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# Unfolding Quantum Computer Readout Noise

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In the current era of noisy intermediate-scale quantum (NISQ) computers, noisy qubits can result in biased results for early quantum algorithm applications. This is a significant challenge for interpreting results from quantum computer simulations for quantum chemistry, nuclear physics, high energy physics, and other emerging scientific applications. An important class of qubit errors are readout errors. The most basic method to correct readout errors is matrix inversion, using a response matrix built from simple operations to probe the rate of transitions from known initial quantum states to readout outcomes. One challenge with inverting matrices with large off-diagonal components is that the results are sensitive to statistical fluctuations. This challenge is familiar to high energy physics, where prior-independent regularized matrix inversion techniques ('unfolding') have been developed for years to correct for acceptance and detector effects when performing differential cross section measurements. We study various unfolding methods in the context of universal gate-based quantum computers with the goal of connecting the fields of quantum information science and high energy physics and providing a reference for future work. The method known as iterative Bayesian unfolding is shown to avoid pathologies from commonly used matrix inversion and least squares methods.

### I. INTRODUCTION

While quantum algorithms are promising techniques for a variety of scientific and industrial applications, current challenges limit their immediate applicability. One significant limitation is the rate of errors and decoherence in noisy intermediate-scale quantum (NISQ) computers [1]. Mitigating errors is hard in general because quantum bits ('qubits') cannot be copied [2–4]. An important family of errors are readout errors. They typically arise from two sources: (1) measurement times are significant in comparison to decoherence times and thus a qubit in the  $|1\rangle$  state can decay to the  $|0\rangle$  state during a measurement, and (2) probability distributions of measured physical quantities that correspond to the  $|0\rangle$  and  $|1\rangle$  states have overlapping support and there is a small probability of measuring the opposite value. The goal of this paper is to investigate methods for correcting these readout errors.

Correcting measured histograms for the effects of a detector has a rich history in image processing, astronomy, high energy physics (HEP), and beyond. In the latter, the histograms represent binned differential cross sections and the correction is called *unfolding*. Many unfolding algorithms have been proposed and are currently in use by experimental high energy physicists (see, e.g., Ref. [5–7] for reviews). One of the goals of this paper is to introduce these methods and connect them with current algorithms used by the quantum information science community.

Quantum readout error correction can be represented as histogram unfolding where each bin corresponds to one of the possible  $2^{n_{\text{qubit}}}$  configurations where  $n_{\text{qubit}}$  is the number of qubits (Fig. 1). Correcting readout noise is a classical problem (though there has been a proposal to do



FIG. 1. A schematic diagram illustrating the connection between binned differential cross section measurements in high energy physics (left) and interpreting the output of repeated measurements from quantum computers (right).

it with quantum annealing [8]), but relies on calibrations or simulations from quantum hardware. Even though discussions of readout errors appear in many papers (see, e.g., Ref. [9–12]), we are not aware of any dedicated study of unfolding methods for quantum information science (QIS) applications. Furthermore, current QIS methods have pathologies that can be avoided with techniques from HEP.

This paper is organized as follows. Sec. II introduces the unfolding setup. Various techniques are described in Sec. III. The core input for unfolding is the matrix of probabilities relating the true and measured spectra that is discussed in Sec. IV. Representative examples are presented in Sec. V and a description of sources of uncertainty in Sec. VI. The discussion in Sec. VII is followed by a summary and outlook in Sec. VIII.

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### II. THE UNFOLDING CHALLENGE

Let t be a vector that represents the true bin counts before the distortions from detector effects (HEP) or readout noise (QIS). The corresponding measured bin counts are denoted by m. These vectors are related by a response matrix R as m = Rt where  $R_{ij} = \Pr(m = i|t = j)$ . In HEP, the matrix R is usually estimated from detailed detector simulations while in QIS, R is constructed from measurements of computational basis states. The response matrix construction is discussed in Sec. IV.

