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Insights from Application of a Hierarchical Spatio-Temporal Model to an Intensive Urban Black Carbon Monitoring Dataset

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Abstract

Existing regulatory pollutant monitoring networks rely on a small number of centrally located measurement sites that are purposefully sited away from major emission sources. While informative of general air quality trends regionally, these networks often do not fully capture the local variability of air pollution exposure within a community. Recent technological advancements have reduced the cost of sensors, allowing air quality monitoring campaigns with high spatial resolution. The 100×100 black carbon (BC) monitoring network deployed 100 low-cost BC sensors across the 15 km² West Oakland, CA community for 100 days in the summer of 2017, producing a nearly continuous site-specific time series of BC concentrations which we aggregated to one-hour averages. Leveraging this dataset, we employed a hierarchical spatio-temporal model to accurately predict local spatio-temporal concentration patterns throughout West Oakland, at locations without monitors (average cross-validated hourly temporal $R^2=0.60$). Using our model, we identified spatially varying temporal pollution patterns associated with small-scale geographic features and proximity to local sources. In a sub-sampling analysis, we demonstrated that fine scale predictions of nearly comparable accuracy can be obtained with our

31 modeling approach by using ~30% of the 100x100 BC network supplemented by a shorter-term
32 high-density campaign.

33

34 **1 Introduction**

35 Short-term and long-term exposure to particulate air pollution, including black carbon
36 (BC), is associated with adverse health effects¹. Studies of short-term pollutant-health
37 associations still often rely on centrally located regulatory monitors to estimate pollutant
38 exposure for each study participant in the region^{2,3}. However, concentrations can vary widely
39 across a given area, such that a single measurement may not best describe population exposures
40 everywhere, leading to possible biases in the estimates of health effects or identification of those
41 most at risk. Our objective is to predict the spatially varying temporal patterns of BC
42 concentrations in West Oakland during the summer months of 2017, a time that corresponds with
43 intensive air pollutant monitoring in the area. Such predictions are of significant interest for use
44 in a wide variety of applications, including epidemiological studies, as they allow researchers to
45 calculate individual-specific short-term and long-term exposures based on finely resolved
46 location information. From the perspective of air quality management and emissions control,
47 more targeted management strategies such as regulatory agencies identifying times of day when
48 areas are most affected by pollution might be possible. Vulnerable residents may be advised of
49 times of day or week when they should be most cautious about spending time outdoors.

50 A number of approaches have been employed to predict intra-urban air pollution levels
51 based on ground-level monitoring data^{4,5}. Land-use regression (LUR) fits the exposure surface to
52 a linear model with a large number of geographic information system (GIS) covariates⁶, often
53 using a combination of scientific and statistical learning techniques to reduce dimension of the
54 covariate space⁷⁻¹¹. Kriging models use a spatial random effect to construct a smooth prediction
55 surface that predicts concentrations at unmonitored locations using an optimal weighted sum of
56 nearby observations¹². Researchers often combine LUR and kriging in a universal kriging (UK)
57 that optimally combines regression and smoothing to improve prediction accuracy^{7,13,14}. Recent
58 advances in mobile monitoring technology and implementation have made comprehensive data-
59 only spatial mapping an option in some areas, and in some cases LUR models have been trained
60 on mobile monitoring data¹⁵⁻¹⁸. Spatiotemporal air pollution models combine LUR or UK with
61 models for spatially varying temporal trends to accommodate temporally sparse data and to

62 predict both spatial and temporal air pollution patterns in an urban environment. These models
63 have been successful, especially at temporal scales of 1-2 weeks, although data availability is a
64 limiting constraint¹⁹⁻²¹.

65 Current reference-grade BC monitors cost ~\$25,000, which places an economic
66 constraint on the number of monitors that can be deployed in a network. To address this barrier,
67 Caubel et al.²² developed a new, low-cost BC sensor with similar precision and accuracy as
68 existing commercial aethalometers based on the filter-based absorption photometry technology.
69 With these low-cost BC sensors, it was possible to monitor BC concentrations with much greater
70 spatial resolution by creating a dense sensor network across the community of West Oakland,
71 California, a neighborhood surrounded by major highways and close to regional seaport and rail
72 facilities. As part of the West Oakland Community Air Quality Study, the 100×100 Network
73 deployed BC sensors across 100 locations in this community for 100 days from May 19 to
74 August 26, 2017²³. This measurement campaign produced a rich dataset of highly resolved BC
75 concentrations in both space and time that we leverage in our modeling effort.

