Lawrence Berkeley National Laboratory

LBL Publications

Title

Reduction of microbial diversity in grassland soil is driven by long-term climate warming

Permalink

https://escholarship.org/uc/item/9v80x6j2

Journal

Nature Microbiology, 7(7)

ISSN

2058-5276

Authors

Wu, Linwei

Zhang, Ya

Guo, Xue

et al.

Publication Date

2022-07-01

DOI

10.1038/s41564-022-01147-3

Copyright Information

This work is made available under the terms of a Creative Commons Attribution-NonCommercial License, available at $\frac{https://creativecommons.org/licenses/by-nc/4.0/}{https://creativecommons.org/licenses/by-nc/4.0/}$

Peer reviewed

1 Reduction of microbial diversity in grassland soil is driven by long-term climate warming

- 3 Linwei Wu^{1,2,*}, Ya Zhang^{2,*}, Xue Guo^{3,*}, Daliang Ning², Xishu Zhou^{2,4}, Jiajie Feng², Mengting
- 4 Maggie Yuan^{2,5}, Suo Liu³, Jiajing Guo^{2,6}, Zhipeng Gao^{2,7}, Jie Ma^{2,8}, Jialiang Kuang², Siyang
- 5 Jian², Shun Han², Zhifeng Yang², Yang Ouyang², Ying Fu², Naijia Xiao², Xueduan Liu⁴, Liyou
- 6 Wu², Aifen Zhou², Yunfeng Yang³, James M. Tiedje⁹, and Jizhong Zhou², 10,11,§
- 8 ¹Institute of Ecology, Key Laboratory for Earth Surface Processes of the Ministry of Education,
- 9 College of Urban and Environmental Sciences, Peking University, Beijing, China;
- 10 ²Institute for Environmental Genomics and Department of Microbiology and Plant Biology,
- 11 University of Oklahoma, Norman, OK, USA;
- 12 ³State Key Joint Laboratory of Environment Simulation and Pollution Control, School of
- 13 Environment, Tsinghua University, Beijing, China;
- ⁴School of Minerals Processing and Bioengineering, Central South University, Changsha, Hunan,
- 15 China;

2

7

- ⁵Department of Environmental Science, Policy, and Management, University of California,
- 17 Berkeley, CA, USA;
- 18 ⁶Hunan Agriculture Product Processing Institute, Hunan Academy of Agricultural Sciences,
- 19 Changsha, Hunan, China;
- ⁷College of Animal Science and Technology, Hunan Agricultural University, Changsha, Hunan,
- 21 China;
- ⁸School of Environmental Studies, China University of Geosciences, Wuhan, Hubei, China;
- ⁹Center for Microbial Ecology, Michigan State University, East Lansing, MI, USA;
- ¹⁰School of Civil Engineering and Environmental Sciences, University of Oklahoma, Norman, OK,
- 25 USA;

26	¹¹ Earth and Environmental Sciences, Lawrence Berkeley National Laboratory, Berkeley, CA,
27	USA.
28	
29	*These authors contributed equally to this work
30	
31	§Corresponding author: Dr. Jizhong Zhou
32	Institute for Environmental Genomics (IEG)
33	Department of Microbiology and Plant Biology
34	University of Oklahoma
35	Norman, OK 73019
36	Phone: 405.325.6073
37	Fax: 405-325-7552
38	E-mail: jzhou@ou.edu
39	

Abstract [No more than 150 words please]

Anthropogenic climate change threatens ecosystem functioning. Soil biodiversity is essential for maintaining the health of terrestrial systems, but how climate change affects the richness and abundance of soil microbial communities remains unresolved. We examined the effects of warming, altered precipitation and annual biomass removal on grassland soil bacterial, fungal and protistan communities over 7 years to determine how these representative climate changes impact microbial biodiversity and ecosystem functioning. We show that experimental warming and the concomitant reductions in soil moisture played the predominant role in shaping microbial biodiversity by decreasing the richness of bacteria (9.6%), fungi (14.5%), and protists (7.5%). Our results also show positive associations between microbial biodiversity and ecosystem functional processes such as gross primary productivity and microbial biomass. We conclude that the detrimental effects of biodiversity loss might be more severe in a warmer world.

MAIN

Biodiversity, the variety of genes, species, and ecosystems which constitute life on our planet¹, is dramatically affected by human alterations of global environment². Biodiversity underscores healthy ecosystem functions and assures the production of essential goods, services, and benefits to society, such as climate regulation, landscape stability, fibers, and food production¹. However, such benefits are threatened by the unprecedented biodiversity loss^{3,4} caused by anthropogenic global environmental changes like climate warming, altered precipitation patterns, and land use changes⁵. Studies demonstrate that biodiversity loss impairs the functioning of natural ecosystems

and diminishes the number and quality of services they provide⁶. Thus, it is imperative to understand how global environmental change affects biodiversity and the underlying mechanisms⁷.

Anthropogenic climate changes are the greatest threats to biodiversity from local to global scales^{5,6}. The effects of climate change on biodiversity include shifts in species' geographical ranges⁴, extinction⁸⁻¹⁰, changes in abundance within species ranges¹¹, loss of phylogenetic community diversity¹², and increased genetic mutation and selection¹³. In contrast to plants and animals, our understanding of the climate change effects on microbial biodiversity are poorly understood. Previous studies demonstrated the effects of climate warming on soil microbial communities in terms of respiratory feedback responses^{14,15}, decomposition¹⁶, microbial biomass¹⁷, community composition^{14,15,18,19}, community succession¹⁸, temporal scaling¹⁹, and network complexity and stability²⁰. However, there is a paucity of information on the effects of warming on below-ground microbial biodiversity (i.e., alpha diversity) due to the lack of well-replicated, long-term time-series observations under realistic field settings that is necessary to discern clear warming impacts. Therefore, despite a longstanding interest in this topic, whether and how climate warming would result in net soil microbial biodiversity gain or loss, and their underlying mechanisms remain unresolved.

Because different species differ greatly in their temperature-dependent metabolic rates, rising temperature would have dramatic effects on resource consumption, growth, reproduction and interactions between species (e.g., competition, predation, parasitism, and symbiosis)⁹. On one hand, certain species with higher fitness at elevated temperature are likely to have a competitive advantage over other species that are less fit²¹. Consequently, warming could trigger extinction

events at local scales and drive biodiversity loss, which may further cause extinction of other species through coextinction cascades⁸. Similarly, warming and associated environmental changes like decreased moisture would act as strong filtering factors against existing microbial species, which could also cause biodiversity loss. On the other hand, in general, warming promotes plant productivity^{14,15}. Such potentially higher plant diversity and/or quantity of resources could support more microbial species by providing more niches with more ways that species can coexist²¹, and result in biodiversity gain. In addition, the effects of these factors could be intertwined, resulting in no change in biodiversity.

To determine whether and how climate warming affects soil biodiversity, we examined the taxonomic and phylogenetic diversity (PD) of grassland soil bacteria, fungi and protists in a multifactor global change experiment¹⁹ over seven consecutive years from 2009 to 2016, which has warming (+3 °C), altered precipitation levels, i.e., -50% (half precipitation) and +100% (double precipitation) of ambient precipitation, as primary factors and clipping (annual removal of aboveground biomass to simulate the land-use practice of mowing for hay²² or bioenergy²³) as a secondary factor. We address the following major questions: whether and how experimental warming, altered precipitation, and clipping affect soil microbial biodiversity over time; whether such effects vary among different microbial lineages; and what are their underlying mechanisms. We hypothesize that warming would reduce the biodiversity of soil bacteria, fungi, and protists via alternation of both environmental filtering and biotic interactions.