The most naive unfolding procedure would be to simply invert the matrix R:  $\hat{t}_{matrix} = R^{-1}m$ . However, simple matrix inversion has many known issues. Two main problems are that  $\hat{t}_{matrix}$  can have unphysical entries and that statistical uncertainties in R can be amplified and can result in oscillatory behavior. For example, consider the case

$$R = \begin{pmatrix} 1 - \epsilon & \epsilon \\ \epsilon & 1 - \epsilon \end{pmatrix}, \tag{1}$$

where  $0 < \epsilon < 1/2$ . Then,  $\operatorname{Var}(\hat{t}_{\text{matrix}}) \propto 1/\operatorname{det}(R) = 1/(1-2\epsilon) \to \infty$  as  $\epsilon \to 1/2$ . As a generalization of this example to more bins (from Ref. [13]), consider a response matrix with a symmetric probability of migrating one bin up or down,

$$R = \begin{pmatrix} 1 - \epsilon & \epsilon & 0 & \cdots \\ \epsilon & 1 - 2\epsilon & \epsilon & \cdots \\ 0 & \epsilon & 1 - 2\epsilon & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{pmatrix}.$$
(2)

Unfolding with the above response matrix and  $\epsilon = 25\%$ is presented in Fig. 2. The true bin counts have the Gaussian distribution with a mean of zero and a standard deviation of 3. The values are discretized with 21 uniform unit width bins spanning the interval [-10, 10]. The leftmost bin correspond to the first index in m and t. The indices are monotonically incremented with increasing bin number. The first and last bins include underflow and overflow values, respectively. Due to the symmetry of the migrations, the true and measured distributions are nearly the same. The significant off-diagonal components result in an oscillatory behavior and the statistical uncertainties are also amplified by the limited size of the simulation dataset used to derive the response matrix.

### III. UNFOLDING METHODS

The fact that matrix inversion can result in unphysical outcomes  $(\hat{t}_i < 0 \text{ or } \hat{t}_i > ||t||_1$ , where  $||x||_1 = \sum_i |x_i|$  is the  $L^1$  norm) is often unacceptable. One solution is to find a vector that is as close to  $\hat{t}$  as possible, but with physical entries. This is the method implemented



FIG. 2. A comparison of unfolding techniques for a Gaussian example and the R matrix from Eq. (2). The symbols t and m denote the true and measured probability mass functions, respectively. Simple matrix inversion is represented by  $\hat{t}_{\text{matrix}}$  (see Sec. II). The **ignis** and IBU methods are represented by  $\hat{t}_{\text{ignis}}$  and  $\hat{t}_{\text{IBU}}$ , respectively (see Sec. III). For this example,  $||m||_1 = 10^4$  and R is assumed to be known exactly. The IBU method uses 10 iterations and a uniform prior. The simulation used in this plot is based on standard Python functions and does not use a quantum computer simulator (see instead Fig. 6).

in qiskit-ignis [14, 15], a widely used quantum computation software package. The ignis method solves the optimization problem

$$\hat{t}_{\text{ignis}} = \arg\min_{t': ||t'||_1 = ||m||_1, t'_i > 0} ||Rm - t'||^2.$$
(3)

Note that the norm invariance is also satisfied for simple matrix inversion:  $||\hat{t}_{\text{matrix}}||_1 = ||m||_1$  by construction because  $||Rx||_1 = ||x||_1$  for all x so in particular for  $y = R^{-1}m$ ,  $||Ry||_1 = ||y||_1$  implies that  $||m||_1 = ||R^{-1}m||_1$ . This means that  $\hat{t}_{\text{ignis}} = \hat{t}_{\text{matrix}}$  whenever the latter is non-negative and so  $\hat{t}_{\text{ignis}}$  inherits some of the pathologies of  $\hat{t}_{\text{matrix}}$ .

Three commonly used unfolding methods in HEP<sup>1</sup> are Iterative Bayesian<sup>2</sup> unfolding (IBU) [22] (also known as Richardson-Lucy deconvolution [23, 24]), Singular Value Decomposition (SVD) unfolding [25], and TUnfold [26].