76 Maintaining a large network of sensors can be difficult in practice, operationally
77 intensive, and susceptible to equipment failure or loss. The 100×100 BC Network achieved an
78 84% success rate at capturing valid hourly BC concentration measurements²³. By the end of the
79 100 days of deployment, over 30 samplers were no longer operating, which enabled us to assess
80 prediction accuracy of our spatio-temporal model subject to realistic maintenance and reliability
81 constraints. A notable strength of our spatio-temporal model is its ability to leverage spatio-
82 temporally sparse observations to improve predictions over the entire modeling period, as was
83 similarly observed with the ability of a spatial only universal kriging model to leverage mobile
84 monitoring data to predict spatial patterns BC concentrations across West Oakland¹⁵.

85 In this paper, we use a spatio-temporal model to predict fine scale variation in BC
86 concentrations across West Oakland, CA during part of the 100×100 Network monitoring period.
87 We also subsample our dataset to evaluate how prediction accuracy is affected by monitor
88 dropout patterns and to evaluate the implications of using a less dense monitoring network,
89 possibly supplemented with a dense network over a shorter period.

90

91 **2 Methods**

92 **2.1 Monitoring Data**

93 The measurement sites, data quality assurance methods, and temporal and spatial
94 variability of observed BC concentrations by different land use types in the 100×100 BC
95 Network have been reported previously²³ and are summarized here. The 100×100 Network
96 deployed 100 monitors throughout West Oakland outside of homes, local businesses, community
97 organizations, and schools, and adjacent to the Port of Oakland. To verify sensor precision, nine
98 network sites had sensors collocated in pairs and all sensors were calibrated based on 2–7 days
99 operation collocated with a commercial BC instrument. Hourly average BC concentrations were
100 calculated by averaging validated 1-minute averages, after correction for a filter loading artifact
101 and errors in sample flow rate measurements. Due to the above-described equipment failure
102 issues, an increasing number of sites were left unmonitored over the course of the 100 days.
103 While 87% of potential hourly BC concentration measurements were successfully collected
104 during the first 74 days of monitoring from mid-May to July, only 66% of were recorded during
105 the last 26 days of monitoring in August. We primarily focused on analyzing data from June and
106 July due to concerns that pollutant patterns at the end of May might have been qualitatively
107 different from the summer seasonal patterns observed in June and July and due to data
108 completeness limitations in August. As described later in this section, we utilized the pattern of
109 data missingness (i.e., missing observations) in August to help assess how well our modeling
110 approach would perform in a scenario with significant monitor dropout.

111

112 **2.2 Hierarchical Spatio-Temporal Model**

113 We used a hierarchical spatio-temporal model to predict time varying concentrations of
114 BC at unmeasured locations in West Oakland^{19,20,24,25}. Since the data tend to have heavy right
115 tails and appear log-normally distributed, all modeling was done on the log-transformed scale to
116 improve model fit. Prediction accuracy evaluation statistics were calculated on the back-
117 transformed concentration scale. The spatio-temporal field is conceptualized as being comprised
118 of location-specific temporal trends, where the trend at each location is the sum of the area-wide
119 average (i.e., a time-series that is spatially constant across the domain) and a linear combination
120 of two temporal basis functions. We included two temporal basis functions, $f_1(t)$ and
121 $f_2(t)$, which were derived from the 100×100 Network data by first filling in missing values using
122 an expectation-maximization-like approach²⁶ and applying cubic smoothing splines to the first

123 two singular vectors. Preliminary exploratory analysis showed that including two temporal basis
124 functions balanced model fit with interpretability of the temporal trends.

125 The full hierarchical spatio-temporal model can be written as

$$\log(Y(s, t)) = \mu(s, t) + v(s, t)$$

126 where $\log(Y(s, t))$ is the log-concentration of black carbon at site s for time t and

$$\mu(s, t) = \eta(t) + \beta_0(s) + \beta_1(s) f_1(t) + \beta_2(s) f_2(t).$$

127 The location-specific coefficients for the temporal basis functions (including the intercept) are
128 $\beta_i(s)$ for $i = 0, 1, 2$, and $\eta(t)$ represents the area-wide average derived by averaging all
129 monitoring data at each time t . The spatial structure of each $\beta_i(s)$ is modeled by universal
130 kriging, with regression on spatial covariates X_i with coefficients α_i in the mean model and
131 normally distributed residuals with exponential covariance structure $\Sigma(\theta_i)$ that accounts for
132 spatial correlation, i.e., $\beta_i(s) \sim N(X_i\alpha_i, \Sigma(\theta_{\beta_i}))$. The β_i -fields are independent of each other, and
133 the exponential covariance function is parameterized by $\theta_i = (\rho_i, \sigma_i^2, \tau_i^2)$ with correlation range
134 ρ_i , partial sill σ_i^2 , and nugget τ_i^2 . Finally, $v(s, t)$ represents temporally independent spatial
135 residual fields with exponential correlation structures that account for short-term events such as
136 meteorology affecting large subsets of the domain at any given time.

