Adjustment for neighborhood context slightly reduced racial/ethnic differences, but differences remained statistically significant. Percent reductions in racial/ethnic differences ranged from 2.3% to 31.4%.

To our knowledge, we are the first study to examine the contribution of neighborhood environments to racial/ethnic disparities in a comprehensive set of CVD risk factors. There have been a few studies that examine neighborhood context in relation to racial/ethnic

differences in specific CVD risk factors. For example, Morenoff et al. (Morenoff et al., 2007), found that black-white differences in hypertension were reduced by 100% and no longer statistically significant after controlling for neighborhood factors using fixed effects and random effects modeling approaches. Mujahid et al. (Mujahid et al., 2011), used MESA data (based on a subset of participants within 3 of 6 sites) to document an age and sex adjusted higher prevalence of hypertension in blacks and Hispanics compared to whites, and a 43% reduction after adjustment for neighborhood social environment indicators (neighborhood social cohesion, social disorder, and safety). In our analyses, we only found 21% reductions in black-white differences in blood pressure and 11% reductions in Hispanic-white differences and significant racial/ethnic differences in our study outcomes remained after adjustment for neighborhood context. Aside from a handful of studies that have examined the contribution of neighborhood context to racial/ethnic differences in obesity (Nicholson and Browning, 2012) and fruit and vegetable intake (Dubowitz et al., 2008a), this area remains under-investigated. Results across these different studies are not directly comparable due to differences in study population characteristics and differences in covariate adjustments. However, the range of reductions in racial/ethnic differences across studies should spur continued investigation into the effects of neighborhoods on racial/ethnic disparities in CVH.

In this study, we also examined whether reductions in racial/ethnic differences in ideal cardiovascular health indicators were robust across three modeling approaches. We compared results of fixed effects, random effects, and hybrid fixed effects models and found that the largest percent reduction was observed most consistently from the hybrid fixed effects approach. The fixed effects model, often considered the gold standard as the most conservative approach, adjusts for all measured and unmeasured neighborhood characteristics related to the study outcomes. However, this approach does not allow one to consider specific neighborhood characteristics that may contribute to the racial/ethnic differences. In our analyses, the fixed effects approach often led to very mild decreases or slight increases in racial/ethnic differences across study outcomes. As an example, blacks had a 60% lower odds of ideal BMI compared to whites in initial models unadjusted for neighborhood context (OR=0.401; 95% C.I. 0.339–0.475) and these differences increased using fixed effects adjustments (OR=0.363, 95% C.I. 0.276–0.478). Our ability to maximize the fixed effects modeling approach may have been compromised by having a limited number of individuals within neighborhoods and retaining only neighborhoods with at least two of the three racial/ethnic groups being compared in analyses. For example, we lost over half of our analytic sample in analyses of overall ideal CVH as there was not at least one white and either black or Hispanic participant with ideal levels of CVH in many neighborhoods. Alternatively, the random effects and hybrid fixed effects model allows for adjustment of neighborhood factors and inclusion of six neighborhood characteristics across socioeconomic, physical and social environments. These models produced larger reductions in racial/ethnic differences particularly in the case of the hybrid fixed effects models. However, reductions were still modest as none of these models fully attenuated results or reduced racial/ethnic differences to being non-statistically significant. Moreover, both the random and hybrid effects approaches have limitations. The random effects model assumes that race/ethnicity is conditionally uncorrelated with the neighborhood random effects, an

assumption difficult to verify. Future research is necessary to compare these models and to identify the ideal circumstances to increase the performance of each model.

Strengths of this study include the use of the MESA study, which provides a racially and ethnically diverse set of individuals and information on a rich set of cardiovascular health indicators and indicators of neighborhood physical, social, and socioeconomic environments to allow for investigations of our research question. However, a few limitations of our study warrant consideration. First, although we had a multi-ethnic cohort, not all racial/ethnic groups are represented across study sites. White participants were the only group recruited across all study sites. Blacks were recruited from all sites except St. Paul, MN, and Hispanics were recruited from 3 of 6 sites, (New York, NY; Los Angeles CA; St. Paul, MN). Moreover, we had to exclude Chinese Americans from analyses due to limited sample sizes. Limited diversity across sites and the fact that MESA participants are free from CVD at baseline means that we may have an underestimate of the underlying racial/ethnic differences in the target populations.