<sup>&</sup>lt;sup>1</sup> There are other less widely used methods such as fully Bayesian unfolding [16] and others [17–21].

 $<sup>^2</sup>$  Even though this method call for the repeated application of Bayes' theorem, this is not a Bayesian method as there is no prior/posterior over possible distributions, only a prior for initialization.



FIG. 3. The set of  $2^{n_{\text{qubit}}}$  calibration circuits.

TUnfold is similar to Eq. (3), but imposes further regularization requirements in order to avoid pathologies from matrix inversion. The SVD approach applies some regularization directly on R before applying matrix inversion. The focus of this paper will be on the widely used IBU method, which avoids fitting and matrix inversion altogether with an iterative approach.

Given a prior truth spectrum  $t_i^0 = \Pr(\text{truth is } i)$ , the IBU technique proceeds according to the equation

$$t_i^{n+1} = \sum_j \Pr(\text{truth is } i|\text{measure } j) \times m_j$$
$$= \sum_j \frac{R_{ji} t_i^n}{\sum_k R_{jk} t_k^n} \times m_j,$$
(4)

where *n* is the iteration number and one iterates a total of *N* times. The advantage of Eq. (4) over simple matrix inversion is that the result is a probability (non-negative and unit measure). The parameters  $t_i^0$  and *N* must be specified ahead of time. A common choice for  $t_i^0$  is the uniform distribution. The number of iterations needed to converge depends on the desired precision, how close  $t_i^0$  is to the final distribution, and the importance of offdiagonal components in *R*. In practice, it may be desirable to choose a relatively small *N* prior to convergence to regularize the result. Typically,  $\leq O(10)$  iterations are needed.

In addition to the  $\hat{t}_{matrix}$  and  $\hat{t}_{ignis}$  approaches discussed in the previous section, Fig. 2 shows the IBU result with N = 10. Unlike the  $\hat{t}_{matrix}$  and  $\hat{t}_{ignis}$  results, the  $\hat{t}_{IBU}$ does not suffer from rapid oscillations and like  $\hat{t}_{ignis}$ , is non-negative. Analogous results for quantum computer simulations will be presented in Sec. V.

### IV. CONSTRUCTING THE RESPONSE MATRIX

In practice, the R matrix is not known exactly and must be measured for each quantum computer and set of operating conditions. One way to measure R is to construct a set of  $2^{n_{\text{qubit}}}$  calibration circuits as shown in Fig. 3. Simple X gates are used to prepare all of the possible qubit configurations and then they are immediately measured.

Fig. 4 shows the R matrix from 5 qubits of the IBM Q Johannesburg machine (see Appendix A for details).



FIG. 4. An example R matrix from 5 qubits of the IBM Q Johannesburg machine using 8192 shots for each of the 2<sup>5</sup> possible states. Note that this matrix depends on the calibration quality, so while it is representative, it is not precisely the version valid for all measurements made on this hardware.

As expected, there is a significant diagonal component that corresponds to cases when the measured and true states are the same. However, there are significant offdiagonal components, which are larger towards the right when more configurations start in the one state. This matrix is hardware-dependent and its elements can change over time due to calibration drift. Machines with higher connectivity have been observed to have more readout noise.

Fig. 5 explores the universality of qubit migrations across R. If every qubit was identical and there were no effects from the orientation and connectivity of the computer, then one may expect that R can actually be described by just two numbers  $p_{0\to 1}$  and  $p_{1\to 0}$ , the probability for a zero to be measured as a one and vice versa. These two numbers are extracted from R by performing the fit in Eq. 5:

$$\min_{\substack{p_{0} \to 1 \\ p_{1} \to 0}} \sum_{ij} |R_{ij} - p_{0 \to 1}^{\alpha} (1 - p_{0 \to 1})^{\alpha'} p_{1 \to 0}^{\beta} (1 - p_{1 \to 0})^{\beta'} |^{2},$$
(5)

where  $\alpha$  is the number of qubits corresponding to the response matrix entry  $R_{ij}$  that flipped from 0 to 1,  $\alpha'$  is the number that remained as a 0 and  $\beta$ ,  $\beta'$  are the corresponding entries for a one state;  $\alpha + \alpha' + \beta + \beta' = n_{\text{qubit}}$ . For example, if the  $R_{ij}$  entry corresponds to the state  $|01101\rangle$  migrating to  $|01010\rangle$ , then  $\alpha = 1$ ,  $\alpha' = 1$ ,  $\beta = 2$ , and  $\beta' = 1$ .