RESULTS

Effects of climate change factors on soil and plant variables

Linear mixed-effects models for determining the sources of variations in hierarchical biological data were first employed to test the effects of treatments and their interactions on soil biogeochemistry and plant communities, in which the regression coefficients represent the directions and magnitudes of the treatment effects, namely effect sizes (β). By comparing the β values, our results revealed that experimental warming, compared to altered precipitation levels and clipping (Fig. 1a, Extended Data Fig. 1), had predominant effects on soil microclimate by increasing temperature, but decreasing moisture (Extended Data Fig. 2a, b; Supplementary note A), and on geochemistry (e.g., decreasing soil pH, increasing NO₃⁻¹) (Extended Data Fig. 2c, d; Supplementary note A). For example, on average, warming decreased soil moisture by 1.5% (absolute) (β = -1.5, p < 0.0001; Extended Data Fig. 2b). In comparison, half precipitation only decreased soil moisture by 0.35% while double precipitation increased soil moisture by 0.7% (β = 0.7, p < 0.0001; Extended Data Fig. 2b). As expected, clipping had significant negative effect on plant biomass, but positive on plant richness (Extended Data Fig. 2f, g; Supplementary note A).

Impacts of warming on microbial biodiversity

It is expected that the alterations in soil microclimate, geochemistry and plant communities would lead to changes in soil microbial biodiversity. Here we define biodiversity²¹ as taxonomic (i.e., species richness and their relative abundance), and phylogenetic¹⁹ diversity in a local community. To test this prediction, all samples were analyzed for bacteria ($56,182\pm27,613$ reads per sample), fungi ($23,569\pm16,323$ reads per sample), and protists ($11,146\pm10,528$ reads per sample) (Extended Data Fig. 3 and S4). Linear mixed-effects models revealed that warming had strong negative effects ($\beta = -0.84 \sim -0.11$, p < 0.007) on richness and other taxonomic diversity indices,

and Faith's phylogenetic diversity (Fig. 1b-f; Table S1; Supplementary note B1). In general, although precipitation alteration or clipping could exert significant effects on richness (for example, precipitation alteration on bacterial richness; Fig. 1b), the effect sizes of warming on richness were $3 \sim 41$ times larger than those of other treatments (Fig. 1b; Table S1; Supplementary note B1). In addition, the effects of treatment interactions were rarely significant except for the positive interactive effects of warming and clipping on fungal and protistan diversity ($\beta = 0.08 \sim 0.91$, p < 0.05) (Table S1), indicating that the warming effect was largely independent of altered precipitation and clipping. Collectively, these results suggest that the diversity of soil bacteria, fungi, and protists is predominantly shaped by experimental warming. A possible explanation is that the changes of microbial biodiversity are mainly driven by soil microclimate and geochemistry such as soil temperature, moisture, and pH²⁴⁻²⁶. As shown above, experimental warming had larger effects on these variables as compared to the effects of the half/double precipitation and clipping treatments. Therefore, we will primarily focus warming-induced treatment effects in subsequent sections.

Overall, warming significantly reduced bacterial richness by 9.6% (β = -0.83, p < 0.0001), fungal richness by 14.5% (β = -0.84, p < 0.0001) and protistan richness by 7.5% (β = -0.99, p < 0.0001). Such negative effects varied yearly with significant impacts on bacteria (β = -1.72 ~ -0.68, p < 0.05) after 2011, and fungi (β = -2.15 ~ -0.36, p < 0.05) after 2013, and protists in 2011, 2013 and 2014 (β = -1.44 ~ -0.60, p < 0.05) (Extended Data Fig. 5, Supplementary note B2). Rarefaction analyses indicated that the observed richness for bacteria, fungi, and protists were always lower under warming than non-warming control except warming & double precipitation & clipping (WDC) versus double precipitation & clipping (DC) for fungi and warming & clipping (WC)

versus clipping (C) for protists (Paired t test, p <0.0001, Extended Data Fig. 4). Warming also significantly decreased the phylogenetic diversity of bacteria by 7.2% (β = -0.49, p < 0.0001), fungi by 9.3% (β = -0.47, p = 0.002), and protists by 4.5% (β = -0.80, p = 0.003) based on Faith's PD, the phylogenetic analogue of taxon richness (Fig. 1f, Table S1). In addition, consistent to warming-induced biodiversity decrease, warming significantly reduced microbial biomass as determined by phospholipid fatty acid analysis (PLFA) (β = -0.83, p = 0.046), and DNA yields (β = -0.72, p = 0.002; Fig. 1g and Table S2; Supplementary note B1). Collectively, all of these results indicate that experimental warming significantly reduced microbial biodiversity.

The negative warming effects on microbial biodiversity varied considerably among different microbial lineages. Warming significantly decreased the richness of most microbial phyla (Fig. 2a), as well as their phylogenetic diversity (Extended Data Fig. 6a; Supplementary note C1). Warming had the largest negative effects on the richness of Acidobacteria, Verrucomicrobia, and Planctomycetes (β = -1.21 to -1.19, p < 0.01), but had a significant positive effect on the richness of Firmicutes (β = 1.52, p < 0.01; Fig. 2a). Similar to species richness, warming significantly decreased the relative abundance of Acidobacteria, Verrucomicrobia and Planctomycetes (β = -0.88 ~ -0.84, p < 0.01). In contrast, warming increased the relative abundance of Actinobacteria, Firmicutes and Gemmatimonadetes (β = 0.52 ~ 1.05, p < 0.05; Extended Data Fig. 6b), which could be due to their preference for drier soils²⁷⁻²⁹. Notably, the increase of Firmicutes and Actinobacteria may be in part due to their spore-forming ability³⁰, which makes them resistant to desiccation stress. In support of this, we examined the characteristics of spore-forming bacteria in more detail. Almost all the families of Firmicutes and Actinobacteria that were increased under warming are known spore-formers³⁰ (Extended Data Fig. 7a, b). In addition, the relative

abundances of the major sporulation genes in Firmicutes (spo0A) and Actinobacteria (bldD), as identified from shotgun sequencing metagenome data, also significantly or marginally significantly increased under warming (p < 0.0001 for bldD; p = 0.08 for spo0A; Extended Data Fig. 7c). Similar patterns were also observed at individual ASV (amplicon sequence variant) level (Fig. 2b). For instance, most taxa in Verrucomicrobia (78.5% of ASVs of Verrucomicrobia, 91.8% of relative abundance) decreased under warming, while most taxa in Firmicutes (88.0% of ASVs, 98.9% of relative abundance) increased (Fig. 2b; Supplementary note C2). In addition, warming effects varied among different fungal guilds as classified by FUNGuild³¹. Warming reduced the richness, phylogenetic diversity, and abundance of arbuscular mycorrhiza fungi (AMF) ($\beta = -1.05$ \sim -0.42, p < 0.01; Fig. 2a; Extended Data Fig. 6a, b; Supplementary note D1), which are beneficial microorganisms capable of forming mutualistic symbiosis with plants. The negative warming effect on AMF abundance was also supported by AMF biomass decreases as determined by PLFA $(\beta = -0.54, p = 0.013; Fig. 1g)$. Interestingly, although warming decreased the richness of putative plant pathogenic fungi (Fig. 2a), it marginally increased their relative abundance ($\beta = 0.43$, p =0.055; Extended Data Fig. 6b), which could have negative effects on plant growth. Moreover, warming significantly reduced the richness, phylogenetic diversity and abundance of Cerozoa and Ochrophyta ($\beta = -1.07 \sim -0.20$, p < 0.002) but increased the richness and phylogenetic diversity of Conosa ($\beta = 0.05 \sim 0.12$, p < 0.02) (Fig. 2a; Extended Data Fig. 6a, b; Supplementary note C1). Similarly, warming significantly decreased the richness and phylogenetic diversity of various functional groups of protists (i.e., consumers, phototrophs, and parasites) ($\beta = -0.98 \sim -0.39$, p <0.04). Warming also reduced the relative abundance of phototrophic protists ($\beta = -0.17$, p = 0.01) (Fig. 2a; Extended Data Fig. 6a, b). These results suggest that warming has differential impacts on various microbial lineages and/or functional guilds, which are consistent with our previous

177

178

179

180

181

182

183

184

185

186

187

188

189

190

191

192

193

194

195

196

197

198

199

observations that warming effects vary greatly among different microbial functional groups¹⁵. Warming-induced diversity decrease on most microbial categories could have significant impacts on ecosystem functioning, as suggested by previous reports in macroecology^{3,4,6} and microbial ecology³². Particularly, since warming decreased beneficial taxa such as AMF, the aboveground plant community could be negatively impacted.

