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Title: COMPARATIVE CIRCULATION DYNAMICS OF THE FIVE MAIN HIV TYPES IN CHINA

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ABSTRACT

The HIV epidemic in China accounts for 3% of the global HIV incidence. We compared the patterns and determinants of interprovincial spread of the five most prevalent circulating types. HIV *pol* sequences sampled across China were used to identify relevant transmission networks of the five most relevant HIV-1 types (B, CRF01_AE, CRF07_BC, CRF08_BC and CRF55_01B) in China. From these, the dispersal history across provinces was inferred. A generalized linear model (GLM) was used to test the association between migration rates among provinces and several measures of human mobility. A total of 10,707 sequences between 2004-2017 across 26 provinces were collected, among which 1,962 newly reported here. A mean of 18 (Min-Max:1-54) independent transmission networks involving up to 17 provinces were identified. Discrete phylogeographic analysis largely recapitulate the documented spread of the HIV types which, in turn, to large extent mirror within-China population migration flows. In line with the different spatio-temporal spread dynamics, the identified drivers thereof were also heterogeneous but are consistent with a central role of human mobility. The comparative analysis of the dispersal dynamics of the five main HIV types circulating in China suggests a key role of large populations centers and developed transportation infrastructures as hubs of HIV dispersal. This advocates for coordinated public health efforts in addition to local targeted interventions.

IMPORTANCE

While traditional epidemiological studies are of great interest in describing the dynamics of epidemics, they cannot fully capture the geospatial dynamics and factors driving the dispersal of pathogens such as HIV as they struggle to capture linkages between infections. To overcome this, we used a discrete phylogeographic approach coupled to a generalized linear model extension to characterize the dynamics and drivers of the across-province spread of the five main HIV types circulating in China. Our results indicate that large urbanized areas with dense population and developed transportation infrastructures are facilitators of HIV dispersal throughout China, and highlight the need to consider harmonized country-wide public policies to control local HIV epidemics.

INTRODUCTION

By the end of 2018, the number of people living with HIV (PWH) in China was close to 1.25 million (1, 2). The distribution of HIV-1 subtypes in China is diverse with over 11 circulating genetic variants (3), each with an evolving geographical distribution, prevalence and modes of transmission (3-6). The first nationwide molecular epidemiological survey in 1996-1998 showed that subtype B'/B (47.5%) and subtype C (34.3%) were the most predominant HIV types in China (7). For subtype B', this in part resulted from its high prevalence among plasma donors in China because of unsanitary commercial plasma collection (8). Surveys conducted in 2002-2003 and 2006 indicated that the circulating recombinant forms (CRF) CRF07_BC, CRF01_AE, and CRF08_BC had become the dominant HIV types in China. Founder effects make that CRF07_BC and CRF08_BC mostly circulated among injecting drug users (IDUs) in North-Eastern and South-Eastern China, respectively (9-12), while subtype B' remains dominant among former plasma donors in Central China (13, 14). Meanwhile, CRF01_AE became the dominant type and replaced subtype B as the principal driver of infection among men reporting having sex with men (MSM) (3). The National Sentinel Surveillance System of China revealed that the proportion of MSM transmission increased from 14.7% in 2009 (15) to 27.6% in 2016 (16) with an increased proportion of CRF01_AE and CRF07_BC infections among MSM, while the proportion of HIV-1 subtype B decreased between 2012 and 2016 (6, 17). In addition to these predominant HIV types, CRF55_01B, generated through recombination between CRF01_AE and subtype B variants, has been first identified among MSM in the city of Shenzhen (18, 19). Circulating primarily among MSM, it has now spread throughout most provinces of China with a prevalence ranging from 1.5% to 12.5% (20). Its prevalence has increased in the past five years, especially in South and East China with higher pooled estimated rate in Guangdong (12.22%, 95% CI 10.34–13.17) and Fujian (8.65%, 95% CI 4.98–13.17)(17). It is now circulating mostly in Guangdong and neighboring provinces in China, and across all risk groups (18).

The burden of HIV is also geographically unevenly spread: whereas HIV is present in all provinces, the top six high-prevalence provinces (Yunnan, Guangxi, Henan, Guangdong and Xinjiang) accounted for over 60% of the national number of PWH (21). The recent upsurge of HIV among MSM in large Chinese cities including Beijing, Chongqing, Chengdu, Guangzhou, Shanghai and Shenyang adds to this imbalance (22).

