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# A Genetic Risk Score to Personalize Prostate Cancer Screening, Applied to Population Data 

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## ABSTRACT

Background: A polygenic hazard score (PHS), the weighted sum of 54 SNP genotypes, was previously validated for association with clinically significant prostate cancer and for improved prostate cancer screening accuracy. Here, we assess the potential impact of PHS-informed screening.

Methods: United Kingdom population incidence data (Cancer Research United Kingdom) and data from the Cluster Randomized Trial of PSA Testing for Prostate Cancer were combined to estimate age-specific clinically significant prostate cancer incidence (Gleason score $\geq 7$, stage T3-T4, PSA $\geq 10$, or nodal/distant metastases). Using HRs estimated from the ProtecT prostate cancer trial, age-specific incidence rates were calculated for various PHS risk percentiles. Risk-equivalent age, when someone with a given PHS percentile has prostate cancer risk equivalent to an average

50-year-old man (50-year-standard risk), was derived from PHS and incidence data. Positive predictive value (PPV) of PSA testing for clinically significant prostate cancer was calculated using PHS-adjusted age groups

Results: The expected age at diagnosis of clinically significant prostate cancer differs by 19 years between the 1st and 99th PHS percentiles: men with PHS in the 1st and 99th percentiles reach the 50 -year-standard risk level at ages 60 and 41, respectively. PPV of PSA was higher for men with higher PHS-adjusted age.

Conclusions: PHS provides individualized estimates of riskequivalent age for clinically significant prostate cancer. Screening initiation could be adjusted by a man's PHS.

Impact: Personalized genetic risk assessments could inform prostate cancer screening decisions.

[^0]
#### Abstract

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## Introduction

Prostate cancer is the second most common malignancy in men worldwide with nearly 1.3 million cases diagnosed globally in 2018 (1). It was the third leading cause of European male cancer-related mortality in 2018, following mortality from lung and colorectal cancers (2). Prostate cancer screening with PSA testing can reduce mortality (3), but universal screening may cause overdetection of cancers that would never become clinically apparent in a man's life-time and overtreatment of indolent disease. Guidelines recommend that individual men participate in informed decision-making about screening, taking into account factors such as their age, race/ethnicity, family history, and preferences (4-6).

Assessment of a man's genetic risk of developing prostate cancer has promise for guiding individualized screening decisions $(7,8)$. We previously developed and validated a polygenic hazard score (PHS), a weighted sum of 54 SNP genotypes, as significantly associated with age at diagnosis of clinically significant prostate cancer, defined as cases where any of the following is applied: Gleason score $\geq 7$, clinical stage T3-T4, PSA $\geq 10$, or where there were nodal or distant metastases (9). Risk stratification by the PHS also improved the screening performance of PSA testing; the positive predictive value (PPV) of PSA testing for clinically significant prostate cancer increased as PHS increased (9).

Here, we apply the prostate cancer PHS to population data to assess its potential impact on individualized screening. Specifically, we combine genetic risk, measured by PHS, and known population incidence rates to estimate a risk-equivalent age, for example, the age at which a man with a given PHS will have the same risk of clinically significant prostate cancer as a typical man at age 50 years Such genetic risk estimates can guide individualized decisions about whether, and at what age, a man might benefit from prostate cancer screening.

[^1]
## Materials and Methods

PHS
Full methodologic details of the development and validation of the prostate cancer PHS have been described previously (9). Briefly, the PHS was developed using PRACTICAL consortium clinical and genetic data from 31,747 men of European ancestry as a continuous survival analysis model (10) and found to be associated with age at prostate cancer diagnosis (9). Validation testing was performed in an independent, separate dataset consisting of 6,411 men from the United Kingdom ProtecT study $(11,12)$. PHS was calculated as the vector product of a patient's genotype $\left(X_{i}\right)$ for $n$ selected SNPs and the corresponding parameter estimates $\left(\beta_{i}\right)$ from a Cox proportional hazards regression (Eq. 1):

$$
\begin{equation*}
P H S=\sum_{i}^{n} X i \beta i \tag{1}
\end{equation*}
$$

The 54 SNPs included in the model, and their parameter estimates, have been published previously (9) and are also shown in Supplementary Table S1.

## Population age-specific incidence

Age-specific prostate cancer incidence data were obtained for men ages 40-70 years from the United Kingdom, 2013-2015 (Cancer Research United Kingdom; ref. 13). Men may be less likely to be screened outside this age range $(3,14)$. The log of the prostate cancer incidence data were fit using linear regression to develop a continuous model of age-specific prostate cancer incidence in the United Kingdom ( $I_{\text {all }}$ ).