There are also several limitations in relation to the neighborhood-level component of our study. Although we included a comprehensive set of neighborhood physical, social, and socioeconomic indicators, we defined neighborhoods as census tracts, and there is debate as to whether these geographic boundaries represent meaningful areas for residents or represent the relevant spatial scale for which processes related to neighborhoods conditions affect various health outcomes (Diez Roux, 2001; Diez Roux and Mair, 2010). Census tracts also vary substantially in spatial size depending on the population density of an area and more dense urban areas (e.g. New York, NY study site) will have smaller census tracts than less dense rural areas (e.g. Forsyth County, NC study site). Future work should consider whether associations vary by study site and whether results are robust to various specifications of neighborhood boundaries. Future work should also consider approaches to capture real-time place-based exposures as individuals move through space and time as well as the multiple spatial contexts individuals experience where they live, work, and play (Kestens et al., 2012; Moore et al., 2013; Zenk et al., 2011).

In this study, we used a variety of measurement tools to capture neighborhood environments including surveys of area residents, and linkages to geospatially referenced databases. Although we had an average of 5 respondents per neighborhood, the range was highly variable and some neighborhoods only included 1 individual. This may have led to bias in our survey-based responses as the reliability of aggregate neighborhood measures is sensitive to the within neighborhood sample size (Mujahid et al., 2007).

In our analyses we compared three statistical modeling approaches to account for neighborhood context, although the estimates across these models are not directly comparable given the assumptions underlying each model. We reported odds ratios to try to improve comparisons across models and to better facilitate comparisons of our study findings to other previous studies. However, odds ratios provide an overestimate of racial/ethnic differences among the highly prevalent components of the ideal cardiovascular health metric (all health factors and health behaviors with the exception of diet). There may also be other analytic approaches such as marginal structural models, that better account for

residential selection especially within the context of a longitudinal study design. Also, controlling for context at the design phase of the study may be a better approach to balance residential selection characteristics of study participants. For example, LaVeist and colleagues explored black-white health differences in the Exploring Health Disparities in Integrated Communities Study (EHDIC), which identified communities in the U.S. that are integrated (having at least 35% blacks and 35% whites living in the census tract), and have comparable socioeconomic status across racial groups (LaVeist et al., 2008). Using a nationally representative data source, they found blacks had higher odds of hypertension as compared to whites, independent of age, gender, marital status, household income, and education level (OR=2.07, 95% C.I. 1.79, 2.40). However, these differences were less pronounced in the integrated communities with the same covariate adjustment (OR=1.45, 95% C.I. 1.12, 1.88; 30% reduction in black-white difference). Similar findings were documented in relation to diabetes and obesity among women.

Racial/ethnic differences in health remain alarming and may result from more than just differences in individual factors. We show that above and beyond individual-level sociodemographic factors, neighborhood context may also contribute to these disparities. Healthy People 2020 emphasizes the importance of creating physical and social environments conducive to healthy lifestyles as having the greatest potential to improve the health of the nation. Our study findings underscore the importance of this overarching goal for racial/ethnic minorities. Future research using longitudinal and life-course study designs is necessary to accurately capture the cumulative burden of living in physically, socially, and economically disadvantaged neighborhoods on racial/ethnic disparities over the lifespan.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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Highlights

- Whites were 3 times more likely to have ideal CVH compared to blacks and Hispanics.
- Neighborhood adjustment resulted in modest reductions in racial/ethnic differences
- Study results were generally robust across analytic approach

