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APOE Effects on Cognition From Childhood To Adolescence

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Abstract

The $\epsilon 4$ allele of *APOE* is a well-established genetic risk factor for cognitive aging and dementia, although its influence on early life cognition is unknown. Consequently, we assessed associations of *APOE* genotypes with cognitive performance during 7, 12 and 16 year-assessments in our ongoing Colorado Adoption/Twin Study of Lifespan behavioral development (CATSLife). In general, *APOE* $\epsilon 4$ was associated with lower Verbal, Performance and Full Scale IQ scores during childhood and adolescence (e.g., Full Scale IQ was lower by 1.91 points per $\epsilon 4$ allele, $d = -.13$), with larger effects in females (e.g., average Full Scale IQ scores were 3.41 points lower in females per each $\epsilon 4$ allele, versus .33 points lower in males). Thus, these results suggest that deleterious effects of the *APOE* $\epsilon 4$ allele are manifested prior to adulthood, especially in females, and support both early origin theories and differential life-course vulnerabilities for later cognitive impairment.

Keywords

early origins; *APOE*; cognitive development

1. Introduction

Although the origins of late-life cognitive health may begin at conception (Barnett et al., 2013), the extent to which cognitive dysfunction in adulthood is presaged during early life is currently unknown. Individual differences in general cognitive ability occur early in development and are stable longitudinally (Plomin et al., 1988; Walhovd et al., 2016);

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Data Statement

Analysis code and outputs are provided at the Open Science Framework platform (https://osf.io/fn9gk/?view_only=f50599adacfc4abdbfda47aa926387eb). We will make available all data used in this manuscript, except where participant directives do not permit us to do so. Data will be made available publicly after the current CATSLife testing is completed and, in the interim, can be re-analyzed in a collaborative effort with our group by qualified investigators.

Disclosure Statement

Declarations of interest: none

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nevertheless, the effects of early life factors may accumulate over the lifespan and impact subsequent cognitive aging (Liu et al., 2010). In order to assess the saliency of some early life genetic factors, we evaluated the associations of *APOE* genotypes with cognitive development from childhood to adolescence in the Colorado Adoption/Twin Study of Lifespan Behavioral Development (CATSLife).

The *APOE* gene encodes the brain's primary cholesterol transporter, apolipoprotein E, which may play additional roles in synaptic plasticity and cell signaling (Holtzman et al., 2012). There are three common *APOE* alleles, $\epsilon 2$, $\epsilon 3$ and $\epsilon 4$ that vary in allele frequency with mean values across populations at about 6.4% (SD = 5.1), 78.3% (SD = 12.1), and 14.5% (SD = 8.5) (Eisenberg et al., 2010). The *APOE* $\epsilon 4$ variant is implicated in A β deposition and tau-pathologies (Shi et al., 2017), respective features in brain plaques and tangles. The *APOE* $\epsilon 4$ allele is an established risk factor for late-onset Alzheimer's Disease (AD) (Liu et al., 2013), with a dose-dependent association between the number of $\epsilon 4$ alleles and AD risk and age of onset (Liu et al., 2013). *APOE* $\epsilon 4$ is also implicated in non-pathological cognitive aging, including general cognitive ability, episodic and working memory, verbal, spatial, and perceptual speed abilities (Davies et al., 2014; Reynolds et al., 2006; Schiepers et al., 2012). The least common *APOE* $\epsilon 2$ allele has been associated with reduced risk of AD and is thought to be possibly neuroprotective (Conejero-Goldberg et al., 2014; Liu et al., 2013). Though of keen interest, the role of *APOE* on cognition in childhood and adolescence is inconclusive (Chang et al., 2016; Ihle et al., 2012; Khan et al., 2014; Weissberger et al., 2018). However, it has been suggested that structural brain differences in $\epsilon 4$ carriers' volumes may appear in infancy, including lower hippocampal, frontal and temporal lobe volumes, as well as gray and white matter (Dean et al., 2014; Knickmeyer et al., 2013). A recent large cross-sectional imaging and neuropsychological study of 1187 children and youth, aged 3 to 20 years, suggested a number of differences in brain volume, fractional anisotropy, or thinning by *APOE* genotypes as well as cognitive ability performance (Chang et al., 2016). For example, smaller hippocampal volumes among $\epsilon 4/\epsilon 4$ individuals were associated with poorer performance on attention and working memory tasks (Chang et al., 2016). A recent comparison of mice from lines generated to overexpress 1N4R human tau via a P301S mutation and to either express human ApoE ($\epsilon 2$, $\epsilon 3$, or $\epsilon 4$ knock-ins) or not to express ApoE (knockouts) suggested that tau-related atrophy was observed in those expressing ApoE $\epsilon 4$ at 9 months of age, approximately equivalent to human middle adulthood (Shi et al., 2017). Indeed, the P301S/ $\epsilon 4$ mice showed greater tau hyperphosphorylation in the hippocampus at 3 months of age, approximately equivalent to human early adulthood, and a bigger loss of neurons in the hippocampal CA1 region at 9 months of age (Shi et al., 2017). Taken together, results of recent studies are consistent with the proposition that early life factors may be associated with later cognitive health, such that the *APOE* $\epsilon 4$ variant is implicated in cognitive health starting in early life.

Whether differential vulnerability is characteristic of all individuals who carry one or two copies of the *APOE* $\epsilon 4$ allele, or whether women are more vulnerable is of interest beyond differential mortality explanations. A recent meta-analysis (Neu et al., 2017) suggests that women with the $\epsilon 3/\epsilon 4$ genotype may be at greater risk than men for developing Mild Cognitive Impairment (MCI) between ages 55 and 70 years and at greater risk of Alzheimer's disease between 65-75 years but are not at differential risk at later ages.

Whether sex differences in the association between *APOE* ϵ 4 and cognitive profiles appear earlier in the lifespan is unclear, particularly in childhood and adolescence, although studies of childhood IQ suggest possibly differential associations for females than males on IQ performance (Calderon-Garciduenas et al., 2016; Taylor et al., 2011).

