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## Gradients in Depressive Symptoms by Socioeconomic Position Among Men Who Have Sex With Men in the EXPLORE Study

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### Abstract

**Objectives**—This study examines gradients in depressive symptoms by socioeconomic position (SEP) (i.e., income, education, employment) in a large, multi-site sample of MSM.

**Methods**—Baseline data were used from EXPLORE, a randomized controlled behavioral HIV prevention trial for HIV-uninfected MSM in 6 U.S. cities (n=4,277). Depressive symptoms were

assessed using the Center for Epidemiologic Studies Depression scale (short form). A series of multiple linear regressions were fitted with interaction terms to assess additive and multiplicative relationships between SEP and depressive symptoms.

**Results**—Depressive symptoms were more prevalent among MSM with lower income, lower educational attainment, and those in the unemployed/other employment category. Income, education and employment all made significant contributions in additive models after adjustment. The employment and income interaction was statistically significant, indicating a multiplicative effect.

**Conclusions**—This study revealed gradients in depressive symptoms across SEP of MSM, pointing to income and employment status, and to a lesser extent education, as key factors for understanding heterogeneity of depressive symptoms.

### Keywords

MSM; depression; gradients; population health; epidemiology

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### Introduction

Men who have sex with men (MSM) and other sexual minority men (e.g., men who identify as gay or bisexual) are at a disproportionate risk for mental health problems (e.g., depressive symptomatology, distress) compared to their heterosexual peers (Mills, 2004). Large population health studies in the U.S. and internationally have demonstrated mental health disparities experienced by various populations of sexual minority men (Cochran & Mays, 2005; King et al., 2008; Meyer, 2003). The minority stress model is the leading theoretical framework for explaining those disparities, positing that sexual minorities experience chronic distal (e.g., discrimination, victimization) and proximal stressors (e.g., expectations of rejection, hyper-vigilance) (Meyer, 1995, 2003). A growing body of evidence has pointed to the importance of socioeconomic position (SEP) (e.g., income, employment, education) as salient influences on the health of sexual minority adults (Institute of Medicine, 2011). Discrimination because of SEP has been found to predict higher depressive and anxious symptoms in gay and bisexual men (Gamarel, Reisner, Parsons, & Golub, 2012). Evidence indicates that sexual minority men experience institutional discrimination in a range of areas, such as employment and housing (Hatzenbuehler, McLaughlin, Keyes, & Hasin, 2010), and structural factors have been implicated in lower incomes of sexual minorities despite equivalent or higher educational levels compared to their heterosexual peers (American Psychological Association, n.d.; Waite & Denier, 2015).

SEP is well known to be associated with inverse gradients across a wide range of health outcomes, including mental health (Adler et al., 1994; Frank, Cohen, Yen, Balfour, & Smith, 2003; Krieger, Williams, & Moss, 1997; Marmot & Wilkinson, 2005; Smith, Smith, Etches, & Mustard, 2012). The social gradient theory of health suggests that a person's health status is a function of their SEP and, through the structured nature of social processes, is liable to selective accumulation of future advantage or disadvantage (Marmot & Wilkinson, 2005). For men in particular, low income has been found to be associated with disproportionately higher rates of ill health, with income gradients more pronounced and non-linear than those

for women (Frank et al., 2003). While there is evidence of a strong association between SEP and health gradients in studies of general populations (Adler et al., 1994; Frank et al., 2003; Krieger et al., 1997; Marmot & Wilkinson, 2005; Smith et al., 2012), there is a paucity of research on gradients among sexual minorities. The consideration of SEP to understand the health of sexual minority men has been largely limited to the field of sexual health (HIV in particular), where they are used as both correlates and moderators of studied relationships (Halkitis & Figueroa, 2013; Janssen, de Wit, Hospers, & van Griensven, 2001; LA McGarrity & Huebner, 2014; Raymond et al., 2014). In the field of mental health, few studies have focused on SEP as correlates of mental health outcomes in this population (Friedman et al., 2014; Irwin, Coleman, Fisher, & Marasco, 2014; Mao et al., 2009). One recent study documented the multiplicative disadvantage of low income and sexual minority status of gay and bisexual Canadian men (Veenstra, 2011). To our knowledge, no studies have examined gradients—in either between-group or within-group analyses—in the prevalence of mental health outcomes among MSM.

To address this knowledge gap, we undertook the current study to examine gradients in depressive symptoms by SEP (income, education, employment) in a sample of MSM who participated in the EXPLORE Study, a randomized behavioral interventional trial of HIV negative men. Although previous examinations of the EXPLORE data have shown that younger participants had elevated prevalence of depressive symptoms compared to their older counterparts (Salomon et al., 2009), the associations of depressive symptoms with income, education, and employment, as well as the interactions between them, have not been studied. The study hypothesizes that: (H1) gradients in depressive symptoms will be observed for MSM at disadvantaged positions (e.g., low vs. high education), and (H2) the largest gradients will be observed for MSM at intersections of disadvantaged socioeconomic positions (e.g., low income *and* low education). The examination of the gradients in mental health may shed light on the basic mechanisms by which higher SEP conveys health benefits; for example, whether they are related to material deprivation, psycho-social factors, or others (Frank et al., 2003). This type of knowledge is needed to inform and advance policies, services, and programs designed to address the mental health disparities of MSM as well as to direct interventions at subgroups of men who experience the highest burden of illness.