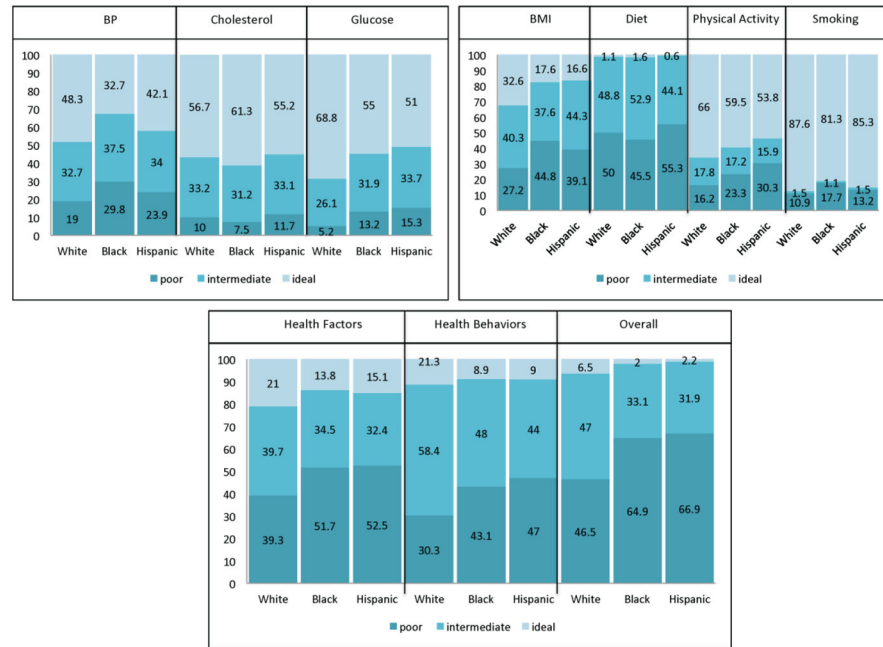


Figure 1.
Race/Ethnic Differences in Indicators of Ideal Cardiovascular Health

Table 1

Indicators of Ideal Cardiovascular Health and Prevalence in Study Population

Component	Score	Definition	% of MESA Participants (N=5263)
Cholesterol			
	Poor	≥ 240 mg/dL	9.6
	Intermediate	200–239 mg/dL or treated to < 200 mg/dL	32.6
	Ideal	< 200 mg/dL, unmedicated	57.8
Glucose			
	Poor	≥ 126 mg/dL fasting	10.1
	Intermediate	100–125 mg/dl fasting unmedicated or treated to < 100 mg/dL	29.7
	Ideal	< 100 mg/dL fasting, unmedicated	60.2
BP			
	Poor	SBP ≥ 140 mmHg or DBP ≥ 90 mmHg	23.5
	Intermediate	SBP 120–139 mmHg or DBP 80–89 mmHg	34.5
	Ideal	SBP < 120 mmHg and DBP < 80 mmHg	42.0
BMI			
	Poor	≥ 30 kg/m ²	35.5
	Intermediate	25 – 29.9 kg/m ²	40.4
	Ideal	< 25 kg/m ²	24.1
Physical Activity			
	Poor	no exercise	21.8
	Intermediate	1–149 minutes of moderate or 1–74 minutes of vigorous exercise/week	17.2
	Ideal	150+ minutes of moderate or 75+ minutes of vigorous exercise/week	61.0
Smoking			
	Poor	Current Smoker	13.5
	Intermediate	Former Smoker, quit ≤ 12 months ago	1.4
	Ideal	Never Smoker or quit > 12 months ago	85.1
Diet			
	Poor	0–1 components of healthy diet	47.7
	Intermediate	2–3 components of healthy diet	51.2
	Ideal	4–5 components of healthy diet	1.1