While the extant literature and early origins theories (Barnett et al., 2013; Liu et al., 2010) of adult-onset disease risk (Barker and Martyn, 1992) suggest that *APOE*-cognition associations might emerge with development, there is little longitudinal data to resolve this question (Ihle et al., 2012). Most childhood studies have been cross-sectional (e.g., Calderon-Garciduenas et al., 2016; Chang et al., 2016; Ihle et al., 2012; Khan et al., 2014) or limited in assessments (i.e., two waves (Taylor et al., 2011)). Indeed, to our knowledge, no childhood studies considering *APOE* and general cognitive ability or IQ have evaluated more than a single occasion, despite knowledge that multiple assessment affords greater reliability to evaluate and observe a possible relationship. We examined associations of *APOE* genotypes with cognitive performance, evaluating whether *APOE* ϵ 4 conferred a disadvantage for IQ performance evaluated at three assessments between childhood and adolescence. We also explore moderation of *APOE* ϵ 4 by sex in associations with IQ performance.

2. Materials and Methods

2.1 Participants

The CATSLife project includes participants from the Colorado Adoption Project (CAP; (Rhea et al., 2013a)), which was initiated in 1975 and enrolled 245 adoptive and 245 control families, and the Longitudinal Twin Study (LTS; (Rhea et al., 2013b)), which was initiated in 1985 and enrolled 483 twin pairs (265 MZ, 218 DZ). These studies used nearly parallel assessments between infancy and adolescence conducted on a nearly annual basis, and with periodic assessments into adulthood. Age descriptives by sample are reported in Table 1. The current analysis sample of 1321 includes individuals (nested within 773 families) who ranged between the ages of 6.50 years at the year 7 assessment and up to 17.99 years of age at the year 16 assessment. Genotyping was conducted on archival DNA samples and samples obtained from the ongoing CATSLife study ($N_{CAP} = 472$, 48.7% male; 44.92% adoptive families; $N_{LTS} = 849$, 48.8% male, 54.42% MZ twins). Self-reported race and ethnicity of the sample is 92.13% white and 7.87% non-white (American Indian/Alaska Native, Asian, Native Hawaiian/Pacific Islander, Black/African-American, More than one race, Unknown/Not reported), with 5.90% of the analysis sample identifying as Hispanic. These numbers are representative of the Colorado Front Range ethnic composition at the time participants were recruited into the foundational CAP and LTS studies.

The protocol was approved by the two respective institutional review boards of the authors' institutions, the University of Colorado at Boulder and the University of California at Riverside.

2.2 Measures

The year 7 and 12 cognitive assessments included WISC IQ measures (WISC-R or WISC-III; Wechsler, 1974, 1991) and the age 16 and CATSLife assessments included WAIS IQ measures (WAIS-R or WAIS-III; Wechsler, 1981, 1997) (see Table 2). The WISC-R and WISC-III tests are similar in item content and subtest composition (12/13 subtests overlap) with equivalent covariance structure among the subtests across the two versions (Dixon and Anderson, 1995). The WAIS-R and WAIS-III tests are likewise very similar, with respective correlations of Verbal IQ (VIQ), Performance IQ (PIQ) and Full Scale IQ scores exceeding .80 across versions (Strauss et al., 2006; Tulskey et al., 1997). Last, IQ scores calculated from child versus adult batteries are comparable and strongly correlated, e.g., WISC-III versus WAIS-III scores among 16-year old's are correlated .78 to .88 (Strauss et al., 2006). IQ scores are scaled by age-based norms, and the general expectation would be, therefore, that age effects ought to be minimal; given the shift in the test batteries associated with age and cohort, however, we included study and age covariates in analyses as described further below.

2.2.1 APOE Genotyping—Taqman assays of *APOE* SNPS, rs7412 and rs429358, were performed using buccal cell derived DNA. The success rate was 97%. *APOE* genotypes were formed from the two SNPs according to common practice (see Table 3). Where MZ twins were un-typed (N=4), they were assigned their cotwin's *APOE* genotype. Where one of the two *APOE* SNP assays failed we took the MZ cotwins genotype for that SNP (N=17). Hardy Weinberg Equilibrium (HWE) was achieved for both SNPs in both samples, selecting one sibling in each twin pair or sibship (all $p > .088$). In addition, the *APOE* genotypes formed from both SNPs achieved HWE in each sample and when combined across the two samples (all $p > .154$). Supplemental Table S1 presents the distribution of independent *APOE* alleles in a selection of one sibling or twin member from each pair/sibship. Overall, the $\epsilon 4$ allele is more frequent (13.95%) than the $\epsilon 2$ allele (7.89%), as expected (see Supplemental Table S1).

2.3 Analyses

Multi-level regression analyses of WISC and WAIS IQ measures were carried out using SAS Proc Mixed 9.4 (SAS Inc, Cary NC), using full maximum likelihood estimation. In Model I, the main effects model, IQ scores were predicted by the number of *APOE* $\epsilon 4$ alleles (0, 1 or 2), adjusting for the number of *APOE* $\epsilon 2$ alleles (0, 1 or 2) given its possible neuroprotective effect. The reference *APOE* genotype was therefore $\epsilon 3/\epsilon 3$. Additional covariates included study, sex, adopted status, and age, with coding as follows. Study was effects-coded as Colorado Adoption Project (CAP = -.5), Longitudinal Twin Study (LTS = .5). Sex was dummy coded and reflected effects for females (Male = 0, Female = 1). Adopted was dummy coded and reflects adoption status (0 = Not, 1 = Adopted). Age was centered at 16 years and reflected any possible age differences within or across longitudinal assessments, but also possible differences due to IQ battery as described above. Thus, in Model I, the fixed effect intercept reflects the expected mean IQ score for males, who were not adopted, at age 16 years, for those who are *APOE* $\epsilon 3/\epsilon 3$, controlling for study. In Model II, interaction terms with sex and *APOE* $\epsilon 4$ alleles were entered to evaluate possible sex moderation of *APOE* $\epsilon 4$ on IQ, adjusting for the interaction of sex and number of $\epsilon 2$ alleles.