## Methods

### Study Population

The EXPLORE Study was a behavioral intervention trial of HIV negative MSM. Recruitment occurred between January 1999 and February 2001 in 6 U.S. cities: Boston, Chicago, Denver, New York, San Francisco, and Seattle. Men were eligible if they were aged 16 years or older and reported having engaged in anal sex with 1 or more men during the past year. Men were excluded if they reported that they had been involved in a mutually monogamous relationship for 2 or more years with a male partner known to be negative for HIV antibodies. The recruitment, sampling, data collection, measures, and other methods of the EXPLORE Study have been described in detail elsewhere (Barresi et al., 2010; Chesney et al., 2003; Koblin et al., 2003; Koblin, Chesney, & Coates, 2004). In brief, participants

were recruited using street- and venue-based outreach in areas where MSM were known to congregate (e.g., dance clubs, bathhouses, health clubs). Project staff also mounted public relations and media campaigns, with participants recruited through Internet sites, community forums, and community-based agencies. Finally, individuals were referred from other study participants and clinics (Marshall et al., 2015). The present study is a cross-sectional examination of the EXPLORE data at baseline (n=4,295), restricted to those who have complete baseline information for the primary outcome of interest (n=4,277).

### Study Measures

The primary outcome in the study was a measure of depressive symptomatology, defined using a short form of the Center for Epidemiologic Studies Depression Scale (CESD-SF). This 7-item measure asked respondents to rate how often during the past week they: felt like they could not shake off the blues even with help from their family or friends, had trouble keeping their minds on what they were doing, felt that everything they did was an effort, had trouble sleeping, felt lonely, felt sad, or felt like they “just couldn’t get going.” Responses corresponded to the number of days the symptoms were experienced [never/rarely (0 days; scale score=0), sometimes (1–2 days; scale score=1), often (3–4 days; scale score= 2), mostly or always (5–7 days; scale score=3)]. The scores for each item were summed to create a composite score, with higher scores representing greater depressive symptomatology. A cut-off point for the CESD-SF has not been established, although a median split has previously been used (Salomon et al., 2009). The use of the short version of the CESD has been previously validated by (Ross & Mirowsky, 2006). The validity of alternative versions of the CESD-SF has been documented in other studies (Andresen, Malmgren, Carter, & Patrick, 1994; Chang & Weng, 2013; Kohout, Berkman, Evans, & Cornoni-Huntley, 1993; Levine, 2013), including in studies with MSM (Salomon et al., 2009), and vulnerable and minority groups (Burton, Marshal, Chisolm, Sucato, & Friedman, 2013; Callahan & Wolinsky, 1994; Kim, Decoster, Huang, & Chiriboga, 2011; Perreira, Deeb-Sossa, Harris, & Bollen, 2005).

Based on examinations of health gradients in the general population (Frank et al., 2003; Krieger, 2001; Myers, 2009; Smith et al., 2012), the following SEP characteristics were selected as independent variables: household annual income, educational attainment, and employment status. Household annual income (<\$11,999, \$12,000 – \$29,999, \$30,000 – \$59,999, \$60,000+), educational attainment (High school or less [including GED], Some College, College Degree, Post-College), and current employment status (Employed full-time, part-time, unemployed, or other) were measured based on categorizations previously used in other studies using EXPLORE data (Salomon et al., 2009). The other employment category included those on disability or “between jobs,” and therefore the unemployed/other categories were collapsed.

The selection of covariates was guided by a review of the literature. The race/ethnicity variable was obtained using a combination of two separate variables on race/ethnicity: (a) Are you white, black, Asian/Pacific Islander, or Other race?; and (b) Are you Latino or Hispanic?, which were used to create the following mutually exclusive categories: White/Not Hispanic, Black/Not Hispanic, Hispanic, Other/Not Hispanic (including Asian/

Pacific Islander). Other covariates in the study included age group (16–25, 26–30, 31–35, and 36+), geographic site (Boston, Chicago, Denver, New York, San Francisco, and Seattle), and having experienced childhood sexual abuse (yes/no). Studies have shown that MSM are more likely to report childhood sexual abuse, and it has been suggested that proximate causes of psychological distress for some MSM may stem from early childhood experiences, such as physical and emotional abuse by family or peers (Mayer et al., 2012). The childhood sexual abuse variable was defined as any sexual experience either with a person 5 years or older before the age of 13, or with a person 10 years or older between the ages of 13 and 17. This definition is consistent with existing literature on sexual abuse and has been used previously (Mimiaga et al., 2009).

## Statistical Analyses

The analytical strategy included modeling the main effects of the SEP variables in additive models (hypothesis 1), and then interactions between them in multiplicative models (hypothesis 2), on the primary outcome of depressive symptoms (CESD-SF) ( $n=4,277$ ). A series of multiple linear regression models were fitted to examine the independent relationships between the SEP variables with depressive symptomatology. Interaction was examined using 2-way interaction terms. All regression models adjusted for age group, race/ethnicity, history of sexual abuse, and geographic site. The results from the various models were compared and contrasted, and their model fit assessed. The models were implemented in Stata 13 using regress procedures. Tests of collinearity between all independent variables showed no substantial collinearity (all correlation coefficients  $< 0.15$ ). All reported  $p$ -values are two-sided.

## Findings

### Sample Description

Table 1 displays the characteristics of the EXPLORE Study participants and their CESD-SF scores at baseline. The mean CESD-SF score in the sample was 6.0 ( $SD = 4.3$ ). The median score was 5.0, and the range represented the full range of possible scale scores (0–21). The majority of participants were white/not Hispanic, young (35 years old or younger), employed full-time, and had college or post college education, with just over 40% of participants reporting household income of less than \$29,999. A large proportion (40%) reported having experienced childhood sexual abuse. Statistical analyses indicated that disadvantaged socioeconomic positions were associated with greater depressive symptomatology compared to non-disadvantaged positions (see Table 1). The highest CESD-SF scores were observed for those with household income of \$11,999 or less (mean score = 7.9), those with educational attainment of high school or less (mean score = 7.2), those who were in the unemployed/other employment category (mean score = 7.2), and those aged 16–25 (mean score = 6.8), and these associations were significant ( $p < 0.05$ ) in unadjusted models (2-tailed tests). Figure 1 shows the distributions of depressive symptoms at intersections of the SEP variables. A consistent gradient of depressive symptoms can be seen, such that CESD-SF scores experienced by MSM are graded within and across socioeconomic positions (e.g., scores are graded by income within each category of

educational attainment, and they are highest for those at intersections of low income and low education).