FIG. 5. Test of universality in the response matrix R. Horizontal dotted lines show the result of a global fit to  $p_{0\to 1}$  and  $p_{1\to 0}$  assuming an equal rate of readout errors for each qubit. Big filled markers represent a fit with independent values for each qubit. Small semi-transparent open markers show the  $2^{n_{\text{qubit}}-1}$  transition probabilities for each qubit when the other  $n_{\text{qubit}} - 1$  qubits are unchanged between truth and measured states.

A global fit to these parameters for the Johannesburg machine results in  $p_{0\to1} \approx 3.2\%$  and  $p_{1\to0} \approx 7.5\%$ . In reality, the qubits are not identical and so one may expect that  $p_{0\to1}$  and  $p_{1\to0}$  depend on the qubit. A fit to  $n_{\text{qubit}}$ values for  $p_{0\to1}$  and  $p_{1\to0}$  (Eq. (5), but fitting 10 parameters instead of 2) are shown as filled triangles in Fig. 5. While these values cluster around the universal values (dotted lines), the spread is a relative 50% for  $p_{0\to1}$  and 60% for  $p_{1\to0}$ . Furthermore, the transition probabilities can depend on the values of neighboring qubits. The open markers in Fig. 5 show the transition probabilities for each qubit with the other  $n_{\text{qubit}} - 1$  qubits held in a fixed state. In other words, the  $2^{n_{\text{qubit}}-1}$  open triangles for each qubit show

$$\Pr(|q_0, \dots, q_i, \dots, q_{n_{\text{qubit}}}\rangle \to |q_0, \dots, q'_i, \dots, q_{n_{\text{qubit}}}\rangle),$$
(6)

where  $(q_i, q'_i) \in \{(0, 1), (1, 0)\}$  with the other qubits  $q_j$ held in a fixed configuration. The spread in these values<sup>3</sup> is smaller than the variation in the solid markers, which indicates that per-qubit readout errors are likely sufficient to capture most of the salient features of the response matrix. However, this is hardware dependent and higher connectivity computers may depend more on the state of neighboring qubits.

Constructing the entire response matrix requires exponential resources. While the measurement of R only needs to be performed once per quantum computer per operational condition, this can be untenable when  $n_{\text{qubit}} \gg 1$ . The tests above indicate that a significant fraction of the  $2^{n_{\text{qubit}}}$  calibration circuits may be required for a precise measurement of R. Sub-exponential approaches may be possible and will be studied in future work.

### V. REPRESENTATIVE RESULTS

For the results presented here, we simulate a quantum computer using qiskit-terra 0.9.0, qiskit-aer 0.2.3, and qiskit-ignis 0.1.1 [14]<sup>4</sup>. We choose a Gaussian distribution as the true distribution, as this is ubiquitous in quantum mechanics as the ground state of the harmonic oscillator<sup>5</sup>. In practice, all of the qubits of a system would be entangled to achieve the harmonic oscillator wave function. However, this is unnecessary for studying readout errors, which act at the ensemble level. The Gaussian is mapped to qubits using the following map:

$$t(b) \propto \exp\left[-\frac{1}{2\sigma} \left(b - 2^{n_{\text{qubit}}-1}\right)^2\right], \qquad (7)$$

where b is the binary representation of a computational basis state, i.e.,  $|00000\rangle \mapsto 0$  and  $|00011\rangle \mapsto 3$ . For the results shown below,  $\sigma = 3.5$ .