Hence in Model II, the *APOE* $\epsilon 4$ effect reflects the $\epsilon 4$ effect for males and the *APOE* $\epsilon 4 * \text{Sex}$ interaction reflects the *APOE* $\epsilon 4$ effect for females.

Additional sensitivity tests included fitting the described Models I and II to IQ data at assessment year 7 alone when the WISC-R battery was given to both CAP and LTS participants. Additional sensitivity analyses of the longitudinal IQ results included covariate adjustments by race and ethnicity, and a sex by study interaction. Race and ethnicity were coded as white (1 = white, 0 = non-white) and Hispanic (1 = Hispanic, 0 = non-Hispanic). Last, we explored possible non-additive effects given hints of non-additivity in the mean IQ patterns; this was done for Model I given sex by *APOE* genotype frequencies would be limited for the rarer genotypes. We recoded $\epsilon 2$ and $\epsilon 4$ additive terms where 0 alleles = -1, 1 allele = 0, and 2 alleles = 1 (Model I.a). We coded dominance effects for $\epsilon 2$ alleles (Model I.b) and $\epsilon 4$ alleles (Model I.c) as follows: 0 alleles = 0, 1 allele = 1 and 2 alleles = 0. Models I.b and I.c respectively allow for heterozygotes to deviate non-additively from respective $\epsilon 2$ and $\epsilon 4$ homozygotes and both differentiate $\epsilon 24$ from $\epsilon 34$.

All analyses adjusted for nesting of individuals within family type, to account for dependencies in the data that could affect the standard errors of the regression parameters of the fixed effect predictors and covariates. Specifically, random effect variances were estimated for the intercept, decomposed into within-sibling and between-sibling variance, as well as residual variance, each by family/sibling type: i.e., siblings were from adoptive (A) or non-adoptive control (C) families, or were dizygotic twins (DZ) or monozygotic twins (MZ). Thus, the analysis accounted for differential dependencies between MZ twins, who share identical genotypes, from DZ twins who share on average 50% of segregating genes in common, as well as for differential dependencies between genetically-unrelated siblings and genetically-related siblings. The between sibling intercept variance represents similarity among siblings in IQ performance that is systematic across longitudinal assessments of IQ and vary by sibling type ($\sigma_{betweenA}^2, \sigma_{betweenC}^2, \sigma_{betweenDZ}^2, \sigma_{betweenMZ}^2$) (see Tables S2, S4, S6). The corresponding within-sibling intercept variance represents differences in IQ among siblings that are systematic across longitudinal assessments of IQ and vary by sibling type ($\sigma_{withinA}^2, \sigma_{withinC}^2, \sigma_{withinDZ}^2, \sigma_{withinMZ}^2$). Last, the residual variance represents within-person variability in performance that is occasion-specific ($\sigma_{residualA}^2, \sigma_{residualC}^2, \sigma_{residualDZ}^2, \sigma_{residualMZ}^2$). No constraints were placed on the magnitudes of random effects estimated by sibling types. In sensitivity analyses of the IQ data at the year 7 assessment, the decomposition of the random effects of the intercept included between pair random effects ($\sigma_{betweenA}^2, \sigma_{betweenC}^2, \sigma_{betweenDZ}^2, \sigma_{betweenMZ}^2$) and residual within pair random effects of IQ (denoted $\sigma_{withinA}^2, \sigma_{withinC}^2, \sigma_{withinDZ}^2, \sigma_{withinMZ}^2$); in these models age was centered at age 7.

Regression parameter tests of significance were as reported via SAS Proc MIXED 9.4 (SAS Inc, Cary NC), which included asymptotically distributed t -statistics formed by taking the parameter estimate over its standard error, with degrees of freedom estimated using the between-within option. For Model I, 1-tailed tests of significance are reported for *APOE* $\epsilon 4$ effects given our hypothesized direction of effect while Model II reported 2-tailed tests of

significance for *APOE* $\epsilon 4$ by sex interactions. We report 95% confidence intervals for all parameters for both Models I and II.

The sample size is appropriate for tests of association. Our expected power for an overall main effect association exceeds .94 assuming an effect size contribution of 1% to the outcome, a sibling correlation of .40, and an $\epsilon 4$ frequency of 15% (Purcell et al., 2003). Under the same conditions, an expected power of .80 is achieved with an effect size contribution of 0.65%. Moreover, the multiple longitudinal assessments of IQ provide increased reliability to evaluate and observe a possible relationship.

3. Results

Unadjusted descriptives of IQ performance at each assessment by *APOE* $\epsilon 4$ status are suggestive of differential performance by *APOE* $\epsilon 4$ (see Table 4), with reduced performance particularly for those carrying one *APOE* $\epsilon 4$ allele in the total sample. When considering sex, a pattern of lower mean IQ performance per *APOE* $\epsilon 4$ allele is observable in females but not for males, though notably the sample sizes for *APOE* $\epsilon 4/\epsilon 4$ are small.