The United Kingdom age-specific proportion of incidence classified as clinically significant prostate cancer was estimated using data from the Cluster Randomized Trial of PSA Testing for Prostate Cancer (CAP). The CAP trial evaluated the impact of a single, low-intensity PSA screening intervention on prostate cancer-specific mortality in the United Kingdom (15). CAP was linked to the ProtecT study, which included men ages 50-69 at randomization (15); ProtecT compared management options including surgery, radiotherapy, and active

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Note: Supplementary data for this article are available at Cancer Epidemiology, Biomarkers \& Prevention Online (http://cebp.aacrjournals.org/).

Additional members from the Prostate Cancer Association Group to Investigate Cancer Associated Alterations in the Genome consortium (PRACTICAL, http:// practical.icr.ac.uk/) are provided in the Supplemental Material.

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Table 1. Risk-equivalent age for clinically significant prostate cancer ${ }^{\text {a }}$, by PHS percentile.

| PHS percentile | Age when man reaches 50-year-standard risk ${ }^{\text {b }}$ (95\% CI) | $\Delta$ Age $^{\text {c }}$ (95\% CI) |
| :---: | :---: | :---: |
| 1 | 60 (59-62) | $-10(-11$ to -8$)$ |
| 5 | 56 (54-58) | -6 (-8 to -4) |
| 20 | 53 (51-55) | $-3(-5$ to -1$)$ |
| 50 | 50 (48-52) | 0 ( -2 to 2) |
| 80 | 47 (45-48) | 3 (1 to 4) |
| 95 | 44 (43-46) | 6 (5 to 8) |
| 99 | 41 (39-43) | 9 (7 to 11) |

Abbreviation: Cl , confidence interval.
${ }^{\text {a }}$ Clinically significant prostate cancer was defined as Gleason score $\geq 7$, clinical stage T3-T4, PSA $\geq 10$, or with nodal/distant metastases.
${ }^{\mathrm{b}}$ Risk of typical 50 -year-old defined as overall population incidence at age 50 . ${ }^{c} \Delta \mathrm{Age}=$ difference between 50 and the age when risk is that of a typical $50-$ year-old man.
surveillance in patients with PSA-detected prostate cancer (12). The clinical and demographic features of the CAP and ProtecT studies have been described previously $(12,15)$. Clinically significant prostate cancer was defined as cases often ineligible for active surveillance (consistent with the definition used in the PHS development). These are cases with Gleason score $\geq 7$, clinical stage T3-T4, PSA $\geq 10$, or with nodal/distant metastases $(9,16,17)$. Men in the intervention arm of the CAP trial who were diagnosed with any prostate cancer were divided into 5 -year age intervals at prostate cancer detection ( $n=8,054$; ref. 15). The proportion of clinically significant disease in each age interval was calculated as the number of clinically significant prostate cancer diagnoses divided by the total number of prostate cancer diagnoses in the CAP cohort for whom PSA and clinical stage information were available ( $n=6,388$; ref. 15). The total (all ages) proportion of clinically significant prostate cancer was similarly calculated from CAP data. The age-specific prostate cancer incidence curve, $I_{\text {all }}$, was multiplied, within each 5 -year age range, by the corresponding age-specific proportion of CAP clinically significant prostate cancer diagnoses, to yield a continuous estimate of age-specific, clinically significant prostate cancer incidence ( $I_{\text {clinically significant }}$ ). A similar calculation was done to estimate age-specific, more aggressive prostate cancer incidence [using a stricter definition of clinically significant disease that corresponds to clinical high risk or above by National Comprehensive Cancer Network (NCCN) guidelines: clinical stage T3-T4, PSA $>20$, Gleason score $\geq 8$, or with nodal/distant metastases refs. 9, 16, 17] as $I_{\text {more-aggressive. }}$. Finally, clinically insignificant prostate cancer incidence ( $I_{\text {clinically }}$ insignificant $)$ was estimated as the difference between $I_{\text {all }}$ and $I_{\text {clinically significant }}$

## Impact of genetic risk on clinically significant prostate cancer incidence

Men in the ProtecT study with genotype data ( $n=6,411$ ) were categorized by their PHS percentile ranges ( $0-2,2-10,10-30,30-70$, $70-90,90-98$, and $98-100$ ) to correspond to percentiles of interest ( 1 , $5,20,50,80,95$, and 99 , respectively). These percentiles refer to the distribution of PHS in the ProtecT dataset within controls aged $<70$ Incidence rates of clinically significant prostate cancer were calculated for each percentile range ( $I_{\text {percentile }}$ ) using Cox proportional hazards regression (parameter estimate, $\beta$ ), following the methods published previously (9). The reference for each HR was taken as the mean PHS
among those men with approximately 50th percentile for genetic risk (i.e., 30th-70th percentile of PHS, called $P H S_{\text {median }}$ ), and this median group was assumed to have an incidence of clinically significant disease matching the overall population ( $I_{\text {clinically significant }}$, calculated above). Incidence rates for the other percentiles of interest $\left(I_{\text {percentile }}\right)$ were then calculated by determining the mean PHS among men in the corresponding percentile range (called $P H S_{\text {percentile }}$ ) and applying Eq. 2:

$$
\begin{equation*}
I_{\text {percentile }}(\text { age })=I_{\text {clinically insignificant }}(\text { age }) e^{\beta(P H S \text { percentile-PHSmedian })} \tag{2}
\end{equation*}
$$