As a first test, the pathological response matrix introduced in Sec. III is used to illustrate a failure mode of the matrix inversion and ignis approaches. Fig. 6 compares  $\hat{t}_{\text{matrix}}, \hat{t}_{\text{ignis}}, \text{ and } \hat{t}_{\text{IBU}}$  for a 4-qubit Gaussian distribution. As the matrix inversion result is already non-negative, the  $\hat{t}_{ignis}$  result is nearly identical to  $\hat{t}_{matrix}$ . Both of these approaches show large oscillations. In contrast, the IBU method with 10 iterations is nearly the same as the truth distribution. This result is largely insensitive to the choice of iteration number, though it approaches the matrix inversion result for more than about a thousand iterations. This is because the IBU method converges to a maximum likelihood estimate [27], which in this case aligns with the matrix inversion result. If the matrix inversion would have negative entries, then it would differ from the asymptotic limit of the IBU method, which is always non-negative.

While Fig. 6 is illustrative for exposing a potential failure mode of matrix inversion and the **ignis** method, it

<sup>&</sup>lt;sup>3</sup> This spread has a contribution from connection effects, but also from Poisson fluctuations as each measurement is statistically independent.

<sup>&</sup>lt;sup>4</sup> Multi-qubit readout errors are not supported in all versions of qiskit-aer and qiskit-ignis. This work uses a custom measure function to implement the response matrices.

<sup>&</sup>lt;sup>5</sup> An alternative distribution that is mostly zero with a small number of spikes is presented in Appendix C.



FIG. 6. The measurement of a Gaussian distribution (ground state of harmonic oscillator) using the pathological response matrix from Sec. III. One million total shots are used both to sample from m and to construct R. The IBU method uses 10 iterations.

is also useful to consider a realistic response matrix. Fig. 7 uses the same Gaussian example as above, but for the IBM Q Johannesburg response matrix described in Sec. IV. Even though the migrations are large, the pattern is such that all three methods qualitatively reproduce the truth distribution. The dip in the middle of the distribution is due to the mapping between the Gaussian and qubit states: all of the state to the left of the center have a zero as the last qubit while the ones on the right have a one as the last qubit. The asymmetry of  $0 \rightarrow 1$  and  $1 \rightarrow 0$ induces the asymmetry in the measured spectrum. For example, it is much more likely to migrate from any state to  $|00000\rangle$  than to  $|11111\rangle$ .

The three unfolding approaches are quantitatively compared in Fig. 8. In particular, averaging over many pseudoexperiments, the spread in the predictions is a couple of percent smaller for  $\hat{t}_{\rm IBU}$  compared to  $\hat{t}_{\rm ignis}$  and both of these are about 10% more precise than  $\hat{t}_{\rm matrix}$ . The slight bias of  $\hat{t}_{\rm ignis}$  and  $\hat{t}_{\rm IBU}$  to the right results from the fact that they are non-negative. Similarly, the sharp peak at zero results from  $\hat{t}_{\rm ignis}$  and  $\hat{t}_{\rm IBU}$  values that are nearly zero when  $t_i \sim 0$ . In contrast, zero is not special for matrix inversion so there is no particular feature in the center of its distribution.

### VI. REGULARIZATION AND UNCERTAINTIES

One feature of any regularization method is the choice of regularization parameters. For the IBU method, these parameters are the prior and number of iterations. Fig. 9



FIG. 7. The measurement of a Gaussian distribution using the response matrix from the IBM Q Johannesburg machine described in Sec. IV. One million total shots are used construct R and  $10^5$  are used for t and m. The significant deviations on the far left and right of the distributions are due in part to large statistical fluctuations, where the counts are low. The IBU method uses 100 iterations.