These multiple and diverse epidemics driven by changing risk factor patterns in part result from the inability of treatment, prevention, and control programs to halt the rapid growth of the HIV epidemics, which now account for ~3% of the global HIV prevalence (23). The epidemic growth of ~80,000 new infections per year (1) coincided with intense rural-to-urban migration flows (24-30)

and considerable investments in land and airway transport infrastructures expediting longer-distance human mobility (31). By the end of 2017, the migrant population, seeking better employment opportunities and living conditions in economically more developed areas, reached 244 million (32, 33), and migrant-workers have become the main driver of within-country migration. Importantly, the labor migrant population is also at higher risk for HIV acquisition and transmission because of poor knowledge about self-protection and the transmission routes of HIV (34), and they have been shown to fuel local epidemics (30, 35).

While traditional epidemiological studies are of great interest in describing the dynamics of epidemics, they struggle to fully capture the geospatial dynamics and factors driving the dispersal of pathogens. By merging virus genetic, geospatial and epidemiological data, phylodynamic models allow investigating the migration history of pathogens and its drivers in the absence of detailed contact tracing data and when linkage among infections is not obvious (36-39). Such analyses have been widely adopted both for human (40, 41) and plant viruses (38, 42), and more recently for HIV (43, 44).

The overall goal of the present study is to characterize the dynamics and drivers of the across-province spread of the main HIV types circulating in China. For this purpose, we capitalize on a discrete phylogeographic approach coupled to a generalized linear model extension.

RESULTS

Population characteristics

A total of 6800, 1578 (822/756), 1158, 957 and 211 available sequences were retrieved for CRF01_AE, subtype B (B/B'), CRF07_BC, CRF08_BC and CRF55_01B respectively. The number of provinces included in each final data set varied from 7 (CRF55_01B) to 17 (CRF01_AE and B/B'). See ***Figure 1*** for the distribution of provinces per data set.

Preliminary phylogenetic analysis and subsampling

For CRF01_AE, an initial set of 6,423 HIV-1 CRF01_AE *pol* sequences from 53 countries across the world between 1990 and 2017 retrieved from the Los Alamos National Laboratory HIV Sequence Database (45) was combined with the CRF01_AE data set of 6,800 sequences to delineate clades that capture the epidemic dynamics in transmission networks that pertain to China. We identified 83 of such clades (n=1876 sequences). To obtain data informative of interprovincial migration patterns, these were reduced to the 54 clades (size 3-24 sequences) that included samples from at least 2 Chinese provinces (totaling 454 sequences from 17 provinces). The same rationale was used for the other data sets. Starting from a total of 1578, 1158 and 957

sequences for subtype B/B', CRF07_BC, and CRF08_BC data sets, we obtained 15, 16, and 7 clades from 17, 10 and 8 provinces respectively. For CRF55_01B, which is circulating in China only, we obtained a single clade of 197 sequences collected across 7 provinces.

Discrete phylogeographic inferences

We used Bayesian phylogeographic inference to evaluate the dispersal history of the five main circulating types across Chinese provinces. This allows, for each subtype and CRF, to identify the significant migration events between Chinese provinces, and to estimate their number and directionality (***Figure 1***). The reconstructed patterns of spread based on the identified clades revealed strong evidence (adjusted Bayes Factor [BF_{adj}] ≥20) of migration between provinces for all sampled HIV populations. The relative contribution of each province as source and sink of HIV dispersal throughout China is summarized in ***Table 1***.

CRF01_AE and CRF07_BC clades have become the two predominant HIV CRF in China with an overall prevalence of 46.34% [95% CI: 40.56–52.17%] and 19.16% (95% CI: 15.02–23.66%), respectively (46). Here, the discrete phylogeographic analysis for CRF01_AE supported a complex migration history, with Beijing (Chinese capital with the second highest population density), Guangdong (southern region, capital Guangzhou), Shanghai (the most populous urban area in China) and Anhui (an important part of the Yangtze River Delta and in the top four provinces of China in labor export) being the provinces most involved in the interprovincial spread of migration events, both as major sources (with 24.9% [95%CI: 24.8-25], 16.6% [95%CI: 16.5-16.7], 15.8% [95%CI: 15.7-15.9] and 10.6% [95%CI: 10.5-10.7] of viral diffusion, respectively) and as major sinks (with 17.2% [95%CI: 17.1-17.3], 11.7% [95%CI: 11.6-11.8], 25.7% [95%CI: 25.5-25.8] and 13.3% [95%CI: 13.2-13.4] of introduction, respectively) (***Figures 1 and 2A***). For CRF_07BC, the second most prevalent type, the main sources were Beijing and Shanghai along with Yunnan (northwest-central region, capital Kunming) (***Figures 1 and 2C***), while for CRF08_BC, the main source was the province of Yunnan. Our model also showed robust evidence of viral migration across China for HIV-1 subtypes B/B' with Hubei (Capital Wuhan) being the major source of viral migration accounting for 69.9% [95%CI: 69.7-70.1] of viral dispersal with predominant diffusion toward the province of Henan (capital Zhengzhou) with 55.8% [95%CI: 55.6-56] of all introduction events, acting as a sink for the B/B' epidemic (***Figures 1 and 2B***). Finally, the southern province of Guangdong with the largest population was the only source of migration for CRF55_01B directed toward Anhui and Hunan that is supported by our data (***Figures 1 and 2E***).

