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Response to Valle and Zorello Laporta: Clarifying the Use of Instrumental Variable Methods to Understand the Effects of Environmental Change on Infectious Disease Transmission.

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MacDonald, Andrew J
Mordecai, Erin A

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1 Response to Valle and Zorello Laporta: Clarifying the use of instrumental variable methods to
2 understand the effects of environmental change on infectious disease transmission

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4 Running head: Response to the critique by Valle & Zorello Laporta

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6 Andrew J. MacDonald^{1*} & Erin A. Mordecai²

7 1. Earth Research Institute and Bren School of Environmental Science and Management,

8 University of California, Santa Barbara, CA, USA

9 2. Department of Biology, Stanford University, Stanford, CA, USA

10 *Corresponding author: Bren School of Environmental Science and Management, University of

11 California, Santa Barbara, CA 93106-5131; andy.j.macdon@gmail.com

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23 Abstract

24 Identifying the effects of environmental change on the transmission of vector-borne and
25 zoonotic diseases is of fundamental importance in the face of rapid global change. Causal
26 inference approaches, including instrumental variable (IV) estimation, hold promise in
27 disentangling plausibly causal relationships from observational data in these complex systems.
28 Valle and Zorello Laporta recently critiqued the application of such approaches in our recent
29 study of the effects of deforestation on malaria transmission in the Brazilian Amazon on the
30 grounds that key statistical assumptions were not met. Here, we respond to this critique by: 1)
31 deriving the IV estimator in order to clarify the assumptions that Valle and Zorello Laporta
32 conflate and misrepresent in their critique; 2) discussing these key assumptions as they relate to
33 our original study and how our original approach reasonably satisfies the assumptions; and 3)
34 presenting model results using alternative instrumental variables that can be argued more
35 strongly satisfy key assumptions, illustrating that our results and original conclusion—that
36 deforestation drives malaria transmission—remain unchanged.

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45 Main Text

46 There is substantial and increasing interest in understanding the role that processes of
47 global change are playing in the ecology and transmission of vector-borne and zoonotic
48 diseases.^{1,2} While these questions are of fundamental importance given the increasing rate of
49 climate and land use change, and the large proportion of emerging infectious diseases that are
50 vector-borne or of zoonotic origin,³ causally linking these two processes is an enormous
51 challenge. Take as an example the case of deforestation impacts on malaria transmission in the
52 Brazilian Amazon, the focus of MacDonald & Mordecai⁴ and the critique by Valle & Zorello
53 Laporta.⁵ The gold standard of a randomized controlled trial in which deforestation is
54 experimentally manipulated and randomly assigned to different regions to assess its impact on
55 malaria transmission presents obvious logistical and ethical barriers that make such an approach
56 largely infeasible. As a result, researchers must rely on observational data and employ statistical
57 approaches to approximate, as closely as possible, the experimental ideal.

58 One promising set of statistical techniques—broadly referred to as causal inference
59 methods, which includes Instrumental Variable (IV) estimation, are increasingly being leveraged
60 to disentangle plausibly causal relationships from observational data in ecology. Due to the
61 challenges described above, these approaches have been employed by researchers assessing
62 global change impacts on infectious disease,⁶⁻¹⁴ including in another recent study investigating
63 the effects of deforestation on malaria transmission in Brazil,¹⁴ with similar results to our own
64 work. Valle and Zorello Laporta⁵ rightly point out that model assumptions are critically
65 important in such approaches, and that causal conclusions should be carefully drawn in these
66 contexts. However, the authors unfortunately conflate the assumptions of IV estimation in their

67 perspective piece. As a relatively new approach in ecology and environmental science,⁶ it is
68 important that the underlying assumptions are clear for appropriate application.

69 IV is a useful approach to overcome what is known as endogeneity bias, which is due to a
70 relationship between the error term and one or more of the explanatory variables, (formally,
71 $E[\varepsilon_i \vee x_i] \neq 0$ where ε and x represent the error term and explanatory variable for observation i).
72 Such a relationship could be due to bidirectional causality where, for example, deforestation may
73 drive malaria transmission but malaria burden may also influence rates of deforestation. In IV, a
74 third variable, known as an instrument (z_i), is used to isolate exogenous variation in explanatory
75 variable x_i and recover a statistically consistent estimator for the true relationship between the
76 exogenous variable and the outcome.