FIG. 8. The distribution of the difference between true and predicted counts from a Gaussian distribution using the response matrix from the IBM Q Johannesburg machine described in Sec. IV. The simulation is repeated 1000 times and each bin contributes to this histogram. The standard deviations of the distributions are given in the legend. The IBU method uses 100 iterations.

shows the average bias in the measurement from the Gaussian example shown in Fig. 6 as a function of the number of iterations. With a growing number of iterations,  $\hat{t}_{\text{IBU}}$ approaches the oscillatory  $\hat{t}_{ignis}$ . The optimal number of iterations from the point or view of the bias is 3. However, the number of iterations cannot be chosen based on the actual bias because the true answer is not known a priori. In HEP, the number of iterations is often chosen before *unblinding* the data by minimizing the total expected uncertainty. In general, there are three sources of uncertainty: statistical uncertainty on m, statistical and systematic uncertainties on R, and non-closure uncertainties from the unfolding method. Formulae for the statistical uncertainty on m are presented in Ref. [22] and can also be estimated by bootstrapping [28] the measured counts. Similarly, the statistical uncertainty on R can be estimated by bootstrapping and then repeating the unfolding for each calibration pseudo-dataset. The sources of statistical uncertainty are shown as dot-dashed and dashed lines in Fig. 10. Adding more iterations enhances statistical fluctuations and so these sources of uncertainty increase monotonically with the number of iterations.

The systematic uncertainty on R and the method nonclosure uncertainty are not unique and require careful consideration. In HEP applications<sup>6</sup>, R is usually determined from simulation, so the systematic uncertainties are simulation variations that try to capture potential sources of mis-modeling. These simulation variations are often estimated from auxiliary measurements with data. In the QIS context, R is determined directly from the data, so the only uncertainty is on the impurity of the calibration circuits. In particular, the calibration circuits are constructed from a series of single qubit X gates. Due to gate imperfections and thermal noise, there is a chance that the application of an X gate will have a different effect on the state than intended. In principle, one can try to correct for such potential sources of bias by the method of iterated X gates where each X gate is replaced with (2m+1) X gates. In the absence of gate noise, these replacements should have no effect on the result. One can fit the *m*-dependence of the result and extrapolate [10, 29]. This method may have a residual bias and the uncertainty on the method would then become the systematic uncertainty on R. A likely conservative alternative to this approach is to modify R by adding in gate noise and taking the difference between the nominal result and one with additional gate noise, simulated using the thermal relaxation error functionality of giskit. This is the choice made in Fig. 10, where the gate noise is shown as a dot-dashed line. In this particular example, the systematic uncertainty on R increases monotonically with the number of iterations, just like the

<sup>6</sup> In HEP, there are additional uncertainties related to the modeling of background processes as well as related to events that fall into or out of the measurement volume. These are not relevant for QIS.



FIG. 9. The difference between  $\hat{t}_{\text{IBU}}$  and t as a function of the number of iterations for the example presented in Fig. 6. By definition, the **ignis** method does not depend on the number of iterations.

sources of statistical uncertainty.

The non-closure uncertainty is used to estimate the potential bias from the unfolding method. One possibility is to compare multiple unfolding methods and take the spread in predictions as an uncertainty. Another method advocated in Ref. [30] and widely used in HEP is to perform a data-driven reweighting. The idea is to reweight the  $t^0$  so that when folded with R, the induced  $m^0$  is close to the measurement m. Then, this reweighted  $m^0$  is unfolded with the nominal response matrix and compared with the reweighted  $t^0$ . The difference between these two is an estimate of the non-closure uncertainty. The reweighting function is not unique, but should be chosen so that the reweighted  $t^0$  is a reasonable prior for the data. For Fig. 10, the reweighting is performed using the nominal unfolded result itself. In practice, this can be performed in a way that is blinded from the actual values of  $\hat{t}_{\rm IBU}$  so that the experimenter is not biased when choosing the number of iterations.

Altogether, the sources of uncertainty presented in Fig. 10 show that the optimal choice for the number of iterations is 2. In fact, the difference in the uncertainty between 2 and 3 iterations is less than 1% and so consistent with the results from Fig. 9. Similar plots for the measurement in Fig. 7 can be found in Appendix B.