205

206

207

208

209

210

211

212

213

214

215

216

217

218

219

220

221

222

200

201

202

203

204

Mechanisms underlying reduced microbial biodiversity

As we posited earlier, warming-induced biodiversity decrease could be due to changes in biotic interactions and abiotic environmental conditions caused by warming. Under warmer conditions, many microbes with adaptive traits (e.g., Firmicutes and Actinobacteria with spore-forming ability) would survive and outcompete other microbes (e.g., Acidobacteria, Verrucomicrobia, and Planctomycetes) (Fig. 2a, b). Consequently, species coexistence patterns would be substantially altered, as revealed by a network analysis showing that the occurrence network was more complex under warming than non-warming control²⁰. The increased positive connections may indicate more microbial cooperations³³, which could be important for their survival under warming. Also there were more negative connections under warming than control²⁰, suggesting there might be more intense competition under warming. Eventually, the warming-induced changes in microbial activities and interactions could trigger various extinction events and ultimate biodiversity decrease due to cascading effects⁸. Alternatively, warming could just act as a deterministic filtering factor to impose significant positive selection on spore-forming microorganisms (e.g., Bacillaceae 2) and/or negative selection on nonspore-forming microorganisms (e.g. Acidothermaceae), which is consistent with the observation that warming enhanced homogeneous selection on Bacillales in Firmicutes³⁴. All of these results suggest that both biotic interactions and

environmental filtering could play important roles in mediating warming-induced biodiversity decrease.

225

226

227

228

229

230

231

232

233

234

235

236

237

238

239

240

241

242

243

244

245

223

224

It is anticipated that soil environmental conditions should also play important roles in driving microbial biodiversity decrease. As shown in Fig. 3a, bacterial, fungal, and protistan richness were highly correlated with soil moisture, temperature and NO_3 -N content (LMM's r = -0.25 ~ 0.24, p < 0.01). Bacterial richness also showed significant correlations with plant richness and biomass (LMM's $r = 0.11 \sim 0.19$, p < 0.05; Supplementary note D1). However, obvious collinearity among these variables also occurred (Fig. 3a) (Supplementary note D1). Thus, to further disentangle the direct and indirect effects of the environmental drivers on microbial biodiversity, structural equation modeling (SEM) analyses were performed with the presumed relationships (Extended Data Fig. 8) among the selected subsets of plant and soil variables which were least-correlated (see Methods for details of model selection). Soil moisture, which was negatively affected by warming (standardized path coefficient, b = -0.69) and half precipitation (b = -0.16), but positively by double precipitation (b = 0.45), played the strongest role in shaping bacterial richness directly (b = 0.43, p = 0.001; Fig. 3b; Supplementary note D2). Soil pH, plant richness and the biomass of C₃ plants were also significantly and positively (b = 0.23-0.31, p < 0.02) correlated to bacterial richness. Furthermore, bacterial richness directly and positively affected protistan richness (b = 0.69, p < 0.001). In comparison, among the variables which directly contribute to fungal richness, only paths of soil moisture (b = 0.44, p = 0.001) and plant richness (b = 0.26, p = 0.015) were significant (Extended Data Fig. 9), suggesting that the environmental drivers appear different between bacteria and fungi. Overall, those variables can explain 61%, 51%, and 50% of the variations in bacterial, fungal and protistan richness (Fig. 3b, Extended Data Fig. 9), respectively.

In addition, SEM analysis revealed that warming played predominant roles in shaping microbial diversity (standardized total coefficient = -0.61 for bacteria, -0.56 for fungi and -0.51 for protists, Fig. 3c) as compared to precipitation or clipping treatments (standardized total coefficients = -0.05 ~ 0.31, Fig. 3c), which was consistent with linear mixed-effects model analysis (Fig. 1a). These results indicate that soil and plant variables, particularly soil moisture, are also important in mediating warming-induced soil microbial diversity decrease directly and indirectly. Since the SEMs could explain over half of the variations in microbial diversity, the environmental filtering effects, especially the induced desiccation stress, could be the main driver for microbial diversity decrease via affecting microbial activities and interactions.

Links between microbial biodiversity and ecosystem functions

A following important question is whether the warming-induced changes in microbial diversity affects ecosystem functional processes. Consistent with the reduced microbial biodiversity, warming also decreased the ecosystem functions of total microbial biomass, bacterial biomass, GPP, and ER (β = -0.17 ~ -0.84, Extended Data Fig. 10). In agreement with various reports in macroecology⁶, the overall bacterial richness had significant positive correlations with total microbial biomass, bacterial biomass, gross primary productivity (GPP), and ecosystem respiration (ER) (r = 0.14 ~ 0.22, p < 0.002) (Fig. 3d; Supplementary note E). Similar positive correlation patterns were also observed for most bacterial groups (e.g., Proteobacteria, Bacteroidetes, Planctomycetes) (Fig. 3d; Supplementary note E) except for Firmicutes, which showed significant negative correlations with total microbial biomass, bacterial and fungal biomass, and ER (r = -0.26 ~ -0.10, p < 0.04). In addition, the overall richness of fungi and most fungal phyla/guilds showed significant positive correlations with GPP and ER (Fig. 3d;

Supplementary note E). The overall protistan richness, major protistan lineages and functional groups also had significant positive correlations with total microbial biomass, bacterial biomass, GPP and ER ($r = 0.08 \sim 0.22$, p < 0.04) (Fig. 3d; Supplementary note E). All of these results indicate that there are significant positive linkages between microbial community diversity and relevant ecosystem functional processes.

DISCUSSION

Understanding how climate change affects microbes and the underlying mechanisms is a critical issue in climate change and microbiology research³⁵. By examining the dynamic changes of soil microbial biodiversity in a well replicated long-term climate change experiment, this study provides explicit evidence that climate warming consistently reduced the taxonomic and phylogenetic diversity of soil bacteria, fungi, and protists across different years. In addition, by examining the interactive effects of warming, precipitation level, and clipping on microbial diversity, this is also the first study to demonstrate that warming plays a predominant role in driving soil biodiversity decrease via altering biotic interactions and soil biogeochemical conditions, particularly soil moisture, which is in agreement with the fact that warming had prevalent effects on accelerating the temporal scaling rates of soil microbial biodiversity¹⁹. Finally, warming-induced diversity decrease could have significant impacts on ecosystem functioning, which augments previous reports in macroecology^{3,4,6}.

Our findings have important implications for predicting ecological consequences of climate change and for ecosystem management. Because warming as a deterministic filtering factor drives

microbial biodiversity decrease¹⁸, the ecosystems under future climate change scenarios would be less diverse. Along with faster biodiversity turnover rates as previously demonstrated¹⁹, it is expected that the linked ecosystem functions and services could become more vulnerable in a warmer world⁶. Particularly, since warming has differential effects on different microbial lineages, such as decrease of beneficial taxa (e.g., AMF), the detrimental effects of biodiversity loss on future ecosystem functioning could be more severe. In addition, since warming-effects on biodiversity is primarily via reduced moisture, it is expected that warming-induced biodiversity decrease could be more severe in drylands (i.e., arid, semi-arid, and dry-subhumid ecosystems), covering 41% of Earth's land³⁶, as compared to wet regions. The future warming-induced precipitation changes could also be important in mediating warming-induced biodiversity decrease. However, further research is necessary to determine whether the warming-induced biodiversity decrease and associated mechanisms are applicable to other ecosystems.