### Multiple Regression Results

As hypothesized (H1), household income, educational attainment and employment status all made significant contributions to additive models after adjustment (see Table 2). In the multiple model (R.4), respondents in the <\$11,999 group had 2.2 higher units of depressive symptoms compared to those in the > \$60,000 group, those with high school education or less had 1.0 higher units of depressive symptoms on average compared to those in the group with post college education, and those in unemployed/other employment category had 0.74 higher units of depressive symptoms compared to those employed full-time (all  $p < 0.001$ ). Table 3 shows results of regression models with 2-way interactions between the SEP variables. As hypothesized (H2), the interaction between employment and income was significant ( $p < 0.01$ ); however, interactions between employment and education (R.2), and education and income (R.3) were not significant. When all cross-product terms were added in model R.4, the employment and income interaction continued to retain significance ( $p = 0.01$ ). Household income remained significant and its regression coefficients stable in all regression models with interaction terms. In our final model with the statistically significant interaction for employment and education (R.1), the individual SEP variables also retained significance. Compared to their respective reference groups, respondents in the <\$11,999 group had 2.4 higher units of depressive symptoms, those in part-time employment category had 1.7 higher units, and those with high school education or less had 0.9 higher units of depressive symptoms (all  $p < 0.05$ ).

### Discussion

This study revealed gradients in depressive symptomatology by SEP amongst MSM participating in the EXPLORE Study. As hypothesized (H1), gradients in depressive symptoms were observed for MSM at more disadvantaged socioeconomic positions across income, educational attainment, and employment status in additive models. Associations between SEP variables with depressive symptoms remained significant in the fully adjusted model (R.4, Table 2). Moreover, the attenuation in the coefficient for income was small, suggesting income may be particularly relevant for understanding heterogeneity in depressive symptomatology observed in the sample. In the models testing statistical interactions to examine gradients at intersections of disadvantaged socioeconomic positions (H2), a multiplicative disadvantage of low income and under-employment was observed, with the individual SEP variables still retaining significance in the model (R.1, Table 3).

These findings are consistent with studies in the general population documenting gradients in mental health by SEP (Frank et al., 2003; Krieger et al., 1997; Marmot & Wilkinson, 2005; Smith et al., 2012). The results of the current study also correspond with the results of a study of sexually active Australian gay men found that major depression was strongly associated with socioeconomic deprivation (Mao et al., 2009). A longitudinal study of MSM (who also have sex with women) also found income to be a predictor of CESD scores, alongside race/ethnicity (Friedman et al., 2014). In the current study, the race/ethnicity

covariate did not retain significance after model adjustment (results not shown); however, race/ethnicity effects (e.g., economic forms of racial/ethnic discrimination, such as in the labour market or educational system) also may be manifested via socioeconomic factors (income; education) (Krieger, 2001; Williams & Mohammed, 2009). In a study of young (ages 18–19) MSM in New York city, the effect of race/ethnicity was no longer significant after controlling for SEP in its association with mental health outcomes (depression, PTSD, and suicidality) (Storholm et al., 2013).

The current study has a number of strengths and limitations. The study utilizes a robust and publicly available secondary dataset to address knowledge gaps in exiting literature. The large sample size of the EXPLORE study provided a novel opportunity to examine gradients in reported depressive symptoms among MSM populations that are typically understudied. However, the results should be interpreted with caution. The data was collected between years 1999–2001, and it is possible that the gradients observed in our analysis may not accurately reflect the current experiences of MSM. Given the recruitment and eligibility criteria for inclusion in the EXPLORE Study, the findings may not be generalizable to the wider population of MSM, particularly HIV-infected populations, and those living outside of urban centers. An abridged version of the CESD scale provided a useful measure of depressive symptoms, but it should not be conflated with clinical diagnoses of depression.

## Conclusions

The findings of the study have important implications for both research and policy. First, social inequalities researchers in studies of general populations have long recommended that SEP be treated as a construct in its own right (as opposed to a control variable) (Adler et al., 1994). This study demonstrates that SEP may be particularly important to understanding mental health outcomes in MSM samples, suggesting that this kind of approach to SEP is relevant in studies of sexual minorities. For sexual minorities, lower SEP may be concomitantly associated with greater risk of discrimination, more barriers to employment, fewer opportunities for connection to the LGB community, and therefore greater stress (Barrett & Pollack, 2005; Gamarel et al., 2012; McGarrity & Huebner, 2014). Future research, and longitudinal studies in particular, are needed to examine the pathways and mechanisms through which SEP influences mental health outcomes among sexual minorities, so that appropriate interventions may be identified and implemented. Second, policy and program efforts targeting socioeconomic barriers facing sexual minorities (e.g., institutional discrimination in employment or housing) may hold promise in reducing depressive symptoms amongst MSM. For example, the introduction or expansion of housing programs (e.g., Ruth Ellis Center, Ali Forney Center), LGBT Centers (e.g., Los Angeles LGBT Center which provides mental health and social support services, including employment support), financial aid programs (e.g., Massachusetts transgender emergency fund; LGBT student financial assistance funds), and the provision of assistance in accessing health and legal services (e.g., Gay City Health Project, California Legal Rural Assistance LGBT Program) may help to reduce the higher rates of depression among MSM of lower SEP. Some preliminary examinations of select programs and health service models for LGBT populations have been conducted (Burwick, Friend, Gates, & Durso, 2014; Sanchez, 2008). A number of best practices for mental health facilities working with LGBT clients



have been recommended on the American Psychological Association website, and an integrated approach to clinical care for MSM has been outlined (Mayer et al., 2012). However, ongoing research and evaluations are needed to identify effective programs and interventions; in particular, those that address and ameliorate the adverse effects of low SEP. Third, because MSM continue to face high levels of stigma-related stress, the development of interventions to address homophobia and heterosexism also should be examined as a means to improve mental health outcomes amongst sexual minority men (Mayer et al., 2012; Meyer, 1995).