### VII. DISCUSSION

This work has introduced a new suite of readout error correction algorithms developed in high energy physics for



FIG. 10. Sources of uncertainty for  $\hat{t}_{\text{IBU}}$  as a function of the number of iterations for the example presented in Fig. 6. Each uncertainty is averaged over all states. The total uncertainty is the sum in quadrature of all the individual sources of uncertainty, except gate noise (which is not used in the measurement simulation, but would be present in practice).

binned differential cross section measurements. These unfolding techniques are well-suited for quantum computer readout errors, which are naturally binned and without acceptance effects (counts are not lost or gained during readout). In particular, the iterative Bayesian method has been described in detail and shown to be robust to a failure mode of the matrix inversion and ignis techniques. When readout errors are sufficiently small, all the methods perform well, with a preference for the ignis and Bayesian methods that produce non-negative results. The ignis method is a special case of the TUnfold algorithm, where the latter uses the covariance matrix to improve precision and incorporates regularization to be robust to the failure modes of matrix inversion. It may be desirable to augment the **ignis** method with these features or provide the iterative method as an alternative approach. In either case, Fig. 7 showed that even with a realistic response matrix, readout error corrections can be significant and must be accounted for in any measurement on near-term hardware.

An important challenge facing any readout error correction method is the exponential resources required to construct the full R matrix as mentioned in Sec. IV. While R must be constructed only once per hardware setup and operating condition, it could become prohibitive when the number of qubits is large. On hardware with few connections between qubits, per-qubit transition probabilities may be sufficient for accurate results. When that is not the case, one may be able to achieve the desired precision with polynomially many measurements. These ideas are left to future studies.

The code for the work presented here can be found at https://github.com/bnachman/QISUnfolding.

### VIII. CONCLUSIONS AND OUTLOOK

With active research and development across a variety of application domains, there are many promising applications of quantum algorithms in both science and industry on NISQ hardware. Readout errors are an important source of noise that can be corrected to improve measurement fidelity. High energy physics experimentalists have been studying readout error correction techniques for many years under the term unfolding. These tools are now available to the QIS community and will render the correction procedure more robust to resolution effects in order to enable near term breakthroughs.

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## Appendix A: Alternative Configurations of the IBM Q Johannesburg Machine

The readout errors of a quantum computer depend on its connectivity. The results presented in Fig. 5 used a particular subset of 5 qubits out of the 20 qubits available. Fig. 11 presents the results for four different configurations, corresponding to taking four rows of the computer qubits. In each row, every qubit is connected to its two neighbors. The first and last qubit in each row is connected to the row above or below. The middle qubits in the middle two rows are additionally connected to each other. Therefore, every qubit in the first and last rows have only two connections while two of the qubits in the middle rows have two connections and the other three qubits have three connections.

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FIG. 11. The same test of universality as in Fig. 5, using various configurations of the IBM Q Johannesburg machine. From top left, clockwise, the plots correspond to rows 1-4 as described in the text.

### Appendix B: Uncertainties when using IBM Q Response Matrix

# Appendix C: Alternative to Harmonic Oscillator Ground State

Fig. 14 presents a spiky distribution instead of a smooth probability density, as typified by the ground state of the harmonic oscillator. As the true distribution is exactly zero for many states, the matrix inversion has large fluctuations because there is nothing special about zero for this method (can give negative results). In contrast, the ignis and IBU methods are always positive and have smaller deviations.

Fig. 12 and 13 are the analogs for Fig. 9 and 10, but for the example shown in Fig. 7, which is based on an IBM Q response matrix.



FIG. 12. The difference between  $\hat{t}_{\text{IBU}}$  and t as a function of the number of iterations for the example presented in Fig. 7. By definition, the **ignis** method does not depend on the number of iterations.



FIG. 14. The measurement of a spiky distribution using the response matrix from the IBM Q Johannesburg machine described in Sec. IV. Five hundred total shots are used both to sample from m and for each state to construct R. The IBU method uses 100 iterations.



FIG. 13. Sources of uncertainty for  $\hat{t}_{\text{IBU}}$  as a function of the number of iterations for the example presented in Fig. 7. Each uncertainty is averaged over all states. The total uncertainty is the sum in quadrature of all the individual sources of uncertainty, except gate noise (which is not used in the measurement simulation, but would be present in practice).