3.1 Multilevel Regression Models I and II

APOE $\epsilon 4$ associations with longitudinal IQ performance at the year 7, 12 and 16 assessments were evaluated via multilevel regression models fitted to longitudinal IQ scores ($N = 1321$, 48.86% male; age range 6.50 to 17.99 years), adjusting for sex, age, adopted status, study sample, and the number of $\epsilon 2$ alleles. Parameters from the full multi-level main effects model fitted with covariates (Model I) are presented in Tables 5 (fixed effects) and S2 (random effects). With respect to prediction by *APOE* genotype, 1-tailed tests were selected for $\epsilon 4$ under Model I given the hypothesized directionality. Results suggested that for each $\epsilon 4$ allele, Full scale IQ scores were lower by 1.91 points compared to $\epsilon 33$ homozygotes ($p = 0.0051/2 = 0.0026$, 1-tailed); this corresponds to an estimated d effect size of $-.13$ using the expected SD for IQ scores of 15 (i.e., $1.91/15$). Consistent effects were observed for Verbal ($B = -1.60$, $p = 0.0224/2 = 0.0112$ 1-tailed, $d = -.11$) and Performance IQ ($B = -1.78$, $p = 0.0118/2 = 0.0059$ 1-tailed, $d = -.12$). False Discovery Rate (FDR) adjusted 1-tailed p -values remain significant: Full scale IQ, $p = 0.0077$; Verbal IQ, $p = 0.0112$; and Performance IQ, $p = 0.0089$.

Figure 1 presents expected mean Full Scale IQ scores as a function of the number of *APOE* $\epsilon 4$ alleles in the total sample.

Next, to evaluate sex differences with respect to *APOE* effects, an interaction term with sex and *APOE* $\epsilon 4$ alleles was entered and evaluated under Model II using a 2-tailed test (see Table 5, Model II), adjusting for the interaction of sex and $\epsilon 2$ alleles. For Verbal IQ, the $\epsilon 4$ by sex interaction was significant ($p = .0324$ 2-tailed), suggesting the $\epsilon 4$ effect may be larger in females than males, while the main effect of $\epsilon 4$, now reflecting male performance, was non-significant. Specifically, the $\epsilon 4$ effect was $-.23$ ($se = .95$; $d = -.02$) in males and -2.95 ($se = 1.38$; $d = -.20$) in females. The $\epsilon 4$ effects for Performance IQ and Full Scale IQ may be especially pronounced for females ($\epsilon 4 \times \text{Sex } p = 0.0132$ 2-tailed), while the main effect of $\epsilon 4$, reflecting differential male performance, was non-significant. For Full Scale IQ, the

$\epsilon 4$ effect was -0.33 ($se = .92$; $d = -.02$) in males and -3.41 ($se = 1.35$; $d = -.23$) in females (see Figure 1). For Performance IQ, the $\epsilon 4$ effect was -0.13 ($se = .97$; $d = -.01$) in males and -3.48 ($se = 1.40$; $d = -.23$) in females. FDR adjusted 2-tailed p -values for the $\epsilon 4 \times$ Sex effects remain significant: Full scale IQ, $p = 0.0198$; Verbal IQ, $p = .0324$, and Performance IQ, $p = 0.0198$.

Figure 1 represents expected performance differences for Full Scale IQ by the number of *APOE* $\epsilon 4$ alleles in males and females.

3.2 Sensitivity Analyses

We performed a sensitivity test to evaluate whether *APOE* effects are determinable at the youngest assessment at year 7 alone ($N = 1176-1177$; see Supplementary Tables S3-S4); a further benefit was that the WISC-R battery was implemented for both study samples. We observed significant main effects of *APOE* for Performance and Full Scale IQ (FDR adjusted 1-tailed $p = .0156$ for both) with consistent, but non-significant, effects for Verbal IQ (FDR adjusted 1-tailed $p = .0749$) (see Table S3, Model I). While the models including sex interactions did not achieve significance (all $p > .0973$; see Table S3, Model II), the pattern of effect sizes was consistent to that observed in the models fitted to longitudinal IQ above.

We also performed a follow-up sensitivity analysis of the longitudinal IQ performance models, with the inclusion of additional covariates, i.e., a sex by study interaction and self-reported race and ethnicity. Adding these covariates did not alter conclusions for longitudinal IQ across assessment years 7, 12 and 16 as detailed above with all tests remaining significant at adjusted FDR p -values: Model I, 1-tailed $p = .0064$ to $p = .0041$; and Model II, 2-tailed $p = .0397$ to $p = .0264$ (see Supplementary Tables S5-S6).

3.2.1 Tests of non-additivity—We explored the possibility of non-additive effects in Model I, given hints of non-additivity in the mean IQ patterns by *APOE* $\epsilon 4$ alleles. We fitted Model I.a, with all covariates, entering recoded additive effects for $\epsilon 2$ and $\epsilon 4$ alleles (see Supplementary Tables S7, Model I.a). Subsequently, we added the dominance effect for $\epsilon 2$ alleles. Results suggested that the $\epsilon 2$ dominance effect was negative and significant for Verbal, Performance and Full Scale IQ ($p = .0003$, $.0247$, and $.0006$, respectively, 2-tailed), suggesting reduced performance for $\epsilon 2$ heterozygotes ($\epsilon 2/\epsilon 3$, $\epsilon 2/\epsilon 4$) than expected from an additive model (see Supplementary Tables S6, Model I.b). FDR adjusted 2-tailed p -values for the $\epsilon 2$ dominance effects remained significant: $p = 0.0247$ to 0.0009 . Moreover, the additive effect for $\epsilon 2$ became significant and positive for Verbal ($p = .0090$) and Full Scale IQ ($p = .0159$), with the same direction for Performance IQ ($p = .1015$), suggesting a benefit in IQ performance for $\epsilon 2/\epsilon 2$ homozygotes. FDR adjusted 2-tailed p -values remained significant for Verbal and Full Scale IQ, both $p = 0.0239$. Last, we added the $\epsilon 4$ dominance effect, which allows $\epsilon 4$ heterozygotes ($\epsilon 2/\epsilon 4$, $\epsilon 3/\epsilon 4$) to deviate from that expected under an additive model. Results suggested that adding the $\epsilon 4$ dominance effect was not significant (all $p > 0.2917$) (see Supplementary Table S7, Model I.c). The more parsimonious Model I.b best represents the contributions of additive ($\epsilon 2$ and $\epsilon 4$) and nonadditive ($\epsilon 2$) influences. Supplementary figure S1 displays Full scale IQ by *APOE* genotypes, based on Models I.a

and Model I.b, as well as covariate-adjusted least-squares means. See supplement for further details.