As described in the original validation of this PHS model for prostate cancer (9), PHS calculated in the ProtecT dataset will be biased by the disproportionately large number of cases included, relative to incidence in the general population. Leveraging the cohort design of the ProtecT study (11), we therefore applied a correction for this bias, using previously published methods (18) and the R "survival" package ( R version 3.2.2; refs. 19, 20). The corrected PHS values were used to update $P H S_{\text {percentile }}$ and $P H S_{\text {median }}$ used in Eq. 2. Then, $95 \%$ confidence intervals for the HRs for each percentile were determined by bootstrapping 1,000 random samples from the ProtecT dataset, while maintaining the same number of cases and controls from the original dataset. The $I_{\text {percentile, }}$ predicted partial hazard (product of $P H S_{\text {percentile }}$ and the estimated $\beta$ ), and SEs (to account for sample weights) were calculated for each bootstrap sample.

Percentile-specific incidence estimates ( $I_{\text {percentile }}$ ) were visualized as the corresponding cumulative incidence curves for clinically significant prostate cancer diagnosis for men ages 50-70 years. Analogous HRs and incidence curves were similarly calculated for the annualized incidence rates of clinically insignificant and more aggressive prostate cancer.

An individualized PHS to aid prostate cancer screening decisions in the clinic might be facilitated by a readily interpretable translation of PHS to terms familiar to men and their physicians. The PHS was, therefore, combined with United Kingdom clinically significant prostate cancer incidence data to give a risk-equivalent age: when a man with a given PHS percentile would have the same risk of clinically significant prostate cancer as, say, that of a typical man at 50 years old (50-year-standard risk). We defined $\Delta$ Age as the difference between age 50 and the age when prostate cancer risk matches that of a typical 50 -year-old man. $95 \%$ confidence intervals for the age when a man reaches 50 -year-standard risk and $\Delta$ Age were determined using the HRs calculated from the 1,000 bootstrapped samples from ProtecT, described above.

Finally, we considered the common clinical scenario of a man presenting to his primary care physician to discuss prostate cancer screening. To illustrate how PHS might influence this discussion, we identified the subset of men in the ProtecT validation dataset who were around the median age of 60 years (55-64), to represent a typical patient. From this subset, we created three groups: those whose prostate cancer risk-equivalent age remained within the selected range (ages 55-64), those whose risk-equivalent age was $<55$, and those whose risk-equivalent age was $\geq 65$. We then calculated the PPV and SE of the mean of PSA testing for development of clinically significant prostate cancer in these three PHS-adjusted (prostate cancer riskequivalent age) groups using methods described previously (9). This was done by taking 1,000 random samples (with replacement) of the subjects with elevated PSA ( $\geq 3.0 \mathrm{ng} / \mathrm{mL}$ ) in the dataset, stratified to ensure each random sample matched the distribution of controls and cases reported for men with elevated PSA in ProtecT $(11,12)$. Stratification was also used to ensure the proportion of clinically significant cases matched the proportion reported in CAP for the age
range of 55-64 (11), such that the PPV for the sample exactly matched the expected value for the linked ProtecT and CAP trials, but the distribution of genetic risk (PHS) was varied at random within each disease status group (control, clinically significant, and clinically insignificant). A similar calculation for PPV of PSA testing for development of any prostate cancer was performed for the three PHS-adjusted age groups.

## Results

Linear regression yielded a model of prostate cancer age-specific incidence rates (Eq. 3; $R^{2}=0.96 ; P=0.001$ ) that was highly consistent with empirical data reported by Cancer Research United Kingdom (Fig. 1).

$$
\begin{equation*}
I_{\text {all }}=0.004 e^{0.203(\text { age }-40)} \tag{3}
\end{equation*}
$$

In the CAP study (15), the overall proportion of prostate cancer incidence classified as clinically significant disease was $72.3 \%$. The proportions of age-specific, clinically significant disease increased with age: $48.0 \%, 55.9 \%, 63.5 \%$, and $79.7 \%$ of men ages $50-54,55-59,60-64$, and 65-69, respectively, were diagnosed with clinically significant prostate cancer. Combining men ages 55-64, the proportion of agespecific, clinically significant prostate cancer was $61.1 \%$.