77 The instrument must meet two conditions for IV to be a consistent estimator, which are
78 sometimes termed “relevance” and “exclusion” criteria. In words, the instrument must be
79 statistically associated with the endogenous variable (“relevance”) and must be related to the
80 outcome only through its relationship with the endogenous variable (“exclusion”). While the
81 wording is easy to remember, it leaves much open to interpretation. For example, does relevance
82 require a causal link? Does exclusion require statistical independence? The derivation makes
83 these key assumptions much more apparent. Before showing the derivation, we will first provide
84 brief background to our original study,⁴ the critique by Valle & Zorello Laporta⁵ and our
85 response.

86 In MacDonald & Mordecai,⁴ we were first interested in predicting annual malaria
87 incidence as a function of annual deforestation, and use aerosol optical depth (AOD) in the
88 month of September from MODIS satellite imagery as our “instrument.” We expand on the

89 methodology and terminology below, but set the context of the argument here. Valle & Zorello
90 Laporta⁵ have two critiques of our IV approach. The first, however, is a misrepresentation of the
91 assumptions of IV, namely that a valid IV requires that the IV has a *causal* effect on the
92 endogenous explanatory variable. They state, “However, it is deforestation that causes aerosol
93 pollution [...] rather than aerosol pollution that causes deforestation [...] As a result, [the
94 relevance] assumption is clearly violated.” As we show below, causality is not required.

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Rather, there must be an “association”, or more specifically, the covariance between the instrument and the endogenous variable must not be zero. However, it is possible that an instrumental variable itself introduces endogeneity bias if it does not meet the exclusion criteria, and this can be particularly problematic in the case of “weak instruments” as we show below. This can occur, for example, in cases where the instrument (e.g., AOD) is strongly driven by the endogenous predictor variable (e.g., deforestation). In our case, we chose AOD as an instrument for deforestation, as it is an indicator of human activity on the landscape.¹⁶ Further, over our

181 study period, AOD was decoupled from deforestation as biomass burning in the Brazilian
182 Amazon—and resulting AOD—was primarily driven by fires intentionally set to keep *existing*
183 pastures and agricultural lands clear¹⁶ and by drought conditions leading to wildfires in already
184 degraded forests,¹⁶⁻¹⁸ rather than by new deforestation activity.

185 Nevertheless, to explore the extent to which our original IV estimates of the effect of
186 deforestation on malaria may have been affected by potential endogeneity introduced by the use
187 of AOD as an IV, we run additional IV models using 1) last year’s AOD as an instrument for this
188 year’s deforestation, and 2) remotely sensed, average municipality soil quality¹⁹ processed in
189 Google Earth Engine,²⁰ interacted with annual international soy and beef commodity prices from
190 the World Bank. We chose last year’s AOD since it is correlated with this year’s deforestation
191 (relevance), but this year’s deforestation could not have caused last year’s AOD. While this
192 addresses the issue of reverse causality, it is plausible that there remain endogeneity issues in this
193 context. For example, if last year’s AOD somehow acts upon this year’s malaria through
194 mechanisms beyond deforestation, then the exclusion criteria would fail. To address these
195 potential lingering concerns, we run additional models using soil quality coupled with
196 international agricultural commodity prices for key Brazilian exports, which may influence a
197 land owners’ decision to clear forest for agricultural production (relevance); in this case,
198 deforestation rates do not cause soil quality and are highly unlikely to shift international
199 commodity prices (exclusion). We run these IV models on our interior Amazon sample of
200 municipalities, where active deforestation rates are highest and where we predict forest clearing
201 should have the strongest effect on malaria transmission,⁴ predicting both total malaria and
202 *Plasmodium falciparum* malaria incidence, following our original study.⁴ Results are presented
203 in the SI (Table S1). In brief, we find significant positive effects of deforestation on malaria

204 transmission in each of these additional model specifications, with coefficients of similar, though
205 slightly larger magnitude than our original study. Our main conclusion, that deforestation
206 increases malaria transmission in the Brazilian Amazon, remains unchanged.