As the results of the current study demonstrate, there are several, intersecting factors implicated in the production of a gradient in mental health outcomes amongst MSM, which reinforces the idea that no single intervention approach is likely to fully resolve this complex problem. As previous research has shown, MSM who are of low SEP may be unable to access services because of resource and financial barriers (Fredriksen-Goldsen et al., 2013; Larissa McGarrity & Huebner, 2014); they may also be unable to benefit from the resources available to gay- and bisexual- identified men due to the lower levels of identification with and access to the LGB community (Barrett & Pollack, 2005). Also, there may be complex and dynamic interactions between experiences of discrimination because of SEP and other forms of discrimination based on sexual orientation, gender identity, and race/ethnicity (Gamarel et al., 2012). While higher SEP (e.g., financial security and educational attainment) may allow some individuals to insulate themselves from exposure to sources of discrimination (McAdams-Mahmoud et al., 2014), reducing disparities in health for all sexual minorities will depend on combinations of universal (e.g., legal protection of human rights) and targeted (e.g., programs and services that are tailored to meet the mental health needs of MSM) intervention approaches.

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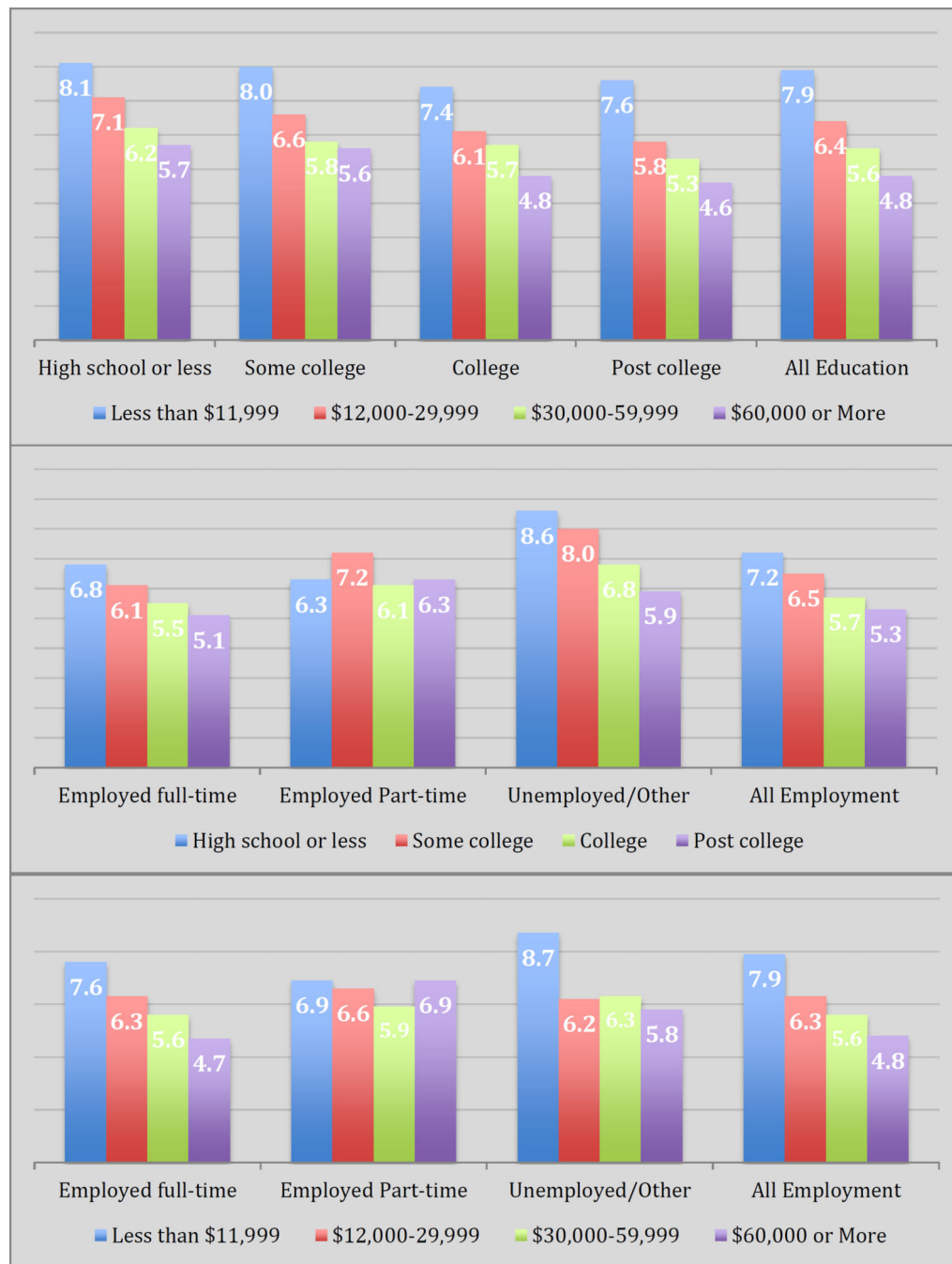
## References