4. Discussion

APOE genotypes may be associated with general cognitive performance earlier than midlife. We observed small detriments in IQ performance across childhood and adolescence with nearly a 2-point decrement for each $\epsilon 4$ allele compared to those with *APOE* $\epsilon 3/\epsilon 3$ genotypes. Performance and Verbal IQ showed consistent effects, with the smaller effect sizes for Verbal IQ. Hence, *APOE* $\epsilon 4$ genotypes may be associated with lower IQ as early as childhood. Moreover, *APOE* may show stronger (or earlier) effects on IQ in females than males, particularly for Performance and Full Scale IQ.

The role of *APOE* on cognition between childhood and early adulthood is inconclusive, with most childhood studies cross-sectional in design (Calderon-Garciduenas et al., 2016; Chang et al., 2016; Ihle et al., 2012; Khan et al., 2014) and limited in assessments of IQ (Taylor et al., 2011). A recent large cross-sectional imaging and neuropsychological study of individuals aged 3 to 20 years (Chang et al., 2016) suggested potential differences in brain and cognitive development for those with particular *APOE* genotypes. For example, those with *APOE* $\epsilon 2/\epsilon 4$ evidenced smaller hippocampal volumes, and those with $\epsilon 4/\epsilon 4$ evidenced lower hippocampal fractional anisotropy at age 8 and younger (Chang et al., 2016). In addition, differential executive functioning and working memory performance were reported for *APOE* $\epsilon 4/\epsilon 4$ and attentional processing for *APOE* $\epsilon 2/\epsilon 4$ (Chang et al., 2016); the authors did not test associations with broader constructs, however. The Lothian Birth Cohort included a single childhood cognitive ability assessment with follow-ups in late adulthood at age 70 and beyond; they report a nonsignificant negative *APOE* $\epsilon 4$ effect for their measure of general verbal cognitive ability and reasoning assessed at age 11 (Luciano et al., 2009). In the current study, the weakest effect we observed was for Verbal IQ with larger *APOE* $\epsilon 4$ effects for Performance and Full Scale IQ. Notably absent from a recent large-scale GWAS of intelligence is implication of the *APOE* region (Savage et al., 2017); only about 5.7% of the nearly 280,000 samples included in this GWAS were from children or young adults. In this, and in similar GWAS, age and sex tend to be treated as covariates, and not leveraged per se to evaluate whether particular variants or gene sets are associated at certain age periods. Cautions in relying on imputations of *APOE* genotypes from GWAS have also been noted, although in recent years imputation has become more reliable (Radmanesh et al., 2014; Roses et al., 2016).

Opportunities to evaluate earlier life cognitive functioning are uncommon in studies of cognitive aging, yet cognitive development across childhood and adolescence may represent a salient period when cognitive reserves are being formed. Cognitive reserve theories suggest that individuals may differ in their capacity to withstand aging-related pathologies due to cognitive processing optimizations that boost functioning, and may allow individuals to weather aging and disease-related neural changes (Barulli and Stern, 2013). Likewise, early origin theories of late-life cognitive health (Barnett et al., 2013; Liu et al., 2010), stimulated in part by the “Barker hypothesis” of prenatal and early-life developmental determinants of adult-onset disease risk, have led to interests in modifiable features and life

course mediators and moderators of cognitive aging and dementia. Our results suggest that cognitive differences associated with *APOE* may emerge early and become magnified later in the life course, and if so, childhood represents a key period of intervention to invest in and boost reserves.

APOE $\epsilon 4$ effects may be larger in females than in males, particularly for Full Scale IQ and Performance IQ. Recent cross-sectional work described in children ages 3-20 years (Chang et al., 2016) treated sex as a covariate but did not evaluate sex as a moderator of observed *APOE* effects. A report of 5995 British 8-year-old children from the ALSPAC study failed to find *APOE* associations with cross-sectional WISC Verbal, Performance or Full Scale IQ (Taylor et al., 2011). However, trends for sex-stratified effects were observed where females with rare *APOE* genotypes, $\epsilon 2/\epsilon 2$ and $\epsilon 4/\epsilon 4$, showed higher average scores than $\epsilon 3/\epsilon 3$ females, while those with $\epsilon 2/\epsilon 4$ and $\epsilon 3/\epsilon 4$ genotypes showed worse scores than $\epsilon 3/\epsilon 3$ females for Verbal ($p = .03$) and Total IQ ($p = .02$). Moreover, a recent cross-sectional study evaluating 105 12-year old children ($SD = 5.4$ years) living in Mexico City reported findings of an increased vulnerability to poorer performance in female carriers of *APOE* $\epsilon 4$ on Total and Performance IQ, consistent with our report (Calderon-Garciduenas et al., 2016). Recent literature suggests that differential risks for MCI and AD in females may appear in proscribed age periods and may not extend over the lifespan (Neu et al., 2017). Our findings suggest that such windows may extend to earlier stages of development, particularly for reasoning traits represented by performance IQ, before any notable cognitive differences raise clinical concerns. It will be important to track whether such early differences may become amplified perhaps due to differential cognitive reserves (Pettigrew et al., 2013; Runge et al., 2014) although early life cognitive differences associated with *APOE* have yet to be directly linked to later cognitive reserves.

Explorations of non-additive effects suggest a possible advantage for *APOE* $\epsilon 2/\epsilon 2$ individuals, followed by $\epsilon 3/\epsilon 3$, and the lowest performance among those who carry one or more $\epsilon 4$ alleles. While the observed patterns by genotype are consistent with patterns of dementia risk (Farrer et al., 1997; Neu et al., 2017), and congruent with possible cognitive advantages and detriments observed in child samples (Chang et al., 2016) (c.f., females, Taylor et al., 2011), we note the $\epsilon 2/\epsilon 2$ and $\epsilon 4/\epsilon 4$ genotypes are relatively rare and necessitates some caution in interpretation of their means relative to the other more common genotypes. Further examination of differential benefits and vulnerabilities of *APOE* genotypes, and sex moderation in larger samples, where it can be fully interrogated, is warranted.