207 The second goal of MacDonald & Mordecai⁴ is to understand whether annual malaria
208 burden feeds back to influence annual rates of deforestation, and we use optimal temperature for
209 malaria transmission in the dry season as our instrument for malaria. Optimal temperature was
210 defined as the sum of days falling within a narrow temperature band that is optimal for malaria
211 transmission (24-26°C) based on earlier mosquito and parasite trait-based mechanistic modeling
212 studies.²¹ Valle & Zorello Laporta's⁵ second critique is that the exclusion assumption may be
213 violated in this model because "it is possible that temperature affects deforestation not only
214 through malaria, but also through other causal paths," particularly the relationship between
215 temperature and agricultural gross domestic production.²² In other words, favorable temperatures
216 for mosquitos and malaria parasites may affect deforestation not just through malaria, but by also
217 being favorable agricultural growing conditions, which increase the potential value of forest
218 clearing. We agree that temperature is important to both agriculture and malaria, and that those
219 clearing land may consider the land's growing potential. However, rather than counting the
220 number of days in a 2°C temperature window during the dry season, we suggest agricultural
221 producers will instead consider the general growing conditions of a region as it relates to
222 commonly grown crops—for example, soil quality, climate, topography, and infrastructure. As
223 land clearing for agriculture is a large and long-term investment, average growing conditions are
224 much more likely to influence clearing decisions than are small deviations in weather from year
225 to year.

226 There are two additional primary reasons that our IV, optimal malaria transmission
227 temperature, is highly unlikely to fail the exclusion criteria. First, we specifically employ
228 municipality “fixed effects” or dummy variables

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to remove roughly time invariant characteristics specific to each municipality through differencing. Thus, average characteristics (e.g., soil quality, average precipitation, average temperature) that are likely to influence the evolution of regional agricultural land use and the location of processing plants and other infrastructure are removed and the model is identified from deviations from the municipality-specific mean. Second, the range of optimal average temperatures for soybean—Brazil’s main crop by area and production²³—cultivation and development in Brazil is from 20°C to 35°C.²⁴ Recall optimal temperature for malaria transmission is 24°C to 26°C, and we use the number of days in the dry season within this narrow temperature band as our instrument. Thus, an additional day at 25°C relative to 27°C would be expected to lead to increases in malaria transmission. However, this same change in temperature

318 would likely have a trivial impact on soy yields, as both temperatures are well within the bounds
319 of optimal soy cultivation. Given the breadth of favorable temperatures for soy, it is unlikely that
320 changes in the number of days between 24°C to 26°C will influence land clearing decisions for
321 agricultural production.

322 We too feel that causal inference approaches hold much promise in disease ecology, and
323 agree that researchers interested in exploring the use of such methods should carefully consider
324 model assumptions. Toward that end, we briefly derive the simplest form of IV to illustrate to
325 potential users what is under the hood of the IV approach and how the exclusion and relevance
326 assumptions function in this technique.

327

328 *Deriving the IV Estimator:* To keep it as intuitive as possible, let us assume a bivariate regression
329 of the form,

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331
$$y_i = \alpha + \beta x_i + \varepsilon_i$$

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333 Where y_i is the outcome variable (e.g., malaria incidence) for observation (e.g., municipality) i ,
334 x_i is the endogenous explanatory variable (e.g., deforestation), ε_i is the error term, α is the
335 intercept, and β is the coefficient of interest.

336

337 To derive the IV estimator, we can take the covariance of each side of equation 1 with respect to
338 the instrument, z_i :

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340 $cov(z_i, y_i) = cov(z_i, \alpha) + cov(z_i, \beta x_i) + cov(z_i, \varepsilon_i)$ 2

341

342 $= 0 + \beta cov(z_i, x_i) + cov(z_i, \varepsilon_i)$ 3

343

344 Since α is a constant, and the covariance of a variable with a constant is 0, the first term drops

345 out. Similarly, because β is a constant, it can be removed from the covariance. The exclusion

346 assumption of IV is that the instrument (z_i) only affects the outcome through changes in the

347 endogenous variable (x_i), which is more formally written as $cov(z_i, \varepsilon_i) = 0$. Thus with basic

348 rearranging, we have derived the IV estimator (β_{IV}),

349

350
$$\beta_{IV} = \frac{cov(z_i, y_i)}{cov(z_i, x_i)}$$

351

352 *Consistency of IV:* If we then want to illustrate that the IV estimator is consistent—in other

353 words, as the sample size gets larger and larger the distribution of the estimator converges to the

354 true parameter value—we can plug the right-hand side of equation 1 into y_i in equation 4. We

355 substitute β_{IV} with $\widehat{\beta}_{IV}$ since we are considering whether the estimated slope from an IV

356 converges in probability to the true slope β .