137 We calculated over 900 geographic information system (GIS) covariates to use in the
138 model including proximity measures (distance to nearest major road, intersection, truck route,
139 railway, railyard, coastline, airport, and port) and buffer measures (major road length, truck route
140 length, land-use category, long-term vegetation index, population density, and emission sources).
141 Following¹⁹, GIS covariates with little to no variation or those that are highly skewed were
142 removed from the modeling process. Specifically, any variables with (a) missing values, (b)
143 >80% identical values, or (c) >2% more than 5 standard deviations (SD) from the mean were
144 removed. Additionally, (d) any variables that describing land use at distances >5 km were
145 removed, since the area of interest is only ~15 km². Because of the high dimensionality of the
146 geographic covariates, we used principal component analysis (PCA) on the GIS covariates and
147 selected the first two principal components to use as spatial covariates in our model.

148 This model is essentially the same model that was developed and applied in the MESA
149 Air study¹⁹ at the hourly rather than two-week timescale, with two significant changes that we
150 made based on preliminary analyses of this dataset. One is that we explicitly include the area-
151 wide average $\eta(t)$ in our model since there is a very strong shared temporal pattern at the hourly

152 time scale. The other change is that we use PCA for dimension reduction of GIS covariates
153 rather than partial least squares (PLS).

154

155 **2.3 Model Estimation and Evaluation**

156 Parameter estimation was performed using the SpatioTemporal package in R and
157 optimization by restricted maximum likelihood (REML)²⁵. Model accuracy was assessed by
158 leave-one-site-out cross-validation. Let $Y(s, t)$ be the observed BC concentration at location s at
159 time t and $\hat{Y}(s, t)$ the associated cross-validated predictions. At each location s , we calculated a
160 measure of cross-validated prediction accuracy as follows:

$$161 \quad R_{CV}^2 = \max(0, 1 - MSE/Var_{obs})$$

162 where

$$163 \quad MSE = \frac{1}{t} \sum_{i=1}^t (Y(s, t_i) - \hat{Y}(s, t_i))^2$$

$$164 \quad Var_{obs} = \frac{1}{T} \sum_{i=1}^T \left(Y(s, t_i) - \frac{1}{T} \sum_{j=1}^T Y(s, t_j) \right)^2$$

165 Like the squared Pearson correlation coefficient, a value of 1 denotes perfect correlation.
166 Additionally, the measure penalizes for bias and scaling errors whereas the Pearson
167 correlation coefficient does not.

168 We calculated these measures for two timescales, namely the hourly values over the
169 entire period of June and July and the consolidated calendar week hourly values averaged over
170 all weeks in June and July; i.e., for the latter, we collapsed the observed and predicted time series
171 at each location to single average values for each hour of each day of the week (see bottom-right
172 panel of Figure 1).

173 Due to the computational burden associated with solving the nonlinear REML
174 optimization problem²⁵, we carried out this step only once on all of the data and used the
175 resulting covariance parameter estimates to compute leave-one-out cross-validated R^2 for each
176 site individually. There is minimal potential for overfitting because only the covariance
177 (smoothing) parameters were estimated outside of the cross-validation loop, while regression
178 coefficients were re-estimated for each cross-validation set. The estimated parameter values are
179 reported in Table S-1.

180

181 **2.4 Simulating Less Intensive Monitoring**

182 Significant data collection dropout occurred during the month of August, so we used only
183 the June and July data for model fitting and evaluation. We conducted a sensitivity analysis to
184 assess the impact on prediction accuracy if we had elected to also include data from August and
185 make predictions during that time period. We started with the fully observed 100×100 dataset in
186 June and July and then created a missingness pattern in July that matched the observed
187 missingness pattern in August. Using this version of the dataset to represent a monitoring
188 campaign with dropouts, we fit the spatio-temporal model and estimated the observed pollutant
189 concentrations at each location using leave-one-out cross-validation. We compared cross-
190 validated prediction accuracy in July using the model with artificially created missingness
191 against the original model, which used all available observations to understand how realistic
192 long-term maintenance and logistical issues would affect prediction accuracy.