Table 2

Distribution of Study Covariates Overall and by Ideal Cardiovascular Health

	Overall (N = 5263)	Poor (N = 3006)	Intermediate (N = 2042)	Ideal (N = 215)	P - value
Individual-level Covariates					
Age, mean (SD)	62.0 (10.1)	62.8 (9.8)	61.2 (10.4)	58.6 (10.0)	<0.001
Gender, n (%)					0.164
Male	2491 (47.3)	1442 (48.0)	960 (47.0)	89 (41.4)	
Female	2772 (52.7)	1564 (52.0)	1082 (53.0)	126 (58.6)	
Education, n (%)					<0.001
Completed high school/GED or less	1807 (34.3)	1244 (41.4)	527 (25.8)	36 (16.7)	
Some college/technical school/associate degree	1563 (29.7)	925 (30.8)	592 (29.0)	46 (21.4)	
Bachelor's/graduate degree	1893 (36.0)	837 (27.8)	923 (45.2)	133 (61.9)	
Income, US \$, n (%)					<0.001
<24,999	1383 (26.3)	943 (31.4)	414 (20.3)	26 (12.1)	
25,000–49,999	1530 (29.1)	903 (30.0)	564 (27.6)	63 (29.3)	
50,000–74,999	1417 (27.0)	727 (24.2)	626 (30.7)	64 (29.8)	
≥75,000	739 (14.0)	296 (9.9)	383 (18.8)	60 (27.9)	
Missing	194 (3.7)	137 (4.6)	55 (2.7)	2 (0.9)	
Study Site, n (%)					<0.001
North Carolina (WFU)	942 (17.9)	556 (18.5)	349 (17.1)	37 (17.2)	
New York City, NY (Columbia)	971 (18.5)	530 (17.6)	392 (19.2)	49 (22.8)	
Baltimore, MD (Johns Hopkins)	930 (17.7)	558 (18.6)	343 (16.8)	29 (13.5)	
Minneapolis, MN (UMN)	957 (18.2)	549 (18.3)	371 (18.2)	37 (17.2)	
Chicago, IL (Northwestern)	775 (14.7)	358 (11.9)	370 (18.1)	47 (21.9)	
Los Angeles, CA (UCLA)	688 (13.1)	455 (15.1)	217 (10.6)	16 (7.4)	
Neighborhood-level Covariates					
Socioeconomic Environment, mean (SD)					
Factor 1	−1 (6.7)	0.9 (6.2)	−3 (7.0)	−3.4 (7.4)	<0.001
Factor 2	1.9 (4.5)	2.0 (4.6)	1.8 (4.4)	1.9 (4.4)	0.155
Factor 3	−0.4 (1.0)	−0.3 (1.0)	−0.4 (1.0)	−0.5 (0.9)	<0.001
Physical Environment, mean (SD)					

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	Overall (N = 5263)	Poor (N = 3006)	Intermediate (N = 2042)	Ideal (N = 215)	P- value
Health foods	-0.2 (1.7)	-0.4 (1.7)	-0.1 (1.9)	0.2 (2.0)	<0.001
Physical activity	-0.1 (1.7)	-0.3 (1.5)	0.2 (1.9)	0.7 (2.2)	<0.001
Social Environment, mean (SD)	-0.1 (2.8)	-0.3 (2.8)	0.0 (2.8)	0.2 (2.9)	<0.001

Table 3

Adjusted Racial/Ethnic Differences in Ideal Cardiovascular Health Factors (N=5263)