- Adler NE, Boyce T, Chesney MA, Cohen S, Folkman S, Kahn RL, Syme SL. Socioeconomic status and health. The challenge of the gradient. *The American Psychologist*. 1994; 49(1):15–24. [PubMed: 8122813]
- American Psychological Association. Lesbian, Gay, Bisexual & Transgender Persons & Socioeconomic Status. Fact Sheet. [Retrieved November 10, 2014] (n.d.). from <http://www.apa.org/pi/ses/resources/publications/factsheet-lgbt.aspx>.
- Andresen EM, Malmgren JA, Carter WB, Patrick DL. Screening for depression in well older adults: evaluation of a short form of the CES-D (Center for Epidemiologic Studies Depression Scale). *American Journal of Preventive Medicine*. 1994; 10(2):77–84. [PubMed: 8037935]
- Barresi P, Husnik M, Camacho M, Powell B, Gage R, LeBlanc D, Koblin B. Recruitment of men who have sex with men for large HIV intervention trials: analysis of the EXPLORE Study recruitment effort. *AIDS Education and Prevention : Official Publication of the International Society for AIDS Education*. 2010; 22(1):28–36. [PubMed: 20166785]

- Barrett DC, Pollack LM. Whose Gay Community? Social Class, Sexual Self-Expression, and Gay Community Involvement. *The Sociological Quarterly*. 2005; 46(3):437–456.
- Burton CM, Marshal MP, Chisolm DJ, Sucato GS, Friedman MS. Sexual minority-related victimization as a mediator of mental health disparities in sexual minority youth: a longitudinal analysis. *Journal of Youth and Adolescence*. 2013; 42(3):394–402. [PubMed: 23292751]
- Burwick, A.; Friend, D.; Gates, G.; Durso, L. Identifying and Serving LGBTQ Youth: Case Studies of Runaway and Homeless Youth Program Grantees. Princeton, NJ: 2014.
- Callahan CM, Wolinsky FD. The effect of gender and race on the measurement properties of the CES-D in older adults. *Medical Care*. 1994; 32(4):341–356. [PubMed: 8139299]
- Chang K, Weng L. Screening for depressive symptoms among older adults in Taiwan: Cutoff of a short form of the Center for Epidemiologic Studies Depression Scale. 2013; 5(3):588–594.
- Chesney MA, Koblin BA, Barresi PJ, Husnik MJ, Celum CL, Colfax G, Coates TJ. An individually tailored intervention for HIV prevention: baseline data from the EXPLORE Study. *American Journal of Public Health*. 2003; 93(6):933–938. [PubMed: 12773358]
- Cochran, SD.; Mays, VM. Estimating Prevalence of Mental Substance-Using Disorders Among Lesbians and Gay Men from Existing National Health Data. In: Omoto, A.; Kurtzman, H., editors. *Sexual Orientation, Mental Health, and Substance Use: Contemporary Scientific Perspectives*. Washington, DC: American Psychological Association; 2005. p. 143–165.
- Frank JW, Cohen R, Yen I, Balfour J, Smith M. Socioeconomic gradients in health status over 29 years of follow-up after midlife: the Alameda county study. *Social Science & Medicine*. 2003; 57(12): 2305–2323. [PubMed: 14572839]
- Fredriksen-Goldsen KI, Emler C, Kim H-J, Muraco A, Erosheva E, Goldsen J, Hoy-Ellis CP. The physical and mental health of lesbian, gay male, and bisexual (LGB) older adults: the role of key health indicators and risk and protective factors. *The Gerontologist*. 2013; 53(4):664–675. [PubMed: 23034470]
- Friedman MR, Stall R, Silvestre AJ, Mustanski B, Shoptaw S, Surkan PJ, Plankey MW. Stuck in the middle: longitudinal HIV-related health disparities among men who have sex with men and women. *Journal of Acquired Immune Deficiency Syndromes (1999)*. 2014; 66(2):213–220. [PubMed: 24662298]
- Gamarel KE, Reisner SL, Parsons JT, Golub S. Association between socioeconomic position discrimination and psychological distress: Findings from a community-based sample of gay and bisexual men in New York City. *American Journal of Public Health*. 2012; 102(11):2094–2101. [PubMed: 22994188]
- Halkitis PN, Figueroa RP. Sociodemographic characteristics explain differences in unprotected sexual behavior among young HIV-negative gay, bisexual, and other YMSM in New York City. *AIDS Patient Care and STDs*. 2013; 27(3):181–190. [PubMed: 23442029]
- Hatzenbuehler ML, McLaughlin KA, Keyes KM, Hasin DS. The impact of institutional discrimination on psychiatric disorders in lesbian, gay, and bisexual populations: A prospective study. *American Journal of Public Health*. 2010; 100(3):452–459. [PubMed: 20075314]
- Institute of Medicine. *The Health of Lesbian, Gay, Bisexual, and Transgender People: Building a Foundation for Better Understanding*. Washington, DC: The National Academies Press; 2011.
- Irwin JA, Coleman JD, Fisher CM, Marasco VM. Correlates of suicide ideation among LGBT Nebraskans. *Journal of Homosexuality*. 2014; 61(8):1172–1191. [PubMed: 24344775]
- Janssen M, de Wit J, Hospers HJ, van Griensven F. Educational status and young Dutch gay men's beliefs about using condoms. *AIDS Care*. 2001; 13(1):41–56. [PubMed: 11177464]
- Kim G, Decoster J, Huang C-H, Chiriboga D. Race/ethnicity and the factor structure of the Center for Epidemiologic Studies Depression Scale: a meta-analysis. *Cultural Diversity & Ethnic Minority Psychology*. 2011; 17(4):381–396. [PubMed: 21988578]
- King M, Semlyen J, Tai SS, Killaspy H, Osborn D, Popelyuk D, Nazareth I. A systematic review of mental disorder, suicide, and deliberate self harm in lesbian, gay and bisexual people. *BMC Psychiatry*. 2008; 8:70. [PubMed: 18706118]
- Koblin B, Chesney MA, Coates TJ. Effects of a behavioural intervention to reduce acquisition of HIV infection among men who have sex with men: the EXPLORE randomised controlled study. *Lancet*. 2004; 364(9428):41–50. [PubMed: 15234855]