From an historical perspective, our analyses showed a higher density of migration events across provinces in the late 1990' and 2000' years for subtype B/B' while migration events in general are more concentrated over the past 15 years for CRF01_AE, CRF07_BC and CRF08_BC. We also found that the historical interprovincial dispersal of CRF55_01B predominantly occurred around 2010 (*data not shown*).

Generalized linear model analyses

We next used the generalized linear model (GLM) extension of the phylogeographic model to evaluate the association between potential predictors and the migration frequencies among provinces (***Figure 3***). For CRF01_AE our model revealed a strong association between migration events and air traffic density as well as connectivity among locations, associations that are robust to randomizing tip-to-location assignments (BF_{adj} >> 100). The conditional effect size for connectivity between major cities over land was negative, meaning that spread between provinces that are easier to travel between over land is less frequent than between provinces that are less well connected over land. For CRF08_BC, a higher HIV prevalence in the province of origin was associated with increased HIV dispersal (BF_{adj}=87.8), which also associates with the number of immigrants at the origin (BF_{adj} =11.3). Whereas a higher HIV prevalence at the province of origin links to more frequent migration from that province, the conditional effect size for the number of immigrants at the origin was negative, implying that for CRF08_BC more immigration towards a province links to less frequent virus migration from that province. The only other predictor that was well-supported is spatial distance for subtype B/B'. For this predictor too, the conditional effect size is negative, indicating that migration is more frequent between closer locations. No other associations are well-supported (i.e. BF_{adj} ≥3) (***Figure 3)***.

DISCUSSION

Starting from >10,000 HIV-1 *pol* sequences from the five main prevalent HIV-1 subtypes and CRFs in China collected between 1996 and 2017, we reconstructed the spatial diffusion of the five most prevalent HIV-1 types across provinces in China. Our reconstructions largely recapitulate their documented spread, which we discuss one by one:

CRF01_AE. CRF01_AE has become the dominant HIV variant in most provinces (3). In line with previous epidemiological studies (3) and molecular analyses (47), our reconstructions capture that most migration events occurred recently between southern and eastern/north-eastern provinces

(see ***Figure 1*** and ***Table 1***). Specifically, we found that Beijing (the political, economic and cultural center of China) was the main source of CRF01_AE dispersal throughout the country and that Guangdong, Shanghai and Anhui are the other major hubs of CRF01_AE dispersal (***Table 1***). This largely matches the geographic scope of within-China population migration flows, which were concentrated within and between the southern and eastern main economic provinces (48, 49). It is of note that CRF01_AE is dominant among MSM (3), and that interprovincial migrants (as compared to intra-provincial migrants) not only are more likely to be male but also tend to be younger and have fewer years of formal education (49), which are factors associated with higher risk behavior (50).

The GLM analyses confirmed a strong association between the intensity of migration events and air traffic density, and an inverse relation of connectivity between major cities over land with the migration intensity. Combined, our results indicate that interprovincial CRF01_AE mobility is driven predominantly by longer-distance migration, possibly MSM-related, a combination that has also been noted in e.g. regions of Canada (51).

HIV Subtype B/B'. After an initial period of dominance, the prevalence of B/B' has declined (17). Consistent with this trend, we found that viral dispersal of HIV-1 subtype B mostly occurred in the 1990s and early 2000s (*data not shown*). Also in line with epidemiological surveys and with previous molecular analyses (52), we found that Hubei and Henan, both with a historically predominant circulation of B/B' among blood donors (53-57), were the major sources of interprovincial dispersal for this HIV-1 type (***Table 1 and Figure 2***).