Methods

Study site and sampling

We conducted the warming experiment at the Kessler Atmospheric and Ecological Field Station (KAEFS) in the US Great Plains in McClain County, Oklahoma (34° 59′ N, 97° 31′ W)^{18-20,22}. The acting director of the site is Meghan Bomgraars (mbomgaars@ou.edu). Detailed site description can be found in the Supplementary note F. In brief, KAEFS is an old-field tall-grass prairie with dominant plants of C₃ forbs (*Ambrosia trifida*, *Solanum carolinense* and *Euphorbia dentate*) and C₄ grasses (*Sorghum halepense* and *Tridens flavus*)²². Based on Oklahoma Climatological Survey data from 1948 to 1999, the air temperature ranges from 3.3 °C in January to 28.1 °C in July with mean annual temperature 16.3 °C, and the precipitation ranges from 82 mm in January and

February to 240 mm in May and June with mean annual precipitation 914 mm¹⁵. The soil type of this site is Port–Pulaski–Keokuk complex, and soil texture class is loam with 51% of sand, 35% of silt and 13% of clay³⁷.

The field site experiment was established on July in 2009 with a blocked split-plot design (Extended Data Fig. 1), in which warming (continuous heating at a target of +3 °C above ambient temperature) and precipitation alteration (targets of -50% and +100% ambient precipitation) were primary factors nested with clipping (annual removal of aboveground biomass in peak growth season) as the secondary factor¹⁸⁻²⁰. The clipping treatment was used to mimic the land-use practice of hay harvest, which is widely practiced in the southern Great Plains of the US²², and biomass harvest for bioenergy²³. In brief, the site has four experimental blocks, each including six plots. Each plot has the size of 2.5×3.5 m², which was further divided into one 2.5×1.75 m² clipped subplot and one 2.5×1.75 m² unclipped subplot, resulting in a total of 48 subplots (Extended Data Fig. 1).

From 2009 to 2016, surface (0–15 cm) soil samples were collected annually from subplots one day before annual clipping. Each sample was mixed from three soil cores (2.5 cm diameter × 15 cm depth) from a soil sampler tube. In the first year (2009), we collected 24 pre-warmed soil samples from the southern subplots. As for the following years, a total of 48 annual soil samples were collected from all subplots in each year. A total of 360 annual soil samples from 2009 to 2016 were collected in this study and stored in a freezer at –80 °C.

Field measurements and soil chemical analyses

Soil temperature was measured every 15 min at depth of 7.5, 20, 45 and 75 cm in the center of each plot using constantan-copper thermocouples wired to a Campbell Scientific CR10x data logger (Campbell Scientific)¹⁸⁻²⁰. Annual average values of temperature at depth 7.5 cm were used to represent soil temperature across experimental years. Volumetric soil water content (%V) was measured using a portable time domain reflectometer (Soil Moisture Equipment Corp.) once or twice a month, and annual average values were used to represent soil moisture¹⁸⁻²⁰. Ecosystem carbon (C) fluxes, including ecosystem respiration (ER), net ecosystem exchange (NEE), soil total respiration (R_s) and heterotrophic respiration (R_h) were measured once or twice a month between 10:00 and 15:00 (local time)¹⁸⁻²⁰. The gross primary productivity (GPP) was then estimated as the difference between net ecosystem exchange and ecosystem respiration.

Above-ground plant community surveys were conducted at peak biomass (usually September) each year. All species within each plot were identified to estimate species richness. From 2009 to 2015, the above-ground plant biomass, separated into C_3 and C_4 species, were estimated by a modified pin-touch method, as described by Sherry et al³⁸. Since 2016, thorough plant survey was conducted, and linear regression models were used to estimate above-ground biomass based on plant height and abundance³⁹. Every individual plant was surveyed in the whole plot, with their species identity and height recorded. For each species, we also measured the height across different individuals off-plot, after which we harvested, dried, and weighed the vegetation for which we had recorded. We then constructed regression relationships between plant height and dry weight biomass for each species. The regression model for one species, *Pseudognaphalium obtusifolium* was insignificant (R^2 =0.07, p = 0.48), and we used the mean biomass across individuals collected off-plot as the biomass estimate for its individual in the plot. The average adjusted R^2 for all other

regression models was 0.76 and the average p value was 0.0056. The total above-ground biomass was then calculated as the sum of estimated biomass for every individual.

Visible stones and plant roots were removed from the soil by metal forceps before chemical and microbial analyses. The chemical properties of all soil samples were analyzed in the Soil, Water, and Forage Analytical Laboratory at Oklahoma State University (Stillwater, OK, USA). Briefly, the total C and total nitrogen (N) contents were determined using a dry combustion C and N analyzer (LECO). Soil nitrate (NO₃⁻) and ammonia (NH₄⁺) were analyzed using a Lachat 8000 flow-injection analyzer (Lachat). Soil pH was determined using a pH meter with a calibrated combined glass electrode at a water-to-soil mass ratio of 2.5:1⁴⁰.

Soil phospholipid fatty acids (PLFA)

Lipids were extracted from the soil samples based on the modified Bligh-Dyer method as described by Buyer and Sasser⁴¹. In brief, soil samples were freeze-dried and sifted to remove any rocks or large debris. 2 g of each freeze-dried soil sample was then incubated in a 2:1:0.8 solution of methanol, chloroform, and K₂HPO₄ buffer. The chloroform phases were collected and the phospholipids were separated from neutral lipids and glycolipids through silicic acid chromatography, subsequently saponified and methylated to fatty-acid methyl esters. The resulting fatty acid methyl esters were separated and identified using gas chromatography (Agilent 6890N, Wilmington, DE). The peak responses were translated into molar responses using an internal standard, and were fitted with a MIDI Sherlock microbial identification system (Version 4.5, MIDI, Newark, NJ). Further, the peak responses were assigned to microbial groups including gramnegative bacteria, gram-positive bacteria, Actinobacteria, anaerobic bacteria, common fungi, and

arbuscular mycorrhizal fungi (AMF) using the Agilent Chemstation software (Agilent Technologies). The total bacterial biomass of the soils was calculated as the total PLFA of all bacterial groups, i.e., the sum of the biomass of gram-negative bacteria, gram-positive bacteria, actinobacteria, and anaerobic bacteria. The total fungal biomass was calculated as the sum of the biomass of common fungi and AMF.

DNA extraction

It is well known that sequence quality is subjected to big variations from DNA extraction, PCR amplification and sequencing. It is critical to control each step to generate the high quality of experimental data. Thus, great cautions were taken in this study to ensure the highest quality of sequencing data with more tedious and laboratory experimental protocols. For DNA extraction, the experimental method with grinding, freeze-thawing and sodium dodecyl sulfate (SDS)-based cell lysis⁴² was used. This method has been most widely used in microbial molecular ecology. In brief, for each soil sample, microbial DNA was extracted from 1.5 g soil using this grinding, freeze-thawing method⁴², and purified with a PowerSoil® DNA isolation kit (MoBio Laboratories) following the manufacturer's protocol. DNA quality was evaluated based on the 260/280 nm and 260/230 nm absorbance ratios using a NanoDrop ND-1000 Spectrophotometer (NanoDrop Technologies). All samples had the 260/230 ratios larger than 1.7 and 260/280 ratios larger than 1.8. DNA concentration was measured by PicoGreen using a FLUOstar Optima fluorescence plant reader (BMG Labtech). DNA samples were stored at -80 °C until use.