193 We also evaluated how the number of continuously operating monitors and their
194 placement impact predictive accuracy of the spatio-temporal model by systematically
195 subsampling our dataset to simulate estimates from a smaller network. We considered having 5,
196 10, 20, and 30 continuously operating monitors and sampled these monitors in three different
197 ways:

- 198 1. Simple random sampling (“Random”): Randomly sampled k monitors with equal
199 probability.
- 200 2. Stratified random sampling by GIS covariates (“GIS Covariates”): Clustered monitors at
201 locations with similar local characteristics by using principal components of their GIS
202 covariates in a k -means algorithm and then from each of the k clusters, randomly
203 selected a monitor with equal probability.
- 204 3. Stratified random sampling by location (“Space Filling”): Clustered monitors spatially by
205 first using a space-filling design to select k centers and then assigned monitors to clusters
206 by distance. Monitors are then randomly selected from each of the k clusters with equal
207 probability.

208

209 For each of these subsampling approaches, we considered one scenario where only the smaller
210 number of continuously functioning monitors is available (“Long-Term Monitors Only”) and

211 another where these monitors were supplemented by a larger number of monitors in June only by
212 including all available data from the 100×100 campaign in June (“Long-Term Monitors +
213 Supplementation in June”). For each of these, we evaluated prediction accuracy of the spatio-
214 temporal model on July data at all 100×100 locations.

215

216 **3 Results**

217 **3.1 Monitoring Data**

218 The spatial and temporal patterns of BC in West Oakland are described extensively
219 elsewhere^{15,23,27}. Briefly, key spatial features that are apparent in the 100x100 and mobile
220 monitoring data include elevated levels near Interstate 880, along the major truck routes of the
221 Port of Oakland, and near industrial clusters, with lower areas within predominantly residential
222 zones. Temporal patterns vary across the domain and differ between weekdays and weekends.
223 Weekday patterns broadly are characterized by a peak concentration during the morning rush
224 hour, declining levels over the afternoon owing to increased atmospheric mixing, a less
225 prominent evening peak, and lower levels in the late night and very early morning.

226

227 **3.2 Model Description**

228 We first look at components of the spatio-temporal models individually to describe the
229 systematically varying spatial and temporal patterns. The area-wide average is shown in red, and
230 two temporal trend functions are shown in green and blue in Figure 1. The top panel shows every
231 hour throughout June and July. The bottom four panels show more interpretable versions
232 consolidated to four different time scales: average diurnal 24-hour weekday, average diurnal 24-
233 hour weekend day, average for each day of the week, and average for each hour during a typical
234 week. In each panel, the red line shows the absolute value of the area-wide average
235 concentration, and the green and blue lines show the relative differences compared to the area-
236 wide average, with values < 0 equal to lower than area-wide average concentrations and values $>$
237 0 equal to greater than area-wide average concentrations.

238 The area-wide average concentration rises throughout the morning until 9 AM on
239 weekdays, then slowly decreases throughout the day reaching its minimum value around 2 AM.
240 On the weekends, there is a different pattern, with low concentrations throughout the day that
241 slightly increase around 6 PM. On average, concentrations are lower on weekends and higher on

242 weekdays, with some variation in diurnal patterns between days of the week. The first temporal
243 basis function describes a pattern where concentrations are higher than the area-wide average
244 from early morning until late in the afternoon on weekdays, with a peak around 9 AM.
245 Concentrations at locations with a positive coefficient for the first temporal basis function are
246 slightly below the area-wide average after 5 PM until midnight on weekdays, and well below the
247 area-wide average all day on weekends. The second temporal basis function also shows an
248 increase on weekdays but shifted later in the day, where concentrations are higher than the area-
249 wide average from 7 AM until 8 PM. The above area-wide average concentration increases
250 throughout the morning until about 1 PM, then slowly decreases over the afternoon. After 8 PM,
251 this temporal basis function shows concentrations as below the area-wide average until the
252 morning, with a minimum around 1 AM. On weekends, concentrations are higher than the area-
253 wide average from 1 PM to 2 AM and are lower than the area-wide average during the other
254 hours of the day.

255 Empirical estimates of the β -fields at locations with monitors are shown in Figure 2
256 along with their predicted values across West Oakland, and supplementary Figure S-1 shows
257 cross-validated predictions of the empirical values of the β -fields. Locations in the southern part
258 of West Oakland, especially the southeast corner downwind of major freeways and port
259 activities, are associated with higher coefficient estimates for the β_0 -field. This suggests that
260 these areas have higher concentrations in general compared to the area-wide average. The β_0 -
261 field models the temporal average of log-transformed BC concentrations, so the difference in
262 predicted values between the highest and lowest areas of the map in Figure 2 corresponds to a
263 seven-fold difference in absolute concentrations.