Zhengzhou (Henan capital) is located at the junction of the major north-south Beijing-Guangzhou and east-west Lanzhou-Lianyungang railways and has evolved into a major national administrative, economic, and transportation hub (58), and Henan and Hubei are among the top five largest 'migration sending areas' (49). This shows that there was ample opportunity for long-distance spread of B/B' in relation to human mobility. In turn, the dominance of shorter-distance spread of B/B' implies that it did not find much fertile ground in highly mobile high risk groups, such as interprovincial migrant workers. This aspect is reflected in the results of our GLM analyses, which showed that migration events occurred more frequently among more nearby provinces (***Figure 3***).

CRF07_BC and CRF08_BC. CRF07_BC was originally reported in the Yunnan province in 1980s and spread quickly among IDUs. In recent years, it has been introduced in MSM populations, which drove its spread to elsewhere in China, particularly to Beijing, Shanghai, Guangdong and Zhejiang (9, 17, 59, 60). CRF08_BC on the other hand was initially reported in Yunnan and Guangxi provinces among IDUs but it has rarely been reported in MSM or other risk populations.

While none of the predictors appear significantly associated with the spread process of CRF07_BC, the origin of CRF08_BC in the south-eastern part of China and its subsequent spread towards economically more developed provinces that are attraction poles for inland migration is reflected in which predictors were found to significantly associate with the migration process, as well as the direction of their effect sizes. Specifically, the migration frequency out of a province increases with increasing HIV prevalence, and the migration intensity is inversely associated with the number of immigrants in the provinces of origin, suggesting that immigration hot spots functioned as a sink for this type (***Figure 3***).

CRF55_01B. This CRF was first identified among MSM in Shenzhen, Guangdong (18, 19). It has now spread throughout most provinces of China (20, 61, 62) although it mostly circulates in Guangdong and neighboring provinces, and across all risk groups (18). In line with this, our results point to Guangdong as main source of the dispersal to the western province of Anhui but also the adjacent province of Hunan (**Table 1**). As Guangdong is an economically well-developed province and attraction pole of migrant workers, it may intuitively seem at odds that it is a source rather than a destination of CRF055_01B spread. This may, however, be explained by return migration, which has become more intense over the years (63).

Prior to the 1980s, rural-urban migration in China was minimal. Since then, China has witnessed an extraordinary internal migration: rural-to-urban migrants increased the urban population by approximately 390 million. Of these rural migrants, approximately 54% were interprovincial migrants, most of which left their home province, but also with many returning after some time, and many visiting their families on a regular basis (e.g. with Chinese New Year). The reconstructed patterns of interprovincial spread for the main HIV types in China support the idea that migrant workers are at least partially involved in their diffusion. Unfortunately, the lack of epidemiological metadata prevented us to more explicitly elucidate the dynamics of spread within and between relevant subpopulations by for example associating epidemiological characteristics with uptake in clusters of closely related viruses (51, 64-66), which can help identify on what aspects to focus screening and prevention efforts. The involvement of migrant workers can be tested more directly within the GLM framework. Regrettably, we could only dispose of the total number of immigrants/emigrants by province instead of more granular pairwise migration flow data. Nonetheless, the identified drivers of HIV dispersal in China are in line with the view that human mobility strongly impacts pathogen epidemic dynamics (67). This combines with the reconstructed patterns of spread that largely reflect within-China population migration flows (that are directed towards and between major population and economic centers), to suggest that the patterns of viral transmission for at least some of the HIV epidemics in China were driven by major

population centers, which can act as gravity attractors before the virus spread to smaller populations (68, 69). This also illustrates that, in the absence of concurrent national prevention efforts with a focus on the most important drivers of ongoing transmission, local epidemics will rapidly be re-seeded, challenging the long-term impact of isolated intervention efforts.

Limitations. One major limitation of our study is that the collection of the HIV-1 *pol* sequences from the five main prevalent HIV-1 types in China has not been performed under a common framework, which may render our analyses prone to sampling bias. To the best of our knowledge, this drawback affects nearly every phylogeographic study of HIV-1 and other viruses. Whereas structured coalescent approaches hold promise for unbiased inferences in the face of biased sampling, inference under high state spaces and large data sets remains challenging for these models. For this reason, we relied on the computationally more efficient discrete trait analysis (70, 71). To counter this model's sensitivity to biased sampling of subpopulations, we (i) adopted a filter based on location state randomizations and (ii) combined the geographical information from different partitions that represent different samples of the same epidemic to minimize the risk of false positive migration linkages and associations with covariates (72, 73). Given that the reconstructed interprovincial spread largely captures the documented spread of the investigated HIV-1 types, we believe that these precautions were effective.