	Blood Pressure ^a	Cholesterol ^a	Glucose ^a
	OR [95% C.I.]	OR [95% C.I.]	OR [95% C.I.]
Black-white			
Unadjusted ^c	0.510 [0.438–0.594] †	1.230 [1.066–1.419] †	0.581 [0.502–0.673] †
Fixed Effect ^{c,d}	0.597 [0.470–0.759] †	1.248 [0.993–1.568]	0.614 [0.485–0.777] †
Random Effect ^{c,e}	0.545 [0.462, 0.642] †	1.200 [1.028–1.401] †	0.605 [0.514–0.713] †
Hybrid Fixed Effect ^{c,f}	0.615 [0.490, 0.772] †	1.211 [0.977–1.502]	0.649 [0.521–0.808] †
Largest % change	21.4%	N/A	16.2%
Hispanic-white			
Unadjusted ^c	0.600 [0.498–0.724] †	1.127 [0.943–1.348]	0.542 [0.450–0.652] †
Fixed Effect ^{c,d}	0.614 [0.470–0.803] †	1.206 [0.932–1.561]	0.640 [0.491–0.835] †
Random Effect ^{c,e}	0.625 [0.516–0.757] †	1.107 [0.923–1.328]	0.555 [0.456–0.676] †
Hybrid Fixed Effect ^{c,f}	0.642 [0.500–0.823] †	1.167 [0.919–1.482]	0.661 [0.516–0.846] †
Largest % change^g	10.5%	N/A	26.0%
P-value^b			
Unadjusted ^c	<0.001	0.017	<0.001
Fixed Effect ^{c,d}	<0.001	0.100	<0.001
Random Effect ^{c,e}	<0.001	0.066	<0.001
Hybrid Fixed Effect ^{c,f}	<0.001	0.149	<0.001

Sample size for fixed effects model: N=4510 for blood pressure, N=4501 for Cholesterol, and N=4505 for glucose

† P-value<0.05 for beta coefficient for black-white or Hispanic-white comparison and CVH indicator

^a Outcome compares ideal vs. intermediate/poor CVH^b p-value for overall association between race/ethnicity and CVH indicator^c All models include the following covariates: study site, gender, race/ethnicity, age, education, income^d Fixed effects model includes: neighborhood tract identifier^e Random effects model includes: neighborhood SES, neighborhood physical environment, neighborhood social environment^f Hybrid fixed effects model includes: neighborhood SES, neighborhood physical environment, neighborhood social environment, neighborhood proportion black, neighborhood proportion Hispanic

N/A listed for largest percent change if estimate was in the unexpected direction or racial/ethnic differences were not statistically significant in unadjusted model

Table 4**Adjusted Racial/Ethnic Differences in Ideal Cardiovascular Health Behaviors**

	BMI^a	Smoking^a	Physical Activity^a	Diet^a
	OR [95% C.I.]	OR [95% C.I.]	OR [95% C.I.]	OR [95% C.I.]
Black-white				
Unadjusted ^c	0.401 [0.339– 0.475] †	0.713 [0.584– 0.872] †	0.894 [0.773– 1.034]	0.446 [0.229– 0.870] †
Fixed Effect ^{c,d}	0.363 [0.276– 0.478] †	0.704 [0.509– 0.974] †	0.954 [0.758– 1.201]	0.319 [0.098– 1.036]
Random Effect ^{c,e}	0.415 [0.344– 0.501] †	0.803 [0.639– 1.007]	0.998 [0.849– 1.173]	0.421 [0.198– 0.893] †
Hybrid Fixed Effect ^{c,f}	0.412 [0.320– 0.531] †	0.747 [0.551– 1.014]	0.969 [0.778– 1.207]	0.370 [0.133– 1.029]
Largest % change^g	2.3%	31.4%	N/A	N/A
Hispanic-white				
Unadjusted ^c	0.417 [0.334– 0.520] †	1.265 [0.985– 1.624]	0.864 [0.720– 1.037]	0.349 [0.136– 0.899] †
Fixed Effect ^{c,d}	0.454 [0.327– 0.628] †	1.188 [0.849– 1.679]	0.910 [0.700– 1.183]	0.211 [0.034– 1.310]
Random Effect ^{c,e}	0.428 [0.340– 0.538] †	1.322 [1.013– 1.724] †	0.920 [0.759– 1.114]	0.349 [0.129– 0.941] †
Hybrid Fixed Effect ^{c,f}	0.494 [0.366– 0.666] †	1.170 [0.844– 1.621]	0.942 [0.737– 1.205]	0.319 [0.088– 1.146]
Largest % change^g	13.2%	N/A	N/A	N/A
P-value^b				
Unadjusted ^c	<0.001	<0.001	0.172	0.017
Fixed Effect ^{c,d}	<0.001	0.035	0.757	0.049
Random Effect ^{c,e}	<0.001	0.002	0.657	0.030
Hybrid Fixed Effect ^{c,f}	<0.001	0.058	0.881	0.068