357

358
$$plim \widehat{\beta}_{IV} = \frac{cov(z_i, \alpha + \beta x_i + \varepsilon_i)}{cov(z_i, x_i)}$$

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360 Following a similar logic as with equation 3, equation 5 becomes:

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362
$$plim \widehat{\beta}_{IV} = \frac{\beta cov(z_i, x_i)}{cov(z_i, x_i)} + \frac{cov(z_i, \varepsilon_i)}{cov(z_i, x_i)}.$$

363

364 From equation 6, the second assumption of IV becomes evident. The second assumption is the
365 relevance assumption, or that the instrument must be statistically associated with the endogenous
366 variable (x_i). As can be seen in equation 6, this means, in mathematical terms,

367 $cov(z_i, x_i) \neq 0$. Covariance does not imply a direction to the relationship, whether AOD (our
368 instrument) determines deforestation or deforestation determines AOD (or neither) is irrelevant,
369 as it is the covariance between the two that is important.

370

371 By these two assumptions of IV, that $cov(z_i, \varepsilon_i) = 0$ and $cov(z_i, x_i) \neq 0$, equation 6
372 simplifies to $plim \widehat{\beta}_{IV} = \beta$, illustrating IV is a consistent estimator of the true relationship.

373

374 *Weak Instruments*: Equation 6 also illustrates another important aspect when considering the
375 application of instrumental variables, and that is a problem known as “weak instruments.” The
376 problem occurs if the exclusion criteria, $cov(z_i, \varepsilon_i) = 0$, fails. Based on the relationship
377 between covariance and correlation (namely, $cov(x, y) = corr(x, y) * \sigma_x \sigma_y$ where σ is the standard
378 deviation of each variable) and assuming $cov(z_i, x_i) \neq 0$, we can rewrite equation 6 to
379 illustrate the problem (omitting subscripts for simplicity).

380

381
$$plim \widehat{\beta}_{IV} = \beta + \frac{corr(z, \varepsilon) * \sigma_z \sigma_\varepsilon}{corr(z, x) * \sigma_z \sigma_x} = \beta + \frac{corr(z, \varepsilon) * \sigma_\varepsilon}{corr(z, x) * \sigma_x}.$$

382

383 If there is a small correlation between the instrument and the error, the last term in equation 7

384 does not drop out and the IV estimator is inconsistent ($plim \widehat{\beta}_{IV} \neq \beta$). If $corr(z, \varepsilon)$ is just slightly

385 different from zero and $corr(z, x)$ is much different than zero, the last term is of minimal

386 influence. However, if the instrument is only weakly correlated with the endogenous covariate,

387 the last term of equation 7 can become large. In practice, weak instruments can cause the IV

388 estimator to be severely biased. Since there is no test to validate the exclusion criteria, the

389 strength of the relationship between the instrument and the endogenous variable is very

390 important in practice, and can be formally tested²⁵ as in the supplementary material from

391 MacDonald and Mordecai.⁴

392

393 *Conclusion:* Understanding the effects of environmental change on infectious disease

394 transmission—from diseases long endemic to the tropics like malaria, to novel emerging

395 pathogens we have yet to discover like SARS-COV-2—is of fundamental and increasing

396 importance. In these complex socio-ecological systems that are difficult to study experimentally,

397 emerging data sources (e.g., high spatio-temporal resolution earth observation data) and causal

398 inference methods (e.g., IV estimation) represent one methodological approach that can help us

399 achieve such clearer understanding.

400

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411 Author Contact Information:

412 Andrew J. MacDonald: Bren School of Environmental Science and Management, University of
413 California, Santa Barbara, CA 93106-5131; andy.j.macdon@gmail.com

414 Erin A. Mordecai: Department of Biology, Stanford University, Stanford, CA 94305;

415 emordeca@stanford.edu

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