Cumulative incidence estimates of clinically significant prostate cancer are shown in Fig. 2 for various levels of genetic risk, as indicated by PHS percentile, showing a difference in age at diagnosis related to PHS strata. Supplementary Figs. S1 and S2 show analogous results for the incidence curves of clinically insignificant and more aggressive prostate cancer, respectively. Table 1 shows risk-equivalent age for each PHS percentile. The expected age at clinically significant prostate cancer diagnosis differs by 19 years between the 1st and 99th PHS percentiles. Specifically, a man with a PHS in the 99th percentile reached a prostate cancer detection risk equivalent to the 50 -year standard at an age of 41 years. Conversely, a man with a PHS in the 1st percentile would not reach the 50-year-standard risk level until age 60 years. Qualitatively, the curves for clinically significant (Fig. 2),
clinically insignificant (Supplementary Fig. S1), and more aggressive (Supplementary Fig. S2) prostate cancer maintain consistent horizontal shifts relative to curves for other PHS percentiles over the age range studied. Quantitatively, this was confirmed by $\Delta$ Age, which remained the same for each PHS percentile across a true age range of 40-70. Thus, $\Delta$ Age was taken to be approximately constant for each PHS percentile and is reported in Table 1.

Figure 3 shows the PPV of PSA testing for clinically significant prostate cancer was 0.21 (SE, 0.01) for men approximately 60 years old (data derived from a total of 1,395 ProtecT men ages 55-64: 283 with clinically significant prostate cancer, 127 with clinically insignificant prostate cancer, and 575 controls with a PSA $\geq 3.0 \mathrm{ng} / \mathrm{mL}$ ). PPV was lower for those with a prostate cancer risk-equivalent age $<55$ years ( 0.12 ; SE, 0.04 ) and higher for those with prostate cancer riskequivalent age $\geq 65$ years ( 0.40 ; SE, 0.03 ).

The PPVs of PSA testing for any prostate cancer were 0.18 (SE, $0.05), 0.37$ (SE, 0.01 ), and 0.61 (SE, 0.03 ) in men with a prostate cancer risk-equivalent age $<55$ years, between 55 and 64 years, and $\geq 65$ years, respectively. These PPVs, in combination with the PPVs of PSA for clinically significant prostate cancer, indicate that in the older prostate cancer risk-equivalent age group ( $\geq 65$ years), $40 \%$ of positive PSA tests are from clinically significant disease, $21 \%$ are from clinically insignificant disease, and $39 \%$ are false positives. The false positive rates for men with a prostate cancer risk-equivalent age $<55$ years and between $55-64$ years are $82 \%$ and $63 \%$, respectively.

## Discussion

We applied the PHS to population incidence data to estimate agespecific risk of clinically significant prostate cancer. The resulting agespecific incidence rates (displayed as incidence curves in Fig. 2) demonstrate clinically meaningful differences across various levels of genetic risk, as estimated by PHS. By combining these population curves with an individual's genetic risk and true age, we demonstrate calculation of a risk-equivalent age at diagnosis of clinically significant prostate cancer. This age relates a man's current prostate cancer risk to


Figure 1
Annual incidence of prostate cancer in the United Kingdom (UK), 2013-2015. Dots represent the raw, agespecific incidence rates of each age range, per 100,000 males. The black line represents the results of linear regression for an exponential curve to give a continuous model of age-specific incidence in the United Kingdom, $R^{2}=0.96 ; P=0.001$.

Figure 2.
Incidence of clinically significant prostate cancer, as derived from application of PHS HRs and population data from the United Kingdom. The overall population incidence is taken as the median risk (50th percentile); this accounts for age-specific proportions of prostate cancer that were clinically significant in the CAP trial (15). HRs were calculated within ProtecT data for various levels of genetic risk ranges (0-2, 2-10, 10-30, 30-70, 70-90, 90-98, and 98-100) to correspond to percentiles of interest ( $1,5,20,50,80,95$, and 99 , respectively), and used to adjust the median incidence curve. Blue lines represent genetic risk lower than the median, while red lines represent genetic risk higher than the median

that of the age-specific population average. The incidence curves for clinically significant prostate cancer are modulated by 19 years between the 1st and 99th percentiles of PHS. Moreover, the PPV of PSA testing in three PHS-adjusted (prostate cancer risk-equivalent age) groups demonstrated that PPV is significantly higher in men with higher risk-equivalent ages of prostate cancer diagnosis. These results have important implications for clinicians considering discussions of whether, and when, to initiate prostate cancer screening in an asymptomatic man.