Sample size for fixed effects model: N=3981 for BMI; N=3687 for smoking; N=4412 for physical activity; N=978 for diet

† P-value<0.05 for beta coefficient for black-white or Hispanic-white comparison and CVH indicator

^a Outcome compares ideal vs. intermediate/poor CVH

^b p-value for overall association between race/ethnicity and CVH indicator

^c All models include the following covariates: study site, gender, race/ethnicity, age, education, income

^d Fixed effects model includes: neighborhood tract identifier

^e Random effects model includes: neighborhood SES, neighborhood physical environment, neighborhood social environment

^f Hybrid fixed effects model includes: neighborhood SES, neighborhood physical environment, neighborhood social environment, neighborhood proportion black, neighborhood proportion Hispanic

N/A listed for largest percent change if estimate was in the unexpected direction or racial/ethnic differences were not statistically significant in unadjusted models

Table 5

Adjusted Racial/Ethnic Differences in Ideal Cardiovascular Health Summary Measures

	Health Factors	Health Behaviors	Overall Ideal CVH
	OR [95% C.I.]	OR [95% C.I.]	OR [95% C.I.]
Black-white			
Unadjusted ^c	0.613 [0.505– 0.744] [†]	0.362 [0.292– 0.448] [†]	0.324 [0.216– 0.487] [†]
Fixed Effect ^{c,d}	0.691 [0.509– 0.939] [†]	0.333 [0.234– 0.473] [†]	0.419 [0.221– 0.797] [†]
Random Effect ^{c,e}	0.622 [0.504– 0.766] [†]	0.398 [0.313– 0.505] [†]	0.349 [0.223– 0.548] [†]
Hybrid Fixed Effect ^{c,f}	0.689 [0.516– 0.919] [†]	0.387 [0.282– 0.533] [†]	0.456 [0.256– 0.812] [†]
Largest % change^g	20.2%	5.6%	19.5%
Hispanic-white			
Unadjusted ^c	0.680 [0.536– 0.862] [†]	0.441 [0.334– 0.581] [†]	0.403 [0.245– 0.664] [†]
Fixed Effect ^{c,d}	0.770 [0.552– 1.075]	0.435 [0.284– 0.668] [†]	0.422 [0.205– 0.870] [†]
Random Effect ^{c,e}	0.681 [0.535– 0.868] [†]	0.477 [0.358– 0.636] [†]	0.430 [0.257– 0.718] [†]
Hybrid Fixed Effect ^{c,f}	0.765 [0.561– 1.044]	0.527 [0.362– 0.769] [†]	0.452 [0.236– 0.867] [†]
Largest % change	28.1%	15.4%	8.2%
P-value^b			
Unadjusted ^c	<0.001	<0.001	<0.001
Fixed Effect ^{c,d}	0.035	<0.001	0.003
Random Effect ^{c,e}	<0.001	<0.001	<0.001
Hybrid Fixed Effect ^{c,f}	0.021	<0.001	0.004

[†]P-value<0.05 for beta coefficient for black-white or Hispanic-white comparison and CVH indicator^aOutcome compares ideal vs. intermediate/poor CVH^bp-value for overall association between race/ethnicity and CVH indicator^cAll models include the following covariates: study site, gender, race/ethnicity, age, education, income^dFixed effects model includes: neighborhood tract identifier^eRandom effects model includes: neighborhood SES, neighborhood physical environment, neighborhood social environment^fHybrid fixed effects model includes: neighborhood SES, neighborhood physical environment, neighborhood social environment, neighborhood proportion black, neighborhood proportion Hispanic