264 Areas where the β_1 -field are highest correspond to sites characterized as *port* and *truck*
265 *routes* in Caubel et al. (2019)²³, and locations with lower coefficients align with locations
266 described as being *residential* or *upwind* in Caubel et al. (2019)²³. This is consistent with the fact
267 that the first temporal basis function indicates concentrations higher than the area average during
268 early morning hours, especially on weekdays when the port area is most active. The β_2 -field is
269 generally higher at industrial and residential sites in the northeast corner of West Oakland and
270 lower in residential sites in the southwest corner of West Oakland. It is also high around the
271 northern section of I-880 that feeds into I-80. The second temporal basis function is associated
272 with higher than average concentrations starting in late-morning on weekdays and a modestly

273 decreasing trend later in the week, potentially influenced by industrial sources during weekday
274 business hours.

275

276 **3.3 Prediction Accuracy**

277 To demonstrate how the spatio-temporal model can improve fine-scale variation
278 estimates across the region using temporal basis functions, we first compare the cross-validated
279 predicted trends and the area-wide average to the actual observed trends for two example sites.
280 Figure 3 shows the case study locations for this analysis, residential monitor 35 (R35) and truck
281 route monitor 75 (TR75). While these monitors are located just three blocks apart and near the
282 same arterial road, the average concentrations and temporal patterns at these two sites are
283 different; monitored concentrations at TR75 are up to $0.5 \mu\text{g m}^{-3}$ higher than the area-wide
284 average, whereas monitored concentrations at R35 are similar to the area-wide average but with
285 a lower weekday morning peak. Predicting concentrations by the area-wide average at these
286 locations results in poor prediction accuracy, as measured by temporal R_{CV}^2 on the consolidated
287 hour of week time scale (0.00 and 0.19 for R35 and TR75, respectively). Predictions from the
288 spatio-temporal model at these two locations are much more accurate, with temporal R_{CV}^2 values
289 on the consolidated hour of week time scale of 0.89 and 0.70 for R35 and TR75, respectively.

290 Overall, prediction accuracy for the spatio-temporal model varies across sites and
291 depending on the temporal scale (Figure 4). Overall, the prediction accuracy is fair to good for
292 locations near the port, with many of the most accurately predicted sites located in the northwest
293 section of West Oakland. The mean temporal R_{CV}^2 is 0.60 for all hourly measurements over June
294 and July and 0.58 for the consolidated hour of week time scale. These compare favorably to the
295 corresponding values of 0.40 and 0.47 for the hourly and consolidated hour of week metrics,
296 respectively, from using just the area-wide average to predict concentrations at all sites rather
297 than predictions from the spatio-temporal model. Using a single well-sited monitoring location
298 to represent the entire area would perform similarly to the area-wide average. On a site-by-site
299 basis, using the spatio-temporal model yields a noticeable overall improvement in prediction
300 accuracy at 72% of the observed sites over the area-wide average, showing the spatio-temporal
301 model can capture fine scale gradients that would not be captured by the area-wide average.

302

303 **3.4 Model Performance with Monitor Dropouts**

304 To evaluate the effect of monitor dropouts like those experienced in August, a 30%
305 dropout rate was simulated for the month of July. Predictions from this “masked” July dataset
306 were compared to the full model run with all available June and July data. Using the full dataset
307 shows a small, but noticeable improvement (Figure S-2, Figure S-3). The magnitude of this
308 improvement is small and suggests that the effect of monitor dropouts, as observed in August, do
309 not significantly impact prediction accuracy when using the spatio-temporal model. Based on
310 these results, we expect that if the spatiotemporal model were fit using the entire 100×100
311 dataset (June–August), predictions for the August period would not be substantially less accurate
312 than those for June and July, despite the higher missingness in monitoring data in August.

313

314 **3.5 Systematic Subsampling Results**

315 Results from the subsampling studies are shown in Figure 5, comparing mean temporal
316 R_{CV}^2 for the consolidated hour of week time scale in July. It does not appear that the method used
317 to subsample monitor locations significantly impacts performance of the model. For a small
318 number of continuously operating monitors such as 5 or 10, the spatio-temporal model
319 predictions are less accurate than predicting based only on only the area-wide average, while
320 with 20 or more continuously operating monitors the spatio-temporal model predictions
321 consistently improve on the area-wide average. Adding more intensive short-term monitoring in
322 June only (i.e., including all 100×100 monitors during that period) leads to an improvement
323 regardless of the number of continuously operating monitors. This indicates that the model
324 leverages data from the more intensive monitoring campaign in June to accurately predict
325 concentrations in July when there were fewer monitors. Overall, our subsampling study suggests
326 that 30 continuous monitors supplemented by a short-term, high-density monitoring campaign
327 would allow us to construct a spatio-temporal model with prediction accuracy approaching that
328 obtained with the full dataset. Performance is similar when the 30 monitors are selected at
329 random compared to selecting them to representatively span the GIS covariate distribution or
330 spatial distribution of the full dataset.