Amplicon sequencing and data preprocessing

We used a two-step PCR amplification method for library preparation of 16S rRNA gene (V4 region), the intergenic region (ITS) between the 5.8S and 28S rRNA genes and 18S rRNA gene (V9 region) to improve sequence representation and quantification^{43,44}. During the first amplification step, 10 ng DNA from each sample was PCR-amplified for 10 cycles in triplicate in 25 µl reaction with the primers without adaptors. The obtained PCR products were purified and dissolved in 50 µl D.I. water. This initial amplification step avoided potential amplification bias caused by long tail of adaptors and other added components. During the second amplification step, 15 μl of the PCR products from each sample were amplified using the primers with all adaptor, barcode, and spacers in triplicate for an additional 15 cycles. The low total cycle numbers (25-30 cycles) ensure that the PCR amplification is not saturated and limits amplification artifacts. Finally, the triplicate amplified products were combined, purified, and quantified for subsequent sequencing using the same MiSeq instrument with 2×250 base pair kits at the Institute for Environmental Genomics, University of Oklahoma. The two-step PCR amplification method with phasing primers in triplicate can help reduce sequencing errors, minimize amplification bias, and preserve semi-quantitative information of PCR amplification⁴³, which is critical for subsequent data analysis, data interpretation, and biological inference⁴⁴.

422

423

424

425

426

427

406

407

408

409

410

411

412

413

414

415

416

417

418

419

420

421

The primer sequences were trimmed from the paired-end sequences, which were then merged using FLASH⁴⁵. Any merged sequences with an ambiguous base or a length of < 245 bp for the 16S rRNA gene, < 220 bp for the ITS, or < 330 bp for 18S rRNA gene were further discarded. An average of $56,182\pm27,613$, $23,569\pm16,323$, and $56,874\pm55,642$ sequence reads were obtained for 16S rRNA gene, and ITS and 18S rRNA gene, respectively (Extended Data Fig. 3a, b). These

high-quality 16S rRNA gene, ITS or 18S rRNA gene sequences were processed to generate amplicon sequence variants (ASVs; also known as unique sequence variants and zero-radius operational taxonomic units) by UNOISE3⁴⁶. Rarefaction analyses (Extended Data Fig. 4) indicated that the sequencing depth was sufficient for assessing the effects of various climate change factors on the diversity of these soil microbial communities.

433

434

435

436

437

438

439

440

441

442

443

444

445

446

447

448

449

450

432

428

429

430

431

The representative 16S rRNA or 18 rRNA gene sequences were aligned using Clustal Omega v1.2.2⁴⁷ for constructing the phylogenetic tree by FastTree2 v2.1.10⁴⁸. The FastTree topology search was constrained with the relatively reliable 16S-based bacterial tree in Silva Living Tree Project⁴⁹ release 132. The fungal phylogenetic tree was constructed using 'constrained topology search' in FastTree v2.1.11⁵⁰. A guide tree was built from the full-length SSU sequences of 511 representative species, one species from each fungal family. Then, the full tree was built from the observed ITS sequences with the constraint alignment converted from the guide tree. The SSU sequences were retrieved from Silva 138.1 Ref NR database. For 16S rRNA gene, the ASVs or OTUs were taxonomically annotated with RDP Classifier using 16S rRNA gene training set 16 with a confidence cutoff of 50%⁵¹, and Chloroplast and Mitochondria were further removed from the bacterial profiles. For ITS, the ASVs or OTUs were taxonomically annotated with RDP Classifier using UNITE Fungal ITS training set (version of August 2018)⁵²; ITS sequences were further assigned into three functional groups—plant pathogens, AMF, and saprotrophs using FUNGuild³¹. Sequences that had multiple function assignments or in FUNGuild were termed as unassigned guild in this analysis. The sequence number in each sample was rarefied to the same depth for 16S rRNA gene (22,599) or ITS sequences (7,761) in subsequent comparative analyses. For 18S rRNA gene, the ASVs were taxonomically annotated with PR2 databases⁵³. We also

assigned the major protistan lineages to their dominant mode of energy acquisition (i.e., trophic functional groups)—either phototrophic, parasitic, or as consumers, following the classifications in Oliverio et al⁵⁴. Sequences annotated as protists were further rarefied to 1,100 for subsequent comparative analyses.

455

456

457

458

459

460

461

462

463

464

465

466

467

468

469

470

471

472

473

451

452

453

454

Sporulation gene profiling from metagenomic dataset

The soil samples under single treatment of warming (warming and normal precipitation and unclipped) and control (ambient temperature and normal precipitation and unclipped) were selected for metagenomic sequencing. That is, 8 subplots \times 8 years = 64 metagenomic samples. Libraries were constructed using genomic DNA with KAPA Hyper Prep Kit (KR0961) based on the manufacturer's instruction, and DNA was sequenced using an Illumina HiSeq 2500 platform. A total of 1100.14 gigabases (Gb) were generated, with an average of 17.19 \pm 2.68 Gb per sample. The quality of the metagenomic data was evaluated using FastOC v0.11.6⁵⁵. CD-HIT⁵⁶ was used to remove duplicates with an identity cutoff of 100%. NGS QC Toolkit (version 2.3.3)⁵⁷ was used for quality filtering, where poor-quality bases with quality score <20 were trimmed from the 3 end until the first base had a quality score ≥ 20 . Trimmed reads with length of > 120 and the average quality score ≥ 20 were kept. In addition, reads with more than one ambiguous base were removed. High-quality reads were then converted to fasta format, split into multiple partitions, and searched against NR database (BLASTx) using DIAMOND⁵⁸ with E value cutoff of 1 × 10⁻⁵, coverage cutoff of 0.5, and maximum target number of 50. The outputs were submitted to MEGAN6 (Ultimate Edition, version 6.6)⁵⁹ for function profiling with parameter of top percent of hits 10%, minimum score 50 and minimum support 1. The annotated functional profiles of SEED Subsystem (3 levels) were exported, and two major sporulation genes annotated at level 3, i.e., spo0A gene of Firmicutes and *bldD* gene of Actinobacteria were selected. Their relative abundances were then determined by dividing the annotated sequence counts by the total number of high-quality sequences of the corresponding metagenomes.

Statistical analyses

This study is based on a well-designed long-term climate change experiment with a blocked split-plot design, with 12 treatment combinations of warming, precipitation levels, and clipping. Each combination has 4 replicated plots (Extended Data Fig. 1). Also, the same plots were repeatedly sampled over 8 years and high quality of experimental data were generated, which greatly increased the power for various robust data analyses to ensure the liability of the statistical inference.

Diversity analyses

Richness was used to measure taxonomic α -diversity, using the *Picante* R package^{60,61}. Other taxonomic α -diversity indices, including Shannon index, inverse Simpson index and Pielou's evenness were also calculated using the *vegan* R package⁶². Faith's index, which is the sum of the total phylogenetic branch length based on the phylogenetic tree constructed, was used to measure phylogenetic α -diversity using the *Picante* R package⁶⁰.

Treatment effects by linear mixed-effects models

Due to block design and repeated measurements, the experimental data are not completely independent. Therefore, linear mixed-effects models (LMMs) were used to assess the effects of

experimental treatments on environmental variables, microbial diversity or the relative abundance of microbial groups. The lme4 R package was used to implement LMMs⁶³. In the linear mixed-effects models, warming (0 for ambient temperature and 1 for warming), precipitation level (0.5 for half, 1 for normal and 2 for double precipitation level) and clipping (0 for unclipped and 1 for clipped) treatments and their interactions were considered as fixed effects, while the sampling time (year) and the block were termed as random intercept effects ($y \sim warming \times precipitation level \times clipping + (1|Block) + (1|year)$). That is, three variables (warming, precipitation level, and clipping) were created to denote the experimental treatments for soil samples. The reason that the precipitation treatments were not treated as categorical variables is because a categorical variable can not capture the gradient of precipitation levels, which is assumed to linearly correlate with soil moisture and other variables.