- Koblin B, Chesney MA, Husnik MJ, Bozeman S, Celum CL, Buchbinder S, Coates TJ. High-risk behaviors among men who have sex with men in 6 US cities: baseline data from the EXPLORE Study. *American Journal of Public Health*. 2003; 93(6):926–932. [PubMed: 12773357]
- Kohout FJ, Berkman LF, Evans DA, Cornoni-Huntley J. Two Shorter Forms of the CES-D Depression Symptoms Index. *Journal of Aging and Health*. 1993; 5(2):179–193. [PubMed: 10125443]
- Krieger N. Theories for social epidemiology in the 21st century: an ecosocial perspective. *International Journal of Epidemiology*. 2001; 30(4):668–677. [PubMed: 11511581]
- Krieger N, Williams DR, Moss NE. Measuring social class in US public health research: concepts, methodologies, and guidelines. *Annual Review of Public Health*. 1997; 18:341–378.
- Levine SZ. Evaluating the seven-item Center for Epidemiologic Studies depression scale short-form: a longitudinal U.S. community study. *Social Psychiatry and Psychiatric Epidemiology*. 2013; 48(9): 1519–1526. [PubMed: 23299927]
- Mao L, Kidd MR, Rogers G, Andrews G, Newman CE, Booth A, Kippax SC. Social factors associated with Major Depressive Disorder in homosexually active, gay men attending general practices in urban Australia. *Australian and New Zealand Journal of Public Health*. 2009; 33(1):83–86. [PubMed: 19236365]
- Marmot, M.; Wilkinson, R., editors. *Social Determinants of Health*. Second. New York: Oxford University Press; 2005.
- Marshall BDL, Shoveller J, Kahler CW, Koblin B, Mayer KH, Mimiaga MJ, Operario D. Heavy Drinking Trajectories Among Men Who Have Sex with Men: A Longitudinal, Group-Based Analysis. *Alcoholism: Clinical and Experimental Research*. 2015; 39(2):380–389.
- Mayer KH, Bekker L-G, Stall R, Grulich AE, Colfax G, Lama JR. Comprehensive clinical care for men who have sex with men: an integrated approach. *The Lancet*. 2012; 380(9839):378–387.
- McAdams-Mahmoud A, Stephenson R, Rentsch C, Cooper H, Arriola KJ, Jobson G, McIntyre J. Minority stress in the lives of men who have sex with men in Cape Town, South Africa. *Journal of Homosexuality*. 2014; 61(6):847–867. [PubMed: 24392722]
- McGarrity L, Huebner D. Behavioral intentions to HIV test and subsequent testing: The moderating role of sociodemographic characteristics. *Health Psychology*. 2014; 33(4):396–400. [PubMed: 23795706]
- McGarrity L, Huebner DM. Is being out about sexual orientation uniformly healthy? The moderating role of socioeconomic status in a prospective study of gay and bisexual men. *Annals of Behavioral Medicine*. 2014; 47(1):28–38. [PubMed: 24307473]
- Meyer IH. Minority Stress and Mental Health in Gay Men. *Journal of Health and Social Behavior*. 1995; 36(1):38. [PubMed: 7738327]
- Meyer IH. Prejudice, social stress, and mental health in lesbian, gay, and bisexual populations: conceptual issues and research evidence. *Psychological Bulletin*. 2003; 129(5):674–697. [PubMed: 12956539]
- Mills TC. Distress and Depression in Men Who Have Sex With Men: The Urban Men’s Health Study. *American Journal of Psychiatry*. 2004; 161(2):278–285. [PubMed: 14754777]
- Mimiaga MJ, Noonan E, Donnell D, Safren S, Koenen KC, Gortmaker S, Mayer KH. Childhood sexual abuse is highly associated with HIV risk-taking behavior and infection among MSM in the EXPLORE Study. *Journal of Acquired Immune Deficiency Syndromes (1999)*. 2009; 51(3):340–348. [PubMed: 19367173]
- Myers HF. Ethnicity- and socio-economic status-related stresses in context: an integrative review and conceptual model. *Journal of Behavioral Medicine*. 2009; 32(1):9–19. [PubMed: 18989769]
- Perreira KM, Deeb-Sossa N, Harris KM, Bollen K. What Are We Measuring? An Evaluation of the CES-D Across Race/Ethnicity and Immigrant Generation. *Social Forces*. 2005; 83(4):1567–1601.
- Raymond HF, Chen Y-H, Syme SL, Catalano R, Hutson M, McFarland W. The role of individual and neighborhood factors: HIV acquisition risk among high-risk populations in San Francisco. *AIDS and Behavior*. 2014; 18(2):346–356. [PubMed: 23760633]
- Ross CE, Mirowsky J. Sex differences in the effect of education on depression: resource multiplication or resource substitution? *Social Science & Medicine (1982)*. 2006; 63(5):1400–1413. [PubMed: 16644077]