331

332 **4 Discussion**

333 By using a flexible, hierarchical spatio-temporal model with monitoring data from the
334 100×100 BC network, we were able to capture fine-scale differences in BC concentrations across

335 West Oakland. Our model predictions are significantly more accurate than what could be
336 obtained by treating the pollution surface as spatially uniform and predicting the time series at
337 each location based on data from a single well-placed regulatory monitor that might be found in
338 a typical urban area, (mean temporal R^2 s 0.60 and 0.40, respectively). We are aware of one other
339 paper that attempted to model urban BC concentrations on an hourly timescale²¹, with mixed
340 success. A direct comparison is not possible because we report spatially varying temporal
341 prediction accuracy, while Dons et al. (2013)²¹ report spatial prediction accuracy for one hour
342 averages. While comparisons with spatiotemporal predictions of other pollutants would also be
343 informative, we are not aware of papers that have reported spatially varying temporal prediction
344 accuracy as in our study. In addition to prediction accuracy, an important strength of the spatio-
345 temporal model is its ability to help identify interpretable spatial patterns in temporal variation.
346 For example, locations with larger positive values of the β_1 coefficient for the first temporal
347 trend have relatively high weekday morning concentrations and tend to include sites identified in
348 Caubel et al. (2019)²³ as associated with port activity. Similarly, locations with larger positive
349 values of the β_2 coefficient for the second temporal trend have relatively high concentrations in
350 the mid-morning through afternoon hours on weekdays and afternoon hours on weekend. These
351 areas include industrial and residential sites in the northeast corner of West Oakland. While our
352 model is specific to West Oakland and the 100×100 campaign, we expect that a similar approach
353 would successful in other locales with similar intensive monitoring data.

354 Recognizing that the 100×100 campaign was a unique opportunity to conduct intensive
355 spatiotemporal monitoring in an urban region, it is important to understand whether similar
356 modeling results can be obtained with less intensive monitoring. The subsampling simulation
357 demonstrates that it is possible to leverage spatially dense observations from a short-term
358 monitoring campaign to make accurate predictions at a time with more limited spatially sparse
359 monitoring. In practice, this suggests that it is feasible to use a short-term monitoring campaign
360 to improve long term predictions in areas that are not near continuously operated monitors. For
361 example, one could design a future monitoring campaign that includes 30 fixed sites over an
362 entire 12 month time period and either another 70 monitors that are only deployed for a few
363 shorter 1 month periods or another 10–20 monitors that are rotated through additional locations
364 for 1 month at a time. Either design would be less operationally expensive than continuous

365 deployment of 100 monitors for the full period and our results suggest they could result in
366 similar model prediction accuracy across the full region.

367 Additional gains in prediction accuracy might be possible if a future campaign takes
368 advantage of optimal monitor location-allocation strategies²⁸, potentially modified to
369 accommodate temporally varying network size. A recently developed algorithm for real-time
370 spatiotemporal monitor allocation²⁹ may be helpful, although with such an approach it will be
371 important to consider the impact of preferential spatial sampling on inference^{30,31}. If the air
372 pollution surface is to be used as the exposure in an epidemiological analysis, then consideration
373 should also be given to compatibility between monitor and main study locations³². Another
374 promising strategy is mobile monitoring, which can cover a much larger number of locations
375 than fixed monitoring over an extended period of time, albeit with temporally sparse coverage¹⁶.
376 Data from mobile monitoring campaigns has been used successfully to fit spatial air pollution
377 models^{15,17,33}. It may be possible to incorporate mobile monitoring data in a spatiotemporal
378 model like the one described here, especially if it can be calibrated and included in a model with
379 continuous fixed site monitoring at a modest number of locations²⁷.

380

381 **4 Conclusion**

382 We have utilized a geostatistical spatiotemporal model applied to data from to 100×100
383 campaign to predict hourly BC concentrations at all locations in West Oakland during the
384 summer of 2017. These predictions provide insights into the complex spatially varying temporal
385 air pollution trends and how they relate to local sources and neighborhood factors. Our
386 subsampling analysis demonstrates that this modeling strategy can be employed to get similar
387 prediction accuracy even with a less intense monitoring campaign in which some monitors are in
388 service for only part of the modeled period. Future research is needed to determine optimal
389 monitor placement strategies that will make it feasible to develop similar high resolution
390 spatiotemporal air pollution predictions in other locations and over longer time periods.