We also tested alternative models in which the effects of sampling year and its interaction with experimental treatments were considered as fixed effects ($y \sim \text{warming} \times \text{precipitation level} \times \text{clipping} \times \text{year} + (1|\text{Block})$; or $y \sim \text{warming} \times \text{precipitation level} \times \text{clipping} + \text{year} (1|\text{Block})$), yet the corresponding model with sampling year as random intercept effect was better, based on lower AIC values. Thus, we decided to use the model with experimental treatments as fixed effects with year and block as random intercept effects. Effect sizes of treatments or treatment interactions were represented by the regression coefficients in the LMMs. Wald type II χ^2 tests were used to calculate the p values from the LMMs using the car R package⁶⁴. Since the precipitation level is considered as a continuous variable in the LMM (0.5 for half precipitation, 1 for normal and 2 for double precipitation), only one regression coefficient of precipitation treatment would be derived by the LMM. The effect size of half precipitation (as compared to ambient precipitation) can be

derived by multiplying the regression coefficient by - 0.5, while the effect size of double precipitation (as compared to ambient precipitation) can be derived by multiplying the regression coefficient by 1. For instance, as shown in Extended Data Fig. 2b, precipitation level was positively correlated to soil moisture with β = 0.7 per fold change (+100%) of precipitation (p < 0.0001; Extended Data Fig. 2b). That is, the double precipitation treatment only caused a 0.7% × 1 = 0.7% (absolute) increase in soil moisture, while half precipitation changed soil moisture by 0.7% × (-0.5) = -0.35%, i.e., half precipitation decreased soil moisture by 0.35% on average.

Predicting microbial diversity with environmental variables

To link the environmental variables to microbial diversity, the correlations between the individual environmental variable and bacterial or fungal diversity was tested by the linear mixed-effects model, in which the sampling year and the block were termed as random intercept effects. Since richness is highly correlated with other diversity indexes (Fig.S1), it is used to represent microbial biodiversity. The marginal coefficient of determination (pseudo-R-squared) was calculated using function 'r.squaredGLMM' of the MuMIn R package⁶⁵, which represents the variance explained by the fixed effect in the linear mixed-effects model.

Because most soil and plant predictor variables were strongly correlated with each other (Fig. 3a), we further performed a model selection analysis to compile sets of variables that are strong predictors of the data, least-correlated, or potentially biologically informative on the basis of *a priori* assumptions. The following tests were performed. First, the contribution of each variable on predicting bacterial or fungal richness was ranked using the method of bootstrap forest partitioning, conducted by the function of 'Predictor Screening' in JMP 15.0 (SAS Institute) (Table S3 & S4).

This method evaluates the relative contribution of predictors on the response⁶⁶, which can identify predictors that might be weak alone, but strong when used in combination with other predictors. Second, the Pearson correlation coefficient between variables were calculated (Table S5) to reveal the colinearty between variables. We then applied different model selection strategies to obtain sets of variables for our linear mixed models, in which the random intercept effects of sampling year and block were included. Three alternative strategies were pursued to select soil and plant variables to be included in candidate models, based on their collinearity and/or informed hypotheses: (i) Model 1 (Table S6 for bacteria and Table S7 for fungi). For highly correlated variables which have similar biological inference, we kept the one with the highest contribution based on Predictor Screening. For example, we kept soil annual mean moisture rather than soil moisture at the sampling month for both LMMs on bacterial and fungal richness. For variables on plant biomass, we kept C₃ plant biomass for the LMM on bacterial richness while we kept C₄ plant biomass for that on fungal richness. Then, we iteratively removed variables with the highest correlation coefficients and kept variables that had high contributions in predictor screening, until the correlation between the remaining variables were less than 0.5. Through this way, each set of six variables were selected for the bacterial and fungal model. (ii). Model 2 (Table S6, S7). We kept six least-correlated variables by removing all variables with a correlation of r > 0.5 using R caret package⁶⁷. (iii). Model 3 (Table S6, S7). The six variables with largest contributions from predictor screening were kept.

562

563

564

565

543

544

545

546

547

548

549

550

551

552

553

554

555

556

557

558

559

560

561

The random intercept effects of sampling year and block were included in Models 1-3. Models 1-3 were compared based on their AIC values, and the model with the lowest AIC chosen as the preferred model. Soil temperature, moisture, pH, and total plant richness were selected for both

bacteria and fungi although slight differences were observed with respect to nitrogen and plant biomass (Table S6, S7). These variables in the preferred model were used in the downstream structural equation modelling analyses.

Structural equation modeling

To further discern the direct and indirect effects of the environmental drivers on microbial biodiversity, structural equation modelling (SEM) analyses were performed to examine the relationships among experimental treatments, soil and plant variables, and microbial diversity. To correct the potential temporal autocorrelation, we used data at the plot level, by averaging the microbial or environmental data across time points of the same plot. We first considered a hypothesized conceptual model (Extended Data Fig. 8) that included all reasonable pathways. Then, we sequentially eliminated non-significant pathways unless the pathways were biologically informative, or added pathways based on the residual correlations. The procedure was repeated until the model showed sufficient fitting with the p values of χ^2 test larger than 0.05 (i.e., the predicted model and observed data are not significantly different) and the root mean square error of approximation (RMSE) < 0.08. The SEM-related analysis was performed using the lavaan R package⁶⁸.

Data availability

The DNA sequences of the 16S rRNA gene, 18S rRNA gene and ITS amplicons were deposited to the National Center for Biotechnology Information (NCBI) under the project accession number PRJNA331185. Raw shotgun metagenomic sequences are deposited in the European Nucleotide Archive (http://www.ebi.ac.uk/ena) under study no. PRJNA533082. Silva 138.1 Ref NR database

is available at https://www.arb-silva.de/documentation/release-138/. Protist Ribosomal Reference database (PR2) databases is available at https://github.com/pr2database/pr2database. The ASV table and ASV representative sequences, soil physical and chemical attributes, and plant biomass and richness are downloadable online at http://www.ou.edu/ieg/publications/datasets.. Source data are provided with this paper.

Code availability

- R scripts for statistical analyses are available on GitHub at https://github.com/Linwei-
- 597 Wu/warming soil biodiversity.

Acknowledgements

We thank numerous former and current members in the Institute for Environmental Genomics for their help in maintaining the long-term field experiment. This work is supported by the US Department of Energy, Office of Science, Genomic Science Program under Award Number DE-SC0004601 and DE-SC0010715, and the Office of the Vice President for Research at the University of Oklahoma to J.Z.. X.G. and X.Z. were generously supported by China Scholarship Council (CSC) to visit the University of Oklahoma. The data analyses performed by X. G. were also supported by the China Postdoctoral Science Foundation (2018M641327 and 2019T120101).

Author contributions

All authors contributed intellectual input and assistance to this study. The original concepts were conceived by J.Z. and J.M.T. Field management was carried out by Linwei W., Y.Z., X.G., J.F., M.Y., J.K., Y.F., A.Z., D.N., J.M., S.J., S.H., Z.Y., Y.O. and Liyou W. Sampling collection, soil

chemical and microbial characterization were carried out by M.Y., X.G., Linwei W., J.G., Z.G., and X.Z. Data analysis were done by Linwei W., Y.Z., X.G., and N.X. with the assistance provided by D.N. and J.Z. All data analysis and integration were guided by J.Z. The manuscript was prepared by J.Z., Linwei W., Y.Z., X.G., with significant input from J.M.T., Y.Y., X.L.. Considering their contributions in terms of site management, data collection, analyses and/or integration, Linwei W., Y. Z., and X.G. were listed as co-first authors.

Competing interests

The authors declare no competing interests.