Prostate cancer can cause considerable mortality and morbidity, but is curable if detected early. Determination of age of clinically significant disease diagnosis is thus highly relevant. Data from the CAP study shown here confirm prior findings of increasing risk of clinically significant prostate cancer as men age (21-24). The proportion of new prostate cancer diagnoses classified as clinically significant in CAP is higher than some older studies that were limited to men with low PSA and normal digital rectal exam (25-27), while another modern population study shows similar or higher proportions with clinically

Figure 3.
Application of prostate cancer risk-equivalent age to the clinical scenario of whether to screen a 60-year-old man (median age from ProtecT). The risk-equivalent age is the patient's true age adjusted by PHS level. This plot shows results for all men from ProtecT aged approximately 60 years old (range, 55-64), grouped by their calculated prostate cancer risk-equivalent age: $<55,55-64$, or $\geq 65$. The PPV of PSA testing for clinically significant prostate cancer and the corresponding SEs of the mean of PSA testing are shown for each of these three groups.

PPV for ProtecT patients 55-64 years old

significant disease (21). Taken together, these results suggest that screening delayed to an older age will yield a higher incidence of clinically significant disease.

The primary screening tool, PSA testing, is associated with a small absolute decreased risk of death from prostate cancer (3), but carries a risk of overdetection and harm from overtreatment in men who would never have experienced clinical manifestations of their prostate cancer (28). Thus, universal screening comes at a high cost, both in burden on health care systems and in the sequelae arising from elevated PSA in men with indolent disease: unnecessary biopsy procedures, overdetection, and treatment-related morbidities $(4,5)$. Conversely, there are some men who will develop clinically significant prostate cancer and would benefit from screening, possibly even at a relatively young age. Screening guidelines recommend individualized decision-making, but the available quantitative or objective data to guide these decisions are insufficient. For instance, family history provides some guidance, but genetic risk has been shown to be more strongly associated with age of clinically significant prostate cancer diagnosis than patient-provided family history $(9,29)$.

PHS, in conjunction with other informative factors such as family history, may help identify men who may develop the highest risk cancers (12). Incorporating a risk-adjusted age in an electronic medical record could reduce burden for general practitioners. The riskadjusted age can be based on whatever threshold of risk for clinically significant prostate cancer is considered optimal. Here, we have used the typical risk at age 50. Waiting until the man whose risk-adjusted screening age reached 60 would be much more likely to avoid overdiagnosis and overtreatment than to miss a clinically significant prostate cancer. This is supported by the clinically significant-specific incidence rates reported here for CAP in the United Kingdom and also by recently reported absolute age-specific incidence rates in Norway (21). One way a risk-stratified approach addresses overdetection is by providing a quantifiable, objective, and accurate rationale to not screen many men until they reach sufficient risk (in which time, their competing risks also have a chance to manifest; these could also inform screening and management decisions, especially if they affect life expectancy). The concern for overtreatment is also a critical consideration. As demonstrated in the ProtecT study, lower risk disease does not need to be treated aggressively at diagnosis and can be monitored with active surveillance and routine PSA checks (12). In addition, other major trials have demonstrated that the risks of biopsy can be mitigated by using multiparametric prostate MRI (30-32). These important mitigating factors are not directly related to polygenic risk, but they do decrease the risks associated with a prostate cancer screening program

The stratification of men based on their genetic risk is of particular interest in the primary care setting, where the majority of prostate cancer screening discussions take place. Shared decision-making between patient and physician has long been recommended in discussions of prostate cancer screening $(5,33)$, and physicians are tasked with determining an individual's risk based on factors such as his family history and ethnicity. However, physicians demonstrate different attitudes toward screening, with some screening all men proactively to avoid underdiagnoses, some screening only those men who request it, and some who attempt to weigh the costs and benefits of PSA screening on a case-by-case basis $(34,35)$. General practitioners, who are already limited by time constraints and their patients' other health issues, must carefully discuss the complex risks and benefits of PSA screening with their patients (36). However, efficiently identifying men at higher risk of clinically significant disease is important because detection of prostate cancer at an early stage allows for definitive treatment to prevent cancer progression or metastases (12).

Quantitative risk stratification could guide physicians in their screening conversations with patients by providing an objective risk-equivalent age for the development of clinically significant disease. This allows for simpler and more standardized informed decisionmaking regarding whether an individual man might benefit from prostate cancer screening. For example, physicians who normally initiate screening discussions at some age (e.g., 50-55) could shift the timing according to the prostate cancer risk-equivalent age. Some men might need to begin prostate cancer screening at a younger age to detect early onset of clinically significant disease. The PHS has previously demonstrated high PPV of PSA testing for clinically significant prostate cancer in men with progressively higher scores (9).