391

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400

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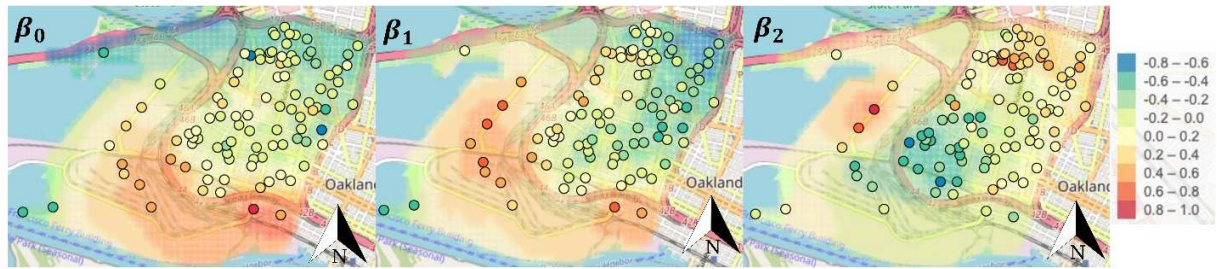
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517 *Figure 1: Systematic time trends from the spatio-temporal model: The area-wide average and both temporal basis*
 518 *functions, $f_1(t)$ and $f_2(t)$, shown on different time scales, (top) hourly scale over the full monitoring period, (middle left)*
 519 *average diurnal trend over weekdays, (middle right) average diurnal trend over weekends, (bottom left) average calendar week*
 520 *summarized by daily averages, and (bottom right) average calendar week summarized by hourly averages. In each panel, the red*
 521 *line shows the absolute value of the area-wide average concentration, and the blue and green lines show the relative differences*
 522 *compared to the area-wide average that are multiplied by the site-specific values of Beta1 and Beta2, respectively.*

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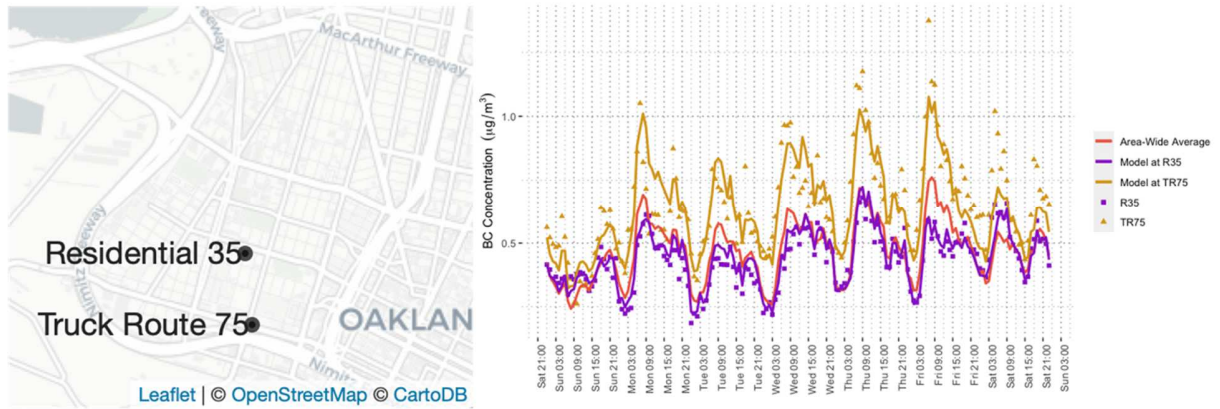
527 *Figure 2 β -fields from the spatio-temporal model, where points indicate estimated values at observed locations. Beta 0*
 528 *corresponds to the difference in BC concentration from the area-wide average at each location. Beta 1 and Beta 2 are*
 529 *coefficients for the respective temporal basis function shown in Figure 1. Note that the relative contribution of the area-wide*
 530 *average and the two temporal basis functions to the overall time series at each location differs across the domain. Coefficients*
 531 *near zero reflect locations where a specific temporal pattern has relatively little influence on the concentration time-series,*
 532 *where coefficients with higher absolute value reflect locales where the relative contribution of a temporal pattern is higher.*