Figure legends

Fig. 1. Effects of experimental warming on soil microbial communities. a, Experimental settings for treatments. b-d, The effect sizes of warming, altered precipitation level and clipping on microbial richness (b), phylogenetic diversity (c) and biomass (d). The estimated effect sizes are regression coefficients based on rescaled response variables (with zero mean and unit standard deviation) in the linear mixed-effects models. Statistical significance is based on Wald type II χ^2 tests (n = 360). Bars represent mean effect sizes and error bars represent standard errors. The effects of treatment interactions and exact p values are indicated in Table S1 & S2. Significances of χ^2 tests are denoted by asterisks: *** p < 0.001, ** p < 0.01, * p < 0.05. PD, Faith's phylogenetic diversity. PLFA, phospholipid fatty acid analysis. AMF, arbuscular mycorrhiza fungi.

Fig. 2. Effects of experimental warming on different microbial taxa. a, Effect sizes of warming on the (rescaled) richness of major microbial groups based on linear mixed-effects models. Data are presented as mean values \pm standard errors of the estimated effect sizes. Statistical significance is based on Wald type II χ^2 tests (n= 360), which is denoted by asterisks: *** p <0.001, ** p < 0.01, * p < 0.05. Insignificant changes are denoted by grey dots. **b**, The phylogenetic relationship of individual bacterial ASVs (amplicon sequence variants, the first ring inside). Only the ASVs with a significant response (adjusted p < 0.05) to warming and their average read numbers ≥2 among warmed or unwarmed samples were included in the tree. The outside and inside bars of the second ring represent the positive and negative effect sizes of warming on rescaled taxon relative abundances. Colors of the branches in the first ring and the bars in the second ring correspond to individual phyla or classes, and the grey color in the tree indicates unclassified or other minor phyla. Colors in the third ring represent ASVs with significant increase (yellow) or decrease (blue) under warming. The area sizes of the pies reflect the total relative abundance of bacterial phyla/classes across all samples, where the yellow and blue parts represent the proportions of the total abundance of ASVs which increased and decreased under warming, respectively.

Fig. 3. Environmental drivers of microbial diversity. a, Correlations between environmental variables and microbial diversity. Edge width corresponds to the absolute value of correlation coefficient determined by the linear mixed-effects models. Colors indicate correlation types. Solid and dashed lines denote significant and insignificant correlations, respectively, based on Wald type II χ^2 tests (n= 360 biologically independent soil samples). Pairwise comparisons of environmental factors are shown in the triangle, with a color gradient denoting Pearson's correlation coefficient. b, Structural equation models (SEMs) showing the relationships among treatments, soil and plant variables, and bacterial and protistan richness. Blue and red arrows indicate positive and negative relationships, respectively. Solid or dashed lines indicate significant (p < 0.05) or nonsignificant

relationships. Numbers near the pathway arrow indicate the standard path coefficients. R² 666 represents the proportion of variance explained for every dependent variable. $\chi^2 = 47.69$, df = 34, 667 p = 0.06 (large p value indicates the predicted model and observed data are equal, i.e., good model 668 669 fitting). Comparative fit index (CFI) = 0.955, and n = 48 independent plots. c, Standardized total effects (direct plus indirect effects) derived from SEMs. d, Correlations between microbial 670 671 richness and ecosystem functioning. The color denotes the correlation coefficient determined by 672 the linear mixed-effects model. Statistical significance is based on Wald type II χ^2 tests with n 673 =360 independent soil samples. The p values were adjusted by false discovery rate and are denoted by asterisks: *** p < 0.001, ** p < 0.01, * p < 0.05. GPP: gross primary productivity; ER: 674 675 ecosystem respiration; NEE: net ecosystem exchange; Rh: heterotrophic respiration; Rs: soil total 676 respiration. 677

References

678

679

- Rands, M. R. *et al.* Biodiversity conservation: challenges beyond 2010. *Science* **329**, 1298-1303, doi:10.1126/science.1189138 (2010).
- Diaz, S., Fargione, J., Chapin, F. S., 3rd & Tilman, D. Biodiversity loss threatens human well-being. *PLoS Biol* 4, e277, doi:10.1371/journal.pbio.0040277 (2006).
- Barnosky, A. D. *et al.* Has the Earth's sixth mass extinction already arrived? *Nature* **471**, 51-57, doi:10.1038/nature09678 (2011).
- Pecl, G. T. *et al.* Biodiversity redistribution under climate change: Impacts on ecosystems and human well-being. *Science* **355**, doi:10.1126/science.aai9214 (2017).
- 5 Stocker, T. *et al.* IPCC, 2013: climate change 2013: the physical science basis.
 Contribution of working group I to the fifth assessment report of the intergovernmental panel on climate change. (2013).
- 691 6 Cardinale, B. J. *et al.* Biodiversity loss and its impact on humanity. *Nature* **486**, 59-67 692 (2012).
- Hautier, Y. *et al.* Eutrophication weakens stabilizing effects of diversity in natural grasslands. *Nature* **508**, 521-525, doi:10.1038/nature13014 (2014).
- Bascompte, J., García, M. B., Ortega, R., Rezende, E. L. & Pironon, S. Mutualistic interactions reshuffle the effects of climate change on plants across the tree of life. *Sci Adv* 5, eaav2539 (2019).
- Blois, J. L., Zarnetske, P. L., Fitzpatrick, M. C. & Finnegan, S. Climate change and the past, present, and future of biotic interactions. *Science* **341**, 499-504 (2013).
- Tylianakis, J. M., Didham, R. K., Bascompte, J. & Wardle, D. A. Global change and species interactions in terrestrial ecosystems. *Ecol lett* 11, 1351-1363 (2008).
- 702 11 Fei, S. *et al.* Divergence of species responses to climate change. *Sci Adv* **3**, e1603055, doi:10.1126/sciadv.1603055 (2017).
- Li, D., Miller, J. E. D. & Harrison, S. Climate drives loss of phylogenetic diversity in a grassland community. *Proc Natl Acad Sci USA* 116, 19989-19994, doi:10.1073/pnas.1912247116 (2019).
- Bay, R. A. *et al.* Genomic signals of selection predict climate-driven population declines in a migratory bird. *Science* **359**, 83-86, doi:10.1126/science.aan4380 (2018).

- Xue, K. *et al.* Annual Removal of Aboveground Plant Biomass Alters Soil Microbial
 Responses to Warming. *MBio* 7, doi:10.1128/mBio.00976-16 (2016).
- 711 15 Zhou, J. *et al.* Microbial mediation of carbon-cycle feedbacks to climate warming. *Nat Clim Change* 2, 106-110 (2012).
- 513 Steidinger, B. S. *et al.* Climatic controls of decomposition drive the global biogeography of forest-tree symbioses. *Nature* **569**, 404 (2019).
- 715 17 Blankinship, J. C., Niklaus, P. A. & Hungate, B. A. A meta-analysis of responses of soil biota to global change. *Oecologia* **165**, 553-565 (2011).
- 717 18 Guo, X. *et al.* Climate warming leads to divergent succession of grassland microbial communities. *Nat Clim Change* **8**, 813-818 (2018).
- 719 Guo, X. *et al.* Climate warming accelerates temporal scaling of grassland soil microbial biodiversity. *Nature Ecology & Evolution*, 1 (2019).
- Yuan, M. M. *et al.* Climate warming enhances microbial network complexity and stability. *Nature Climate Change*, doi:10.1038/s41558-021-00989-9 (2021).
- Thakur, M. P. *et al.* Climate warming promotes species diversity, but with greater taxonomic redundancy, in complex environments. *Sci Adv* **3**, e1700866, doi:10.1126/sciadv.1700866 (2017).
- Xu, X., Sherry, R. A., Niu, S., Li, D. & Luo, Y. Net primary productivity and rain-use
 efficiency as affected by warming, altered precipitation, and clipping in a mixed-grass
 prairie. Global Change Biology 19, 2753-2764 (2013).
- Luo, Y., Sherry, R., Zhou, X. & Wan, S. Terrestrial carbon-cycle feedback to climate warming: experimental evidence on plant regulation and impacts of biofuel feedstock harvest. *Gcb Bioenergy* 1, 62-74 (2009).
- 732 Chen, M.-M. *et al.* Effects of soil moisture and plant interactions on the soil microbial community structure. *European Journal of Soil Biology* **43**, 31-38 (2007).
- Zhou, J. et al. Temperature mediates continental-scale diversity of microbes in forest soils. *Nature communications* 7, 1-10 (2016).
- Rousk, J. *et al.* Soil bacterial and fungal communities across a pH gradient in an arable soil. *The ISME journal* **4**, 1340-1351 (2010).
- DeBruyn, J. M., Nixon, L. T., Fawaz, M. N., Johnson, A. M. & Radosevich, M. Global
 biogeography and quantitative seasonal dynamics of Gemmatimonadetes in soil. *Applied and environmental microbiology* 77, 6295-6300 (2011).
- 741 28 Van Horn, D. J. *et al.* Soil microbial responses to increased moisture and organic resources along a salinity gradient in a polar desert. *Applied and environmental microbiology* **80**, 3034-3043 (2014).
- 744 29 Van Nuland, M. E. *et al.* Warming and disturbance alter soil microbiome diversity and function in a northern forest ecotone. *FEMS Microbiol Ecol* (2020).
- Reimer, L. C. *et al.* Bac Dive in 2022: the knowledge base for standardized bacterial and archaeal data. *Nucleic Acids Research* **50**, D741-D746 (2022).
- Nguyen, N. H. *et al.* FUNGuild: an open annotation tool for parsing fungal community datasets by ecological guild. *Fungal Ecology* **20**, 241-248 (2016).
- Delgado-Baquerizo, M. *et al.* Microbial diversity drives multifunctionality in terrestrial ecosystems. *Nature Communications* 7, 10541, doi:10.1038/ncomms10541 (2016).
- Banerjee, S. *et al.* Agricultural intensification reduces microbial network complexity and the abundance of keystone taxa in roots. *The ISME journal* **13**, 1722-1736 (2019).