The point estimates and d -effect size estimates we observed indicate that the effects on IQ are small per *APOE* $\epsilon 4$ allele, up to a few points. In terms of clinical relevance, a difference of a few points, while small, may be potentially relevant in terms of cognitive reserve for which any disadvantages could become magnified later in adulthood (c.f., Barulli and Stern, 2013; Liu et al., 2010). This bears further study. Moreover, childhood IQ is predictive of biological age as well as the pace of aging at midlife (Belsky et al., 2017; Schaefer et al., 2016), with lower childhood IQ associated with increased biological aging, as well as increased cardiovascular disease (CVD) risks before age 65 (Hart et al., 2004) even after adjusting for covariates, suggesting that there may be physical health pathways to consider.

With respect to *APOE*, this would include cholesterol-lipid pathways. Finally, for further perspective, the small differences in IQ that we observed are similar to the effect sizes noted in a recent meta-analysis of over 600,000 individuals from 42 studies across the lifespan suggesting that each additional year of achieved education may be causally associated with a 1 to 5 IQ point gain (Ritchie and Tucker-Drob, 2018). Echoing Ritchie and Tucker-Drob (2018), here in the context of *APOE* ϵ 4, a few points may not be consequential for any one individual, but could be meaningful in terms of early interventions to allay accelerated cognitive aging or dysfunction in a population sense.

Cognitive decline is a devastating and feared aspect of aging, yet its developmental origins have become a focus only recently. Cognitive development during childhood and adolescence may contribute to the formation of crucial cognitive reserves that may hold a unique saliency to later cognitive functioning. Understanding the emergence and phenomena of differential cognitive growth in early life and differential maintenance in functioning in adulthood is critical to evaluating the promise of interventions. Additional longitudinal studies are warranted to consider early origins of cognitive health, and the possibility of developmental effects that emerge in this age span. To this end, we are in the process of collecting additional Full Scale IQ and specific cognitive abilities data on the entire CATSLife sample between 28–46 years of age with an expected completion in 2020.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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Highlights

- *APOE* ε4 associated with lower intelligence (IQ) scores from childhood to adolescence
- Females may be at greater risk to show these early life effects
- Effects of the ε4 allele in childhood support early origin theories
- Distinctive age-related profiles may be important to later cognitive impairment

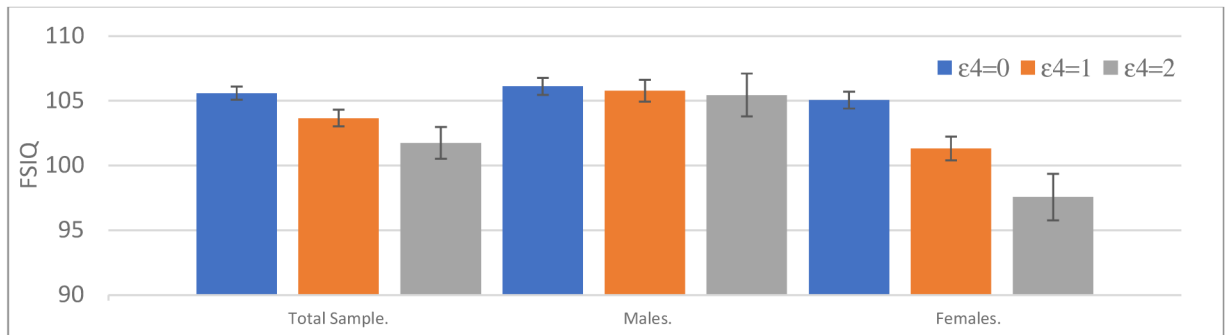


Figure 1. Mixed Model estimates: *APOE* $\epsilon 4$ effects on Full Scale IQ based on child and adolescent assessments (years 7, 12, and 16).

Note. Analyses adjusted for nesting of individuals within family type, number of *APOE* $\epsilon 2$ alleles, sex, age, adoption status, and study (CAP or LTS). *APOE* $\epsilon 3/\epsilon 3$ was the referent group. The ‘Total Sample’ estimates were averaged across sex (i.e., sex = .5), with all other effects at centered values. Standard errors shown.

Table 1.

Analysis sample age descriptives by study.

| Assessment | CAP N=472 | M_{AGE} | SD_{AGE} | Min | Max | LTS N=849 | M_{AGE} | SD_{AGE} | Min | Max |
|-------------------|----------------------|-----------------------------|------------------------------|------------|------------|----------------------|-----------------------------|------------------------------|------------|------------|
| Year 7 | 397 | 7.42 | 0.37 | 6.50 | 8.42 | 780 | 7.43 | 0.36 | 6.67 | 8.50 |
| Year 12 | 403 | 12.45 | 0.41 | 11.67 | 14.17 | 738 | 12.43 | 0.37 | 11.33 | 14.00 |
| Year 16 | 451 | 16.34 | 0.48 | 16.00 | 17.99 | 724 | 16.37 | 0.41 | 16.00 | 17.92 |

Note. CAP = Colorado Adoption Project; LTS = Longitudinal Twin Study.

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Table 2.

IQ Assessments by study.

| Assessment | Year 7 | Year 12 | Year 16 |
|------------|---------|---------|---------|
| WISC - R | CAP/LTS | CAP | |
| WISC - III | | LTS | |
| WAIS - R | | | CAP |
| WAIS - III | | | LTS |

Note. CAP = Colorado Adoption Project; LTS = Longitudinal Twin Study.