- [1] J. Preskill, Quantum 2, 79 (2018).
- [2] J. L. Park, Found. Phys. 1, 23 (1970).
- [3] W. K. Wootters and W. H. Zurek, Nature 299, 802 (1982).
- [4] D. Dieks, Phys. Lett. A **92**, 271 (1982).
- [5] G. Cowan, Conf. Proc. C0203181, 248 (2002).
- [6] V. Blobel, in *PHYSTAT 2011 Proceedings* (2011) pp. 240– 251.
- [7] V. Blobel, in *Data Analysis in High Energy Physics* (John Wiley & Sons, 2013) pp. 187–225.
- [8] K. Cormier, R. Di Sipio, and P. Wittek, "Unfolding as quantum annealing," arXiv (2019), arXiv:1908.08519 [physics.data-an].
- [9] A. Kandala, A. Mezzacapo, K. Temme, M. Takita, M. Brink, J. M. Chow, and J. M. Gambetta, Nature 549, 242 (2017).
- [10] E. F. Dumitrescu, A. J. McCaskey, G. Hagen, G. R. Jansen, T. D. Morris, T. Papenbrock, R. C. Pooser, D. J. Dean, and P. Lougovski, Phys. Rev. Lett. **120**, 210501 (2018).
- [11] N. Klco and M. J. Savage, "Minimally-entangled state preparation of localized wavefunctions on quantum computers," arXiv (2019), arXiv:1904.10440 [quant-ph].
- [12] K. Yeter-Aydeniz, E. F. Dumitrescu, A. J. McCaskey, R. S. Bennink, R. C. Pooser, and G. Siopsis, Phys. Rev. A 99, 032306 (2019).
- [13] V. Blobel, "Unfolding methods in high-energy physics experiments," DESY-84-118 (1984).
- [14] IBM Research, "Qiskit," https://qiskit.org (2019).
- [15] IBM Research, "Qiskit Ignis," https://qiskit.org/ignis (2019).
- [16] G. Choudalakis, "Fully Bayesian unfolding," arXiv (2012), arXiv:1201.4612 [physics.data-an].

- [17] N. D. Gagunashvili, "Machine learning approach to inverse problem and unfolding procedure," arXiv (2010), arXiv:1004.2006 [physics.data-an].
- [18] A. Glazov, "Machine learning as an instrument for data unfolding," arXiv (2017), arXiv:1712.01814 [physics.dataan].
- [19] K. Datta, D. Kar, and D. Roy, "Unfolding with generative adversarial networks," arXiv (2018), arXiv:1806.00433 [physics.data-an].
- [20] G. Zech and B. Aslan, in *PHYSTAT 2003 Proceedings*, Vol. C030908 (2003) p. TUGT001.
- [21] L. Lindemann and G. Zech, Nucl. Instrum. Methods Phys. Res. A 354, 516 (1995).
- [22] G. D'Agostini, Nucl. Instrum. Methods Phys. Res. A 362, 487 (1995).
- [23] L. B. Lucy, Astron. J. 79, 745 (1974).
- [24] W. H. Richardson, J. Opt. Soc. Am. 62, 55 (1972).
- [25] A. Hocker and V. Kartvelishvili, Nucl. Instrum. Methods Phys. Res. A 372, 469 (1996).
- [26] S. Schmitt, J. Instrum. 7, T10003 (2012).
- [27] L. A. Shepp and Y. Vardi, IEEE Trans. Med. Imaging 1, 113 (1982).
- [28] B. Efron, Ann. Statist. 7, 1 (1979).
- [29] A. Kandala, K. Temme, A. D. Córcoles, A. Mezzacapo, J. M. Chow, and J. M. Gambetta, Nature 567, 491 (2019).
- [30] B. Malaescu, "An iterative, dynamically stabilized method of data unfolding," arXiv (2009), arXiv:0907.3791 [physics.data-an].