Several factors can explain that only few of the tested predictors associated with the spread patterns. The high-level resolution of our phylogeographic reconstructions makes that potential predictors can only be evaluated against a limited number of migration events between locations, in particular for CRF07_BC, CRF08_BC and CRF55_01B. Also, when only few migration events are observed, the impact of imperfect representations of the location-specific diversity on the ancestral reconstructions will increase and can obfuscate the relevance of potential predictors. Furthermore, our models did not capture potential time-varying dynamics of the selected predictors over the study period. This is particularly important for longstanding epidemics, such as HIV-1 subtype B/B'. Unfortunately, we could not test this hypothesis as we did not dispose of time-variable predictors.

CONCLUSION

The rapid increase of HIV-1 prevalence among migrant populations and the lack of effective intervention strategies is one of the current challenges for China (74, 75). In this study, the combined use of phylogeographic reconstructions and generalized linear model provides insights into the spatial viral dynamics of various HIV epidemics across provinces in China. The role of large urbanized areas with dense population and developed transportation infrastructures as

facilitator of HIV dispersal throughout China illustrates the need to consider harmonized country-wide public policies to control local HIV epidemics.

MATERIALS AND METHODS

Ethics statement

The study was approved by the ethics committee of the First Affiliated Hospital of China Medical University in Shenyang and Wuhan University of Bioengineering.

Data set compilation

We retrieved all publicly available HIV partial *pol* sequences (HXB2 position 2253-3554) of CRF01_AE, CRF07_BC, CRF08_BC, B/B' and CRF55_01B with known sampling date and sampling province of China from the Los Alamos National Laboratory HIV Sequence Database (45).

We additionally collected 1,962 CRF01_AE and HIV-1 subtype B partial *pol* sequences from the NHC Key Laboratory of AIDS Immunology, China Medical University (GenBank accession numbers MT336741:MT336811; MT368039:MT369927). We also retrieved publicly available HIV *pol* sequences from other countries along with sampling time and related geographical information. When multiple sequences were available for one participant, only the closest sequence from the estimated time of infection was kept.

Identification of Chinese clades

The geospatial unit in all phylogeographic analyses was the Chinese Province (first subnational administrative level). For all five subtypes except CRF55_01B, for which only isolates from China are available (19), we applied the step-by-step approach described below.

1. Following the approach of Cuypers et al. (76) and using an as complete as possible background data set (77), we first identified clades that likely correspond to distinct HIV introductions in China. To this end, the sequences for each subtype were first complemented with the publicly available location-annotated HIV *pol* sequences from the same subtype and aligned to a *pol* reference sequence (HXB2, GenBank accession K03455 (78). AliView (79) was used for manually editing the alignments.
2. Next, phylogenetic trees were inferred using FastTree2 (80) under a general GTR+Γ substitution model. These served to identify strongly supported Chinese clades, i.e. clades only including Chinese sequences and associated with a Shimodaira Hasegawa (SH) support of at least 0.9 (81-83).
3. Within these monophyletic clades, well-supported clusters of sequences sampled from the same administrative area (province) were identified. These were downsampled by randomly selecting one sequence from each cluster. This step reduces computational burden while

preserving estimation accuracy for the migration flow quantities between Chinese provinces (40).

Time scale for the evolutionary histories

When sequence data sets lack a clear temporal signal, it is common practice to use empirical evolutionary rate estimates for specifying a suitable prior distribution on the evolutionary rate parameter (e.g.(84, 85).

Subtype B/B'. To obtain plausible priors for the evolutionary rate for HIV-1 subtype B, we considered that various evolutionary rate have been reported for *pol*, varying from ~0.001 to ~0.003 substitutions/site/year (s/s/y) (86-88). For this reason, we specified a normal distribution as prior on the mean clock rate with mean 0.002 s/s/y and standard deviation such that the 95% confidence interval is bound at 0.001 s/s/y and 0.003 s/s/y.

CRF01_AE. We considered data from the literature (mean rate estimate ~0.0015 (87)) as well as the population-level substitution rate estimate of ~0.0027 [95%HPD: 0.0013- 0.0032] s/s/y obtained from clade-based specific analyses of the CRF01_AE data set with a Bayesian hierarchical phylogenetic model (HPM) approach (data not shown, (88)). This led us to specify a normal distribution as the prior on the mean clock rate with mean 0.002 s/s/y and standard deviation of 0.0005.