- Salomon E, Mimiaga MJ, Husnik MJ, Welles SL, Manseau MW, Montenegro AB, Mayer KH. Depressive symptoms, utilization of mental health care, substance use and sexual risk among young men who have sex with men in EXPLORE: implications for age-specific interventions. *AIDS and Behavior*. 2009; 13(4):811–821. [PubMed: 18709453]
- Sanchez, GA. Evaluating effective health service models for the LGBT community: the healthcare scope of LGBT community centers in the United States. Amherst: University of Massachusetts; 2008.
- Smith B, Smith P, Etches J, Mustard C. Overqualification and Risk of All-cause and Cardiovascular Mortality: Evidence From the Canadian Census Mortality Follow-up Study (1991–2001). *Can J Public Health*. 2012 Aug.:1–3.
- Storholm ED, Siconolfi DE, Halkitis PN, Moeller RW, Eddy J, Bare MG. Sociodemographic Factors Contribute to Mental Health Disparities and Access to Services Among Young Men Who Have Sex with Men in New York City. *Journal of Gay & Lesbian Mental Health*. 2013; 17(3):37–41.
- Veenstra G. Race, gender, class, and sexual orientation: intersecting axes of inequality and self-rated health in Canada. *International Journal for Equity in Health*. 2011; 10(1):3. [PubMed: 21241506]
- Waite S, Denier N. Gay Pay for Straight Work: Mechanisms Generating Disadvantage. *Gender & Society*. 2015; XX(X):1–28.
- Williams DR, Mohammed S. Discrimination and racial disparities in health: evidence and needed research. *Journal of Behavioral Medicine*. 2009; 32(1):20–47. [PubMed: 19030981]



**Figure 1.**  
CESD-SF Scores at Intersections of SEP Variables among EXPLORE Study Participants

**Table 1**  
 Characteristics of the EXPLORE Study Participants and Baseline CESD-SF Scores across Strata (n=4,277)

| Sample characteristics        | CESD-SF |                |      |                      |       |
|-------------------------------|---------|----------------|------|----------------------|-------|
|                               | n       | % <sup>1</sup> | Mean | 95% CI               |       |
|                               |         |                |      | p-value <sup>2</sup> |       |
| <b>Income group</b>           |         |                |      |                      |       |
| Less than \$11,999            | 562     | 13.1           | 7.9  | 4.5–8.3              | <0.01 |
| \$12,000–29,999               | 1166    | 27.2           | 6.4  | 6.1–6.6              | <0.01 |
| \$30,000–59,999               | 1656    | 38.6           | 5.6  | 5.4–5.8              | <0.01 |
| \$60,000 or More              | 904     | 21.1           | 4.8  | 4.6–5.1              | Ref.  |
| <b>Educational attainment</b> |         |                |      |                      |       |
| High school or less           | 407     | 9.5            | 7.2  | 6.8–7.7              | <0.01 |
| Some college                  | 1129    | 26.3           | 6.5  | 6.3–6.8              | <0.01 |
| College degree                | 1534    | 35.7           | 5.7  | 5.5–5.9              | 0.01  |
| Post college                  | 1223    | 28.5           | 5.3  | 5.0–5.5              | Ref.  |
| <b>Employment Status</b>      |         |                |      |                      |       |
| Unemployed/Other              | 622     | 14.5           | 7.2  | 6.8–7.5              | <0.01 |
| Employed part-time            | 426     | 9.9            | 6.6  | 6.2–7.0              | <0.01 |
| Employed full-time            | 3247    | 75.6           | 5.6  | 5.5–5.8              | Ref.  |
| <b>Race/ ethnicity</b>        |         |                |      |                      |       |
| Black/ Not Hispanic           | 281     | 6.5            | 6.1  | 5.6–6.6              | 0.61  |
| Hispanic                      | 652     | 15.2           | 6.5  | 6.2–6.9              | <0.01 |
| Other/ Not Hispanic           | 218     | 5.1            | 6.4  | 5.8–7.0              | 0.13  |
| White/ Not Hispanic           | 3112    | 72.5           | 5.8  | 5.6–5.9              | Ref.  |
| <b>Age group</b>              |         |                |      |                      |       |
| 16–25                         | 817     | 19.0           | 6.8  | 6.5–7.1              | <0.01 |
| 26–30                         | 915     | 21.3           | 6    | 5.7–6.3              | 0.90  |

| Sample characteristics |      | CESD-SF        |      |         |                      |
|------------------------|------|----------------|------|---------|----------------------|
|                        | n    | % <sup>1</sup> | Mean | 95% CI  | p-value <sup>2</sup> |
| 31–35                  | 911  | 21.2           | 5.7  | 5.5–6.0 | 0.09                 |
| >=36                   | 1652 | 38.5           | 5.6  | 5.4–5.8 | Ref.                 |
| <b>Sexual Abuse</b>    |      |                |      |         |                      |
| Yes                    | 1723 | 40.1           | 6.5  | 6.3–6.8 | <0.01                |
| No                     | 2558 | 59.6           | 5.5  | 5.4–5.7 | Ref.                 |
| <b>Geographic site</b> |      |                |      |         |                      |
| Seattle                | 743  | 17.3           | 5.8  | 5.5–6.1 | 0.37                 |
| Chicago                | 624  | 14.5           | 5.9  | 5.6–6.3 | 0.88                 |
| Boston                 | 729  | 17.0           | 6.1  | 5.8–6.4 | 0.24                 |
| New York               | 737  | 17.2           | 6.4  | 6.1–6.7 | <0.01                |
| Denver                 | 726  | 16.9           | 5.6  | 5.3–5.9 | 0.02                 |
| San Francisco          | 736  | 17.4           | 5.8  | 5.5–6.1 | Ref.                 |