The potential utility of prostate cancer risk-equivalent age in the clinic is additionally demonstrated by its impact on PPV of PSA testing for clinically significant prostate cancer. Suppose a 60 -year-old man presents to his physician to inquire about prostate cancer screening. If this man has a prostate cancer risk-equivalent age close to his true age (55-64), the PPV of a PSA test (for prediction of clinically significant prostate cancer) for him will be approximately $24 \%$. If his riskequivalent age is $<55$, the PPV decreases to $13 \%$, and he might be reassured in foregoing PSA testing. Postponing, or even forgoing, screening in men with low PHS percentiles to when they reach their risk-equivalent age could decrease the harms associated with screening, or early detection and treatment of prostate cancer $(4,5)$. Other men may choose to delay the initiation of PSA testing until they are older and have increased risk. Conversely, if this same man has a riskequivalent age $\geq 65$, the PPV of PSA testing increases substantially to $45 \%$, implying that screening may be more informative for him. Of note, the increase in PPV in this illustration exceeds that of the reported effect of carrying a mutation in BRCA1 or BRCA2 (37).

Cost-effectiveness is another concern regarding prostate cancer screening. Use of PHS, a one-time test valid for a man's entire life, can improve screening efficiency while reducing overall costs. The genotyping chip assay requires only a saliva sample and can be run for costs similar to those for single-gene testing (e.g., the BRCA mutation). Genotyping also informs genetic risks for other diseases, possibly allowing multiple tests to be run on the same genotype results $(38,39)$. PSA screening (and subsequent prostate biopsy) could be offered only to those men at higher risk of clinically significant disease. PHS might increase the efficiency of any prostate cancer screening program by incorporating knowledge that there are some men with higher baseline genetic risks of developing clinically significant prostate cancer, even at a younger age, while others have a low baseline genetic risk.

Limitations of this work include that the PHS did not incorporate genotypic data from men of non-European ancestry during its development (9), a reflection of the available data, which may affect the potential use of the PHS for screening decision-making in men from other ethnic groups. This is noteworthy, as disparities in prostate cancer incidence and survival show that in the United States, men with African ancestry are more likely to develop prostate cancer and to die from their disease (40). Our group and others are studying the application of genetic scores to non-European ancestry groups. In addition, we used incidence data from a single country (the United Kingdom) with relatively low rates of screening. While the epidemiologic data used in this work are of high quality and drawn from the same United Kingdom population as was previously used for the validation of the PHS model (9), further work should evaluate the PHS in other populations. Finally, there are now over 140 SNPs reported to have associations with prostate cancer, identified using a meta-analysis that included ProtecT data (41), but not all of these SNPs are represented on the custom array used to develop the original PHS.

Furthermore, the PHS model was validated using independent data from ProtecT; the inclusion of those other SNPs associated with prostate cancer would have introduced circularity into the validation. Adding more SNPs to further improve the model is an area of active investigation. If we, or others, succeed in developing a further optimized PHS, we expect the range of $\Delta$ Age to expand.

We conclude that clinically meaningful risk stratification can be achieved through application of a PHS that is associated with age at clinically significant prostate cancer diagnosis to United Kingdom population data. PHS can also be used to calculate estimates of riskequivalent age for the development of clinically significant prostate cancer for individual men. The PPV of PSA was higher for men with higher PHS-adjusted prostate cancer-equivalent ages. Assessing personalized genetic risk via PHS could assist patients and physicians, alike, with the important decision of whether, and when, to initiate prostate cancer screening.

## Disclosure of Potential Conflicts of Interest

C.C. Fan is a data scientist at HealthLytix. F.C. Hamdy is an editor-in-chief at British Journal of Urology International and has a consulting/advisory board relationship with Intuitive. R.A. Eeles reports receiving speaker honoraria from the GU-ASCO meeting (January 2016), RMH FR meeting (November 2017; support from Janssen), and University of Chicago (May 2018) and educational honorarium paid by Bayer \& Ipsen to attend GU Connect at a venue at ESMO. O.A. Andreassen is a consultant for HealthLytix. A.M. Dale reports receiving a commercial research grant from GE Healthcare, has ownership interest (including patents) in CorTechs Labs, Inc., and has a consulting/advisory board relationship with Human Longevity, Inc. T.M. Seibert is a consultant for Multimodal Imaging Services Corporation and WebMD, Inc. No potential conflicts of interest were disclosed by the other authors

## Disclaimer

The content is solely the responsibility of the authors and does not necessarily represent the official views of the NIH, NIHR, or the Department of Health and Social Care.