CRF07_BC, CRF08_BC and CRF55_01B. For these subtypes, a normal distribution was specified as the prior on the mean clock rate of ~0.001 s/s/y and standard deviation of 0.0005 according to clade-based estimates (data not shown).

Many of the clades that represent the HIV epidemics in China are limited in size. As this precludes reliable inference under the parameter-rich uncorrelated relaxed clock model (e.g. (89)), we opted to model the rate of evolutionary change in clades with ≥10 taxa with a relaxed clock model (90) while for the smaller clades a strict clock model was assumed.

Phylogeographic inference

Phylogeographic inference was performed using the discrete diffusion model (70, 91) implemented in the software package BEAST 1.10 (92). To promote estimation accuracy and precision of the transition rates among locations, the substitution model (GTR+Γ) and spread process were shared among clades of the same type (40, 72). A constant size coalescent prior was assumed for all clades of CRF01_AE, subtype B and CRF55_01B, and a non-parametric Bayesian skygrid tree prior was for clades with ≥20 taxa for CRF07_BC and CRF08_BC (93, 94).

To identify the subset of transition rates that was most informative to reconstruct the dispersal history, we used a model averaging procedure (Bayesian stochastic search variable selection – BSSVS) (70). In this procedure, the level of support depends on the *a priori* expected and *a posteriori* noted fraction of time during the Markov Chain Monte Carlo (MCMC) integration that a migration link or predictor helps explain the migration history. In the default setup, however, the *a priori* expectation only depends on the number of locations but does not account for the relative abundance of samples by location. This can bias inference in the presence of uneven sampling. The adjusted Bayes factor (BF_{adj}) (73) improves on this by incorporating information on the relative abundance of samples by location. It also relies on the *a priori* expected and *a posteriori* noted inclusion frequencies under BSSVS but relative to the original test, it requires two analyses: a first one where the trait values remain associated with their respective taxa, and a second one during which the trait values are randomized over the tips of the tree during the MCMC sampling. The latter provides the expectation in the absence of structure in the population, akin to the date randomization test when evaluating the presence of temporal signal(42, 95). As before, support for the significance is calculated as a ratio with the posterior odds as the enumerator, but as denominator we consider the inclusion frequency from the randomized analysis instead of the prior odds. Bayes factor (BF) support for all possible types of location exchanges was calculated with SpreaD3 (96). BF and BF_{adj} between 3-10, 10-20, and above 20 were considered to be substantial, positive and strong supports respectively for the observed transition rates between sampled locations (97). Estimates of the posterior probability of expected number of migration events between all pairs of locations (Markov jumps) were computed through stochastic mapping techniques (98, 99).

MCMC chains were run to ensure adequate mixing. Maximum clade credibility (MCC) trees were obtained with TreeAnnotator 1.10 (92) and convergence and mixing properties were inspected using Tracer 1.7 (100).

Generalized linear model analyses

We used the GLM extension of the discrete phylogeographic model implemented in BEAST 1.10 (39) to investigate the contribution of a series of location-associated variables to the migration rates among Chinese provinces. These variables included socio-demographic indicators (population size, number of emigrants and immigrants), HIV prevalence, sample size, and variables related to connectivity between locations (i.e. air traffic density, travel time by railways, the presence of shared borders, a measure of connectivity based on an accessibility model to major cities, and a proxy for spatial distance). Predictors were considered both at the origin and destination location. The population size, the number of emigrants and immigrants and HIV