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Table 3.*APOE* genotype frequencies in analysis sample by study.

| <i>APOE</i> | rs429358 | rs7412 | <i>CAP</i> | | <i>LTS</i> | |
|------------------|----------|--------|------------|----------------|------------|----------------|
| | | | <i>N</i> | <i>Percent</i> | <i>N</i> | <i>Percent</i> |
| e22 | T/T | T/T | 6 | 1.27 | 9 | 1.06 |
| e23 | T/T | C/T | 50 | 10.59 | 104 | 12.25 |
| e24 | C/T | C/T | 13 | 2.75 | 16 | 1.88 |
| e33 | T/T | C/C | 283 | 59.96 | 524 | 61.72 |
| e34 | C/T | C/C | 110 | 23.31 | 189 | 22.26 |
| e44 | C/C | C/C | 10 | 2.12 | 7 | 0.82 |
| Current N | 1321 | | 472 | | 849 | |

Note. CAP = Colorado Adoption Project; LTS = Longitudinal Twin Study.

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Table 4.Age and IQ test performance by *APOE* ϵ 4 alleles across assessments at Years 7, 12, and 16.

| Total Sample | Variable | <i>APOE</i> ϵ 4=0 | | | <i>APOE</i> ϵ 4=1 | | | <i>APOE</i> ϵ 4= 2 | | |
|----------------|-----------------|----------------------------|-------------|-----------|----------------------------|-------------|-----------|-----------------------------|-------------|-----------|
| | | N | Mean | SD | N | Mean | SD | N | Mean | SD |
| Year 7 | Verbal IQ | 865 | 106.31 | 14.32 | 296 | 105.72 | 14.65 | 15 | 107.40 | 10.44 |
| | Performance IQ | 866 | 110.78 | 12.54 | 296 | 108.68 | 13.62 | 15 | 110.53 | 10.33 |
| | Full Scale IQ | 865 | 109.24 | 12.99 | 296 | 107.77 | 13.50 | 15 | 109.60 | 7.87 |
| | Age | 871 | 7.42 | 0.37 | 296 | 7.45 | 0.36 | 15 | 7.38 | 0.40 |
| Year 12 | Verbal IQ | 839 | 105.89 | 13.19 | 288 | 105.58 | 12.76 | 14 | 108.07 | 14.07 |
| | Performance IQ | 839 | 105.85 | 13.75 | 287 | 104.36 | 13.59 | 13 | 109.62 | 8.44 |
| | Full Scale IQ | 839 | 106.38 | 12.95 | 287 | 105.40 | 12.74 | 13 | 109.08 | 10.87 |
| | Age | 841 | 12.44 | 0.39 | 288 | 12.45 | 0.36 | 14 | 12.40 | 0.42 |
| Year 16 | Verbal IQ | 873 | 103.70 | 12.58 | 288 | 102.94 | 11.58 | 14 | 103.43 | 10.91 |
| | Performance IQ | 873 | 104.25 | 12.44 | 288 | 103.22 | 11.88 | 14 | 105.86 | 14.20 |
| | Full Scale IQ | 873 | 104.17 | 11.75 | 288 | 103.20 | 10.63 | 14 | 104.71 | 10.31 |
| | Age | 873 | 16.36 | 0.44 | 289 | 16.36 | 0.43 | 14 | 16.55 | 0.41 |
| Males | Variable | N | Mean | SD | N | Mean | SD | N | Mean | SD |
| Year 7 | Verbal IQ | 430 | 105.81 | 14.49 | 146 | 107.03 | 14.17 | 8 | 106.88 | 11.31 |
| | Performance IQ | 430 | 111.14 | 12.72 | 146 | 110.82 | 13.08 | 8 | 115.38 | 10.93 |
| | Full Scale IQ | 430 | 109.13 | 13.29 | 146 | 109.72 | 13.11 | 8 | 111.88 | 7.88 |
| | Age | 432 | 7.47 | 0.39 | 146 | 7.51 | 0.36 | 8 | 7.54 | 0.40 |
| Year 12 | Verbal IQ | 407 | 106.61 | 13.36 | 142 | 107.55 | 13.41 | 7 | 114.57 | 11.31 |
| | Performance IQ | 407 | 104.89 | 13.80 | 142 | 106.12 | 13.49 | 7 | 111.71 | 6.32 |
| | Full Scale IQ | 407 | 106.29 | 12.84 | 142 | 107.56 | 13.33 | 7 | 114.71 | 7.95 |
| | Age | 409 | 12.48 | 0.40 | 142 | 12.51 | 0.38 | 7 | 12.58 | 0.45 |
| Year 16 | Verbal IQ | 421 | 104.49 | 13.05 | 147 | 104.46 | 11.13 | 6 | 109.50 | 11.36 |
| | Performance IQ | 421 | 103.97 | 12.94 | 147 | 104.99 | 12.40 | 6 | 114.83 | 14.96 |
| | Full Scale IQ | 421 | 104.54 | 12.05 | 147 | 104.98 | 10.86 | 6 | 112.33 | 8.96 |
| | Age | 421 | 16.39 | 0.48 | 148 | 16.41 | 0.44 | 6 | 16.64 | 0.47 |
| Females | Variable | N | Mean | SD | N | Mean | SD | N | Mean | SD |
| Year 7 | Verbal IQ | 435 | 106.81 | 14.14 | 150 | 104.44 | 15.04 | 7 | 108.00 | 10.21 |
| | Performance IQ | 436 | 110.42 | 12.37 | 150 | 106.60 | 13.86 | 7 | 105.00 | 6.53 |
| | Full Scale IQ | 435 | 109.34 | 12.70 | 150 | 105.87 | 13.64 | 7 | 107.00 | 7.57 |
| | Age | 439 | 7.37 | 0.33 | 150 | 7.39 | 0.35 | 7 | 7.19 | 0.35 |
| Year 12 | Verbal IQ | 432 | 105.20 | 13.01 | 146 | 103.66 | 11.83 | 7 | 101.57 | 14.22 |
| | Performance IQ | 432 | 106.75 | 13.66 | 145 | 102.64 | 13.52 | 6 | 107.17 | 10.48 |
| | Full Scale IQ | 432 | 106.45 | 13.06 | 145 | 103.28 | 11.81 | 6 | 102.50 | 10.56 |
| | Age | 432 | 12.40 | 0.38 | 146 | 12.39 | 0.33 | 7 | 12.21 | 0.31 |
| Year 16 | Verbal IQ | 452 | 102.96 | 12.10 | 141 | 101.36 | 11.87 | 8 | 98.88 | 8.58 |
| | Performance IQ | 452 | 104.51 | 11.96 | 141 | 101.38 | 11.07 | 8 | 99.13 | 9.67 |
| | Full Scale IQ | 452 | 103.82 | 11.46 | 141 | 101.34 | 10.10 | 8 | 99.00 | 7.27 |
| | Age | 452 | 16.32 | 0.40 | 141 | 16.31 | 0.40 | 8 | 16.47 | 0.39 |