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## References

1. Bray F, Ferlay J, Soerjomataram I, Siegel RL, Torre LA, Jemal A. Global cancer statistics 2018: GLOBOCAN estimates of incidence and mortality worldwide for 36 cancers in 185 countries. CA Cancer J Clin 2018;68:394-424.
2. Ferlay J, Colombet M, Soerjomataram I, Dyba T, Randi G, Bettio M, et al. Cancer incidence and mortality patterns in Europe: estimates for 40 countries and 25 major cancers in 2018. Eur J Cancer 2018;103:356-87.
3. Schröder FH, Hugosson J, Roobol MJ, Tammela TLJ, Ciatto S, Nelen V, et al. Screening and prostate-cancer mortality in a randomized European study. N Engl J Med 2009;360:1320-8.
4. Grossman DC, Curry SJ, Owens DK, Bibbins-Domingo K, Caughey AB, Davidson KW, et al. Screening for prostate cancer: US Preventive Services Task Force recommendation statement. JAMA 2018;319:1901-13.
5. Wolf AMD, Wender RC, Etzioni RB, Thompson IM, D'Amico AV, Volk RJ, et al. American Cancer Society guideline for the early detection of prostate cancer: update 2010. CA Cancer J Clin 2010;60:70-98.
6. National Health Service (NHS). Prostate cancer - PSA testing; [about 5 screens]. Available from: https://www.nhs.uk/conditions/prostate-cancer/psa-testing/.
7. Witte JS. Personalized prostate cancer screening: improving PSA tests with genomic information. Sci Transl Med 2010;2:62ps55.
8. Pashayan N, Duffy SW, Chowdhury S, Dent T, Burton H, Neal DE, et al. Polygenic susceptibility to prostate and breast cancer: implications for personalised screening. Br J Cancer 2011;104:1656-63.
9. Seibert TM, Fan CC, Wang Y, Zuber V, Karunamuni R, Parsons JK, et al. Polygenic hazard score to guide screening for aggressive prostate cancer: development and validation in large scale cohorts. BMJ 2018;360:1-7.
10. Desikan RS, Fan CC, Wang Y, Schork AJ, Cabral HJ, Cupples LA, et al. Genetic assessment of age-associated Alzheimer disease risk: development and validation of a polygenic hazard score. PLoS Med 2017;14:e1002258.
11. Lane JA, Donovan JL, Davis M, Walsh E, Dedman D, Down L, et al. Active monitoring, radical prostatectomy, or radiotherapy for localised prostate cancer: Study design and diagnostic and baseline results of the ProtecT randomised phase 3 trial. Lancet Oncol 2014;15:1109-18.
12. Hamdy FC, Donovan JL, Lane JA, Mason M, Metcalfe C, Holding P, et al. 10-year outcomes after monitoring, surgery, or radiotherapy for localized prostate cancer. N Engl J Med 2016;375:1415-24.
13. Cancer Research UK. Prostate cancer incidence statistics; [about 4 screens]. Available from: https://www.cancerresearchuk.org/health-professional/cancer-statistics/statistics-by-cancer-type/prostate-cancer/incidence.
14. Parker C, Gillessen S, Heidenreich A, Horwich A, ESMO Guidelines Committee. Cancer of the prostate: ESMO Clinical Practice Guidelines for diagnosis, treatment and follow-up. Ann Oncol 2015;26:v69-v77.
15. Martin RM, Donovan JL, Turner EL, Metcalfe C, Young GJ, Walsh EI, et al. Effect of a low-intensity PSA-based screening intervention on prostate cancer mortality: the CAP randomized clinical trial. JAMA 2018;319:883-95.
16. National Comprehensive Cancer Network. The NCCN Clinical Practice Guidelines in Oncology. Prostate cancer. Version 1.2019. Plymouth Meeting (PA): NCCN; 2019. Available from: https://www.nccn.org/professionals/ physician_gls/pdf/prostate.pdf.
17. American College of Radiology. PI-RADS ${ }^{\mathrm{TM}}$ prostate imaging-reporting and data system 2015 version 2. Reston (VA): American College of Radiology; 2015. Available from: https://www.acr.org/-/media/ACR/Files/RADS/Pi-RADS/PIR ADS-V2.pdf.
18. Therneau TM, Li H. Computing the Cox model for case cohort designs. Lifetime Data Anal 1999;5:99-112.
19. Therneau TM, Grambsch PM. Modeling survival data: extending the Cox model. New York: Springer; 2000.
20. R Core Team. R: a language and environment for statistical computing. Vienna (Austria): R Foundation for Statistical Computing; 2015.
21. Huynh-Le MP, Myklebust TÅ, Feng CH, Karunamuni R, Johannesen TB, Dale AM, et al. Age dependence of modern clinical risk groups for localized prostate cancer-a population-based study. Cancer 2020;126:1691-9.