444 prevalence were obtained from the National Bureau of Statistics of China (101) and the China
445 National Center for Disease Control and Prevention (102).
446
447 The numbers of sequences sampled at the origin/destination were included in the GLM to account
448 for the potential impact of sampling biases within the analysis (39). The air traffic density was
449 approximated by an air passenger flux matrix that quantifies the number of passengers traveling
450 between each pair of administrative areas (39). We use a data set provided by the OAG (Official
451 Airline Guide; www.oag.com) and containing the annual average number of seats on scheduled on
452 commercial flights between pairs of airports between 2014 and 2016 (103), assuming that the
453 number of seats represents a reasonable proxy for the number of passengers traveling between
454 airports. Travel time by railways was represented by the shortest travel time between the capitals
455 of each province obtained from 12306 China Railway website (104).
456
457 Geospatial connectivity measures included in the GLM were the following: a binary determination
458 if administrative areas share a common border, the average travel time by railways between
459 locations as well as two measures of connectivity among administrative areas computed using an
460 algorithm based on circuit theory and implemented in the program Circuitscape 4.0.5 (105): a
461 measure of connectivity and a proxy of spatial distance, obtained by computing pairwise
462 resistances on an inaccessibility grid and a uniform grid, respectively. For a given pair of locations,
463 Circuitscape computes the pairwise electric resistance based on a geo-referenced grid (or “raster”)
464 covering the study area and defining the local electric resistance values. To compute the proxy of
465 spatial distance, we simply used a homogeneous raster file with cell values uniformly set to “1”,
466 and for the pairwise connectivity measures, we used the inaccessibility raster as in (106) to define
467 the local values of electric resistance. Cell values of this inaccessibility raster indicate the travel
468 time required to reach the nearest urban center, with an urban center defined as a contiguous
469 area with 1,500 or more inhabitants per square kilometer or a population center of at least 50,000
470 inhabitants (106). For computational tractability, the resolution of both the uniform and
471 inaccessibility raster were decreased to ~5 arcmin (original resolution: ~0.5 arcmin). There are
472 several advantages to use pairwise Circuitscape distances computed on a uniform raster alone or
473 in complement to great-circle distances as proxies of spatial distance. First, pairwise Circuitscape
474 distances constitute more realistic measures because the underneath path model does not
475 assume straight-line movements and it also prevents movement through inaccessible areas.
476 Furthermore, given that the uniform raster is the homogeneous version of the inaccessibility raster,
477 pairwise Circuitscape resistances computed on the uniform raster also represent a proper
478 negative control (38, 107) for the inclusion, in the GLM analysis, of pairwise Circuitscape

resistances computed on an heterogeneous raster like the inaccessibility one. Indeed, the inclusion of a GLM predictor that does not have an impact on the dispersal, but for which pairwise distances have been computed using an advanced path model (like the one implemented in Circuitscape), can yield a false positive result in the absence of an appropriate negative control (38). Circuitscape computes pairwise electric resistance between two points or between two sets of points that all have to be associated with precise geographic coordinates. Given that such precise sampling coordinates were not available for sampled sequences, we randomly assign geographic coordinates to each sampled sequence. While this assignment was stochastic, we still used a human population density raster (resolution of ~5 arcmin) to define the sampling probability of all the raster cells within an administrative area. Hence, for each sequence originated from a given administrative area, its probability of being sampled from a particular raster cell was proportional to human population density value assigned to this cell. As this is a stochastic procedure, the sampling coordinate assignment and subsequent Circuitscape analyses were repeated 100 times. Final matrices of pairwise resistances computed on the uniform and inaccessibility rasters were obtained by averaging the 100 matrices computed after each repetition of the above procedure. Note that we used the same procedure to compute the averaged great-circle distances among locations.

To protect against a potential impact of sampling imbalances on the GLM results, support for the need for a predictor to help explain the variation in migration rates across locations was obtained after accounting for the relative abundance of the involved trait states (73).

Data availability. HIV-1 subtype B partial *pol* sequences are available in GenBank under accession numbers MT336741 to MT336811 and MT368039 to MT369927.

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FIGURES AND TABLES

Figure 1. Migration events between Province in China. The thickness of the arrows corresponds to the average number of inferred migration events, their curvature indicates the migration direction, and their colors reflect the support for each link (green, orange, purple for respectively $3 \leq BF_{adj} < 10$ (substantial), $10 \leq BF_{adj} < 20$ (positive) and $BF_{adj} \geq 20$(strong)). Provinces are colored according to the number of sequences included in the clusters for each HIV type.

Figure 2. Migration events between provinces in China. Sankey plot showing the proportion of migration events from each source province toward the recipient provinces. Left side of the plots shows the source of migration events. Right side of the plot shows the destination of migration events. Only results with adjusted Bayes Factors (BF_{adj}) ≥3 are shown. Panel A to E for types CRF_01AE, B, CRF_07BC, CRF_08BC and CRF55_01B respectively.