- Ning, D. *et al.* A quantitative framework reveals ecological drivers of grassland microbial community assembly in response to warming. *Nature communications* **11**, 1-12 (2020).
- 756 35 Tiedje, J. M. *et al.* Microbes and Climate Change: a Research Prospectus for the Future. *Mbio* (2022).
- 758 36 Maestre, F. T. *et al.* Increasing aridity reduces soil microbial diversity and abundance in global drylands. *Proc Natl Acad Sci USA* **112**, 15684-15689 (2015).
- Li, D., Zhou, X., Wu, L., Zhou, J. & Luo, Y. Contrasting responses of heterotrophic and autotrophic respiration to experimental warming in a winter annual-dominated prairie.
 Global Change Biology 19, 3553-3564 (2013).
- Sherry, R. A. *et al.* Lagged effects of experimental warming and doubled precipitation on annual and seasonal aboveground biomass production in a tallgrass prairie. *Global Change Biology* 14, 2923-2936 (2008).
- 766 39 Catchpole, W. & Wheeler, C. Estimating plant biomass: a review of techniques.
 767 Australian Journal of Ecology 17, 121-131 (1992).
- 768 40 McLean, E. Soil pH and lime requirement. *Methods of soil analysis. Part 2. Chemical and microbiological properties*, 199-224 (1982).
- High throughput phospholipid fatty acid analysis of soils. Applied Soil Ecology **61**, 127-130 (2012).
- Zhou, J., Bruns, M. A. & Tiedje, J. M. DNA recovery from soils of diverse composition.
 Appl. Environ. Microbiol. 62, 316-322 (1996).
- Wu, L. *et al.* Phasing amplicon sequencing on Illumina Miseq for robust environmental microbial community analysis. *BMC microbiology* **15**, 125 (2015).
- Zhou, J. *et al.* Reproducibility and quantitation of amplicon sequencing-based detection.
 ISME J 5, 1303-1313 (2011).
- Magoč, T. & Salzberg, S. L. FLASH: fast length adjustment of short reads to improve genome assemblies. *Bioinformatics* **27**, 2957-2963 (2011).
- 780 46 Edgar, R. C. Updating the 97% identity threshold for 16S ribosomal RNA OTUs.
 781 Bioinformatics 34, 2371-2375 (2018).
- 782 47 Sievers, F. *et al.* Fast, scalable generation of high-quality protein multiple sequence 783 alignments using Clustal Omega. *Molecular Systems Biology* 7, doi:10.1038/msb.2011.75 784 (2011).
- 785 48 Price, M. N., Dehal, P. S. & Arkin, A. P. FastTree 2 Approximately maximum-likelihood trees for large alignments. *PLOS ONE* 5, e9490,
 787 doi:10.1371/journal.pone.0009490 (2010).
- 788 49 Munoz, R. *et al.* Release LTPs104 of the all-species living tree. *Systematic and applied microbiology* **34**, 169-170 (2011).
- Nuccio, E. E. *et al.* Climate and edaphic controllers influence rhizosphere community assembly for a wild annual grass. *Ecology* **97**, 1307-1318 (2016).
- Wang, Q., Garrity, G. M., Tiedje, J. M. & Cole, J. R. Naive Bayesian classifier for rapid assignment of rRNA sequences into the new bacterial taxonomy. *Applied and environmental microbiology* **73**, 5261-5267 (2007).
- Nilsson, R. H. *et al.* The UNITE database for molecular identification of fungi: handling dark taxa and parallel taxonomic classifications. *Nucleic acids research* 47, D259-D264 (2018).

- Guillou, L. et al. The Protist Ribosomal Reference database (PR2): a catalog of
 unicellular eukaryote small sub-unit rRNA sequences with curated taxonomy. Nucleic
 Acids Research 41, D597-D604 (2013).
- Oliverio, A. M. *et al.* The global-scale distributions of soil protists and their contributions to belowground systems. *Science Advances* **6**, eaax8787, doi:10.1126/sciadv.aax8787 (2020).
- Andrews, S. FastQC: a quality control tool for high throughput sequence data. (Babraham Bioinformatics, Babraham Institute, Cambridge, United Kingdom, 2010).
- Li, W. & Godzik, A. Cd-hit: a fast program for clustering and comparing large sets of protein or nucleotide sequences. *Bioinformatics* **22**, 1658-1659 (2006).
- Patel, R. K. & Jain, M. NGS QC Toolkit: a toolkit for quality control of next generation sequencing data. *PloS one* **7**, e30619 (2012).
- Buchfink, B., Xie, C. & Huson, D. H. Fast and sensitive protein alignment using DIAMOND. *Nat Methods* **12**, 59-60, doi:DOI 10.1038/nmeth.3176 (2015).
- Huson, D. H., Auch, A. F., Qi, J. & Schuster, S. C. MEGAN analysis of metagenomic data. *Genome Res* 17, 377-386, doi:10.1101/gr.5969107 (2007).
- Kembel, S. W. *et al.* Picante: R tools for integrating phylogenies and ecology. *Bioinformatics* **26**, 1463-1464 (2010).
- R Core Team. R: A language and environment for statistical computing. *Vienna, Austria: R Foundation for Statistical Computing* (2014).
- Oksanen, J. et al. Package 'vegan'. Community ecology package, version 2 (2013).
- Bates, D., Mächler, M., Bolker, B. & Walker, S. Fitting linear mixed-effects models using lme4. *arXiv preprint arXiv:1406.5823* (2014).
- Fox, J. & Weisberg, S. *An R companion to applied regression*. (Sage Publications, 2018).
- 823 65 Barton, K. & Barton, M. K. Package 'mumin'. Version 1, 18 (2015).
- 824 66 Carver, R. *Practical data analysis with JMP*. (SAS Institute, 2019).

829

- Kuhn, M. Building predictive models in R using the caret package. *Journal of statistical* software **28**, 1-26 (2008).
- Rosseel, Y. lavaan: an R package for structural equation modeling and more Version 0.4-9 (BETA). (Ghent University, 2011).

Fig 1