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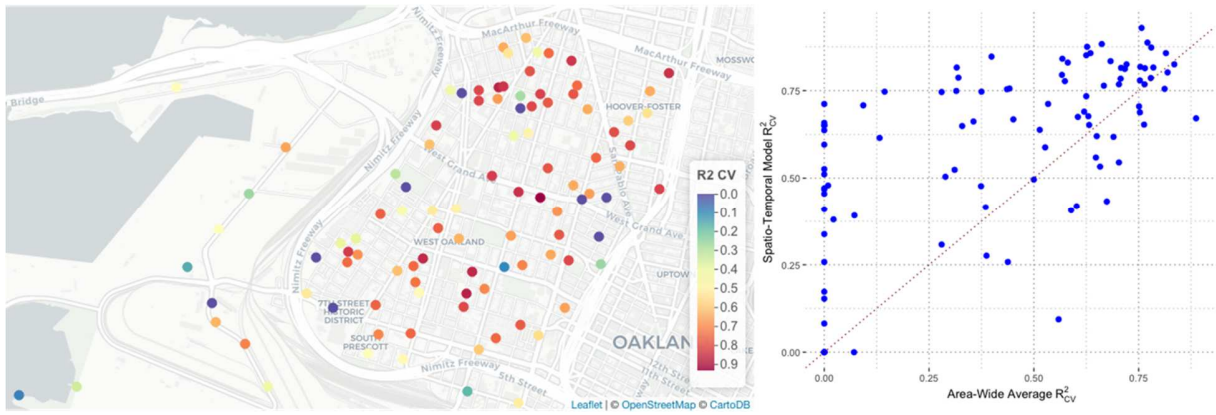
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Figure 3: BC concentrations at example community sites Residential 35 (R35) and Truck Route 75 (TR75). Left: Locations of Residential 35 and Truck Route 75, Right: Points show observed averages for each hour over a calendar week, while lines show estimates using the spatio-temporal model against the area-wide average.

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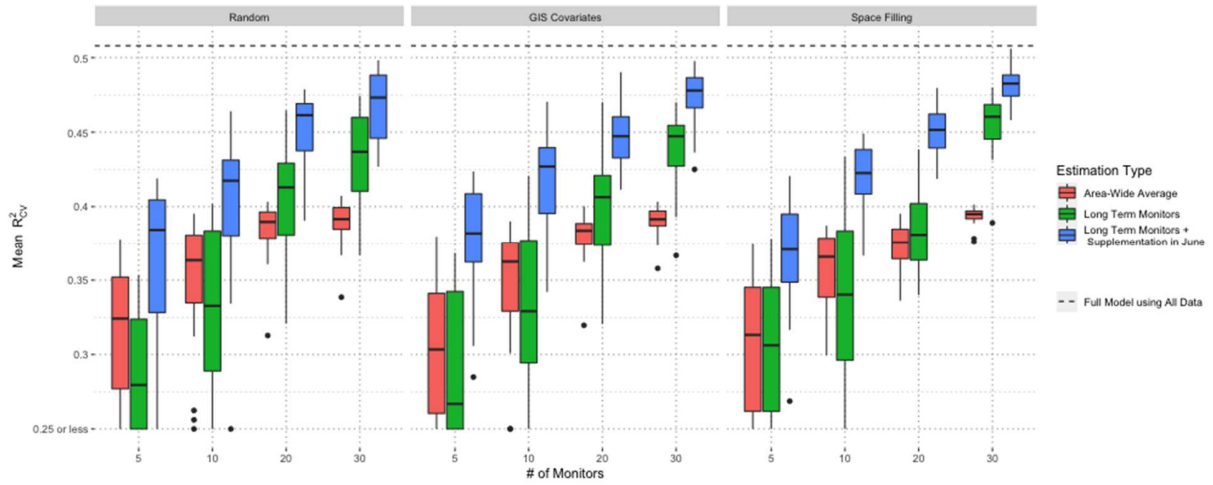


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Figure 4: Prediction accuracy for each site from the 100x100 Network using the spatio-temporal model and the area-wide average. Reported R_{CV}^2 is for a typical week. Left: R_{CV}^2 for each site from the 100x100 Network. Right: R_{CV}^2 compared at each site using the spatio-temporal model and the area-wide average.

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552 *Figure 5: Cross-validated prediction accuracy on July data for each sub-sampled set of monitors. Area-wide average (orange)*
553 *show prediction accuracy if no modeling was done and the surface was assumed to be constant across West Oakland. Long term*
554 *monitors (Green) shows prediction accuracy from the spatio-temporal model without the supplementary monitoring in June.*
555 *Long term monitors + supplement in June (Blue) shows prediction accuracy if short-term sampling is done at a large number of*
556 *sites in June. The dotted line represents the best possible spatio-temporal model, where all data (including July) is used.*