Figure 3. Predictors of transition rates among locations. A. The boxplots report the posterior distribution of each GLM coefficient, i.e. the contribution of each predictor to the model, when included in the model (the conditional effect size). The adjusted BFs after accounting for sampling heterogeneity are reported when ≥3, and the corresponding conditional effect sizes are plotted in darker grey. **B.** Map of the Chinese Provinces. **C.** Variables tested as predictors of dispersal transition rates across locations. (A) Number of HIV Cases per provinces were obtained from the National Bureau of Statistics of China (101) and the China National Center for Disease Control and Prevention(102); (B) the population size (in million), (C) Number of Emigrants and (D) Immigrants were obtained from the National Bureau of Statistics of China (101). The numbers of sequences sampled at the origin/destination (See **Figure 1**) were also included in the GLM to account for the potential impact of sampling biases within the analysis (39).

Table 1. Relative importance of provinces in the interprovincial spread of the main HIV types in China. For each province, the percentage refers to the average proportion that the province is as the source (from) or recipient (to) if the emigration event. For each type, the three most relevant provinces are listed.

Type	From	Mean [95%HPD]	To	Mean [95% HPD]
CRF_01AE	Beijing	24.9% [24.8-25]	Shanghai	25.7% [25.5-25.8]
	Guangdong	16.6% [16.5-16.7]	Beijing	17.2% [17.1-17.3]
	Shanghai	15.8% [15.7-15.9]	Anhui	13.3% [13.2-13.4]
	Anhui	10.6% [10.5-10.7]	Guangdong	11.7% [11.6-11.8]
	Shandong	5.6% [5.5-5.6]	Jiangsu	10.6% [10.5-10.7]
	Zhejiang	5.4% [5.4-5.5]	Henan	7.5% [7.4-7.6]
	Liaoning	5.3% [5.3-5.4]	Guangxi	4.4% [4.4-4.5]
	Guangxi	3.9% [3.8-3.9]	Liaoning	3.1% [3.1-3.2]
	Sichuan	3.3% [3.3-3.4]	Shandong	2.4% [2.4-2.5]
	Jiangsu	3.1% [3.1-3.2]	Zhejiang	2.3% [2.3-2.3]
	Henan	2.3% [2.3-2.3]	Sichuan	1.1% [1-1.1]
	Fujian	1.8% [1.8-1.8]	Chongqing	0.5% [0.5-0.5]
	Hebei	1.2% [1.2-1.2]		
CRF_07BC	Beijing	43.4% [43.1-43.8]	Guangdong	43.4% [43.1-43.8]
	Shanghai	29.9% [29.5-30.3]	Zhejiang	39.2% [38.8-39.6]
	Yunnan	15.7% [15.4-16]	Beijing	10.2% [10-10.5]
	Xinjiang	7.2% [7-7.4]	Xinjiang	6.4% [6.2-6.6]
	Liaoning	3.7% [3.6-3.9]	Ningxia	0.7% [0.7-0.8]
CRF_08BC	Yunnan	85.2% [84.8-85.7]	Guangxi	44.2% [43.6-44.9]
	Guangdong	14.8% [14.3-15.2]	Guangdong	32.4% [31.8-33]
			Hebei	11.3% [10.8-11.7]
			Shanghai	8.6% [8.2-9]
			Sichuan	3.5% [3.3-3.7]
B	Hubei	69.9% [69.7-70.1]	Henan	55.8% [55.6-56]
	Henan	16.7% [16.6-16.9]	Guangdong	10.9% [10.8-11]
	Liaoning	5.4% [5.3-5.5]	Zhejiang	9.8% [9.7-10]
	Zhejiang	3% [2.9-3.1]	Hubei	7.5% [7.4-7.6]
	Hebei	2.1% [2-2.1]	Beijing	7.4% [7.2-7.5]
	Beijing	1.3% [1.3-1.4]	Anhui	5.6% [5.5-5.7]
	Guangdong	0.9% [0.9-0.9]	Shandong	1% [0.9-1]
	Anhui	0.3% [0.3-0.4]	Hebei	0.5% [0.5-0.6]
	Yunnan	0.2% [0.2-0.2]	Guangxi	0.4% [0.3-0.4]
	Shandong	0.1% [0-0.1]	Jiangsu	0.3% [0.3-0.4]
	Fujian	0% [0-0]	Shanghai	0.3% [0.3-0.3]
	Jilin	0% [0-0]	Jilin	0.2% [0.2-0.2]
			Liaoning	0.2% [0.1-0.2]
			Ningxia	0.2% [0.1-0.2]
			Fujian	0% [0-0]
			Yunnan	0% [0-0]
CRF_5501B	Guangdong	100% [99.8-100]	Anhui	73.9% [71.8-75.8]
			Hunan	26.1% [24.2-28.2]