Table 5.Multilevel fixed effects: *APOE* ε4 effects on IQ across Year 7, Year 12 and Year 16 assessments.

| Verbal IQ Fixed Effects | Model I | | | | Model II | | | |
|---------------------------------|---------|---------------------|--------|--------|----------|---------------------|--------|--------|
| | B | se | LL | UL | B | se | LL | UL |
| Intercept | 106.42 | 0.64 ^c | 105.16 | 107.67 | 105.89 | 0.69 ^c | 104.54 | 107.23 |
| Study | -6.28 | 0.88 ^c | -8.01 | -4.55 | -6.21 | 0.88 ^c | -7.95 | -4.48 |
| Adopted | -3.64 | 0.98 ^c | -5.55 | -1.72 | -3.56 | 0.98 ^c | -5.47 | -1.64 |
| Sex | -2.36 | 0.71 ^c | -3.75 | -0.96 | -1.30 | 0.87 | -3.00 | 0.39 |
| Age | -0.36 | 0.03 ^c | -0.43 | -0.30 | -0.36 | 0.03 ^c | -0.43 | -0.30 |
| ε2 | -0.42 | 0.86 | -2.12 | 1.27 | 0.30 | 1.22 | -2.09 | 2.69 |
| ε4 | -1.60 | 0.70 ^{a,f} | -2.97 | -0.23 | -0.23 | 0.95 | -2.09 | 1.63 |
| Sex*ε2 | -- | -- | -- | -- | -1.37 | 1.60 | -4.50 | 1.77 |
| Sex*ε4 | -- | -- | -- | -- | -2.95 | 1.38 ^{a,t} | -5.65 | -0.25 |
| Performance IQ Fixed Effects | Model I | | | | Model II | | | |
| | B | se | LL | UL | B | se | LL | UL |
| Intercept | 105.76 | 0.61 ^c | 104.56 | 106.96 | 105.33 | 0.66 ^c | 104.02 | 106.63 |
| Study | -8.17 | 0.81 ^c | -9.76 | -6.58 | -8.11 | 0.81 ^c | -9.70 | -6.51 |
| Adopted | -1.56 | 1.06 | -3.63 | 0.51 | -1.47 | 1.06 | -3.55 | 0.60 |
| Sex | -1.16 | 0.70 | -2.53 | 0.22 | -0.30 | 0.86 | -1.99 | 1.38 |
| Age | -0.79 | 0.03 ^c | -0.86 | -0.72 | -0.79 | 0.03 ^c | -0.86 | -0.72 |
| ε2 | -0.17 | 0.84 | -1.82 | 1.48 | -0.64 | 1.24 | -3.06 | 1.79 |
| ε4 | -1.78 | 0.71 ^{a,f} | -3.16 | -0.39 | -0.13 | 0.97 | -2.03 | 1.77 |
| Sex*ε2 | -- | -- | -- | -- | 0.85 | 1.65 | -2.38 | 4.08 |
| Sex*ε4 | -- | -- | -- | -- | -3.48 | 1.40 ^{a,t} | -6.23 | -0.73 |
| Full Scale IQ Fixed Effects | Model I | | | | Model II | | | |
| | B | se | LL | UL | B | se | LL | UL |
| Intercept | 106.60 | 0.61 ^c | 105.41 | 107.80 | 106.11 | 0.65 ^c | 104.82 | 107.39 |
| Study | -7.81 | 0.83 ^c | -9.44 | -6.17 | -7.75 | 0.83 ^c | -9.38 | -6.11 |
| Adopted | -2.97 | 0.99 ^b | -4.92 | -1.02 | -2.90 | 0.99 ^b | -4.84 | -0.95 |
| Sex | -2.05 | 0.69 ^b | -3.41 | -0.70 | -1.06 | 0.84 | -2.71 | 0.59 |
| Age | -0.63 | 0.03 ^c | -0.69 | -0.58 | -0.63 | 0.03 ^c | -0.69 | -0.58 |
| ε2 | -0.45 | 0.83 | -2.09 | 1.18 | -0.35 | 1.19 | -2.67 | 1.98 |
| ε4 | -1.91 | 0.68 ^{b,f} | -3.25 | -0.58 | -0.33 | 0.92 | -2.14 | 1.48 |
| Sex*ε2 | -- | -- | -- | -- | -0.22 | 1.56 | -3.28 | 2.84 |
| Sex*ε4 | -- | -- | -- | -- | -3.41 | 1.35 ^{a,t} | -6.05 | -0.77 |

^a $p < .05$,

^b
 $p < .01,$

^c
 $p < .001;$

^f
 $p < .05$ 1-tailed, FDR corrected,

^t
 $p < .05$ 2-tailed, FDR corrected.

Note. $N = 1321$. Study (CAP = $-.5$, LTS = $.5$), Adopted (0=Not, 1=Adopted), Sex (Male = 0, Female = 1); Age was centered at 16 years; e2 = number of alleles (0,1,2); e4 = number of alleles (0,1,2). Model I refers to main effects models with *APOE* and Model II includes interaction effects with sex and *APOE*. LL, UL = lower and upper 95% confidence interval.

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