22. Muralidhar V, Ziehr DR, Mahal BA, Chen YW, Nezolosky MD, Viswanathan VB, et al. Association between older age and increasing Gleason score. Clin Genitourin Cancer 2015;13:525-30.
23. Draisma G, Postma R, Schröder FH, Van Der Kwast TH, De Koning HJ. Gleason score, age and screening: modeling dedifferentiation in prostate cancer. Int J Cancer 2006;119:2366-71.
24. Shao Y-H, Demissie K, Shih W, Mehta AR, Stein MN, Roberts CB, et al. Contemporary risk profile of prostate cancer in the United States. J Natl Cancer Inst 2009;101:1280-3.
25. Thompson IM, Goodman PJ, Tangen CM, Parnes HL, Minasian LM, Godley PA, et al. Long-term survival of participants in the Prostate Cancer Prevention Trial. N Engl J Med 2013;369:603-10.
26. Thompson IM, Goodman PJ, Tangen CM, Lucia MS, Miller GJ, Ford LG, et al. The influence of finasteride on the development of prostate cancer. N Engl J Med 2003;349:215-24.
27. Lippman SM, Klein EA, Goodman PJ, Lucia MS, Thompson IM, Ford LG, et al. Effect of selenium and vitamin $E$ on risk of prostate cancer and other cancers. JAMA 2009;301:39.
28. Ilic D, Neuberger MM, Djulbegovic M, Dahm P. Screening for prostate cancer. Cochrane Database Syst Rev 2013;(1):CD004720
29. Chen H, Liu X, Brendler CB, Ankerst DP, Leach RJ, Goodman PJ, et al. Adding genetic risk score to family history identifies twice as many high-risk men for prostate cancer: results from the Prostate Cancer Prevention Trial. Prostate 2016; 76:1120-9.
30. Kasivisvanathan V, Rannikko AS, Borghi M, Panebianco V, Mynderse LA, Vaarala MH, et al. MRI-targeted or standard biopsy for prostate-cancer diagnosis. N Engl J Med 2018;378:1767-77.
31. Rouvière O , Puech P , Renard-Penna R, Claudon M, Roy C, Mège-Lechevallier F, et al. Use of prostate systematic and targeted biopsy on the basis of multiparametric MRI in biopsy-naive patients (MRI-FIRST): a prospective, multicentre, paired diagnostic study. Lancet Oncol 2019;20:100-9.
32. Ahmed HU, El-Shater Bosaily A, Brown LC, Gabe R, Kaplan R, Parmar MK, et al. Diagnostic accuracy of multi-parametric MRI and TRUS biopsy in prostate cancer (PROMIS): a paired validating confirmatory study. Lancet 2017;389: 815-22.
33. Carter HB, Albertsen PC, Barry MJ, Etzioni R, Freedland SJ, Greene KL, et al. Early detection of prostate cancer: AUA guideline. Linthicum (MD): American Urological Association; 2018. Available from: https://www.auanet.org/guidelines/ prostate-cancer-early-detection-(2013-reviewed-for-currency-2018).
34. Ilic D, Murphy K, Green S. What do general practitioners think and do about prostate cancer screening in Australia? Aus Fam Phys 2013;42:904-8.
35. Pickles K, Carter SM, Rychetnik L. Doctors' approaches to PSA testing and overdiagnosis in primary healthcare: a qualitative study. BMJ Open 2015;5:e006367.
36. Dunn AS, Shridharani K V, Lou W, Bernstein J, Horowitz CR. Physician-patient discussions of controversial cancer screening tests. Am J Prev Med 2001;20:130-4.
37. Page EC, Bancroft EK, Brook MN, Assel M, Al Battat MH, Thomas S, et al. Interim results from the IMPACT study: evidence for prostate-specific antigen screening in BRCA2 mutation carriers. Eur Urol 2019;76:831-42.
38. Michailidou K, Hall P, Gonzalez-Neira A, Ghoussaini M, Dennis J, Milne RL, et al. Large-scale genotyping identifies 41 new loci associated with breast cancer risk. Nat Genet 2013;45:353-61.
39. Torkamani A, Wineinger NE, Topol EJ. The personal and clinical utility of polygenic risk scores. Nat Rev Genet 2018;19:581-90.
40. DeSantis CE, Siegel RL, Sauer AG, Miller KD, Fedewa SA, Alcaraz KI, et al. Cancer statistics for African Americans, 2016: progress and opportunities in reducing racial disparities. CA Cancer J Clin 2016;66:290-308.
41. Schumacher FR, Al Olama AA, Berndt SI, Benlloch S, Ahmed M, Saunders EJ, et al. Association analyses of more than 140,000 men identify 63 new prostate cancer susceptibility loci. Nat Genet 2018;50:928-36.

## A Genetic Risk Score to Personalize Prostate Cancer Screening, Applied to Population Data

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