<sup>1</sup>Percentages may not add up due to rounding and missing values

<sup>2</sup>P-value for t-test comparisons with reference category

**Table 2**

Results of Linear Regressions Modeling Depressive Symptomatology (CESD-SF) among EXPLORE Study Participants

|   | Regression coefficient<br>(Standard error) |                              |                              |                              |
|---|--|------------------------------|------------------------------|------------------------------|
|   | Model R.1                                  | Model R.2                    | Model R.3                    | Model R.4                    |
| <b>Employment status</b>  |  |                              |                              |                              |
| Employed part-time  | 0.7 <sup>***</sup><br>(0.22)               |                              |                              | 0.05<br>(0.23)               |
| Unemployed/Other  | 1.4 <sup>***</sup><br>(0.19)               |                              |                              | 0.7 <sup>***</sup><br>(0.20) |
| Employed full-time  | 1.0  |                              |                              | 1.0                          |
| <b>Income level</b>   |  |                              |                              |                              |
| Less than \$11,999  |  | 2.8 <sup>***</sup><br>(0.24) |                              | 2.2 <sup>***</sup><br>(0.28) |
| \$12,000–29,999   |  | 1.4 <sup>***</sup><br>(0.20) |                              | 1.1 <sup>***</sup><br>(0.20) |
| \$30,000–59,999   |  | 0.8 <sup>***</sup><br>(0.17) |                              | 0.7 <sup>***</sup><br>(0.18) |
| \$60,000 or More  |  | 1.0                          |                              | 1.0                          |
| <b>Educational attainment</b>   |  |                              |                              |                              |
| High school or less   |  |                              | 1.7 <sup>***</sup><br>(0.25) | 1.0 <sup>***</sup><br>(0.26) |
| Some college  |  |                              | 1.0 <sup>***</sup><br>(0.19) | 0.7 <sup>***</sup><br>(0.19) |
| College degree  |  |                              | 0.3 <sup>*</sup><br>(0.16)   | 0.3<br>(0.16)                |
| Post college  |  |                              | 1.0                          | 1.0                          |
| Root MSE  | 4.2  | 4.1                          | 4.2                          | 4.1                          |
| Model <i>F</i> -test<br>( <i>df</i> <sub>1</sub> , <i>df</i> <sub>2</sub> ) | 13.0<br>(14, 4221)                         | 17.9<br>(15, 4213)           | 11.8<br>(15, 4218)           | 15.2<br>(20, 4206)           |

Notes. All models adjust for age group, race/ethnicity, history of sexual abuse, geographic site

<sup>\*\*\*</sup>  
p < .001,

<sup>\*\*</sup>  
p < .01,

<sup>\*</sup>  
p < .05



**Table 3**

Results of Linear Regressions with Statistical Interactions Modeling Depressive Symptomatology (CESD-SF) among EXPLORE Study Participants

|   | Regression coefficient<br>(Standard error) |                    |                    |                   |
|---|--|--------------------|--------------------|-------------------|
|   | Model R.1                                  | Model R.2          | Model R.3          | Model R.4         |
| <b>Employment status</b>                              |  |                    |                    |                   |
| Employed part-time                                    | 1.7*<br>(0.76)                             | 0.5<br>(0.45)      | 0.04<br>(0.24)     | 1.8*<br>(0.86)    |
| Unemployed/Other                                      | 1.0*<br>(0.51)                             | 0.1<br>(0.35)      | 0.8***<br>(0.21)   | 0.6<br>(0.57)     |
| Employed full-time                                    | 1.0  | 1.0                | 1.0                | 1.0               |
| <b>Income level</b>                                   |  |                    |                    |                   |
| Less than \$11,999                                    | 2.4***<br>(0.42)                           | 2.2***<br>(0.28)   | 2.5***<br>(0.55)   | 2.5***<br>(0.56)  |
| \$12,000–29,999                                       | 1.3***<br>(0.22)                           | 1.2***<br>(0.21)   | 0.9**<br>(0.36)    | 1.4***<br>(0.37)  |
| \$30,000–59,999                                       | 0.8***<br>(0.19)                           | 0.7***<br>(0.18)   | 0.7*<br>(0.27)     | 0.7**<br>(0.28)   |
| \$60,000 or More                                      | 1.0  | 1.0                | 1.0                | 1.0               |
| <b>Educational attainment</b>                         |  |                    |                    |                   |
| High school or less                                   | 0.9***<br>(0.26)                           | 1.1***<br>(0.31)   | 1.0<br>(0.89)      | 1.0*<br>(0.42)    |
| Some college  | 0.6***<br>(0.19)                           | 0.5*<br>(0.22)     | 0.9*<br>(0.45)     | 0.6<br>(0.46)     |
| College degree  | 0.3<br>(0.16)                              | 0.3<br>(0.18)      | 0.2<br>(0.30)      | 0.1<br>(0.30)     |
| Post college  | 1.0  | 1.0                | 1.0                | 1.0               |
| <i>Interaction terms (p-values for omnibus tests)</i> |  |                    |                    |                   |
| Employment * Income.....                              | p<0.01                                     | .....              | .....              | p=0.01            |
| Employment * Education...                             | .....                                      | p=0.08             | .....              | p=0.31            |
| Income * Education.....                               | .....                                      | .....              | p=0.92             | p=0.97            |
| Root MSE  | 4.1  | 4.1                | 4.1                | 4.1               |
| Model F-test<br>(df <sub>1</sub> , df <sub>2</sub> )  | 12.5<br>(26, 4200)                         | 12.1<br>(26, 4200) | 10.6<br>(29, 4197) | 8.6<br>(39, 4187) |

Notes. All models adjust for age group, race/ethnicity, history of sexual abuse, geographic site

\*\*\*  
p < .001,

\*\*  
p < .01,

\*  
p < .05